Stern's Replication-Final Version

June 27, 2024

1 Replication of 'Modelling the emissions-income relationship using long-run growth rates' by David I. Stern, Reyer Gerlagh, Paul J. Burke

The idea of this work is to replicate the regression output tables presented in the paper. The sequence of the tables and it's corresponding regression equations are given below for your convenience.

1.0.1 Data

The data file and the RATS (Regression Analysis of Time Series) code have been downloaded from Stern's official website (http://www.sterndavidi.com/datasite.html).

```
[12]: import numpy as np
import matplotlib.pyplot as plt
import pandas as pd
import statsmodels.api as sm
import statsmodels.formula.api as smf
from stargazer.stargazer import Stargazer
from IPython.core.display import HTML

df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='CO2 1971-2010')

# Display the first few rows of the DataFrame
print(df.head())
```

```
ISO3CODE
                                    GDPPC1971 LNGDPPC1971
                                                            DMLNGDPPC1971
                          CNAME
                         Angola
                                  2390.478101
                                                  7.779249
                                                                 -0.319958
0
       AGO
1
       ALB
                        Albania
                                  3728.588935
                                                  8.223785
                                                                  0.124578
2
       ARG
                      Argentina
                                  3047.466797
                                                  8.022066
                                                                 -0.077141
3
       ATG Antigua and Barbuda
                                  4391.901990
                                                  8.387518
                                                                  0.288311
                      Australia 17143.668520
4
       AUS
                                                  9.749384
                                                                  1.650177
    GRGDPPC DMMULGDP
                         MULGDP
                                      EPC1971 LNEPC1971 ... DMLNFFPC1971
0 0.006774 -0.002167 0.052697
                                   563.898623
                                                6.334874 ...
                                                                  4.464793
  0.019730 0.002458 0.162253
                                  1988.773445
                                                7.595273 ...
                                                                  3.001371
                                                8.202107 ...
  0.011180 -0.000862 0.089685
                                  3648.630633
                                                                  3.234411
```

```
3 0.028881 0.008327 0.242239 6391.381435 8.762706 ... -4.230234
4 0.017044 0.028126 0.166172 11762.885257 9.372705 ... 6.174614
```

	PWTPOP1971	AREA	1NPOPD1971	DMLNPOPD1971	MAF	REF	LAM	OECD90	\
0	6.047736	1246700.0	1.579184	-2.043424	1	0	0	0	
1	2.188650	27400.0	4.380497	0.757889	0	1	0	0	
2	24.376109	2736690.0	2.186854	-1.435754	0	0	1	0	
3	0.066554	440.0	5.018994	1.396386	0	0	1	0	
4	12.987847	7682300.0	0.525095	-3.097513	0	0	0	1	

ASIA

- 0
- 1 0
- 2 0
- 3 0
- 4 0

[5 rows x 33 columns]

[6]: pip install openpyxl

Collecting openpyxl

Downloading openpyx1-3.1.4-py2.py3-none-any.whl (251 kB)

Collecting et-xmlfile

Downloading et_xmlfile-1.1.0-py3-none-any.whl (4.7 kB)

Note: you may need to restart the kernel to use updated packages.

Installing collected packages: et-xmlfile, openpyxl

Successfully installed et-xmlfile-1.1.0 openpyxl-3.1.4

1.1 Table 2 : General model (CO2 emission)

The regression shown below is of the general model for CO2 emission as in regression equation 2 in the original paper.

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_1 \hat{G}_i G_{i,0} + \beta_2 E_{i,0} + \beta_3 G_{i,0} + \sum_{j=4}^k X_{j,i} + \epsilon_i$$

where hats indicate long-run growth rates, \$ E\$ is the log of emissions per-capita, and G is the log of GDP per capita. X_i is a vector of explanatory and control variables. The sample mean has been deducted from each continuous level variable in this vector.

