

Експоненциально распределена НСВ

Введем, если $X \sim \text{Exp}(\lambda)$, $\lambda > 0$, то X имеет плотность вида

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

Проверка:

$$1 = \int_{-\infty}^{\infty} f_X(x) dx = \int_0^{\infty} \lambda e^{-\lambda x} dx \stackrel{\substack{u=\lambda x \\ \frac{du}{dx}=\lambda}}{\int_0^{\infty}} e^{-u} du = [-e^{-u}]_0^{\infty} = 1$$

$$F_X(x) = \int_0^x \lambda e^{-\lambda y} dy \stackrel{\substack{u=\lambda y \\ \frac{du}{dy}=\lambda}}{\int_0^x} e^{-u} du = [-e^{-u}]_0^x = 1 - e^{-\lambda x}, \quad x > 0$$

$$\bar{F}_X(x) = P(X > x) = e^{-\lambda x}, \quad x > 0$$

$$E[X] = \int_0^{\infty} x \cdot \lambda e^{-\lambda x} dx \stackrel{\substack{x=\frac{u}{\lambda} \\ \frac{du}{dx}=\lambda}}{\int_0^{\infty}} \frac{u}{\lambda} e^{-u} du = \frac{1}{\lambda} \int_0^{\infty} u e^{-u} du = \frac{1}{\lambda} \left[[-u e^{-u}]_0^{\infty} + \int_0^{\infty} e^{-u} du \right] = \frac{1}{\lambda} [-e^{-u}]_0^{\infty} = \frac{1}{\lambda}$$

$$E[X^2] = \int_0^{\infty} x^2 \cdot \lambda e^{-\lambda x} dx \stackrel{\substack{x=\frac{u}{\lambda} \\ \frac{du}{dx}=\lambda}}{\int_0^{\infty}} \frac{u^2}{\lambda^2} e^{-u} du = \frac{1}{\lambda^2} \int_0^{\infty} u^2 e^{-u} du = \frac{1}{\lambda^2} \left[[-u^2 e^{-u}]_0^{\infty} + \int_0^{\infty} 2u e^{-u} du \right] = \frac{1}{\lambda^2} \int_0^{\infty} 2u e^{-u} du = \frac{2}{\lambda^2} \int_0^{\infty} u e^{-u} du = \frac{2}{\lambda^2}$$

$$D[X] = E[X^2] - (E[X])^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}$$

Безусловно! $\forall x > 0, P(X > x+y | X > y) = P(X > x)$

Докажем! $P(X > x+y | X > y) = \frac{P(X > x+y \cap X > y)}{P(X > y)} = \frac{P(X > x+y)}{P(X > y)} = \frac{e^{-\lambda(x+y)}}{e^{-\lambda y}} = e^{-\lambda x} \stackrel{x > 0}{=} P(X > x)$

Съвместно непрекъснат разпределение

Деф. Казваме, че $X = (X_1, \dots, X_n)$ е вектор от Н.С.В., ако $f_X = f_{X_1, \dots, X_n}: \mathbb{R}^n \rightarrow [0, \infty)$ Такава, че:

а) $f_X(x_1, \dots, x_n) \geq 0, \forall (x_1, \dots, x_n) \in \mathbb{R}^n$

б) $\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_X(x_1, \dots, x_n) dx_1 \dots dx_n = 1$

в) $D \subseteq \mathbb{R}^n, P(X \in D) = \int_D f_X(x_1, \dots, x_n) dx_1 \dots dx_n$

Деф. (Носител на сл. вел.)

Нека f_X е свм. плътност на X . Тогава $D_{f_X} = \{(x_1, \dots, x_n) \in \mathbb{R}^n : f_X(x_1, \dots, x_n) > 0\}$ се нарича носител на X

Деф. (Маргинална плътност)

Нека f_X е свм. плътност на $X = (X_1, \dots, X_n)$. Тогава:

$$f_{X_j}(x_j) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_X(x_1, \dots, x_{j-1}, x_j, x_{j+1}, \dots, x_n) dx_1 \dots dx_{j-1} dx_{j+1} \dots dx_n$$

е маргиналната плътност на X_j за някое $j = \overline{1, n}$

Деф. (Условна плътност, $n=2$)

Нека f_X е свм. плътност на $X = (X_1, X_2)$. Тогава, ако $f_{X_1}(x_1) > 0$,

$$f_{X_2|X_1}(x_2|x_1) = \frac{f_X(x_1, x_2)}{f_{X_1}(x_1)}$$
 е условна плътност на X_2 при X_1

Деф. Нека $X = (X_1, \dots, X_n)$ е в-р от сл. вел. Тогава $F_X(x_1, \dots, x_n) = P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_n \leq x_n)$ е свм. ф-я на разпределение.

