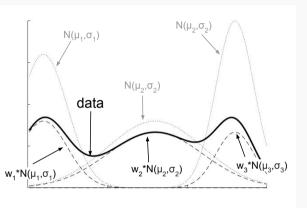
K-means, Gaussian Mixture Models, UBM-GMM

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Gaussian mixture models (GMM)



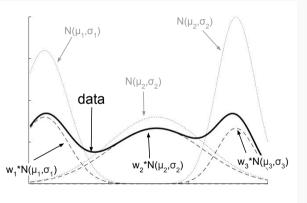
► A GMM is the weighted sum of individual Gaussian distributions

$$P(\vec{x}|\lambda_s) = \sum_{k=1}^{M} w_k \mathcal{N}(\vec{x}|\vec{\mu}_k, \Sigma_k)$$

$$\lambda_s = \{w_k, \vec{\mu}_k, \Sigma_k\}_{k=1}^M$$
$$\sum_{k=1}^K w_k = 1$$

$$\mathcal{N}(\vec{x}|\vec{\mu}_k, \Sigma_k) = (2\pi)^{-d/2} |\Sigma_k|^{-1/2} \exp\{-\frac{1}{2}(\vec{x} - \vec{\mu}_k)^T \Sigma_k^{-1} (\vec{x} - \vec{\mu}_k)\}$$

Gaussian mixture models (GMM)



► The problem of fitting a GMM is a incomplete data problem. Hence, the mixtures needs to be estimated iteratively using Expectation Maximization (E-M) algorithm.

Parameter estimation of GMM using E-M algorithm

- \triangleright Estimate parameters $(\bar{\mu}_k, \Sigma_k)$.
- Define one more quantity 'responsibility'.

$$P(z_k = 1|x) = \frac{P(z_k = 1)P(\bar{x}|z_k = 1)}{\sum_{j=1}^{K} P(\bar{x}|z_j)P(z_j = 1)}$$
(1)

 $P(z_k = 1|x) = \gamma_k$, which is the responsibility of k_{th} mixture in describing a point x.

$$P(z_{k} = 1 | x) = \frac{P(z_{k} = 1)P(\bar{x}|z_{k} = 1)}{\sum_{j=1}^{K} P(\bar{x}|z_{j})P(z_{j} = 1)}$$
$$= \frac{\pi_{k}N(\bar{x}; \bar{\mu}_{k}, \Sigma_{k})}{\sum_{i=1}^{K} \pi_{i}N(\bar{x}; \bar{\mu}_{i}, \Sigma_{i})}$$

$$N(\bar{x}, \bar{\mu}_j, \Sigma_j) = \frac{-1}{(\sqrt[2]{2\pi})^d |\Sigma|} e^{(\frac{-1}{2}(\bar{x} - \bar{\mu}_j)^t \Sigma^{-1}(\bar{x} - \bar{\mu}_j))}$$

A single point will be described completely by all the mixtures together.

$$\sum_{k=1}^{K} \gamma_k = 1$$

$$\gamma_{nk} = P(z_k = 1|\bar{x}_n)$$
$$\sum_{n=1}^{N} \gamma_{nk} = N_k$$

 N_k is the effective number of points that belongs to k^{th} cluster

$$\begin{array}{lll} \gamma(z1=1|\bar{x}2)=.125, & \gamma(z1=1|\bar{x}1)=.75, & \gamma(z1=1|\bar{x}3)=.05\\ \gamma(z2=1|\bar{x}2)=.75, & \gamma(z1=1|\bar{x}1)=.20, & \gamma(z2=1|\bar{x}3)=.20\\ \gamma(z3=1|\bar{x}2)=.125, & \gamma(z1=1|\bar{x}1)=.05, & \gamma(z3=1|\bar{x}3)=.75 \end{array}$$