The third term, $\hat{G}_I G_{i,0}$, is the interaction between the rate of economic growth and the initial level of log income per capita. The EKC turning point is calculated with the assumption that $\alpha_1 > 0$ and $\beta_1 < 0$ when $\frac{\partial \hat{E}}{\partial \hat{G}} = 0$ and $\tau = exp\left(-\frac{\alpha_1}{\beta_1} + \mu_G\right)$, where μ_G is the cross-country mean of initial GDP per capita variable prior to the estimation

The following variables are contained in $X_{j,i}$: Dummy variable for countries that are centrally planned or market economy, dummies for countries that have legal origins in the UK, France,

Germany or any Scandinavian countries, average summer and winter temperatures, de-meaned log of fossil fuel consumption in 1971 and de-meaned log of population density in 1971.

```
[5]: import numpy as np
     import matplotlib.pyplot as plt
     import pandas as pd
     import statsmodels.api as sm
     import statsmodels.formula.api as smf
     from stargazer.stargazer import Stargazer
     from IPython.core.display import HTML
     from statsmodels.stats.diagnostic import het breuschpagan, het white
     from sklearn.linear model import LinearRegression
     from scipy.stats import chi2
     df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='CO2 1971-2010')
     # Run OLS regression
     reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE_
     →+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + L
     →DMLNPOPD1971", data=df).fit(cov_type='HC3')
     print(reg1.summary())
     #Different Types of Robust Standard Errors
     #'HCO': The original White (1980) heteroskedasticity-consistent standard errors.
     #'HC1': A small sample correction to HCO.
     \#'HC2': Another small sample correction, which is more reliable when the sample \sqcup
      \rightarrow size is small.
     #'HC3': MacKinnon and White's (1985) heteroskedasticity-consistent standard
     →errors, which is more robust in small samples.
     #print(reg1.resid())
     \#Create a matrix of independent varibales which can be used for the residual \sqcup
     \rightarrow tests
     exog het = reg1.model.exog
     # Perform the Breusch-Pagan Test for Heteroskedasticity. Note exog_het is the_
     →matrix of explanatory (independent) variables
     bp_test = het_breuschpagan(reg1.resid,exog_het)
     print("Breusch-Pagan Test for Heteroskedasticity:") # HO: Variance of the ⊔
      →residuals is constant (homoskedastic)
```

```
print (f"LM Statistics: {bp_test[0]}") # The Lagrange Multiplier (LM) test⊔
→ follows a chi-sq distribution and the degree of freedom is the number of
→independent variables. The LM statistic is calculated based on the residuals⊔
→ from the regression model and the independent variables
print (f"LM Test p-value: {bp_test[1]}") # If p-value is less than 0.05 (5%)
→ level) then we reject the null and vice-versa
\#print\ (f"F-Statistics:\ \{bp\_test[2]\}")\ \#\ F-stat\ is\ the\ ratio\ of\ the\ mean
→squared error of the model and the mean squared error of the residual.
#print (f"F-Statistics p-value: {bp_test[3]}")
# Performing the White test for heteroskedasticity
white_test = het_white(reg1.resid, exog_het) # HO: Homoskedasticity
# Extracting the test statistic and p-value
white_stat = white_test[0]
white_pvalue = white_test[1]
print(f"White test statistic: {white_stat}")
print(f"White test p-value: {white_pvalue}")
#print("White Test for Heteroskedasticity:") # Another way of writing the above
\rightarrow command of white test and p value
#print(f"T Statistics: {white test[0]}")
#print(f"P-value: {white_test[1]}")
#print(f"F-statistics: {white_test[2]}")
#print(f"F-test p-value: {white_test[3]}")
# Interpretation
if white pvalue < 0.05: # Typically, we use a significance level of 0.05
   print("Reject the null hypothesis of homoskedasticity. Evidence of_{\sqcup}
⇔heteroskedasticity.")
else:
   print("Fail to reject the null hypothesis of homoskedasticity. No evidence⊔
r1s= reg1.resid ** 2 # This command enables python to store the squared of □
\rightarrow residuals in r1s
# Assuming df is your DataFrame and PWTPOP1971 is the column you want to invert
df['invpop'] = 1 / df['PWTPOP1971']
df['lpop1971'] = np.log (df['PWTPOP1971']) # To create a new varible with log
df['rls']=reg1.resid**2 # stored the residual square in the data file.
# To check the addition of the new variable invop: print(df.head())
```

```
# Trying to run an auxiliary regression following the codes of Stern:

#reg2 = smf.ols("r1s ~ invpop", data=df).fit() (Error: Number of rows mismatchusbetween data argument and r1s (136 versus 134))

#print(reg2.summary())

# Test for the EKC turning point.
alpha1 = reg1.params['GRGDPPC'] # Storing the estimates of GRGDPPC
beta1 = reg1.params['DMMULGDP']

turning_point= np.exp(- alpha1/beta1 + 8.099206859) # Deriving the EKC turninguspoint level: From the paper (Note \mu_G = 8.09 = ln(6340.99); where 6340.99_usis cross-country mean of GDPPC1971 (initial GDP per capita 1971))

print(f"EKC turning point:{turning_point}")

OLS Regression Results
```

