# STA410HW1

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## 1 DCT to Denoise an Image

#### 1.1 Matrix Transformation

*Proof.* Let  $\hat{Z} = A_m Z A_m^T$ , a  $m \times n$  matrx be given. Let  $\{A_n\}_n$  be a family of matrices satisfying  $A_n^T A_n = D_n$ , where  $D_n$  is diagonal. Consider the following:

$$D_{m}^{-1} A_{m}^{T} \hat{Z} A_{n} D_{n}^{-1} = D_{m}^{-1} A_{m}^{T} A_{m} Z A_{m}^{T} A_{n} D_{n}^{-1}$$

$$= D_{m}^{-1} (A_{m}^{T} A_{m}) Z (A_{m}^{T} A_{n}) D_{n}^{-1}$$

$$= D_{m}^{-1} D_{m} Z D_{n} D_{n}^{-1}$$

$$= Z$$

$$(1)$$

1.2 Threshold Transformation

The following function are the r function for threshold transformation, denoiceH is the hard-thresholding and denoiceS stands for soft-thresholding:

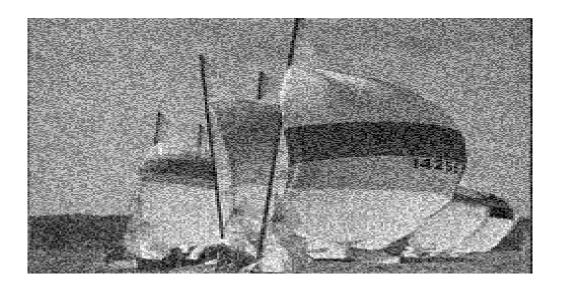
```
denoiseH <- function(dctmat, quant) {</pre>
  # Do the DCT on matrix
  dctmatT <- mvdct(dctmat)</pre>
  # if quant is missing, set it to the 0.8
  if(missing(quant)) {lambda <- quantile(abs(dctmatT), 0.8)}</pre>
  else {lambda <- quantile(abs(dctmatT),quant)}</pre>
  # hard-thresholding
  a <- dctmatT[1,1]
  dctmat1 <- ifelse(abs(dctmatT)>lambda,dctmatT,0)
  dctmat1[1,1] \leftarrow a
  # inverse DCT to obtain denoised image "clean"
  clean <- mvdct(dctmat1,inverted=T)</pre>
  clean <- ifelse(clean<0,0,clean)</pre>
  clean <- ifelse(clean>1,1,clean)
  clean
}
denoiseS <- function(dctmat,lambda) {</pre>
  # Do the DCT on matrix
  dctmatT <- mvdct(dctmat)</pre>
  # if lambda is missing, set it to the 0.8 quantile of abs(dctmat)
  if(missing(lambda)) lambda <- quantile(abs(dctmatT), 0.8)</pre>
  # soft-thresholding
  a <- dctmatT[1,1]
  dctmat1 <- sign(dctmatT)*pmax(abs(dctmatT)-lambda, 0)</pre>
  dctmat1[1,1] <- a
  # inverse DCT to obtain denoised image "clean"
```

```
clean <- mvdct(dctmat1,inverted=T)
clean <- ifelse(clean<0,0,clean)
clean <- ifelse(clean>1,1,clean)
clean
}
```

### 1.3 Denoise Methods and Results

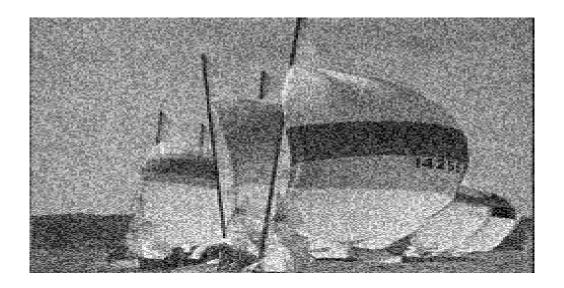
First of all, take a look at the default image.

```
image(boat, axes=F, col=grey(seq(0,1,length=256)))
```

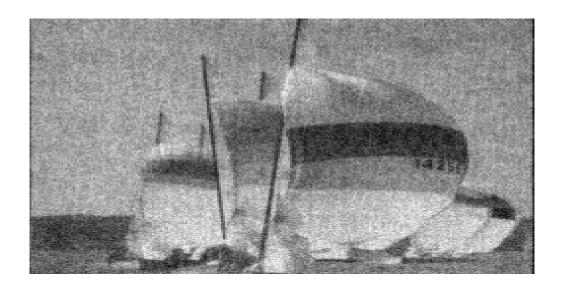


Now, try the ones we have above with different threshold  $\lambda$ 

```
# Hard-threshold with quantile 0.4
K <- denoiseH(boat, 0.4)
image(K, axes=F, col=grey(seq(0,1,length=256)))</pre>
```



```
# Soft-threshold with lambda = 12
K <- denoiseS(boat, 12)
image(K, axes=F, col=grey(seq(0,1,length=256)))</pre>
```



