

Chapter 1

CLT for inid Sequences

1.1 Notations and Definitions

A random variable z is r th *integrable* if $E[|z|^r] < \infty$. Equivalently,

$$\lim_{M \rightarrow \infty} E[|z|^r \mathbb{I}\{|z|^r > M\}] = 0,$$

where $\mathbb{I}\{\cdot\}$ is the indicator function. Without referring explicitly to the r th moment, we say z is *integrable* if $r = 1$, and *square integrable* if $r = 2$.

A *triangular array* $\{(x_{1n}, x_{2n}, x_{3n}, \dots, x_{r_n n})\}_{n \in \mathbb{N}}$ stacks like a triangular:

$$\begin{pmatrix} x_{11} & x_{21} & \cdots & x_{r_1 1} & & & \\ x_{12} & x_{22} & \cdots & \cdots & x_{r_2 2} & & \\ x_{13} & x_{23} & \cdots & \cdots & \cdots & x_{r_3 3} & \\ \vdots & & & & & & \ddots \\ x_{1n} & x_{2n} & \cdots & \cdots & \cdots & \cdots & \cdots & x_{r_n n} \end{pmatrix}$$

Here r_n is an increasing number in n , and $r_n \rightarrow \infty$ as $n \rightarrow \infty$. Suppose for each n , the elements in $(x_{in})_{i=1}^{r_n}$ are independently non-identically distributed (inid). (Please keep a liberal mind and consider “identically distributed” as a special case of “non-identically distributed”).

Without loss of generality, assume $E[x_{in}] = 0$ for all i and n and denote $\sigma_{in}^2 = E[x_{in}^2]$. Define the *partial sum* (up to n) as $S_n = \sum_{i=1}^{r_n} x_{in}$ and (the n th) *aggregate variance* as $\tilde{\sigma}_n^2 = \sum_{i=1}^{r_n} \sigma_{in}^2$.

1.2 Lindeberg Condition

Lindeberg-Lévy Central Limit Theorem is for independently and identically distributed (iid) sequences. In this lecture we consider independent but heterogeneous sequences.

Definition 1.1. Lindeberg Condition:

$$\lim_{n \rightarrow \infty} \frac{1}{\tilde{\sigma}_n^2} \sum_{i=1}^{r_n} E[x_{in}^2 \mathbb{I}\{x_{in}^2 \geq \varepsilon \tilde{\sigma}_n^2\}] = 0$$

for all $\varepsilon > 0$.

Theorem 1.1 (Lindeberg-Feller CLT). *If the triangular array $\{(x_{in})_{i=1}^{r_n}\}_{n \in \mathbb{N}}$ satisfies the Lindeberg condition, then*

$$\frac{S_n}{\tilde{\sigma}_n} \xrightarrow{d} N(0, 1)$$

Lindeber-Feller CLT allows heterogeneity across $i = 1, \dots, r_n$. It includes *Lindeberg-Lévy CLT* as a special case. To see this fact, under iid let us use z to represent the homogeneous distribution. Denote $\text{var}(z) = \sigma_z^2 \in (0, \infty)$, and equivalently $\lim_{M \rightarrow \infty} E[z^2 \mathbb{I}\{z^2 \geq M\}] = 0$ (square integrability). Set $r_n = n$, and thus $\tilde{\sigma}_n^2 = n\sigma_z^2$. As a result,

$$\begin{aligned} \frac{1}{\tilde{\sigma}_n^2} \sum_{i=1}^n E[x_{in}^2 \mathbb{I}\{x_{in}^2 \geq \varepsilon \tilde{\sigma}_n^2\}] &= \frac{1}{n\sigma_z^2} \times n E[z^2 \mathbb{I}\{z^2 \geq n\sigma_z^2 \varepsilon\}] \\ &= \text{const} \times E[z^2 \mathbb{I}\{z^2 \geq n\sigma_z^2 \varepsilon\}] \rightarrow \infty \end{aligned}$$

since $n\sigma_z^2 \varepsilon \rightarrow \infty$ as $n \rightarrow \infty$.

With iid and $r_n = n$, we can drop the subscript n and write $z_i = x_{in}$. The ratio

$$\frac{S_n}{\tilde{\sigma}_n} = \frac{\sum_{i=1}^n z_i}{\sqrt{n\sigma_z^2}} = \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{z_i}{\sigma_z}$$

retains its familiar form in CLT.