=======================================		=========	=======	========	=======	=====
Dep. Variable: GREPC			R-squared:		0.727	
Model:	odel: OLS			quared:	0.700	
Method:	Method: Least Squares		F-statis	tic:		19.56
Date:	Thu,	27 Jun 2024	Prob (F-	statistic):	5.	22e-23
Time:		11:29:36	Log-Like	lihood:		390.24
No. Observations:		134	AIC:			-754.5
Df Residuals:		121	BIC:			-716.8
Df Model:		12				
Covariance Type:		HC3				
=		========	:======	========	========	======
	coef	std err	z	P> z	[0.025	
0.975]						
Intercept	-0.0025	0.003	-0.949	0.343	-0.008	
0.003						
GRGDPPC	0.8901	0.107	8.282	0.000	0.679	
1.101						
DMMULGDP	-0.1330	0.075	-1.780	0.075	-0.280	
0.013						
	0.0167	0.004	4.659	0.000	0.010	
0.024	0.0454		5 005			
DMLNEPC1971	-0.0154	0.002	-7.005	0.000	-0.020	
-0.011						
CPE	-0.0054	0.006	-0.917	0.359	-0.017	
0.006						
legorfr	0.0007	0.003	0.238	0.812	-0.005	
0.006						
legorge	0.0006	0.005	0.127	0.899	-0.009	

Omnibus: Prob(Omnibus): Skew: Kurtosis:		6.565 0.038 -0.364 3.845	Durbin-Wa Jarque-Be Prob(JB): Cond. No.	era (JB):		2.078 6.948 0.0310 896.
DMLNPOPD1971 0.002	-0.0003	0.001	-0.245	0.806	-0.003	
6.42e-05 DMLNFFPC1971 0.002	0.0011	0.000	2.889	0.004	0.000	
0.002 DMWINT	-0.0004	0.000	-1.674	0.094	-0.001	
0.007 DMSUMT	0.0008	0.000	2.223	0.026	0.0001	
0.010 legorsc	-0.0038	0.005	-0.707	0.479	-0.014	

[1] Standard Errors are heteroscedasticity robust (HC3)

Breusch-Pagan Test for Heteroskedasticity:

LM Statistics: 13.513843530901335 LM Test p-value: 0.33282345059096746 White test statistic: 99.90689760380853 White test p-value: 0.013479168535001177

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

EKC turning point: 2648714.923934659

1.2 Table 3 : Restricted model: CO2 emission $1971-2010 \sim \text{Ordas Criado et al. (2011) (Column 2)}$

By applying restrictions to the general model shown above, Ordas Criado et al. (2011) is estimated. Equation (3) depicts the regression equation:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_2 E_{i,0} + \beta_3 G_{i,0} + \sum_{j=4}^k \beta_j X_{j,i} + \epsilon_i$$

The restrictions on the general model is: $\beta_1 = 0$

The variables in the above regression equation have the same meaning as mentioned above. This regression has considered all the explanatory variables (X_i) , but the results are not specified in the table

