# STAT 2509 Assignment 4:

1 Run the Procedures on the provided Regression Data:

a) Use the Bonnard Selection Procedure using Fo\*=4.2 (to-add-variable); We must find the Max SSIR for our one term models: XI is higest at 14829 thus, we can see that !

 $F = \frac{SSR(X_1)}{MSE(X_1)} = \frac{SSR(X_1)}{MSE(X_1)} = \frac{14829}{119.45210} = 124.14 > 6=4.2$ 

. We keep the XI term

WE Continue the Same thing on the two term model:

- XI Was Kest, So we must use a model w/ XI in it:

XIIK4 has the highest SSR at 15543

thus, we can see that!

$$F_{\lambda} = \frac{MSR(x_{4}|x_{1})}{MSE(x_{1}|x_{4})} = \frac{SSR(x_{4}|x_{1})/(dSs_{12}(x_{1}|x_{4}) - dSs_{12}(x_{1})}{MSE(x_{1}|x_{4})} - dSs_{12}(x_{1}|x_{4})$$

$$= \underbrace{[SSR(x_{1}) - SSIR(x_{1}|x_{4})]/(12 - 13)}_{69.93|49}$$

$$= \frac{14629 - 15543}{69.93149} = \frac{-714}{69.93149} = 10.21 > F_0 = 4.2$$

. We keep the X1, X4 term

We Continue for the three term model, Keeping X1, X4: XyX2,X4 has the highest SSR at 15848

$$X_{1}X_{2}X_{4}$$
 has the highest SSR at 15848  
thus, we can see that:  

$$F_{3} = \frac{MSR(X_{2}|X_{1}X_{4})}{MSE(X_{2}|X_{1}X_{4})} = \frac{SSR(X_{2}|X_{1}X_{4})}{MSE(X_{1}X_{2}X_{4})} - dfssr(x_{1}X_{2}X_{4})$$

$$= \frac{MSR(X_{1}X_{4})}{MSE(X_{1}X_{2}X_{4})} - dfssr(x_{1}X_{2}X_{4})$$

$$= \frac{SSR(X_{1}X_{4})}{MSE(X_{1}X_{2}X_{4})} - \frac{16543 - 15848}{48.59743}$$

$$= \frac{305}{48.59743} = \frac{6.28}{48.59743} > F_{0} = 4.2$$

 $= \frac{305}{48.57743} = 6.28 > F_0 = 4.2$ 

. We keep the Kilkajky term

```
We Continue for the four term model, Keeping X,, Xa, X4:
                              XIIX2, X3, X4 is the only model with ssie at 15857,8
                              thus, we can see that:
                                   F_{4} = \frac{MSR(X_{3}|X_{1}X_{2}X_{4})}{MSE(X_{1}X_{2}X_{3}X_{4})} = \frac{SSR(X_{3}|X_{1}X_{2}X_{4})/(dss_{SR}(x_{1}X_{2}X_{3}X_{4}) - dss_{SR}(x_{1}X_{2}X_{3}X_{4})}{MSE(X_{1}X_{2}X_{3}X_{4})}
                                                                                  MSE(XIX2X3X4)
                                      _ [SSR(X, XQX4)-SSIR(XXXXX)/(10-11)
                                      - [[5848-15857.8]/(-1)
52.47557
                                    =0.1868 7 < Fo=4.2
                            .. We don't add the X3 term So the best model is (X1X2X4)
                 b) Use the Backward Elimination Procedure using Fo=4.1 (to-delete-variable):

Note that (to) = Fo* | let ?=15857.8, 2=52.47557
                        F_{1} = \frac{MSR(x_{1}|x_{2}x_{3}x_{4})}{MSE_{5}} = \frac{(12277 - 7)/(10-11)}{3} = 68.20 > F_{0}
F_{2} = \frac{MSR(x_{2}|x_{1}x_{3}x_{4})}{MSE_{5}} = \frac{(15561 - 7)/(10-11)}{3} = 5.66 > F_{0}
         AS:
 MSR (An. Wazkus Xny)
                           F3 = MSR (X3/X1/2X4) = (15848 = P )/(0-11) = 0.19 < Fo
      MSEG
                           Fy = MSR(X4/X,X2X3) = (14897 - P)/(10-11) = 16.40 > Fo
        MSEA
                         4) (F3 = 0,19) 84 is the Smallest So eliminate X3 from
                                the model. let 7=48,59743
                          FI - MSIR (X1/X2/4) = (11/10-15848)/(11-12) = 97.49>Fo
     As:
   MSR(Xni |Xmaxna)
                        F_ = MSR(x2|X14) = (15543-15848)/(11-12) = 6.28>Fo
       MSES
_ (SSRg-SSEp)/(dbsse-
                         Fy = MSR(X4/X/2) = (14990-15848)/(11-12) = 17.66 > Fo
   dossar)
     MSER
                       All F values are greater then Fo So no variables can be
                     deleted.
                           Lowe remove only the X3 term from the model thus the
                               best model is XIXXX4
```