1.3 Lyapunov Condition

Lindeberg condition is difficult to interpret. Lyapunov condition is a more interpretable sufficient condition.

Definition 1.2. Lyapunov Condition: There exists some $\delta > 0$ such that

$$\lim_{n \rightarrow \infty} \frac{1}{\tilde{\sigma}_n^{2+\delta}} \sum_{i=1}^{r_n} E[|x_{in}|^{2+\delta}] = 0$$

Lyapunov condition implies Lindeberg condition. To see this fact, we use the quantify in the Lindeberg condition as a starting point:

$$\begin{aligned} \frac{1}{\tilde{\sigma}_n^2} \sum_{i=1}^{r_n} E[x_{in}^2 \mathbb{I}\{x_{in}^2 \geq \varepsilon \tilde{\sigma}_n^2\}] &= \frac{1}{\tilde{\sigma}_n^2} \sum_{i=1}^{r_n} E\left[|x_{in}|^2 \mathbb{I}\left\{|x_{in}|^\delta \geq \varepsilon^{\frac{\delta}{2}} \tilde{\sigma}_n^\delta\right\}\right] \\ &= \frac{1}{\tilde{\sigma}_n^2} \sum_{i=1}^{r_n} E\left[\frac{|x_{in}|^{2+\delta}}{|x_{in}|^\delta} \mathbb{I}\left\{|x_{in}|^\delta \geq \varepsilon^{\frac{\delta}{2}} \tilde{\sigma}_n^\delta\right\}\right] \\ &\leq \frac{1}{\tilde{\sigma}_n^2} \times \frac{1}{\varepsilon^{\delta/2} \tilde{\sigma}_n^\delta} \sum_{i=1}^{r_n} E\left[|x_{in}|^{2+\delta} \mathbb{I}\left\{|x_{in}|^\delta \geq \varepsilon^{\frac{\delta}{2}} \tilde{\sigma}_n^\delta\right\}\right] \\ &\leq \frac{1}{\varepsilon^{\delta/2}} \times \frac{1}{\tilde{\sigma}_n^{2+\delta}} \sum_{i=1}^{r_n} E\left[|x_{in}|^{2+\delta}\right] \rightarrow 0 \end{aligned}$$

as $n \rightarrow \infty$, where the limit follows from Lyapunov condition.

1.3.1 Sufficient Condition for Lyapunov Condition

- Condition1: $\sup_{i \leq r_n} E[|x_{in}|^{2+\delta}] \leq B < \infty$ for all sufficiently large n .

Let $\bar{\sigma}_n^2 = \tilde{\sigma}_n^2 / r_n$ be the *average variance*.

- Condition2: $\liminf_{n \rightarrow \infty} \bar{\sigma}_n^2 > b > 0$.

Under Condition 1 and Condition 2 we have

$$\begin{aligned} \frac{1}{\bar{\sigma}_n^{2+\delta}} \sum_{i=1}^{r_n} E[|x_{in}|^{2+\delta}] &\leq \frac{1}{(\sqrt{r_n b})^{2+\delta}} \times r_n \max_{i \leq r_n} E[|x_{in}|^{2+\delta}] \\ &\leq \frac{r_n B}{(\sqrt{r_n b})^{2+\delta}} = \text{const} \times r_n^{-\delta/2} \rightarrow 0 \end{aligned}$$

since $r_n \rightarrow \infty$ as $n \rightarrow \infty$.

If we further assume $\bar{\sigma}_n^2 \rightarrow \sigma_*^2$ as $n \rightarrow \infty$, then under Condition 1 we have $\frac{S_n}{\sqrt{n}} \xrightarrow{d} N(0, \sigma_*^2)$.

1.3.2 Uniform CLT

If $E[|z|^{2+\delta}] \leq B < \infty$ and $\text{var}(z) \geq b > 0$ for all $f \in \mathcal{F}$, then

$$\sup_{f \in \mathcal{F}} \left| P_f \left(\frac{\sqrt{n}(\bar{z}_n - E(z))}{\sqrt{\text{var}(z)}} \leq a \right) - \Phi(a) \right| \rightarrow 0.$$

This is a uniform CLT over a class of distributions in \mathcal{F} , instead of a single distribution f . Here P_f means that the probability is computed under a specific distribution f .

In a direct proof, the approximation error is controlled by B and b . The textbook uses a counter-positive argument: If the statement is false, then there is a sequence $f_1, f_2, \dots \in \mathcal{F}$ that violates the convergence. That contradicts with Lyapunov CLT.