# 2 Hermite Distribution and FFT

### 2.1 PGF

*Proof.* Let U, V be independent Poisson r.v. with means  $\lambda_u$  and  $\lambda_v$ . Define X = U + 2V Let  $g_U(s)$ ,  $g_V(s)$  and  $g_X(s)$  denote the pgf of U, V, X respectively.

$$P(U=k) = \frac{\lambda_u^k e^{-\lambda}}{k!} \tag{2}$$

$$P(V=k) = \frac{\lambda_v^k e^{-\lambda}}{k!} \tag{3}$$

Now, using equation 2 and 3 together with taylor expansion of e

$$g_{U}(s) = E(s^{U})$$

$$= \sum_{j=0}^{\infty} P(U = j)s^{j}$$

$$= \sum_{j=0}^{\infty} \frac{\lambda_{u}^{j} e^{-\lambda_{u}}}{j!} s^{j}$$

$$= e^{-\lambda_{u}} \sum_{j=0}^{\infty} \frac{(\lambda_{u}s)^{j}}{j!}$$

$$= e^{-\lambda_{u}} \cdot e^{\lambda_{u}s}$$

$$= e^{\lambda_{u}(s-1)}$$

$$(4)$$

Similarly,

$$g_{2V}(s) = E(s^{2V})$$

$$= \Sigma_{j=0}^{\infty} P(V = j) s^{2j}$$

$$= \Sigma_{j=0}^{\infty} \frac{\lambda_{v}^{j} e^{-\lambda_{v}}}{j!} s^{2j}$$

$$= e^{-\lambda_{v}} \Sigma_{j=0}^{\infty} \frac{(\lambda_{v} s^{2})^{j}}{j!}$$

$$= e^{-\lambda_{v}} \cdot e^{\lambda_{v} s^{2}}$$

$$= e^{\lambda_{v}(s^{2}-1)}$$

$$(5)$$

Now we put everything together, and the fact that U and V are independent.

$$g_X(s) = E(s^X)$$

$$= E(s^{(U+2V)})$$

$$= E(s^U) \cdot E(s^{2V})$$

$$= e^{\lambda_u(s-1)} \cdot e^{\lambda_v(s^2-1)}$$

$$= e^{\lambda_u(s-1) + \lambda_v(s^2-1)}$$
(6)

2.2 Finding M

*Proof.* Fix some  $\epsilon > 0$ . By Markov's Inequality, we have

$$P(X \ge M) = P(s^X \ge s^M) \le \frac{E(s^X)}{s^M} = \frac{\exp[\lambda_u(s-1) + \lambda_v(s^2 - 1)]}{s^M}$$
 (7)

Now, to ensure  $P(X \ge M) \le \epsilon$ , we can first determine a  $M^*$  for each s > 1 such that,

$$\epsilon = \frac{E(s^X)}{s^{M^*}} \ge P(X \ge M^*)$$

$$\epsilon = \frac{\exp(\lambda_u(s-1) + \lambda_v(s^2 - 1))}{s^{M^*}}$$

$$\Rightarrow s^{M^*} = \frac{\exp(\lambda_u(s-1) + \lambda_v(s^2 - 1))}{\epsilon}$$

$$\Rightarrow M^* \ln(s) = \lambda_u(s-1) + \lambda_v(s^2 - 1) - \ln(\epsilon)$$

$$\Rightarrow M^* = \frac{\lambda_u(s-1) + \lambda_v(s^2 - 1) - \ln(\epsilon)}{\ln(s)}$$
(9)

In fact, we can view  $M^*$  as a function of s, ie.  $M^*(s)$ . Then we take

$$M = \inf_{s>1} M^*(s) = \inf_{s>1} \frac{\lambda_u(s-1) + \lambda_v(s^2 - 1) - \ln(\epsilon)}{\ln(s)}$$
(10)

To complete the proof, fix  $\delta > 0$ , there exists an s > 1 for  $M^*(s)$  such that

$$M + \delta > M^*(s) \tag{11}$$

Then we have:

$$\frac{\exp(\lambda_u(s-1) + \lambda_v(s^2 - 1))}{s^{M+\delta}} < \frac{\exp(\lambda_u(s-1) + \lambda_v(s^2 - 1))}{s^{M^*(x)}}$$

$$\Rightarrow \frac{\exp(\lambda_u(s-1) + \lambda_v(s^2 - 1))}{s^{M+\delta}} < \epsilon$$
(12)

$$\Rightarrow \frac{\exp(\lambda_u(s-1) + \lambda_v(s^2 - 1))}{s^{M+\delta}} < \epsilon \tag{13}$$