Since:

F>R.R => 75.55 > 7.074 the full model fits & is significant

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C) Use the Stepwise regression Procedure using Fox = 4.2 (to-add) & F*=4.1
        (to-delete):
Checkis Sit all one term model: forward Selection we really a LD We Keep X, fit all 2-term models: X, -> X, X4
        is X1 redundent when X4 is added into the model

F = MSIR (X1X4) = [SSR(X1)-SSIR(X1X4)]/(12-13)

MSE(X1X4) 69,73/49

= [11085-15543][-1] = 63.75
         Since Fi>Fo, we keep both X1, X4
         For 3 term models, from forward Selection, we also Keep X2 (X, 2 X4)
         IS X, redundant When X2 & X4 are in the model!
              F_{1} = \frac{MSIR(X_{1}|X_{2}X_{4})}{MSE(X_{1}X_{2}X_{4})} = \frac{\left[SSR(X_{2}X_{4}) - SSR(X_{1}X_{2}X_{4})\right]/(|I-12)}{48.69743}
= \frac{\left[11100 - 15848\right]/(-1)}{48.59743} = 97.70
              Since Fi> Fo We Keep X,
      Since Fi>Fo we keep X4
        So for we have X1/2/X4
       From Langed Selection Procedure we already Know X3 is not in the model
           11 Don't test for X3
       .. The nest Set is XIXXX3
```

2) Run the tests on the Avoided Site data;

CRD, three Steps:
Parametric test - must have Normal distribution Steps: test all three Conditions I) Harteley's Test -> tests Assumption of Constant varience I) Main test -> tests differences between treatments III) Tukeys test -> Which groups are different C.R.D Tese: Assumptions: 1) Plots are randomly assigned to 4 independent Swampy Sites 2) taken from 4 normally distributed populations 3) With equal varience, of I) To check the Assumption of equal varience use Hartey's test, we need Si's for i=1,2,3,4 where  $n=1,2=n_3=n_4=6$ , K=4,  $\pi=6$ , End=6, N=24:  $S_1^2 = \frac{\sum_{i=1}^n y_{i,i}^2 - (\sum_{i=1}^n y_{i,i}^2)^2/n_1}{\sum_{i=1}^n y_{i,i}^2} = 217.47, \sum_{i=1}^n y_{i,i}^2$ - (5.72+6.32+6.12+62+5.82+6.22)-[5.7+6.3+6.1+6+5.8+6.2]  $= \frac{217.47 - (36.1)/6}{5} = 0.0536666666$  $S_{2}^{2} = \frac{\sum_{j=1}^{n} y_{2j}^{2} - (\sum_{j=1}^{n} y_{2j})^{2}/n_{2}}{\sum_{j=1}^{n} y_{2j}^{2}} = 192.31, \sum_{j=1}^{n} y_{2j}^{2} = 33.9 = \sqrt{1 - y_{2j}^{2}}$ = 192.31-(33.9)3/6 = 0.155  $S_{3}^{2} = \frac{\sum_{i=1}^{n_{3}} y_{3i}^{2} - (\sum_{i=1}^{n_{3}} y_{3i}^{2})^{2}/n_{3}}{\sum_{i=1}^{n_{3}} y_{3i}^{2} = |72.57|} \sum_{j=1}^{n_{3}} y_{3j} = 32.1 = T_{3} = \sqrt{5}$  $= \frac{172.57 - (32.1)^2/6}{6-1} = 0.167 + Max$  $S_{4}^{2} = \frac{\sum_{i=1}^{N} j_{i} j_{i}^{2} - (\sum_{i=1}^{N} y_{i}^{2})^{2} / h_{4}}{n-1} \left| \sum_{i=1}^{N} y_{i}^{2} = 80.35, \sum_{j=1}^{N} y_{4j} = 21.9 = T_{4} = y_{4} \right|$  $=\frac{80.35-79.935}{C-1}=0.083$ 