[43]: # Table 3 CO2 Regression Results

import numpy as np
import matplotlib.pyplot as plt

```
import pandas as pd
import statsmodels.api as sm
import statsmodels.formula.api as smf
from stargazer.stargazer import Stargazer
from IPython.core.display import HTML
from statsmodels.stats.diagnostic import het_breuschpagan, het_white
from sklearn.linear_model import LinearRegression
from scipy.stats import chi2
df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='CO2 1971-2010')
# Ordas-Criado Model
reg2= smf.ols("GREPC ~ GRGDPPC + DMLNGDPPC1971 + DMLNEPC1971 + CPE + legorfr +
→legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).
→fit(cov_type='HCO')
print(reg2.summary())
# For conducting different statistical tests we have to follow the following
# 1. Create a matrix of exogenous variables
exog_het = reg2.model.exog
# 2. Run the white test: HO: Homoskedasticity
white_test = het_white(reg2.resid, exog_het)
white_pvalue = white_test[1]
# 3. Print the results
print(f"White test for heteroskedasticity:")
print(f"T-stats for white test: {white_test[0]}")
print(f"P-value: {white_test[1]}")
print(f" F statistics: {white_test[2]}")
print(f" p_value: {white_test[3]}")
if white_pvalue < 0.05:</pre>
   print("Reject the null hypothesis of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
else:
   print("Fail to reject the null hypothesis of homoskedasticity. No evidence⊔
→of heteroskedasticity")
# 3. Run the Breusch-Pagan test: HO: Homoskedasticity
```

```
# The het breuschpagan function returns a tuple with four elements:
# a) Lagrange Multiplier (LM) Statistic: The first element, accessible with \Box
\rightarrow bp_test[0], is the test statistic for the Breusch-Pagan test.
# b) p-value: The second element, accessible with bp test[1], is the p-value
\hookrightarrow associated with the LM statistic.
# c) f-value: The third element, accessible with bp test[2], is the f-value of
\rightarrow the test.
# d) f p-value: The fourth element, accessible with bp test[3], is the p-value
\rightarrow associated with the f-statistic.
bp_test = het_breuschpagan(reg2.resid, exog_het)
bp_pvalue= bp_test[1]
print(f"BP test:")
print(f"LM Statistic (Chi-square): {bp_test[0]}")
print(f"LM Test p-value: {bp_test[1]}")
print(f" F statistics: {bp_test[2]}")
print(f" p_value: {bp_test[3]}")
if bp pvalue < 0.05:
      print("Reject the null hypothesis of homoskedasticity. Evidence of,
⇔heteroskedasticity.")
else:
      print("Fail to reject the null hypothesis of homoskedasticity. Evidence⊔
→of no heteroskedasticity.")
# 4. Run the Harvey test (Harvey test doesn't exist in any statistical packagesu
→ in python. Thus it has to do it manually)
    # Step 1: Calculate the log of the squared residuals
df['log_sq_residuals'] = np.log(reg2.resid**2)
    # Step 2: Regress log_squared_residuals on the independent variables
harvey_model = smf.ols("log_sq_residuals ~ GRGDPPC + DMLNGDPPC1971 +__
→DMLNEPC1971 + CPE + legorfr + legorge + legorsc + DMSUMT + DMWINT + L
→DMLNFFPC1971 + DMLNPOPD1971", data=df).fit()
    # Step 3: Obtain the R-squared value from this regression
r_sq = harvey_model.rsquared
    # Step 4: Calculate the LM statistic
```

```
n = len(df) # this command helps in identifying the number of obs in the
→dataset. len() ~ determines the length of various objects.
lm_stat_harvey = n * r_sq # The LM statistics is calulated as N * r_squared
    # Step 5: Calculate the p-value
p_value_harvey = chi2.sf(lm_stat_harvey, df=11) # degree of freedom (df) = 11__
→ for eleven independent variables
    # Step 6: Print the results
print("Harvey-Godfrey test for heteroskedasticity:") # HO: Homoskedasticity
print(f" LM Statistics: {lm_stat_harvey}")
print(f"LM test p-value: {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of \Box
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of_{\sqcup}
→no hoteroskedasticity.")
# 5. Performing the Likelihood test (LR test) against the general model.
→ (follows a chi-sq distribution)
# The Likelihood Ratio (LR) test is used to compare the goodness-of-fit of two_{\sqcup}
→nested models. The HO: The restricted model provides a good fit.
reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE
→+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +
→DMLNPOPD1971", data=df).fit(cov_type='HCO')
    # Calculating the log-likelihood values of both the models
llf_restricted = reg2.llf
llf_unrestricted = reg1.llf
    # Calculating the LR statistics
lr_statistic = -2 * (llf_restricted - llf_unrestricted)
    # Degrees of freedom (df): number of restrictions (difference in number of \Box
→parameters). In this case it should be 1
df_diff = reg1.df_model - reg2.df_model # Note: reg1 is the unrestricted model_
\rightarrow and reg2 is a restricted model
    # Calculating the p_value
```

```
p_value = chi2.sf(lr_statistic, df_diff)

print("LR test against the general model:")
print(f"LR Statistic: {lr_statistic}")
print(f"p-value: {p_value}")

if p_value < 0.05:
    print("Reject the null hypothesis of restricted model provides a good fit:")

The unrestricted model fits significantly better.")
else:
    print("Fail to reject the null hypothesis of restricted model provides a_U

Good fit: No significant improvement in fit for the unrestricted model.")</pre>
```

