

1 Introduction

Understanding stock price dynamics is crucial for informed investment decisions. This study models and forecasts Walmart Inc.'s (WMT) daily adjusted closing prices from 2020-01-01 to 2023-12-06 with the following objectives: (1) Compute log returns, (2) Find the optimal ARIMA and GARCH models to capture volatility clustering and conditional heteroskedasticity, (3) Fit an ensemble Validate models using AIC, BIC, and residual diagnostics, and (4) Generate a 10-step ahead forecast to provide actionable insights for investors and risk managers by modeling and forecasting stock price movements.

Let P_t be the price of an asset at time t, then the log returns is defined as:

$$r_t = \log P_t - \log P_{t-1}$$

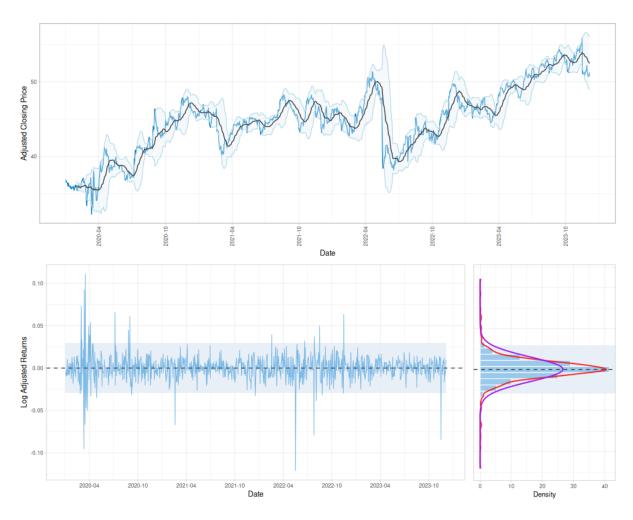


Figure 1: Walmart Inc.'s adjusted closing prices with Bollinger Bands, log returns with a 95% confidence interval, and a histogram of log returns with KDE and theoretical normal distribution.

From Figure 1, there seems to exist some volatility clusters that need to be addressed. Furthermore, the log return distribution seems to be Leptokurtic in nature, illustrating heavy tails and deviations from normality. We now proceed with rigorously testing for normality and stationarity.

2 Fitting ARIMA Model

The forecast package in R allows us to use auto.arima which automatically selects the best model based on AIC/BIC.

Model	ARIMA(0,0,1) with zero mean
MA1 Coefficient	-0.0666
Standard Error	0.0312
σ^2	0.0002266
Log-Likelihood	2694.43
AIC	-5384.85
AICc	-5384.84
BIC	-5375.1

Table 1: Summary of fitted ARMA model.

The fitted ARIMA(0,0,1) model for Walmart Inc.'s log returns is summarized in Table 1, with a moving average coefficient of -0.0666 (standard error = 0.0312), indicating weak short-term autocorrelation. The residual variance is estimated as $\sigma^2 = 0.0002266$, and model selection criteria, including AIC (-5384.85) and BIC (-5375.1), confirm its suitability.

In mathematical terms, we can write the ARIMA model as:

$$r_t = a_t - (-0.0666) \cdot a_{t-1}$$

where $a_t \sim N(0, 0.0312)$ and $\mathbb{E}[r_t] = 0$.

3 Fitting ARCH/GARCH Models

The fGarch library in R allows us to fit various GARCH models and compare performances between the them. For the purposes of keeping the models simple and explainable, the focus is on GARCH(1,0) and GARCH(1,1), using conditional distributions from c("norm", "ged", "std", "snorm", "sged", "sstd").

Model	GARCH(1,0), cond.dist="std"
mu	6.090e-04
omega	1.505e-04
alpha1	2.988e-01
shape	3.436e+00
Log-Likelihood	2908.097
AIC	-5.987829
BIC	-5.967716
SIC	-5.987863
HQIC	-5.980173

Model	GARCH(1,1), cond.dist="std"
mu	5.220e-04
omega	1.560e-05
alpha1	1.291e-01
beta1	7.905e-01
shape	4.062e+00
Log-Likelihood	2929.534
AIC	-6.029967
BIC	-6.004826
SIC	-6.030020
HQIC	-6.020398

Table 2: Summary of fitted ARCH and GARCH model.