```
Qaim! The variences are not all the same
                                                        Ho: 512=53=53=53
                                                      Ha: At least one of the or does not equal one of the others
                                                       0 = 0.01 OR 0.05
                     Test - Statistic:
                                                    F_{\text{max}} = \frac{S_{\text{max}}^2}{S_{\text{min}}^2} = \frac{0.167}{0.0536666666} = 3.1118
                 Rejection-Region:
We raject the if Frax > Frax (x,15)-1) ia = (Frax (4,5); 0.01 = 28 } for a = 0.1 we here the if Frax > Frax (x,15)-1) ia = (Frax (4,6); 0.05 = 13.7) tests
               9=0.1 So find 4> K=#sites=4, [7]=#Samples=6
           both 0,01 & 0,05 | thus, Fmox(4,6-1)ia = Fmox(4,5)ia
                                       We Can now See!
                                                                        Fmax (4,5);0.05 (Fmax (4,5);0.01
                                                                          4) Thus, we cannot Reject to => the varience of each group is equal
                                       ". We Conclude at 90% Confidence our assumption of equal varience is
                                                 not violated; Assumption is valid.
TSS = \( \frac{4}{2} \sqrt{\frac{1}{2}} \quad \frac{1}{2} \quad \quad \frac{1}{2} \quad \quad \frac{1}{2} \quad \quad \frac{1}{2} \quad \quad \frac{1}{2} \quad \quad \quad \quad \frac{1}{2} \quad \q
                                                 =\sum_{j=1}^{n}y_{1j}^{2}+\sum_{j=1}^{n}y_{2j}^{2}+\sum_{j=1}^{n}y_{3j}^{2}+\sum_{j=1}^{n}y_{3j}^{2}+\sum_{j=1}^{n}y_{3j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}y_{2j}+\sum_{j=1}^{n}
                                             = 217.47+192.31+172.57+80.35 - \frac{36.1+33.9+32.1+21.97^2}{24}
                                           =662.7 - \frac{(124)^2}{24} = 662.7 - 640\frac{2}{3} = 22\frac{1}{30} = 22.0333
                          SST_r = \sum_{i=1}^{4} \frac{T_i^2}{n_i} - \frac{\left(\sum_{i=1}^{4} \sum_{j=1}^{4} J(i)\right)^2}{n_i}
                                                                      = \frac{36.1^2}{6} + \frac{33.9^2}{6} + \frac{32.1^2}{6} + \frac{21.9^2}{6} - \left[ \frac{1}{361} 2.1^{\frac{1}{2}} \frac{1}{362} + \frac{1}{362} 2.1^{\frac{1}{2}} \frac{1}{362} + \frac{1}{362} 2.1^{\frac{1}{2}} \right]
                                                                    =660\frac{61}{150} - \frac{124^2}{24} = 19.74
```

$$SSE = TSS - SST_r = 22.0333 - 19.74 = 2.293333$$

$$MST_r = \frac{SST_r}{K-1} = \frac{19.74}{4-1} = \frac{19.74}{3} = 6.58$$

$$MSE = \frac{SSE}{n-K} = \frac{2.293333}{24-4} = 0.114666$$

$$F = \frac{MST_r}{MSE} = \frac{6.58}{0.11466666} = 57.3837 = 57.38$$

$$ANOVA Table:$$

Source	df	Sum of Squares	Mean Squares	F-Value	Pr>F	df=K-1=4-1=3