	old regression results							
Dep. Variable: GREPC Model: OLS Method: Least Squares Date: Mon, 17 Jun 2024 Time: 21:09:35 No. Observations: 134 Df Residuals: 122 Df Model: 11 Covariance Type: HCO			Prob (F-s Log-Like AIC: BIC:	quared: tic: statistic): lihood:	0.718 0.692 24.79 3.85e-26 388.02 -752.0 -717.3			
0.975]	coef		z	P> z	[0.025			
0.001 GRGDPPC	-0.0039 0.9670	0.002		0.086	-0.008 0.804			
1.130 DMLNGDPPC1971 0.022	0.0156	0.003	4.976	0.000	0.009			
DMLNEPC1971 -0.012 CPE 0.005	-0.0157 -0.0038	0.002	-8.108 -0.813	0.000	-0.019 -0.013			
legorfr 0.006	0.0008	0.003	0.322	0.748	-0.004			
legorge 0.009	0.0015	0.004	0.376	0.707	-0.006			
legorsc 0.005 DMSUMT	0.0010	0.004	-0.779 3.213	0.436	0.000			
0.002								

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Kurtosis:		3.995	Cond. No			747.
Skew:		-0.392	Prob(JB)	:		0.0113
Prob(Omnibus):		0.020	Jarque-B	era (JB):		8.966
Omnibus:		7.830	Durbin-W	atson:		2.052
0.002	========	.=======	=======	========	=======	======
DMLNPOPD1971	-0.0006	0.001	-0.573	0.567	-0.003	
0.002						
DMLNFFPC1971	0.0014	0.000	4.283	0.000	0.001	
2.83e-05						
DMWINT	-0.0004	0.000	-1.820	0.069	-0.001	

[1] Standard Errors are heteroscedasticity robust (HCO)

White test for heteroskedasticity:

T-stats for white test: 92.3245807849109

P-value: 0.007504469440336944 F statistics: 2.5369039604491936 p value: 8.793027341162336e-05

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

BP test:

LM Statistic (Chi-square): 15.490049093711212

LM Test p-value: 0.16114301505547435 F statistics: 1.4496565478110688

p_value: 0.1595272495911336

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

Harvey-Godfrey test for heteroskedasticity:

LM Statistics: 11.957825790720378 LM test p-value: 0.3668216687373419

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

hoteroskedasticity.

LR test against the general model: LR Statistic: 4.432836971044708 p-value: 0.035253863973579394

Reject the null hypothesis of restricted model provides a good fit: The

unrestricted model fits significantly better.

1.3 Table 3 : Restricted model: CO2 emission $1971-2010 \sim \text{Green Solow Model}$ (Column 3)

By applying restrictions to the general model shown above, Green Solow model (GSM) is estimated. Equation (4) shows the regression equation:

$$\hat{E}_{i} = \alpha_{0} + \beta_{2} E_{i,0} + \sum_{j=4}^{k} \beta_{j} X_{j,i} + \epsilon_{i}$$

