

Chapter 23

stevenjin8

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Comments and Proofs

Another way to think of importance sampling is that we are trying to estimate p , but some regions of p are more important than others (e.g. when $f\dot{p}$ is large for equation 23.19). Thus sampling from q lets us estimate p , but with a high resolution in regions that are more important

Initially, I thought that it would be best to sample from regions where $|\frac{d}{d\mathbf{x}} f(\mathbf{x})p(\mathbf{x})|$, since those are the regions points aren't representative of their neighbours. But then I realized that if the Hessian of $f\dot{p}$ is small, then $f(\mathbf{x})p(\mathbf{x})$ will be representative of its neighbourhood, despite not being necessarily representative of its neighbours. In that case, it might be best case to sample from regions where $|\frac{d^2}{d\mathbf{x}d\mathbf{x}^T} f(\mathbf{x})p(\mathbf{x})|$ might also be better.

Exercises

Exercise 2

The optimal M is given by

$$\begin{aligned} M &= \max \left[\frac{\text{Ga}(x|\alpha, \beta)}{\mathcal{T}(x|0, 1, 1)} \right] \\ &\propto \max \left[x^{\alpha-1} e^{-\beta x} (1 + x^2) \right]. \end{aligned}$$

Deriving the expression inside the max and setting to 0 tells us that

$$M = \left[\frac{\text{Ga}(x^*|\alpha, \beta)}{\mathcal{T}(x^*|0, 1, 1)}, \right]$$

where x^* is the solution to

$$-\beta x^3 + (\alpha + 1)x^2, -\beta x + \alpha - 1 = 0.$$

I'm not sure which of the roots to pick, but in the demo, there was only one real solution, so I just used that one.

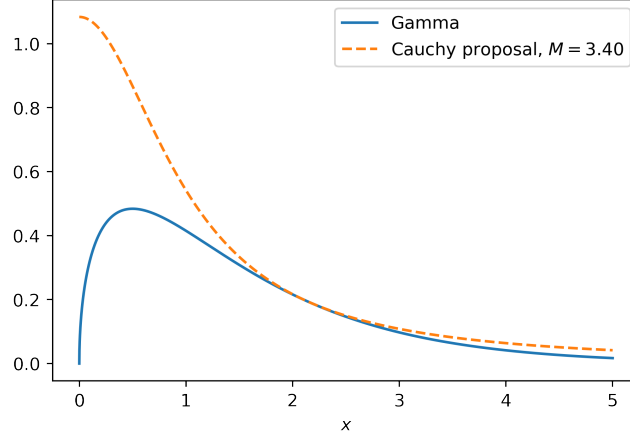


Figure 1: True gamma distribution vs Cauchy proposal with optimal M

It's pretty easy to see that rejection sampling of a gamma distribution with a Cauchy proposal is fairly inefficient. One reason is that the gamma distribution has no support for negative numbers, thus at least half of the samples will be rejected. A better idea might be to use a student T distribution with $\mu = \frac{\alpha-1}{\beta}$, $\sigma^2 = \frac{\alpha}{\beta^2}$ and $\nu = 1$. The idea here is to have both distributions have the same mode and scale (note that the Cauchy distribution does not have a variance). An even better solution would be to sample as in section 23.2 since the gamma cdf has a closed form.

Exercise 3

When the transition dynamics of a model is non-linear, exact inference is intractable, even if the observation model is linear. We can see this when looking at the distribution for \mathbf{z}_2 given $\mathbf{y}_1, \mathbf{y}_2$:

$$p(\mathbf{z}_2 | \mathbf{y}_1, \mathbf{y}_2) = \int p(\mathbf{z}_2 | f(\mathbf{z}_1), \mathbf{y}_2) p(\mathbf{z}_1 | \mathbf{y}_1) d\mathbf{z}_1.$$

This is intractable because we have f in the integral, which could be any function. However, if we use particle filtering, we no longer have to calculate the integral.

The optimal proposal distribution q is given by equation 23.54. Given the form of our model, this is equivalent to having a prior of $\mathcal{N}(f(\mathbf{z}_{t-1}), \mathbf{Q}_{t-1})$ and a likelihood of $\mathcal{N}(\mathbf{H}_t \mathbf{z}_t, \mathbf{R}_t)$. Plugging this into equation 4.125, we have

$$\begin{aligned} p(\mathbf{z}_t | \mathbf{z}_{t-1}, \mathbf{y}_t) &= \mathcal{N}(\mathbf{z}_t | \boldsymbol{\mu}, \boldsymbol{\Sigma}) \\ \boldsymbol{\Sigma}^{-1} &= \mathbf{Q}_{t-1}^{-1} + \mathbf{H}_t^T \mathbf{R}_t^{-1} \mathbf{H}_t \\ \boldsymbol{\mu} &= \boldsymbol{\Sigma} [\mathbf{H}_t^T \mathbf{R}_{t-1}^{-1} \mathbf{y}_t + \mathbf{Q}_{t-1}^{-1} f(\mathbf{z}_{t-1})]. \end{aligned}$$