Model	3	19.74	6.58	57,38	<0.0001	df=11-K=24-4
Error	20	2.2933333	0.11466666			= 20
Corrected	23	22.033333				Total of = df+df=
Total 1	2	001100000				= 20+3=23

$$F_{\alpha(K-t,n-K)} = F_{0.10(4-1,24-4)} = F_{0.10(3,20)} = 2.38$$

LD: Since FT> Fa = D Reject Ho

Note: Our Caim for the main test

Claim: there is a difference in the mean plant growth for the Swamp Sites Ho: KI = K2=K3=K4 => the means are the same

Ha: at least one of the 16's does not equal another K

thus, As per above, we can conclude as per a 10% level of Significence there's a difference in the means of the plant growth of the Swamp Sites.

II) Turkeys Test -> Which groups are different:

1) Calculate (=)=(=)=6 pairs of []i-Ji for

Ho: Ki = Hi - D Sites are the Same I Ceaim:

Ha: Hi + Hi - Sites are different I the means are different

For i=1, j=2: |36.1-33.9| |6=2.2/6=0.366667. i=1, j=3: |36.1-32.1| |6=4/6=0.6666667. i=1, j=4: |36.1-21.9| |6=14.2/6=2.366667. i=2, j=3: |33.9-32.1| |6=1.8/6=0.3

i=2, j=4: |33.9-21.9|/6=12/6=2

i=3, i=4: 32.1-21.9 6=10.2/6=1.7

Hypothesis test:

I) Qaim: At least one median differs from the rest
Ho: Md, = Md2 = Md4
Ha: At least one of the Md's =

I) Test-Statistics:
$$\begin{aligned}
&H = \frac{12}{n(n+1)} \left[ \sum_{i=1}^{4} \frac{T_{R_{i}^{2}}}{n_{i}} - 3(n+1) \right] \\
&= \frac{12}{24(25)} \left[ \frac{119^{2}}{6} + \frac{90.5^{2}}{6} + \frac{67.5^{2}}{6} + \frac{21^{2}}{6} \right] - 3(25) \\
&= \frac{12}{600} \left[ 4603.75 \right] - 75 \\
&= 90.075 - 75 = 17.075
\end{aligned}$$

II) Rejection-Region:

We reject the if  $H > X_{a;(K-1)}^2 = X_{a,lo;(3)}^2 = 6.25$   $H = 17.075 > X_{a,lo;(3)}^2 = 6.25 - 17.075 > 6.25$  So  $H > X_{a,lo;(3)}^2 = 6.25$ 

45 We reject to Since H > x 0.10;(3)

IV) Conclusion:

We have rejected to, so we accept that that atleast one of the Md's = another of the Md's.

The attached SAS Code verifies the above results.

450 1 01 /

# Qa SAS output

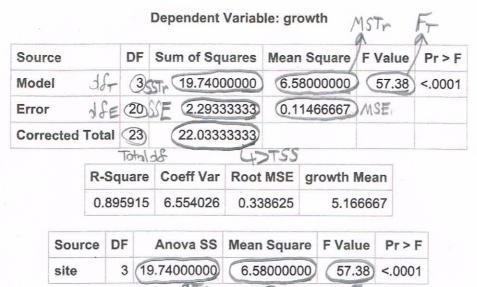
## The SAS System

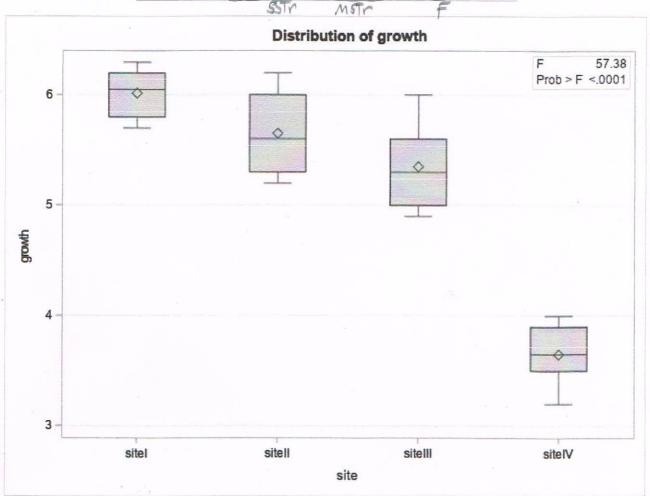
### The ANOVA Procedure

	Class Le	vel Information			
lass	Levels	Values			
ite	4	sitel sitell sitelll s	itelV		
	4	>4 sites		-l;	
Numb	er of Ob	servations Read	24	3#1-1	0.4-6-
Numb	er of Ob	servations Used	24	p# data	enjngs.

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#### The ANOVA Procedure

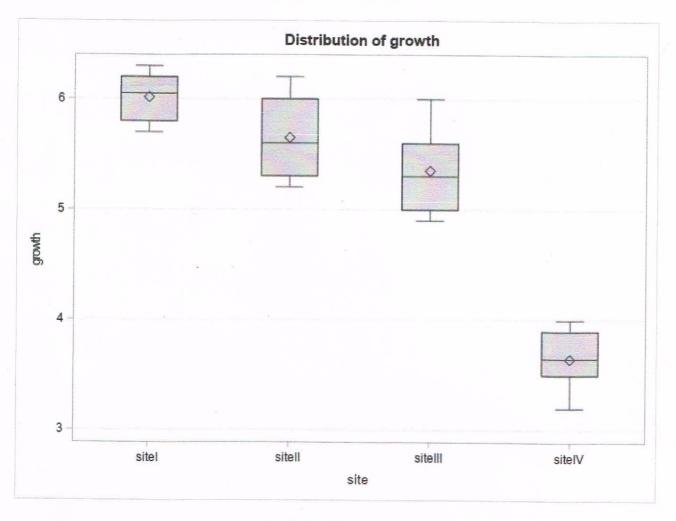




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The SAS System

### The ANOVA Procedure



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#### The ANOVA Procedure

## Tukey's Studentized Range (HSD) Test for growth

Note: This test controls the Type I experimentwise error rate.

Alpha	Q.D -D 0 = 0.1
Error Degrees of Freedom	20 Dn-K
Error Mean Square	0.114667 MSE
Critical Value of Studentized Range	(3.46154) 17 Pack, n-K)
Minimum Significant Difference	0.4785 >h.S.d.