The restrictions on the general model are: $\alpha_1 = \beta_1 = \beta_3 = 0$

The variables in the above regression equation have the same meaning as mentioned above. This regression has considered all the explanatory variables (X_i) , but the results are not specified in the table.

```
[13]: import numpy as np
      import matplotlib.pyplot as plt
      import pandas as pd
      import statsmodels.api as sm
      import statsmodels.formula.api as smf
      from stargazer.stargazer import Stargazer
      from IPython.core.display import HTML
      from statsmodels.stats.diagnostic import het_breuschpagan, het_white
      from sklearn.linear_model import LinearRegression
      from scipy.stats import chi2
      df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='CO2 1971-2010')
      # Table 3 CO2 Regression Results
      # Green Solow model
      reg3= smf.ols("GREPC ~ DMLNEPC1971 + CPE + legorfr + legorge + legorsc + DMSUMT⊔
       →+ DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).fit(cov_type = 'HCO')
      print(reg3.summary())
      # White test for heteroskedasticity
       # Step 1: Create a matrix of exogenous variables
      exog_het = reg3.model.exog
       # Follow the above steps for the White test.
      white_test = het_white(reg3.resid, exog_het)
      print(f"White test for heteroskedasticity: {white_test[0]}")
      print(f"p-value: {white_test[1]}")
      print(f" F stat: {white_test[2]}")
      print(f"p-value : {white_test[3]}")
      if white_test[1] < 0.05:</pre>
          print("Reject the null hypothesis of homoskedasticity. Evidence of_{\sqcup}
       ⇔heteroskedasticity.")
      else:
```

```
print("Fail to reject the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.")
# Breusch-Pagan Test for heteroskedasticity
bp_test = het_breuschpagan(reg3.resid,exog_het)
print(f"Breusch-Pagan test - LM (chi-sq) stat: {bp_test[0]}")
print(f"p-value: {bp test[1]}")
print(f"F stat: {bp_test[2]}")
print(f"p-value of F stat: {bp_test[3]}")
if bp_test[1] < 0.05:</pre>
    print("Rejecting the null of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
    print("Fail to reject the null of homoskedasticity. Evidence of no⊔
⇔heteroskedasticity.")
# Harvey-Godfrey test. Note the python packages donot have the harvey-test_{\sqcup}
→ module, so have to do it manually.
# Step 1: Calculate the log of squared residuals
df['log_sq_resid'] = np.log(reg3.resid ** 2)
# Step 2: Regress the log of residuals on the independent variables
harvey_reg = smf.ols("log_sq_resid ~ DMLNEPC1971 + CPE + legorfr + legorge + ⊔
→legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).
→fit(cov_type='HCO')
# Step 3: Calculate the r-sq value from the regression
r_sq = harvey_reg.rsquared
# Step 4: Calculate the LM statistics:
n=len(df) # Explain the number of obs
lm_stat = n * r_sq # the harvey test statistic is calculated as N * R ~2
# Step 5: Calculate the p_value
p_value_harvey = chi2.sf(lm_stat,df=9)
print("Harvey test:")
print(f"LM (chi-sq) test statistics: {lm_stat}")
```

```
print(f"p_value: {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
   print("Reject the null hypothesis of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
else:
   print("Fail to reject the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.")
# LR statistics: It is used to find the goodness of fit of restricted model
\rightarrow against the unrestricted model.
# The unrestricted model:
reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE_
\hookrightarrow+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +_{\sqcup}
→DMLNPOPD1971", data=df).fit(cov_type='HCO')
# Calculate the log likelihood ratio of both the models
llf_unrestricted = reg1.llf
llf restricted = reg3.llf
# Caculating the LR Statistics
lr_stat = - 2 * (llf_restricted - llf_unrestricted)
→restricted and unrestricted model)
df_diff = reg1.df_model - reg3.df_model
# LR stat follows a chi-sq distribution
p_value = chi2.sf(lr_stat, df_diff)
print (f"LR test statistics against the general model: {lr_stat}")
print(f"p_value: {p_value}")
if p_value < 0.05:</pre>
   print("Reject the null hypothesis of restricted model provides a good fit: u
→The unrestricted model fits significantly better.")
else:
   print("Fail to reject the null hypothesis of restricted model provides a⊔
 →good fit: No significan timprovement in fit for the unrestricted model.")
```

OLS Regression Results

=======================================			
Dep. Variable:	GREPC	R-squared:	0.348
Model:	OLS	Adj. R-squared:	0.301
Method:	Least Squares	F-statistic:	6.711
Date:	Tue, 18 Jun 2024	<pre>Prob (F-statistic):</pre>	8.53e-08
Time:	21:59:33	Log-Likelihood:	331.95
No. Observations:	134	AIC:	-643.9
Df Residuals:	124	BIC:	-614.9
Df Model:	9		
Covariance Type:	HCO		

