

# Chapter 1

## Time Series - Exercises

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**Exercise 14.1** For a scalar time series  $Y_t$  define the sample autocovariance and autocorrelation

$$\hat{\gamma}(k) = n^{-1} \sum_{t=k+1}^n (Y_t - \bar{Y})(Y_{t-k} - \bar{Y})$$
$$\hat{\rho}(k) = \frac{\hat{\gamma}(k)}{\hat{\gamma}(0)} = \frac{\sum_{t=k+1}^n (Y_t - \bar{Y})(Y_{t-k} - \bar{Y})}{\sum_{t=1}^n (Y_t - \bar{Y})^2}.$$

**Assume the series is strictly stationary, ergodic, strictly stationary, and  $E[Y_t^2] < \infty$ . Show that  $\hat{\gamma}(k) \xrightarrow{p} \gamma(k)$  and  $\hat{\rho}(k) \xrightarrow{p} \rho(k)$  as  $n \rightarrow \infty$ . (Use the Ergodic Theorem.)**

**Solution. 14.1**

(1) Since  $Y_t$  is strictly stationary, its transformation  $(Y_t - \bar{Y})(Y_{t-k} - \bar{Y})$  is also strictly stationary.

Use the Ergodic Theorem,

$$(n-k-1)^{-1} \sum_{t=k+1}^n (Y_t - \bar{Y})(Y_{t-k} - \bar{Y}) \xrightarrow{p} E[(Y_t - \bar{Y})(Y_{t-k} - \bar{Y})] = \gamma(k)$$

as  $n \rightarrow \infty$ , and therefore

$$\hat{\gamma}(k) = n^{-1} \sum_{t=k+1}^n (Y_t - \bar{Y})(Y_{t-k} - \bar{Y}) = \frac{n-k-1}{n} \times \frac{1}{n-k-1} \sum_{t=k+1}^n (Y_t - \bar{Y})(Y_{t-k} - \bar{Y}) \xrightarrow{p} \gamma(k).$$

(2)

$$\hat{\rho}(k) = \frac{\hat{\gamma}(k)}{\hat{\gamma}(0)} \xrightarrow{p} \frac{\gamma(k)}{\gamma(0)} = \rho(k)$$

**Exercise 14.2** Show that if  $(e_t, \mathcal{F}_t)$  is a MDS and  $X_t$  is  $\mathcal{F}_t$ -measurable then  $u_t = X_{t-1}e_t$  is a MDS.

**Solution. 14.2**

Since  $X_t$  is  $\mathcal{F}_t$ -measurable, we have  $u_t$  adapted to  $\mathcal{F}_t$  in that

$$E[u_t | \mathcal{F}_t] = E[X_{t-1}e_t | \mathcal{F}_t] = X_{t-1}e_t.$$

as  $x_{t-1}$  is  $\mathcal{F}_t$  measurable. Moreover,

$$E[u_t \mid \mathcal{F}_{t-1}] = E[X_{t-1}e_t \mid \mathcal{F}_{t-1}] = X_{t-1}E[e_t \mid \mathcal{F}_{t-1}] = 0$$

We thus have verified  $u_t = X_{t-1}e_t$  being a MDS.

**Exercise 14.3** Let  $\sigma_t^2 = E[e_t^2 \mid \mathcal{F}_{t-1}]$ . Show that  $u_t = e_t^2 - \sigma_t^2$  is a MDS.

**Solution. 14.3**

- (1)  $\sigma_t^2 = E[e_t^2 \mid \mathcal{F}_{t-1}]$  is adapted to  $\mathcal{F}_t$ .
- (2)  $E[\sigma_t^2] = E[E[e_t^2 \mid \mathcal{F}_{t-1}]] < \infty$ .
- (3)

$$\begin{aligned} E[u_t \mid \mathcal{F}_{t-1}] &= E[e_t^2 - \sigma_t^2 \mid \mathcal{F}_{t-1}] \\ &= E[e_t^2 - E[e_t^2 \mid \mathcal{F}_{t-1}] \mid \mathcal{F}_{t-1}] \\ &= E[e_t^2 \mid \mathcal{F}_{t-1}] - E[e_t^2 \mid \mathcal{F}_{t-1}] = 0 \end{aligned}$$

Therefore,  $u_t = e_t^2 - \sigma_t^2$  is a MDS.

**Exercise 14.8** Suppose  $Y_t = Y_{t-1} + e_t$  with  $e_t$  i.i.d. (0,1) and  $Y_0 = 0$ . Find  $\text{var}[Y_t]$ . Is  $Y_t$  stationary?

**Solution. 14.8**

- (1)

$$\begin{aligned} Y_t &= Y_{t-1} + e_t \\ &= Y_{t-2} + e_{t-1} + e_t \\ &= Y_{t-3} + e_{t-2} + e_{t-1} + e_t \\ &\vdots \\ &= Y_0 + e_1 + \dots + e_{t-2} + e_{t-1} + e_t \\ &= \sum_{i=1}^t e_i \end{aligned}$$

$$\text{var}[Y_t] = \text{var}\left[\sum_{i=1}^t e_i\right] = \sum_{i=1}^t \text{var}[e_i] = \sum_{i=1}^t 1 = t$$

- (2)  $Y_t$  is not stationary because  $\text{var}[Y_t]$  varies with  $t$ .