site Comparison	Difference Between Means	en Simultaneous 90% Confidence			
sitel - sitell	[5] 0.3667	-0.1119	0.8452		MI=1
sitel - sitelll	19-43 0.6667	0.1881	1.1452	***	KI =
sitel - sitelV	151-54 2.3667	1.8881	2.8452	***	KI #
sitell - sitel	192-91 -0.3667	-0.8452	0.1119		H2=
sitell - sitelll	\(\bar{y}_2-\bar{y}_3 \) 0.3000	-0.1785	0.7785		K2=
sitell - sitelV	192-94 2.0000	1.5215	2.4785	***	Kst
sitelll - sitel	写五 -0.6667	-1.1452	-0.1881	***	K3#
sitelll - sitell	13-42 -0.3000	-0.7785	0.1785		H3=
sitellI - sitelV	73-74 1.7000	1.2215	2.1785	***	H3#
sitelV - sitel	2.3667	-2.8452	-1.8881	***	K4#
sitelV - sitell	4-41 -2.0000	-2.4785	-1.5215	***	K4 #
siteIV - siteIII	1-1.7000 July 1.7000	-2.1785	-1.2215	***	H47

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### The NPAR1WAY Procedure

XOI	Class	(Ran sified	k Sum by Va	s) for Varial riable site	ole growth
N	Sum of Scores			Std Dev Under H0	Mean Score
6	£19.00	Tirl	75.0	14.977157	19.833333
6	90.50	TRZ	75.0	14.977157	15.083333
6	69.50	TR3	75.0	14.977157	11.583333
6	21.00	TRH	75.0	14.977157	3.500000
	<b>N</b> 6 6	Class  Sum of Scores  6 (19.00) 6 (90.50) 6 (69.50)	Classified  Sum of Exp Scores Und  6 419.00 TRI  6 90.50 TR3	Classified by Value of N Scores	N Scores Under H0 Under H0 6 19.00   75.0   14.977157 6 90.50   78.3   75.0   14.977157 6 69.50   78.3   75.0   14.977157

Kruskal-Walli	s Test	
Chi-Square	17.1271	4> ∞H
DF	(3)	-DK-1
Pr > Chi-Square	0.0007	



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(3) Run the tests on the Provided Plot-treatment data:

RBD, three Steps:

I) Harteley's Test

II) Main Test

III) Turkeys Test

Parametric test

## RBD Test:

Assumptions !

1) three independent insecticides assigned randomly to Sour plots
2) Populations Correspond to each Combination of insecticide - Plot are normally distributed

3) With Equal varience, 52

4) No interactions between treatment-group Combination on the insecticide-Plot Combination

I) To check the assumption of equal varience use Hartley's test, we need  $S_i^2$ 's for i=1,2,3, where  $n_1=n_2=n_3=4$ , K=3, b=4,  $\overline{n}=b=4$ ,  $\overline{L}\overline{n}=4$ , n=bK=12:

$$S_1^2 = \frac{\sum_{i=1}^{b} y_i^2 - \left(\sum_{i=1}^{b} y_i\right)^2}{b-1} = \frac{13362 - (280)^2/4}{4-1} = 45.66$$

$$S_{2}^{2} = \frac{\sum_{j=1}^{b} y_{2j}^{2} - \left(\sum_{j=1}^{b} y_{2j}^{2}\right)^{2}}{b} = \frac{30625 - (342)^{2}/4}{4 - 1} = 58.25 \, \text{A}^{\text{max}}.$$

$$S_3^2 = \frac{\sum_{j=1}^{2} y_{3j}^2 - \sum_{b=1}^{2} y_{3j}^2}{b} = \frac{25698 - (320)^2/4}{4 - 1} = 32.66 \, \text{A}^{\text{min}}$$

$$\sum_{j=1}^{6} y_{ij} = \frac{1}{3} \frac{1}{3} \frac{1}{3} \frac{1}{3} \frac{1}{3} = \frac{1}{3} \frac{1}{3} \frac{1}{3} \frac{1}{3} = \frac{1}{3} \frac{1}{3} = \frac{1}{3} \frac{1}{3} \frac{1}{3} = \frac{1}{3} \frac{1}{3} \frac{1}{3} = \frac{1}{3} \frac{1}{3} \frac{1}{3} = \frac{1}{3} = \frac{1}{3} \frac{1}{3} = \frac$$