Table 2 summarizes the fitted GARCH(1,0) and GARCH(1,1) models using a Student-t conditional distribution, which among all other conditional distribution options best maximizes log-likelihood with reasonable AIC, BIC, SIC, HQIC values. Although GARCH(1,0) achieves slightly lower AIC (-5.99 vs. -6.03) and BIC (-5.97 vs. -6.00), indicating a marginally better balance between model fit and complexity, we find that GARCH(1,1) achieves a higher log-likelihood (2929.534 vs. 2908.097) and captures long-term volatility persistence through the additional β_1 parameter, which is crucial for financial time series exhibiting volatility clustering. Therefore, we prefer GARCH(1,1) over GARCH(1,0).

In mathematical terms, we can write the GARCH(1,1) model as:

$$a_t = \sigma_t \epsilon_t$$
, $\sigma_t^2 = (1.560 \text{e-} 05) + (1.291 \text{e-} 01) a_{t-1}^2 + (7.905 \text{e-} 01) \sigma_{t-1}^2$

where $\epsilon_t \sim \text{std}(0,1)$ with shape parameter 4.0624. The relatively large value of β_1 (0.7905) relative to α_1 (0.1291) reflects the long memory in volatility, which is consistent with financial time series exhibiting volatility clustering.

4 Fitting Ensemble ARIMA(0,0,1) + GARCH(1,1) Model

A preliminary ARIMA(0,0,1) + GARCH(1,1) model with a Student-t conditional distribution is fitted to the log returns to capture volatility clustering and heavy tails. We iteratively detect and remove influential points (standardized residuals exceeding a threshold of ± 3) by fitting until the process continues until no new influential points are identified, or the changes in residual diagnostics become negligible.

Model	ARIMA(0,0,1) + GARCH(1,1)
mu	4.9198e-04
omega	6.3505e-06
alpha1	1.0224e-01
beta1	8.5001e-01
shape	$1.0000\mathrm{e}{+01}$
Log-Likelihood	2972.232
AIC	-6.246803
BIC	-6.221243
SIC	-6.246858
HQIC	-6.237065

Table 3: Summary of fitted ARMA(0,0,1)+GARCH(1,1) model.

From Table 3, we can see that there is a substantial increase in log-likelihood compared to individual GARCH(1,1) or ARIMA(0,0,1) models, with only slight increase in AIC and BIC. Furthermore, the statistically significant β_1 value (0.8501) indicates strong long-term persistence in volatility, while the moderate statistically significant α_1 value (0.1022) captures short-term shocks effectively.

5 Residual Diagnostics

Figure 2 show that the fitted ensemble model is satisfactory. The residuals over time are random around zero, the histogram aligns with a normal distribution, the Q-Q plot shows minimal deviation from normality, and the ACF plots confirm no significant autocorrelation in residuals, squared residuals, or absolute residuals, indicating no remaining structure or volatility clustering.

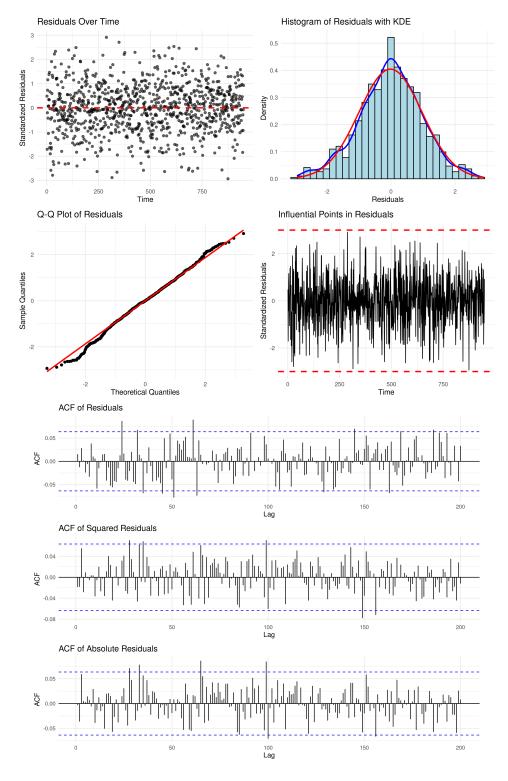


Figure 2: Residual diagnostics for the ARIMA(0,0,1) + GARCH(1,1) model.