Ако f_X съществува, $F_X(x_1, \dots, x_n) = \int_{y_1=0}^{x_1} \int_{y_2=0}^{x_2} \dots \int_{y_n=0}^{x_n} f_X(y_1, \dots, y_n) dy_1 \dots dy_n$

$$\frac{\partial^n F_X}{\partial x_1 \dots \partial x_n} = f_X$$

Деф. Казваме, че $X_1 \perp X_2$, ако $F_X(x_1, x_2) = F_{X_1}(x_1) F_{X_2}(x_2)$ ($X = (X_1, X_2)$)

$$P(X_1 \leq x_1, X_2 \leq x_2) = P(X_1 \leq x_1) P(X_2 \leq x_2), \forall (x_1, x_2) \in \mathbb{R}^2$$

Също така, ако $\forall (x_1, x_2) \in \mathbb{R}^2: f_X(x_1, x_2) = f_{X_1}(x_1) f_{X_2}(x_2)$, то $X_1 \perp X_2$

Опред. (Независимост в совокупности)

Нека $X = (X_1, \dots, X_n)$ е в-р от непр. сл. вел. Тогава казваме, че X_1, \dots, X_n са независими в совокупности, ако $f_X(x_1, \dots, x_n) = \prod_{j=1}^n f_{X_j}(x_j), \forall (x_1, \dots, x_n)$

Нека X има f_X и $g: \mathbb{R}^n \rightarrow \mathbb{R}$, то

$$E[g(X)] = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(x_1, \dots, x_n) f_X(x_1, \dots, x_n) dx_1 \dots dx_n$$

Следствие Нека $X = (X_1, X_2)$ е вектор от НСВ, т.е. $E[X_1] < \infty$ и $E[X_2] < \infty$

Тогава $E[X_1 + X_2] = E[X_1] + E[X_2]$

Доказателство $g(x_1, x_2) = x_1 + x_2$

$$\begin{aligned} E[X_1 + X_2] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x_1 + x_2) f_X(x_1, x_2) dx_2 dx_1 = \int_{-\infty}^{\infty} x_1 \underbrace{\int_{-\infty}^{\infty} f_X(x_1, x_2) dx_2}_{f_{X_1}(x_1)} dx_1 + \int_{-\infty}^{\infty} x_2 \underbrace{\int_{-\infty}^{\infty} f_X(x_1, x_2) dx_1}_{f_{X_2}(x_2)} dx_2 \\ &= \int_{-\infty}^{\infty} f_{X_1}(x_1) x_1 dx_1 + \int_{-\infty}^{\infty} f_{X_2}(x_2) x_2 dx_2 = E[X_1] + E[X_2] \end{aligned}$$

Следствие Нека $X = (X_1, X_2)$ е в-р от НСВ, т.е. $X_1 \perp\!\!\!\perp X_2 \Rightarrow D[X_1 + X_2] = D[X_1] + D[X_2]$

Незав. в совокупности $(X_1, \dots, X_n) \Rightarrow D[X_1 + \dots + X_n] = \sum_{j=1}^n D[X_j]$

Теорема (Смяна на променливите)

Нека $X=(X_1, X_2)$ е в-р от НСВ и има $g: \mathbb{R}^2 \rightarrow \mathbb{R}^2$

Нека f_X е сввм. плътност.

Нека $Y=(Y_1, Y_2)=g(X)=(g_1(X_1, X_2), g_2(X_1, X_2))$

Нека $g: \mathcal{D}_{f_X} \rightarrow \mathbb{R}^2$ е взаимно еднозначна

Нека $g(\mathcal{D}_{f_X})=\{y \in \mathbb{R}^2 : y=g(x) \text{ за } x \in \mathcal{D}_{f_X}\}$

Нека $X=h(Y)=(h_1(Y_1, Y_2), h_2(Y_1, Y_2))$, където $h=g^{-1}$

Нека h и g са непр. и диф. съответно в $g(\mathcal{D}_{f_X})$ и \mathcal{D}_{f_X} .