==========	=======		=======			
	coef	std err	z	P> z	[0.025	0.975]
Intercept	0.0142	0.004	3.823	0.000	0.007	0.021
DMLNEPC1971	-0.0098	0.002	-6.299	0.000	-0.013	-0.007
CPE	-0.0053	0.006	-0.887	0.375	-0.017	0.006
legorfr	-0.0018	0.004	-0.411	0.681	-0.010	0.007
legorge	-0.0032	0.007	-0.448	0.654	-0.017	0.011
legorsc	0.0021	0.009	0.242	0.809	-0.015	0.019
DMSUMT	0.0010	0.000	2.526	0.012	0.000	0.002
DMWINT	-0.0009	0.000	-3.333	0.001	-0.001	-0.000
DMLNFFPC1971	0.0016	0.001	3.197	0.001	0.001	0.003
DMLNPOPD1971	0.0047	0.001	3.875	0.000	0.002	0.007
=========	=======		=======	========		
Omnibus:		11.030	Durbin-	Watson:		2.023
<pre>Prob(Omnibus):</pre>		0.004	Jarque-	Bera (JB):		14.594
Skew:		0.473	Prob(JE	3):		0.000677
Kurtosis:		4.311	Cond. N	lo.		78.8
	========					=======

[1] Standard Errors are heteroscedasticity robust (HCO) White test for heteroskedasticity: 40.282801456426604

p-value: 0.6316941370388863
 F stat: 0.8694361589339198
p-value : 0.6917665958023151

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

Breusch-Pagan test - LM (chi-sq) stat: 22.50092776639823

p-value: 0.007419972477546516
F stat: 2.780406835222417

p-value of F stat: 0.005313026608224283

Rejecting the null of homoskedasticity. Evidence of heteroskedasticity.

Harvey test:

LM (chi-sq) test statistics: 16.874183517906896

p_value: 0.05072309299337895

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

```
LR test statistics against the general model: 116.57928678402527 p_value: 4.207836015582442e-25 Reject the null hypothesis of restricted model provides a good fit: The unrestricted model fits significantly better.
```

1.4 Table 3 : Restricted model: CO2 emission 1971-2010 \sim Basic EKC Model (Column 4)

By applying restrictions to the general model shown above, the basic EKC model is estimated. Equation (5) shows the regression equation:

$$\hat{E}_{i} = \alpha_{0} + \alpha_{1}\hat{G}_{i} + \beta_{1}\hat{G}_{i}G_{i,0} + \sum_{j=4}^{k} \beta_{j}X_{j,i} + \epsilon_{i}$$