$$\sum_{j=1}^{5} y_{ij}^{2} = |3362| \sum_{j=1}^{5} y_{2j}^{2} = 30625| \sum_{j=1}^{5} y_{3j}^{2} = 25698$$

(Raim: The Variences are not all the Same

Ho: 
$$\sigma_1^2 = \sigma_2^2 = \sigma_3^2$$
:

Ha: At least one of the  $\sigma^2 \neq$  the others

 $\alpha = 0.01$ 

Test-Statistic:

Fmax =  $\frac{S_{max}^2}{32.66} = \frac{58.25}{32.66} = \frac{1.307}{332} = \frac{1.783163265}{1.7832}$ 

Rejection-Region:

We reject the if Fmax > Fmax(4.04-0)/a Fmax(3.3)/0.01 = 85

We reject the if Fmax > Fmax(4.04-0)/a Fmax(3.3)/0.05 = 27.8

thus, Fmax(3.4-1)/a = Fmax(3.3)/a

 $\alpha = 0.1$  So find both 0.01 2 0.05 for the fmax test:

Fmax < Fmax(3.3)/0.05 < Fmax(3.3)/0.01

Let Thus, we cannot Reject to  $\Rightarrow$  the varience of each group is equal in we conclude at 97% Conditions our assumption of equal varience is not Violated; Assumption is valid.

II) Move on to Main Test:

TSS =  $\frac{3}{3}$ 

Tile =  $\frac{1}{12}$ 
 $\frac{1}$ 

$$SSE = TSS - SST_r - SSB = 2334.9166 - 1925.166 - 386.25$$

$$SSE = 23.5 | SST = TSS - SSE = 2334.9166 - 23.5 = 2311.4166$$

$$MST_r = \frac{SST_r}{K-1} = \frac{1925.166}{3-1} = \frac{962.5833}{3-1}$$

$$MSB = \frac{8SB}{b-1} = \frac{386.25}{4-1} = \frac{128.75}{2.3}$$

$$MSE = \frac{SSE}{df} = \frac{23.5}{(K-1)(b-1)} = \frac{23.5}{2.3} = \frac{3.9166}{3.9166}$$

$$MST = \frac{SST}{df_r} = \frac{2311.4166}{5} = \frac{12-1}{462.2833}$$

$$Total df = Kb-1 = 4.3-1 = 12-1 = 11, df_E = (K-1)(b-1) = 2.3 = 6$$

$$df_T = Total df - df_E = 11-6=5$$

$$F = \frac{MST}{MSE} = \frac{462.2633}{3.9166} = \frac{118.03}{3.9166}$$

ANOVA Table:

Source	1	Sum of Sources	Mean Square	F-Value	PASF
Model	5	2311.4166	462.2833	118.03	(0.000)
Emor	6	23.5	3.9166		
Correctes		2334.9166			

	Source	36	ANOVA SS	MS .	F-Value	Pr>F		
	P106	3	386.25	128.75	32.87	0.0004		
1	in sect	2	1925.166	962.5833	245.77	40.0001		
					= 245.7 : 32.87			
	For LK-	€, bK	)=Fo.1	0(3,12)=	2.61			
	4	>F	> Fo.106	3,12) QS	118.03 >	2.61	Reject	Но

Note: Our Claim for the main test

legim: there's a difference in the mean effectiveness of the insecticides on the Plots

Ho: K1=K2=K3=> the means are the same

Ha: At least one Ws doesn't equal another K

thus, as per above, we can conclude as per a 1% level of Significence there's a difference in the means of the insecticitie effectiveness on the Plots.

III) Turkeys Test:

1) Calculate (1) = (3) = 3 Pairs of | Fi-Ji | for Ho: Ki = Ki vs. Ha: Ki + Ki for i, i=1,2,3 ; i +i

Ho: Sites are same J Claim:
tta: Sites are different J the treatments all are the Same

For i=1, i=2: 230-349/4=29.75

j=1, j=3: |230-320|/4=22.5

1=2,5=3: |349-320|14= 7.25

2) h.s.d. =  $9a(K_1(b-1)(K-1))\sqrt{MSE(1+1)}$ =  $9(3.6)(47/48)^{1/2}$  $= 3.558 (47/48)^{1/2} = 3.5207$ 

thus:

1 1- 12 = 29.75 > 3.5207 = D Ky + K2

1-13 = 22.5>3.5207 => KI = K3

13-53 = 7.25 > 3.5207 => K2 + K3