1.4 Uniform Integrability

Definition 1.3. The sequence of random variables z_n is *uniformly integrable* as $n \rightarrow \infty$ if

$$\lim_{M \rightarrow \infty} \limsup_{n \rightarrow \infty} E[|z_n| \mathbb{I}\{|z_n| > M\}] = 0.$$

Uniform integrability requires that the triangular array z_n is asymptotic tight.

Example 1.1. Consider a counterexample

$$z_n = \begin{cases} -\sqrt{n} & \text{with probability } 1/n \\ 0 & \text{with probability } 1 - 2/n \\ \sqrt{n} & \text{with probability } 1/n. \end{cases}$$

Each z_n is square integrable in that $\text{var}[z_n] = 2n \times (1/n) = 2$. However, this z_n is NOT uniformly square integrable because for any finite M we have

$$\limsup_{n \rightarrow \infty} E[z_n^2 \mathbb{I}\{z_n^2 > M\}] = 2$$

and thus

$$\lim_{M \rightarrow \infty} \limsup_{n \rightarrow \infty} E[z_n^2 \mathbb{I}\{z_n^2 > M\}] = 2 \neq 0.$$

Definition 1.4. A triangular array of random variables is *uniformly integrable* as $n \rightarrow \infty$ if

$$\lim_{M \rightarrow \infty} \limsup_{n \rightarrow \infty} \max_{i \leq r_n} E[x_{in}^2 \mathbb{I}\{x_{in}^2 > M\}] = 0. \quad (1.1)$$

Proposition 1.1. *If Condition 2 and the triangular array $\{(x_{in})_{i=1}^{r_n}\}_{n \in \mathbb{N}}$ is uniform square integrable, then Lindeberg condition holds.*

Proof. For any $\varepsilon > 0$, we have

$$\begin{aligned} \frac{1}{\tilde{\sigma}_n^2} \sum_{i=1}^{r_n} E [x_{in}^2 \mathbb{I} \{x_{in}^2 \geq \varepsilon \tilde{\sigma}_n^2\}] &\leq \frac{1}{r_n b} \sum_{i=1}^{r_n} E [x_{in}^2 \mathbb{I} \{x_{in}^2 \geq \varepsilon r_n b\}] \\ &\leq \frac{1}{r_n b} \times r_n \max_{i \leq r_n} E [x_{in}^2 \mathbb{I} \{x_{in}^2 \geq r_n \varepsilon b\}] \\ &\leq \text{const} \times \max_{i \leq r_n} E [x_{in}^2 \mathbb{I} \{x_{in}^2 \geq r_n \varepsilon b\}] \rightarrow 0 \end{aligned}$$

by the definition of uniform integrability since $r_n \varepsilon b \rightarrow \infty$ as $n \rightarrow \infty$. □

1.4.1 Uniform Stochastic Bound

Theorem 1.2. *If (1.1) holds, then*

$$r_n^{-1/r} \max_{i \leq r_n} |x_{in}| \xrightarrow{p} 0.$$

Remark 1.1. As a special case, if we set $r_n = n$, then $\max_{i \leq n} |x_{in}| = o_p(n^{1/r})$ if x_{in} is r th uniformly integrable.

Proof. We start with the definition of convergence in probability:

$$\begin{aligned} P \left(r_n^{-1/r} \max_{i \leq r_n} |x_{in}| > \varepsilon \right) &= P \left(\max_{i \leq r_n} |x_{in}|^r > r_n \varepsilon^r \right) \\ &\leq \sum_{i \leq r_n} P (|x_{in}|^r > r_n \varepsilon^r) \\ &= \sum_{i \leq r_n} E [\mathbb{I} \{|x_{in}|^r > r_n \varepsilon^r\}] \\ &\leq r_n \max_{i \leq r_n} E [\mathbb{I} \{|x_{in}|^r > r_n \varepsilon^r\}] \\ &\leq r_n \times \frac{1}{r_n \varepsilon^r} \max_{i \leq r_n} E [|x_{in}|^r \mathbb{I} \{|x_{in}|^r > r_n \varepsilon^r\}] \\ &= \text{const} \times \max_{i \leq r_n} E [|x_{in}|^r \mathbb{I} \{|x_{in}|^r \geq r_n \varepsilon^r\}] \\ &\rightarrow 0 \end{aligned}$$

under (1.1), since $r_n \varepsilon^r \rightarrow \infty$ as $n \rightarrow \infty$. □