$$\mu_k = \frac{1}{N_k} \sum_{n=1}^{N_k} \gamma_{nk} \bar{x}_n$$

Estimation of parameters

$$\theta_{ML} = rg \max_{\theta} \log p(D|\theta)$$

$$\bar{\theta} = [\bar{\mu}_1, \bar{\mu}_2, ..., \bar{\mu}_k, \Sigma_1, \Sigma_2, ... \Sigma_k, \pi_1, \pi_2, ..., \pi_k]$$

$$\ln P(D|\theta) = \ln \prod_{n=1}^{N} P(\bar{x}_n|\theta) = I(\theta)$$

$$InP(D|\theta) = \sum_{n=1}^{N} \ln \sum_{j=1}^{K} \pi_{j} N(\bar{x}_{n}; \bar{\mu}_{j}, \Sigma_{j})$$
$$\frac{\partial I(\bar{\theta})}{\partial \bar{\theta}} = 0$$
$$\frac{\partial I(\bar{\theta})}{\partial \bar{\mu}_{k}} = 0; \frac{\partial I(\bar{\theta})}{\partial \bar{\Sigma}_{k}} = 0$$

 π_k requires a constraint $\sum_{k=1}^K \pi_k = 1$

$$\frac{\partial I(\bar{\theta})}{\partial \bar{\mu}_k} = \sum_{n=1}^N \frac{\partial}{\partial \bar{\mu}_k} \ln \sum_{j=1}^K \pi_j N(\bar{x}_n; \bar{\mu}_j, \Sigma_j)$$

$$= \sum_{n=1}^N \frac{\pi_k}{\sum_{j=1}^K \pi_j N(\bar{x}_n; \bar{\mu}_j, \Sigma_j)} \frac{\partial}{\partial \bar{\mu}_k} N(\bar{x}_n; \bar{\mu}_k, \Sigma_k)$$

$$\frac{\partial}{\partial \bar{\mu}_k} N(\bar{x}_n; \bar{\mu}_k, \Sigma_k) = \frac{-1}{(\sqrt[2]{2\pi})^d |\Sigma|^{\frac{1}{2}}} e^{(\frac{-1}{2}(\bar{x}_n - \bar{\mu}_k)^t \Sigma_k^{-1}(\bar{x}_n - \bar{\mu}_k))}$$

$$\frac{\partial}{\partial \bar{x}} \bar{x}^t M \bar{x} = 2M \bar{x}$$

$$\frac{\partial I(\bar{\theta})}{\partial \bar{\mu}_{k}} = -N(\bar{x}_{n}; \bar{\mu}_{k}, \Sigma_{k}) \Sigma_{k}^{-1} (\bar{x}_{n} - \bar{\mu}_{k})$$

$$= \Sigma_{n=1}^{N} \frac{\pi_{k} N(\bar{x}; \bar{\mu}_{k}, \Sigma_{k})}{\sum_{j=1}^{K} \pi_{j} N(\bar{x}; \bar{\mu}_{j}, \Sigma_{j})}$$

$$\frac{\partial}{\partial \bar{\mu}_{k}} = 0 \implies \sum_{n=1}^{N} \gamma_{nk} \Sigma_{k}^{-1} (\bar{x}_{n} - \bar{\mu}_{k}) = 0$$

$$\sum_{n=1}^{N} \gamma_{nk} \bar{x}_{n} = \sum_{n=1}^{N} \gamma_{nk} \bar{\mu}_{k}$$

$$\bar{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma_{nk} \bar{x}_n$$

To find Σ_k , take derivative with respect to Σ_k and equate it to 0.

$$\bar{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma_{nk} (\bar{x}_n - \hat{\mu}_k) (\bar{x}_n - \hat{\mu}_k)^t$$

Identities

Estimation of π_k such that $\sum_{i=1}^K \pi_i = 1$

This is a constraint optimization problem.

