

## Advanced Econometrics II:

June 8, 2024

### Assignment 2

(This assignment is due on June 14, 2024 at noon. Please submit answers to your TA on time. The total marks are indicated in each question.)

1. In a classical linear regression model  $y_i = x_i' \beta + u_i$  with *iid* data (assuming both  $x_i$  and  $u_i$  are random and  $x_i$  is independent of  $u_i$ ;  $u_i \sim iid(0, \sigma^2)$ , and the 4-th moment of  $u_i$  is finite), we estimate the variance of  $u_i$  by  $\hat{\sigma}^2 = \frac{1}{n-k} \sum_{i=1}^n \hat{u}_i^2$ . Please derive that:

- (a) [5 points]  $\hat{\beta} - \beta = O_p(n^{-0.5})$
- (b) [10 points]  $\hat{\sigma}^2 - \sigma^2 = O_p(n^{-0.5})$

2. [15 points] Suppose  $\{X_t\}_{t=1}^n$  is generated by a classical *MA*(1) process:

$$X_t = Z_t + \theta_1 Z_{t-1} \quad Z_t \sim WN(0, \sigma^2), t = 1, 2, \dots, n$$

( $\theta_1$  and  $\sigma^2$  are treated as known). If a standard *AR*(1) process,  $X_t = \phi X_{t-1} + Y_t$  is mistakenly fitted to  $\{X_t\}$  ( $\phi$  is treated as unknown). Please derive the autocorrelation function (ACF) of  $\{Y_t\}$ . The ACF that you derived must be a function of  $\theta_1$  and  $\sigma^2$ .

3. [10 points] For the classical stationary *AR*(1) model:  $y_t = \rho y_{t-1} + u_t$ , where  $u_t \sim iid(0, \sigma^2)$  and  $E y_t^4$  is finite, we have

$$\sqrt{n}(\hat{\rho} - \rho) = \frac{\frac{1}{\sqrt{n}} \sum_{t=1}^n y_{t-1} y_t}{\frac{1}{n} \sum_{t=1}^n y_{t-1}^2}$$

Please derive that the denominator has  $\frac{1}{n} \sum_{t=1}^n y_{t-1}^2 \xrightarrow{P} E(y_{t-1}^2) = \sigma_y^2$

4. For the *VAR*(2) model, where the estimated results of each equation are as follows:

$$a_t = 0.5a_{t-1} + 0.7b_{t-1} + 0.3a_{t-2} + e_{1t}$$

$$b_t = -0.2a_{t-1} + 0.4b_{t-1} + e_{2t}$$

- (a) [5 points] Defining  $y_t$  as a 2-element vector containing  $a_t$  and  $b_t$ , write the above estimated single-equation form into the *VAR*(2) model form that we have learned in the class.
- (b) [5 points] What is the effect on  $a_{t+2}$  of a one-unit shock in  $b_t$  ( $b_t$  increases by one-unit)

5. [15 points] Consider a linear regression model:

$$y_t = x_{1t} \beta_1 + x_{2t} \beta_2 + u_t \equiv X_t' \beta + u_t,$$

where  $x_{1t}$  is *iid* with  $E x_{1t} = \mu_1 \neq 0$ ,  $x_{2t} = x_{2,t-1} + v_t$ ,  $v_t$  is *iid*(0,  $\sigma_v^2$ ),  $u_t$  is *iid*(0,  $\sigma_u^2$ ) with  $v_t$  is independent of  $u_t$ , and  $v_t$  is independent of  $x_{1t}$ .

Please derive the joint distribution of  $\begin{pmatrix} \sqrt{n}(\hat{\beta}_1 - \beta_1) \\ n(\hat{\beta}_2 - \beta_2) \end{pmatrix}$ , where  $\hat{\beta} = \begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix} = (X'X)^{-1}X'Y$  is the OLS estimator of  $\beta = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix}$ .

Hint:  $\begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix} = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} + (X'X)^{-1}X'u$ . Define  $D_n = \begin{pmatrix} \sqrt{n} & 0 \\ 0 & n \end{pmatrix}$ . Consider  $\begin{pmatrix} \sqrt{n}(\hat{\beta}_1 - \beta_1) \\ n(\hat{\beta}_2 - \beta_2) \end{pmatrix} \equiv D_n \left( \begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix} - \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} \right) = D_n(X'X)^{-1}X'u = D_n(X'X)^{-1}D_n D_n^{-1}X'u = [D_n^{-1}(X'X)D_n^{-1}]^{-1}D_n^{-1}X'u$ . The limiting distribution will involve integrations of Brownian motions.

6. [5 points] For the panel data, please describe how to construct the test for testing the existence of time-invariant unobservable (individual) heterogeneity?

7. [5 points] Similar as the two-way fixed effect model, we can modify the one-way random effect model by adding time specific disturbance to get the two-way random effect model in the form of:

$$y_{it} = \alpha + x'_{it}\beta + \epsilon_{it} + u_i + v_t$$

where

$$\begin{aligned} E[\epsilon_{it}|x] &= E[u_i|x] = E[v_t|x] = 0 \\ E[\epsilon_{it}u_j|x] &= E[\epsilon_{it}v_s|x] = E[u_i v_t|x] = 0 \text{ for all } i, t, j, s \\ \text{Var}[\epsilon_{it}|x] &= \sigma_\epsilon^2, \text{Cov}[\epsilon_{it}, \epsilon_{js}|x] = 0 \text{ for all } i, t, j, s \\ \text{Var}[u_i|x] &= \sigma_u^2, \text{Cov}[u_i, u_j|x] = 0 \text{ for all } i, j \\ \text{Var}[v_t|x] &= \sigma_v^2, \text{Cov}[v_t, v_s|x] = 0 \text{ for all } t, s \end{aligned}$$

Please write out the full disturbance covariance matrix for a data set with  $n=2$  and  $T=2$

8. [10 points] For the classical random-effect model that we have learned in the class:

$$y_{it} = x'_{it}\beta + \alpha + \epsilon_{it} + u_i$$

where,  $x'_{it}$  is strictly exogenous;  $\text{Var}(\epsilon_{it}) = \sigma_\epsilon^2$  and  $\sigma_\epsilon^2$  is unknown. We can estimate the  $\sigma_\epsilon^2$  by:

$$\hat{\sigma}_\epsilon^2 = \frac{1}{n(T-1)-k} \sum_{i=1}^n \sum_{t=1}^T \ddot{e}_{it}^2 \quad (*)$$

where,  $\ddot{e}_{it} = e_{it} - \bar{e}_i$ .

Please prove that  $\hat{\sigma}_\epsilon^2$  is a consistent estimator of  $\sigma_\epsilon^2$ .

[Hints: Please show that  $\ddot{e}_{it} = \ddot{e}_{it} - \ddot{x}'_{it}(b - \beta)$  first and plug it into equation (\*).

Secondly, calculate  $\hat{\sigma}_\epsilon^2$ .]