So We Know that all the sites are the Same within Limits to a 97% Considerce So, We Cannot reject Ho

45 Accept to: the Sites are all the Same means for treatment

## Non-Parametric Test:

Assume:

1) R.B.D.

2) In each insecticise-Plot Combination we have population with approximatly the Same Shape & Spread

3) No interactions between flots & insecticides

Hypothesis Test:

I) (egim: At least one median differs from the rest

Ho: Md, = Md, = Mda = Md3 = Md4

Ha: At least one of the Md's = the others

II) Test-Statistics:

$$H = \frac{12}{bK(K+1)} \left[ \sum_{i=1}^{4} T_{Ri}^{2} - 3b(K+1) \right]$$

$$= \frac{12}{3k(4)} \left[ 4^{2} + 12^{2} + 8^{2} \right] - 3(4)(3+1)$$

$$= \frac{1}{4} \left[ 224 \right] - 48 = \boxed{8}$$

IV) Conclusion:

We don't reject to & Conclude at 1% level of Significence there's an evidence to say that the medians of the insecticides does not differ.

Sits output

Q3 SAS Code

age I of 4

The SAS System

The ANOVA Procedure

	Class L	evel Info	rmation		,		
	Class	Levels	Values	>60	lumn	S	
	plot	4	1234				
	insect	3	123				
		J	> Rows				
Nui	mber of O	bservation	ons Read	12	7	# 1-4-	C. 100
Nui	mber of O	bservation	ons Used	12	1	m Uqla	Samples

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## The ANOVA Procedure

Dependent Variable: seedlings

Sour	ce		DF	Sum of So	quares	Mea	n S	quare	F Va	alue	Pr > F
Mode	el	1/7	(5	BST 2311.4	116667	46	32.28	33333	(118	3.03	<.0001
Erro	r	48	E (6	SE 23.5	500000	MSE	3.91	16667		5	F
Corr	ected	Total	(11	2334.9	16667						
			4	Total of	TSS						
	1	R-Squ	ıare	Coeff Var	Root N	ISE	see	dlings	Mea	n	
		0.989	935	2.641678	1.979057		74.9166		7		
				SUS				MSIS		7 Fiz	,
	Sour	ce D	F 4	Anova SS	Mean	Squa	are	F Valu	ıe/ F	r >	F
	<u> </u>		386.250000	128.	750000		00 32.8	0.000	.0004	4	
			925.166667	962.583		583333 245.7		77) <	.000	1	
				459	STr	C	F>/	ISTr	DE	i.	
				Conno	or 10104	1125	ř.			,	

#### The ANOVA Procedure

## Tukey's Studentized Range (HSD) Test for seedlings

Note: This test controls the Type I experimentwise error rate.

Alpha	0.01
Error Degrees of Freedom	(6
Error Mean Square	3.916667
Critical Value of Studentized Range	6.33032
Minimum Significant Difference	6.264

	. Co	mparisons s are i	ignificant at the 0.0 ndicated by ***.	1 level	
	insect Comparison	Difference Between Means	Simultaneous 99% Limits		
-2,5=3	2 - 3	7.250	0.986	13.514	***
2,0=1	2 - 1	29.750	23.486	36.014	***
3/3=2	3 - 2	-7.250	-13.514	-0.986	***
3,5=1	3 - 1	22.500	16.236	28.764	***
13=2	1 - 2	-29.750	-36.014	-23.486	***
j=3	1 - 3	-22.500	-28.764	-16.236	***

Same as Worksheet 23, Turklys Test Result

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Hi + His for all listed Companisons

## The FREQ Procedure

# Summary Statistics for insect by seedlings Controlling for plot

Statistic	Alternative Hypothesis	DF	Value	Prob
1	Nonzero Correlation	1	2.0000	0.1573
