

I PROBABILITY

1. Given

$$\text{Mean } (\mu) = 0$$

$$\text{Variance } (\sigma^2) = 0.1$$

For a Gaussian (Normal) distribution with mean μ and variance σ^2 , the density at point x is

$$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\left(\frac{(x-\mu)^2}{2\sigma^2}\right)}$$

Density at $x=0$

$$p(x=0) = \frac{1}{\sqrt{2\pi(0.1)}} e^{-\left(\frac{0}{2(0.1)}\right)}$$

$$p(x=0) = \frac{1}{\sqrt{0.2\pi}} \approx \frac{1}{0.79} \approx 1.26$$

The density of p at $x=0 \approx 1.26$
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2.

Let T = the (random) time the student finishes exam in the interval $[0,1]$.

$$\therefore P(T < x) = \frac{x}{2}; \quad 0 \leq x \leq 1 \quad \rightarrow \textcircled{1}$$

We want

$$P(T=1 \mid T > 0.75)$$

i.e., the probability that the student uses the full hour given that by time 0.75 they have not yet finished

We know

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

$$A: \{T > 1\}$$

$$B: \{T > 0.75\}$$

If the student uses the full hour (i.e., finishes exactly at 1 or later), that means they were still working at 0.75.

$$\therefore A \cap B = A$$

$$\Rightarrow P(A|B) = \frac{P(A)}{P(B)}$$

from ①

$$P(A) = P(T \geq 1)$$

$$\text{Since, } P(T < 1) = \frac{1}{2}$$

$$\text{we get } P(T \geq 1) = 1 - P(T < 1) = \frac{1}{2} \rightarrow \textcircled{2}$$

$$P(B) = P(T > 0.75) = 1 - P(T \leq 0.75)$$

$$= 1 - \frac{0.75}{2} = 1 - 0.375 \rightarrow \textcircled{3}$$

$$\Rightarrow P(A|B) = \frac{P(A)}{P(B)} = \frac{0.5}{1 - 0.375} \quad (\text{from } \textcircled{2}, \textcircled{3})$$

$$\boxed{P(A|B) = \frac{0.5}{0.625} = 0.8}$$

\Rightarrow The conditional probability that the student will use the full hour, given that they are still working after 0.75 hr is 0.8 (80%)

3.

Sick?	Sunny	rainy	cloudy	snow
yes	0.144	0.02	0.016	0.02
no	0.576	0.08	0.064	0.08

$$P(\text{sick}=\text{yes} \mid \text{weather}=\text{rainy}) = ?$$

From table,

$$P(\text{sick}=\text{yes}, \text{rainy}) = 0.02 ; P(\text{sick}=\text{yes no}, \text{rainy}) = 0.08$$

Marginal probability of Rainy weather

$$\begin{aligned}
 P(\text{rainy}) &= P(\text{sick}=\text{yes}, \text{rainy}) + P(\text{sick}=\text{no}, \text{rainy}) \\
 &= 0.02 + 0.08 \\
 &= 0.10
 \end{aligned}$$

\therefore By Bayes rule,

$$P(\text{sick}=\text{yes} \mid \text{rainy}) = \frac{P(\text{sick}=\text{yes}, \text{rainy})}{P(\text{rainy})} = \frac{0.02}{0.10} = 0.2$$

The probability that you are sick given the weather is rainy is 0.2 (or 20%)

II. CALCULUS AND LINEAR ALGEBRA

1. Given $f(z) = \frac{1}{1+e^{-z}}$

$$f(z) = (1+e^{-z})^{-1}$$

$$f'(z) = (-1) \cdot (1+e^{-z})^{-2} \cdot \frac{d}{dz}(1+e^{-z}) \left| \left(\frac{d}{dz}(1+e^{-z}) = -e^{-z} \right) \right.$$

$$f'(z) = \frac{(-1) \cdot e^{-z}}{(1+e^{-z})^2}$$

$$\Rightarrow f'(z) = \frac{e^{-z}}{(1+e^{-z})^2}$$

This can be written as

$$f'(z) = \frac{1}{(1+e^{-z})} \left[1 - \frac{1}{(1+e^{-z})} \right]$$

(or)

~~$$f'(z) = f(z) \left(1 - \frac{1}{f(z)} \right)$$~~

$$f'(z) = f(z)(1 - f(z))$$

2.