6 10-Day Forecast

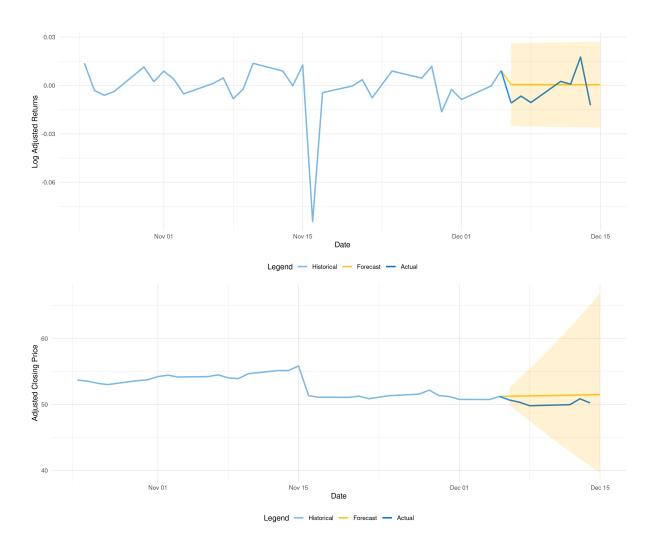


Figure 3: 10-day forecast for log returns and adjusted closing prices, including 95% confidence intervals, alongside historical and actual data for comparison. (last 30+10 days only). Blue represents historical data, yellow indicates the forecast, and dark blue corresponds to actual observed values.

Figure 3 allows us to compare our 10-day forecast with the true log returns and adjusted closing price. It is crucial to note that the model only fits data from 2020-01-01 to 2023-12-05. The actual log returns from 2023-12-06 to 2023-12-15 is used to compare model performance on unseen data.

The forecasted log returns align well with the actual values, with all observations falling within the 95% confidence interval. This indicates that the ARIMA(0,0,1) + GARCH(1,1) model effectively captures the short-term dynamics of the stock. However, the increasing width of the confidence interval for closing prices reflects the compounding effect of forecast uncertainty over time. The RMSE for the 10-day forecast, actual log returns, and adjusted prices were \$ 0.0103 and \$ 1.1693, respectively.

Given the residual diagnostics in Figure 2 affirmation that the ARIMA(0,0,1) + GARCH(1,1) model effectively captures the underlying structure of Walmart's stock returns, we can conclude that this is a good and effective model for forecasting.

Appendix

All codes and images associated with this project can be found in the GitHub repository below:

https://github.com/Stochastic1017/Walmart-Stock-Forecasting

Note: The repository will be made public **December 18**, 2024 (00:00 AM) to ensure academic integrity by preventing public access to these materials before the due date.

Figure/Table	Code
Figure 1	Code for Figure 1
Table 1	Code for Table 1
Table 2	Code for Table 2
Table 3	Code for Table 3
Figure 2	Code for Figure 2
Figure 3	Code for Figure 3

For a more comprehensive analysis and access to all the code used in this project, the following files are available in the repository. Note that each file exceeds the 5-page limit and consolidates all R code used for this project:

File	URL
Consolidated-Walmart-Stock-Forecasting.Rmd	Rmd File Link
${\tt Consolidated-Walmart-Stock-Forecasting.pdf}$	PDF File Link

Finally, the LaTeX file and the rendered pdf file associated to this submission write-up is available in this repository:

File	URL
Walmart-Stock-Forecasting.tex	TEX File Link
Walmart-Stock-Forecasting.pdf	PDF File Link

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