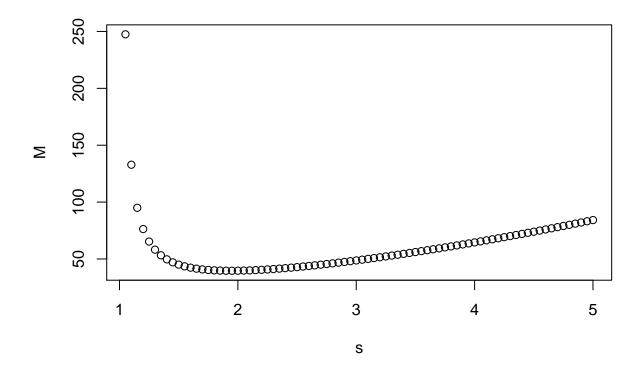
Now shrink  $\delta \to 0$ , we can conclude that

$$P(X \ge M) \le \frac{E(s^X)}{s^M} = \frac{\exp(\lambda_u(s-1) + \lambda_v(s^2 - 1))}{s^M} \le \epsilon$$
(14)

#### FFT for Distribution 2.3

Before we determine value of M, we first make some plots of Mvs.s with fixed  $\epsilon, \lambda_u, \lambda_v$ 

Mplot(20, 1, 5)

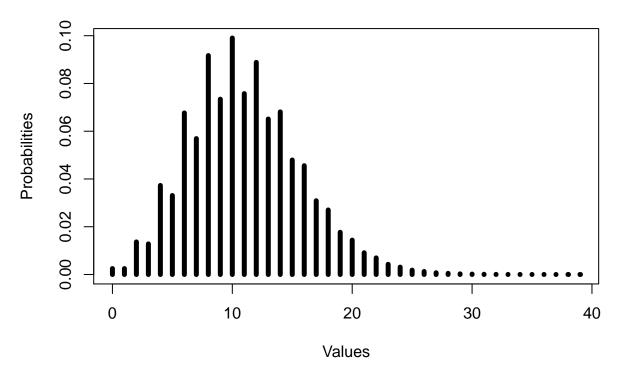


We can observe that the min of M is reached when s is approximately 2 Now we evalute the value M and do

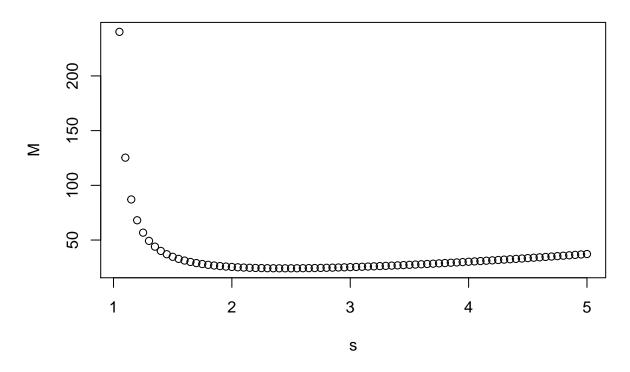
## [1] 40

```
p1 <- eval(M, 1, 5)
p1
## $x
                        5
                             7 8
                                   9 10 11 12 13 14 15 16 17 18 19 20 21 22
## [24] 23 24 25 26 27 28 29 30 31 32 33 34 35 36 37 38 39
##
## $prob
   [1] 2.479113e-03 2.478928e-03 1.363323e-02 1.280693e-02 3.728459e-02
    [6] 3.307070e-02 6.765273e-02 5.690851e-02 9.167947e-02 7.341829e-02
## [11] 9.902129e-02 7.574583e-02 8.882990e-02 6.509909e-02 6.809986e-02
## [16] 4.793939e-02 4.555863e-02 3.087956e-02 2.702588e-02 1.767481e-02
## [21] 1.439668e-02 9.102134e-03 6.957679e-03 4.259957e-03 3.076531e-03
## [26] 1.827044e-03 1.253552e-03 7.231109e-04 4.735226e-04 2.656769e-04
## [31] 1.666968e-04 9.107956e-05 5.493897e-05 2.926468e-05 1.701925e-05
## [36] 8.847602e-06 4.973335e-06 2.525658e-06 1.375237e-06 6.828672e-07
To clearly illistrate this result, let's have a plot.
plot(p1$x, p1$prob, type='h', lwd=5,xlab= 'Values',
     ylab = 'Probabilities', main ='Approximation to a Hermite distribution 1')
```

# Approximation to a Hermite distribution 1



Now we do the same thing for part (ii)



```
M \leftarrow Mval(20, 0.1, 2)
М
## [1] 25
p2 <- eval(M, 0.1, 2)
p2
## $x
               2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22
## [1] 0 1
## [24] 23 24
##
## $prob
## [1] 1.224565e-01 1.224581e-02 2.455252e-01 2.451172e-02 2.461379e-01
## [6] 2.453212e-02 1.645008e-01 1.636836e-02 8.245502e-02 8.190995e-03
## [11] 3.306392e-02 3.279125e-03 1.104863e-02 1.093951e-03 3.164566e-03
## [16] 3.128173e-04 7.930966e-04 7.826936e-05 1.766785e-04 1.740765e-05
## [21] 3.542274e-05 3.484422e-06 6.456336e-06 6.340574e-07 1.078698e-06
plot(p2$x, p2$prob, type='h', lwd=5, xlab= 'Values',
    ylab = 'Probabilities', main ='Approximation to a Hermite distribution 2')
```

# **Approximation to a Hermite distribution 2**