2	Row Mean Scores Differ	2	8.0000	0.0183

Total Sample Size = 12

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```
siteIII 5.4 siteIII 5.0 siteIII 6.0 siteIII 5.6 siteIII 4.9 siteIII 5.2
                                                                                                                                                                                     siteII 6.2 siteII 5.3 siteII 5.7 siteII 6.0 siteII 5.2 siteII 5.5
                                                                                                                          cards;
siteI 5.7 siteI 6.3 siteI 6.1 siteI 6.0 siteI 5.8 siteI 6.2
footnote 'Connor 101041125'
                                                                                         input site$ growth @@;
                            ods graphics off;
                                                           data ecology;
```

siteIV 3.7 siteIV 3.2 siteIV 3.9 siteIV 4.0 siteIV 3.5 siteIV 3.6 proc anova; class site; run;

proc NPAR1WAY WILCOXON; run;

means site/tukey cldiff alpha=0.10;

model growth=site;

class site;

run;

DJ SAS Code

Name: Connor Raymond Stewart ID: 101041125

```
tables plot*insect*seedlings/CMH2 scores=rank noprint;
                                                                                                                                                                                                                                                      class plot insect;
model seedlings=plot insect;
means insect/tukey cldiff alpha=0.01;
                                                              input insect plot seedlings @@;
footntoe 'Connor 101041125';
                                                                                                       1 1 56 1 2 49 1 3 65 1 2 1 84 2 2 78 2 3 94 2 3 1 80 3 2 72 3 3 83 3
                      ods graphics off;
                                         data beans;
                                                                                                                                                                                                                                                                                                                         run;
proc freq;
                                                                                                                                                                                                                                      proc anova;
                                                                                    cards;
                                                                                                                                                                                                                 run;
```

run;

Q3 SAS Coole Name: Connor Raymond Stewart

ID: 101041125