$$L(\pi_k, \lambda) = \sum_{n=1}^K \ln \sum_{j=1}^K \pi_j N(\bar{x}_n; \bar{\mu}_j, \Sigma_j, \pi_j) - \lambda (\sum_{j=1}^K \pi_j - 1)$$

$$\frac{\partial L(\pi_k, \lambda)}{\partial \pi_k} = \sum_{n=1}^N \frac{\pi_k N(\bar{x}; \bar{\mu}_k, \Sigma_k)}{\sum_{j=1}^K \pi_j N(\bar{x}; \bar{\mu}_j, \Sigma_j)}$$

$$\lambda \pi_k = \sum_{n=1}^N \gamma_{nk}$$

$$\lambda \sum_{k=1}^K \pi_k = \sum_{n=1}^N \sum_{k=1}^K \gamma_{nk}$$

$$\pi_k = \frac{N_k}{N}$$

GMM: EM algorithm (Recap)

1) Random initialization of $\vec{\mu}_k, \Sigma_k$ and w_k

2) Expectation-Step

Align vectors to model

$$\gamma_{nk} = \frac{w_k \mathcal{N}(\vec{x}_n | \vec{\mu}_k, \Sigma_k)}{\sum\limits_{j=1}^{M} w_j \mathcal{N}(\vec{x}_n | \vec{\mu}_j, \Sigma_j)}$$

$$N_k = \sum\limits_{n=1}^{N} \gamma_{nk}$$

 $\rightarrow \gamma_{nk}$ is the responsibility of k-th component towards *n*-th feature vector

3) Maximization-Step

Update model parameters by maximum likelihood estimation (MLE)

$$\hat{\vec{\mu}}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma_{nk} \vec{x}_n$$

$$\hat{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma_{nk} (\vec{x}_n - \hat{\vec{\mu}}_k) (\vec{x}_n - \hat{\vec{\mu}}_k)^T$$

$$\hat{w}_k = \frac{N_k}{N}$$

4) Repeat step 2 and 3 until convergence.

GMM-proof of convergence

- Let $\bar{\Theta}^{\mathrm{old}}$ be parameters used at the start of any EM-iteration and $\bar{\Theta}$ be the updated parameters.
- Let \mathcal{D} be the given set of data points and \bar{z} be the latent variable defined by $\bar{\Theta}^{\mathrm{old}}$.
- ▶ We need to prove that $P\left(\mathcal{D} \mid \bar{\Theta}\right) \geq P\left(\mathcal{D} \mid \bar{\Theta}^{\mathrm{old}}\right)$ for every iteration.
- ▶ In every iteration we need to maximize $E\left[\log P\left(\mathcal{D}, \bar{z} \mid \bar{\Theta}\right) \mid \mathcal{D}, \bar{\Theta}^{\mathrm{old}}\right]$
- Lets define the auxiliary function as follows:

$$\begin{split} A\left(\bar{\Theta},\bar{\Theta}^{\mathrm{old}}\right) &= \sum_{z} P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}^{\mathrm{old}}\right) \ \log \ P\left(\mathcal{D},\bar{z} \mid \bar{\Theta}\right) \\ &= \sum_{z} P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}^{\mathrm{old}}\right) \ \log \ \left(P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}\right) P\left(\mathcal{D} \mid \bar{\Theta}\right)\right) \\ &= \sum_{z} P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}^{\mathrm{old}}\right) \ \log \ \left(P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}\right)\right) \\ &+ \sum_{z} P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}^{\mathrm{old}}\right) \ \log \ \left(P\left(\mathcal{D} \mid \bar{\Theta}\right)\right) \end{split}$$

GMM-proof of convergence (Contd..)