$$f(w) = \frac{1}{1 + e^{-w^T x}}$$

Let $z = w^T x$; $f(w) = \sigma(z)$; where $\sigma(z) = \frac{1}{1 + e^{-z}}$

we know

$$\sigma'(z) = \sigma(z)(1 - \sigma(z))$$

↳ from prev problem.

derivative of inner function.

$$z = w^T x = \sum_{j=1}^D w_j x_j \Rightarrow \frac{\partial z}{\partial w_i} = x_i$$

By chain rule

$$\frac{\partial f}{\partial w_i} = \frac{d\sigma}{dz} \frac{\partial z}{\partial w_i} = \sigma(z)(1 - \sigma(z)) x_i$$

Since $\sigma(z) = f(w)$

$$\boxed{\frac{\partial f}{\partial w_i} = f(w)(1 - f(w)) x_i}$$

3.

$$J(w) = \frac{1}{2} \left[\sum_{i=1}^m (w^T x^i - y^i)^2 \right] \neq \text{~~other expression~~}$$

let $\eta_i(w) = w^T x^i - y^i$

for $u \neq 0$, $\frac{d}{du} |u| = \text{sign}(u)$

at $u=0$ the subgradient is any value in $[-1, 1]$

$$\frac{\partial}{\partial w} | \eta_i(w) | = \text{sign}(\eta_i(w)) \frac{\partial \eta_i}{\partial w}$$

And, $\frac{\partial \eta_i}{\partial \omega} = \frac{\partial}{\partial \omega} (\omega^T x^{(i)} - y^{(i)}) = x^{(i)}$

$$\Rightarrow \boxed{\nabla_{\omega} J(\omega) = \frac{1}{2} \sum_{i=1}^m \text{sign}(\omega^T x^{(i)} - y^{(i)}) x^{(i)}}$$

Equivalently, in matrix form with $X = [x^{(1)}, x^{(2)}, \dots, x^{(m)}]^T$
 $\& S_i = \text{sign}(\eta_i)$

$$\Rightarrow \boxed{\nabla_{\omega} J(\omega) = \frac{1}{2} X^T S_i} \quad ; \quad \begin{array}{l} S_i \in \{-1, 1\} \text{ for } \eta_i \neq 0 \\ S_i \in [-1, 1] \text{ if } \eta_i = 0 \end{array}$$

4. $J(\omega) = \frac{1}{2} \left[\sum (\omega^T x^{(i)} - y^{(i)})^2 \right] + \lambda \|\omega\|_2^2$
 $\& \text{ let } \eta_i(\omega) = \omega^T x^{(i)} - y^{(i)}$

Differentiate the squared term
 for each i ,

$$\frac{\partial}{\partial \omega} \frac{1}{2} \eta_i(\omega)^2 = \frac{1}{2} \cdot 2 \eta_i(\omega) \frac{\partial \eta_i}{\partial \omega} = \eta_i(\omega) \cdot x^{(i)}$$

(Since $\frac{\partial \eta_i}{\partial \omega} = \frac{\partial (\omega^T x^{(i)} - y^{(i)})}{\partial \omega} = x^{(i)}$)

Summing over i gives

$$\frac{\partial}{\partial \omega} \left[\frac{1}{2} \sum_{i=1}^m \eta_i(\omega)^2 \right] = \sum_{i=1}^m \eta_i(\omega) x^{(i)}$$

Differentiate with L_2 regularizer.

$$\frac{\partial}{\partial \omega} \lambda \|\omega\|_2^2 = \frac{\partial}{\partial \omega} \lambda \omega^T \omega = 2\lambda \omega$$

Combining

$$\boxed{\nabla_{\omega} J(\omega) = \sum (\omega^T x^{(i)} - y^{(i)}) x^{(i)} + 2\lambda \omega}$$

Matrix form (with $X \in \mathbb{R}^{m \times d}$ whose i -th row $x^{(i)T}$, $y \in \mathbb{R}^m$):

$$\boxed{\nabla_{\omega} J(\omega) = X^T (X\omega - y) + 2\lambda \omega}$$

5.