Нека $J(y)=\left| \det \begin{pmatrix} \frac{\partial h_1(y)}{\partial y_1} & \frac{\partial h_1(y)}{\partial y_2} \\ \frac{\partial h_2(y)}{\partial y_1} & \frac{\partial h_2(y)}{\partial y_2} \end{pmatrix} \right| \neq 0, \forall y \in g(\mathcal{D}_{f_X})$

Тогав сввм. пл. на Y е $f_Y(y)=f_X(h(y))|J(y)|, \forall y \in g(\mathcal{D}_{f_X})$

Доказателство $A \subseteq g(\mathcal{D}_{f_X})$ и искаме да потвърдим, че $P(Y \in A) = \iint_A f_Y(y_1, y_2) dy_1 dy_2$

$$P(Y \in A) = P(g(X) \in A) = P(X \in h(A)) = \int_{X \in h(A)} f_X(x) dx \stackrel{x=h(y)}{=} \int_{y \in A} f_X(h(y)) |J(y)| dy \Rightarrow$$

$\Rightarrow f_X(h(y))|J(y)|$ е плътността на Y .

⊕ Нека V_1, V_2, \dots, V_n са независими в съвкупност НСВ, $V_i \sim N(\mu_i, \sigma_i^2)$. Тогав $\sum_{i=1}^n V_i \sim N(\sum_{i=1}^n \mu_i, \sum_{i=1}^n \sigma_i^2)$. Ще докажем, че е изпълнено за $n=2$. От мат. индукция следва $\forall n \in \mathbb{N}$.

$V_1 + V_2 = \mu_1 + \sigma_1 Z_1 + \mu_2 + \sigma_2 Z_2$, където $Z_1, Z_2 \sim N(0, 1)$ и $Z_1 \perp Z_2 \Rightarrow V_1 + V_2 = (\mu_1 + \mu_2) + \sigma_1 Z_1 + \sigma_2 Z_2$

Поставяме $X_1 = \sigma_1 Z_1 \sim N(0, \sigma_1^2)$ и $X_2 = \sigma_2 Z_2 \sim N(0, \sigma_2^2) \Rightarrow V_1 + V_2 = \mu_1 + \mu_2 + X_1 + X_2$

Ако $X_1 + X_2 \sim N(0, \sigma_1^2 + \sigma_2^2)$, то $V_1 + V_2 \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$

$$f_{X_1, X_2}(x_1, x_2) = f_{X_1}(x_1) f_{X_2}(x_2) = \frac{1}{\sqrt{2\pi}\sigma_1} \cdot e^{-\frac{x_1^2}{2\sigma_1^2}} \cdot \frac{1}{\sqrt{2\pi}\sigma_2} \cdot e^{-\frac{x_2^2}{2\sigma_2^2}} > 0$$

$$\begin{cases} y_1 = x_1 + x_2 \\ y_2 = x_2 \end{cases} \Rightarrow y = g(x) \Leftrightarrow x = h(y) = (y_1 - y_2, y_2) \quad J(y) = \left| \det \begin{pmatrix} 1 & -1 \\ 0 & 1 \end{pmatrix} \right| = 1$$

$$f_Y(y) = f_X(h(y)) \cdot 1 = \frac{1}{\sqrt{2\pi}\sigma_1\sigma_2} e^{-\frac{(y_1 - y_2)^2}{2\sigma_1^2} - \frac{y_2^2}{2\sigma_2^2}}, \forall y \in \mathbb{R}^2$$

$$f_{Y_1}(y_1) = \frac{1}{\sqrt{2\pi}\sigma_1\sigma_2} \int_{-\infty}^{\infty} e^{-\frac{(y_1 - y_2)^2}{2\sigma_1^2} - \frac{y_2^2}{2\sigma_2^2}} dy_2 = \frac{\frac{(y_1 - y_2)^2}{\sigma_1^2} + \frac{y_2^2}{\sigma_2^2} = cy_1^2 + (by_2 - ay_1)^2}{\sqrt{2\pi}\sigma_1\sigma_2} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(by_2 - ay_1)^2} dy_2 \stackrel{by_2 = w}{=} \int_{-\infty}^{\infty} e^{-\frac{1}{2}w^2} dw$$