• Substituting $\bar{\Theta}^{\text{old}}$ for $\bar{\Theta}$ in the previous equation we get

$$A\left(\bar{\Theta}^{\mathrm{old}},\bar{\Theta}^{\mathrm{old}}\right) = \sum_{z} P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}^{\mathrm{old}}\right) \log \left(P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}^{\mathrm{old}}\right)\right) + \sum_{z} P\left(\bar{z} \mid \mathcal{D},\bar{\Theta}^{\mathrm{old}}\right) \log \left(P\left(\mathcal{D} \mid \bar{\Theta}^{\mathrm{old}}\right)\right)$$

Now $\log P(\mathcal{D} \mid \bar{\Theta}) - \log P(\mathcal{D} \mid \bar{\Theta}^{\text{old}})$ can be written as:

$$\log P(\mathcal{D} \mid \bar{\Theta}) - \log P(\mathcal{D} \mid \bar{\Theta}^{\text{old}}) = A(\bar{\Theta}, \bar{\Theta}^{\text{old}}) - A(\bar{\Theta}^{\text{old}}, \bar{\Theta}^{\text{old}}) + \sum_{z} P(\bar{z} \mid \mathcal{D}, \bar{\Theta}^{\text{old}}) \log \left(\frac{P(\bar{z} \mid \mathcal{D}, \bar{\Theta}^{\text{old}})}{P(\bar{z} \mid \mathcal{D}, \bar{\Theta})}\right)$$
(2)

GMM-proof of convergence (Contd..)

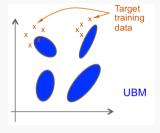
In Equation 2

- ▶ The last part of the equation is Kullback-Leibler divergence which is always positive or null.
- ▶ The update equation are derived such that the A increases (See https://courses.iitm.ac.in/mod/resource/view.php?id=789 for derivation based on auxiliary function)

Therefore, the likelihood always increases from iteration to iteration.

- ▶ UBM is a GMM trained on huge data pooled together from all the available classes.
- ▶ To overcome the huge data requirement for training GMM for individual classes.
- ▶ UBM is a GMM built with all class data using MLE algorithm.
- ▶ Maximum A-Posteriori (MAP) adaptation is used to train target models.
- ▶ UBM does not discriminate different speaker but acts as a reference for all class models.

UBM-GMM: MAP adaptation (contd...)



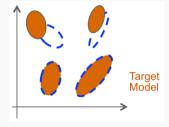
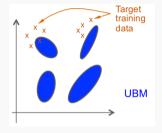


Figure 1: Adaptation using MAP

▶ Target model mean is updated using sufficient statistics, mixing co-efficient (α_k) and relevance factor r as

$$egin{aligned} lpha_k &= rac{ extsf{N}_k}{ extsf{N}_k + r} \ \hat{ec{u}}_k^{ extsf{new}} &= lpha_k \hat{ec{\mu}}_k + (1 - lpha_k) ec{\mu}_k^{ extsf{ubm}} \end{aligned}$$

UBM-GMM: MAP adaptation (contd...)



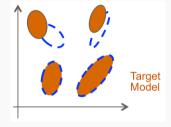


Figure 2: Adaptation using MAP

► The covariance can also be adapted using the same formule. But for most of the application we limit ourselves to adapting the mean.

Scoring in UBM-GMM

UBM acts as a reference for all class models. Hence the scores are calcluated with refence to the UBM.

Likelihood scoring (in the context of speech)

The average log-likelihood ratio score for a test utterance $\mathcal{X} = \{\bar{\mathbf{x}}_1, \bar{\mathbf{x}}_2, ..., \bar{\mathbf{x}}_T\}$ with claim is calculated as

$$LogLikelihood_{avg}(\mathcal{X}, \lambda_{claim}, \lambda_{UBM}) = \frac{1}{T} \sum_{t=1}^{I} \{log \ P(\bar{\mathbf{x}}_t | \lambda_{claim}) - log \ P(\bar{\mathbf{x}}_t | \lambda_{UBM})\}$$

Advantages of MAP in UBM-GMM

- Mixtures of UBM and target GMM have one-one correspondence
- Only few mixture components contribute to the a paticular class's feature vectors
- LogLikelihood_{avg} can be computed with maximum contributing C mixture components of the UBM.