$$J(\omega) = \sum_{i=1}^m \left[y^{(i)} \log \left(\frac{1}{1 + e^{-\omega^T x^{(i)}}} \right) + (1 - y^{(i)}) \log \left(1 - \frac{1}{1 + e^{-\omega^T x^{(i)}}} \right) \right]$$

$$\text{let } \sigma(z) = \frac{1}{1 + e^{-z}} \quad \& \quad z_i = \omega^T x^{(i)}$$

$$\Rightarrow J(\omega) = \sum_{i=1}^m \left[y^{(i)} \log \sigma(z_i) + (1 - y^{(i)}) \log (1 - \sigma(z_i)) \right]$$

for each i ,

$$g_i(\omega) = y^{(i)} \log \sigma(z_i) + (1 - y^{(i)}) \log (1 - \sigma(z_i))$$

$$\text{using chain rule } \sigma'(z) = \sigma(z)(1 - \sigma(z))$$

$$\nabla_{\omega} \log \sigma(z_i) = \frac{\sigma'(z_i)}{\sigma(z_i)} x^{(i)} = (1 - \sigma(z_i)) x^{(i)},$$

$$\nabla_{\omega} \log (1 - \sigma(z_i)) = \frac{-\sigma'(z_i)}{1 - \sigma(z_i)} x^{(i)} = -\sigma(z_i) x^{(i)}$$

Thus

$$\begin{aligned} \nabla_{\omega} g_i(\omega) &= y^{(i)} (1 - \sigma(z_i)) x^{(i)} + (1 - y^{(i)}) (-\sigma(z_i)) x^{(i)} \\ &= (y^{(i)} - \sigma(z_i)) x^{(i)} \end{aligned}$$

Sum over i

$$\nabla_{\mathbf{w}} J(\mathbf{w}) = \sum_{i=1}^m (y^{(i)} - \sigma(\mathbf{w}^T \mathbf{x}^{(i)})) \mathbf{x}^{(i)}$$

6.



$$f(\mathbf{w}) = \tanh[\mathbf{w}^T \mathbf{x}]$$

$$\text{let } z = \mathbf{w}^T \mathbf{x}$$

outer function

$$\frac{d}{dz} \tanh z = 1 - \tanh^2 z$$

inner function

$$\frac{\partial z}{\partial \mathbf{w}} = \mathbf{x}$$

chain rule

$$\nabla_{\mathbf{w}} f(\mathbf{w}) = \frac{d}{dz} \tanh z \cdot \frac{\partial z}{\partial \mathbf{w}}$$

$$= (1 - \tanh^2(\mathbf{w}^T \mathbf{x})) \mathbf{x}$$

$$\Rightarrow \boxed{\nabla_{\mathbf{w}} f(\mathbf{w}) = (1 - \tanh^2(\mathbf{w}^T \mathbf{x})) \mathbf{x}}$$

7. $A = \begin{bmatrix} 2 & 1 & -1 \\ -3 & -1 & 2 \\ -2 & 1 & 2 \end{bmatrix} \quad b = \begin{bmatrix} 8 \\ -11 \\ -3 \end{bmatrix}$

$$R_2 \leftarrow 2R_2 + 3R_1$$

$$R_3 \leftarrow R_3 + R_1$$

$$[A|b] = \left[\begin{array}{ccc|c} 2 & 1 & -1 & 8 \\ 0 & -2+3 & 4-3 & 2 \\ 0 & 2 & 1 & 5 \end{array} \right]$$

$$\left[\begin{array}{ccc|c} 2 & 1 & -1 & 8 \\ 0 & 1 & 1 & 2 \\ 0 & 2 & 1 & 5 \end{array} \right]$$

$$R_3 \leftarrow R_3 - 2R_2$$

$$\left[\begin{array}{ccc|c} 2 & 1 & -1 & 8 \\ 0 & 1 & 1 & 2 \\ 0 & 0 & -1 & 1 \end{array} \right]$$

from row 3; $-x_3 = 1$

$$\Rightarrow x_3 = -1$$

sub $x_3 = -1$ in row 2

$$x_2 + x_3 = 2 \Rightarrow x_2 - 1 = 2 \Rightarrow x_2 = 3$$

Sub x_2, x_3 in row 1

$$2x_1 + x_2 + x_3 = 8$$

$$2x_1 + 3 - (-1) = 8 \Rightarrow 2x_1 + 4 = 8$$

$$\boxed{x_1 = 2}$$

$$\Rightarrow \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} = \begin{bmatrix} 2 \\ 3 \\ -1 \end{bmatrix}$$