$$= \frac{e^{-\frac{y_1^2}{2}}}{\sqrt{2\pi}\sigma_1\sigma_2 b} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(w - ay_1)^2} dw \stackrel{v=w-ay_1}{=} \frac{e^{-\frac{y_1^2}{2}}}{\sqrt{2\pi}\sigma_1\sigma_2 b} \int_{-\infty}^{\infty} e^{-\frac{1}{2}v^2} dv = \frac{e^{-\frac{y_1^2}{2}}}{\sqrt{2\pi}\sigma_1\sigma_2 b} \stackrel{\text{Гauss} \Rightarrow \sqrt{2\pi}}{=} \frac{1}{\sqrt{2\pi}\sigma_1\sigma_2}$$

Стега да намерим c и b

$$\frac{(y_1 - y_2)^2}{\sigma_1^2} + \frac{y_2^2}{\sigma_2^2} = cy_1^2 + (by_2 - ay_1)^2 \Leftrightarrow \frac{y_1^2}{\sigma_1^2} - \frac{2y_1y_2}{\sigma_1^2} + \frac{y_2^2}{\sigma_1^2} + \frac{y_2^2}{\sigma_2^2} = y_2^2 \underbrace{\left(\frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)}_{b^2} - \frac{2y_1y_2}{\sigma_1^2} + \frac{y_1^2}{\sigma_1^2}$$

$$\left(y_2^2 b^2 - \frac{2y_1y_2}{\sigma_1^2} \cdot \frac{b}{b} + \frac{y_1^2}{\sigma_1^4 b^2} \right) - \frac{y_1^2}{\sigma_1^4 b^2} + \frac{y_1^2}{\sigma_1^2} = \left(by_2 - \frac{y_1}{\sigma_1^2 b} \right)^2 - \frac{y_1^2}{\sigma_1^4 b^2} + \frac{y_1^2}{\sigma_1^2}$$

$$= \underbrace{\left(by_2 - \frac{y_1}{\sigma_1^2 b} \right)^2}_{a \cdot y_1} + y_1^2 \underbrace{\left(\frac{1}{\sigma_1^2} - \frac{1}{\sigma_1^4 b^2} \right)}_c$$

$$b^2 = \frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \Rightarrow \frac{\sigma_1^2 + \sigma_2^2}{\sigma_1^2 \sigma_2^2}$$

$$a^2 = \frac{1}{\sigma_1^4 b^2} = \frac{\sigma_2^2}{\sigma_1^2 (\sigma_1^2 + \sigma_2^2)}$$

$$c = \frac{1}{\sigma_1^2} - \frac{1}{\sigma_1^4 b^2} \cdot \frac{1}{\sigma_1^2} = \frac{\sigma_2^2}{\sigma_1^2 (\sigma_1^2 + \sigma_2^2)} = \frac{\sigma_1^2}{\sigma_1^2 (\sigma_1^2 + \sigma_2^2)} = \frac{1}{\sigma_1^2 + \sigma_2^2}$$

$$\Rightarrow f_{y_1}(y_1) = \frac{e^{-\frac{y_1^2}{2}} \cdot e}{\sqrt{2\pi} \sigma_1 \sigma_2 b} = \frac{e^{-\frac{y_1^2}{2}} \cdot \frac{1}{\sigma_1^2 + \sigma_2^2}}{\sqrt{2\pi} \sigma_1 \sigma_2 \frac{\sigma_1^2 + \sigma_2^2}{\sigma_1^2 \sigma_2^2}} = \frac{1}{\sqrt{2\pi} (\sigma_1 + \sigma_2)} \cdot e^{-\frac{y_1^2}{2(\sigma_1^2 + \sigma_2^2)}} \quad \text{— ПЛБННДСН НА } N(0, \sigma_1^2 + \sigma_2^2)$$

$$\Rightarrow y_1 = x_1 + x_2 \sim N(0, \sigma_1^2 + \sigma_2^2)$$

□