The restrictions on the general model are: $\beta_2 = \beta_3 = 0$

The variables in the above regression equation have the same meaning as mentioned above. This regression has considered all the explanatory variables (X_i) , but the results are not specified in the table.

```
[7]: # Table 3 EKC model
     import numpy as np
     import matplotlib.pyplot as plt
     import pandas as pd
     import statsmodels.api as sm
     import statsmodels.formula.api as smf
     from stargazer.stargazer import Stargazer
     from IPython.core.display import HTML
     from statsmodels.stats.diagnostic import het_breuschpagan, het_white
     from sklearn.linear_model import LinearRegression
     from scipy.stats import chi2
     df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='CO2 1971-2010')
     reg4 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + CPE + legorfr + legorge + legorsc∟
     → + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).
     →fit(cov_type='HCO')
     print(reg4.summary())
     # Finding the EKC income per capita turning point:
     # Step 1: Storing the estimates from the regression
     reg4.alpha1 = reg4.params['GRGDPPC']
     reg4.beta1 = reg4.params['DMMULGDP']
     EKC_turning_point = np.exp( - reg4.alpha1 / reg4.beta1 + 8.099206859)
```

```
# Note: The EKC turning point level : From the paper (Note \mu G = 8.09 = 1
\rightarrow ln(6340.99); where 6340.99 is cross-country mean of GDPPC1971 (initial GDP<sub>L</sub>
→per capita 1971))
#print(reg4.alpha1)
#print(req4.beta1)
print(f"EKC turning point:{EKC turning point}")
# White test for heteroskedasticity
reg4_exog = reg4.model.exog
white_test = het_white(reg4.resid, reg4_exog)
print(f"White test for heteroskedasticity: {white_test[0]}")
print(f"P_value (of t-stat): {white_test[1]}")
print(f"F Test stat: {white test[2]}")
print(f"P-value (of F-test): {white_test[3]}")
if white_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedsaticity. Evidence of,
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of_{\sqcup}
→no heteroskedasticity.")
# Breusch-Pagan test for heteroskedasticity
bp test = het breuschpagan(reg4.resid, reg4.model.exog)
print(f"BP Test statistics: {bp_test[0]}")
print(f"p_value (Chi-sq distribution): {bp_test[1]}")
print(f"F stat: {bp_test[2]}")
print(f"p_value: {bp_test[3]}")
if bp_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedsaticity. Evidence of ⊔
⇔heteroskedasticity.")
else:
    print("Fail to rejet the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.")
# Harvey-Godfrey test for heteroskedasticity (n * R^2)
# Step 1: Calculate the log of squared residuals from reg4
df['log_sq_resid'] = np.log(reg4.resid ** 2)
# Step 2: Regress the log of squared residuals on the independent variables in
\rightarrow the model
```

```
harvey_reg = smf.ols(" log_sq_resid ~ GRGDPPC + DMMULGDP + CPE + legorfr + L
→legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).
→fit(cov_type='HCO')
# Step 3: Store the r-sq value from the regression
r_sq_harvey = harvey_reg.rsquared
# calculate the number of obs
n = len(df)
# Calculate the Harvey test statistics: n * R^2
harvey_test_stat = n * r_sq_harvey
# Calculate the chi_sq stat:
p_value = chi2.sf(harvey_test_stat , df = 10) # 10 independent variables in the
→model (in harvey_reg)
print(f"harvey_test (LM stat): {harvey_test_stat}")
print(f"p-value: {p_value}")
if p_value < 0.05:
   print("Reject the null hypothesis of heteroskedasticity. Evidence of ⊔
→heteroskedasticity.")
else:
   print("Fail to reject the null hypothesis of heteroskedasticity. Evidence⊔
# LR test stat against the general model:
# Estimating the unrestricted model here again to obtain the number of \Box
→paprameters from the unrestricted model
reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE,
→+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + L
→DMLNPOPD1971", data=df).fit(cov_type='HCO')
llf_unrestricted = reg1.llf
llf_restricted = reg4.llf # Here llf stands for log-likelihood function.
lr_stat = 2 * (llf_unrestricted - llf_restricted)
df_diff = reg1.df_model - reg4.df_model
p_value_lr = chi2.sf(lr_stat, df_diff)
print(f"LR statistics: {lr_stat}")
print(f"p_value (chi-sq): {p_value_lr}")
if p_value_lr < 0.05:</pre>
```

print("Reject the null hypothesis of restricted model provides a good fit. $_\sqcup$ $_\to The$ unrestricted model fits the model better.")

else:

print("Fail to reject the null hypothesis of restricted model provides a_{\sqcup} \rightarrow good fit. No significant improvement in fit for the unrestricted model.")

OLS Regression Results

==========	ULS Regression Results								
Dep. Variable	e:	GREPC	R-squared:			0.478			
Model:		OLS	Adj. R-	-		0.435			
Method: Least Squares			F-stati:			10.77			
Date:	Thu,	27 Jun 2024	Prob (F	-statistic):		5.73e-13			
Time:		11:30:27	Log-Like	elihood:		346.81			
No. Observat:	ions:	134	AIC:			-671.6			
Df Residuals	:	123	BIC:			-639.7			
Df Model:		10							
Covariance Ty	ype: 	HCO							
========	coef	std err	z	P> z	[0.025	0.975]			
Intercept	-0.0011	0.004	-0.324	0.746	-0.008	0.006			
GRGDPPC	0.7964	0.125	6.386	0.000	0.552	1.041			
DMMULGDP	-0.2648	0.091	-2.921	0.003	-0.442	-0.087			
CPE	-0.0167	0.009	-1.863	0.062	-0.034	0.001			
legorfr	0.0036	0.004	0.958	0.338	-0.004	0.011			
legorge	-0.0057	0.007	-0.824	0.410	-0.019	0.008			
legorsc	-0.0072	0.008	-0.885	0.376	-0.023	0.009			
DMSUMT	0.0003	0.000	0.689	0.491	-0.001	0.001			
DMWINT	-8.524e-05	0.000	-0.340	0.734	-0.001	0.000			
DMLNFFPC1971	-0.0005	0.000	-1.103	0.270	-0.001	0.000			
DMLNPOPD1971	0.0001	0.001	0.113	0.910	-0.002	0.002			
Omnibus:	========	15.185	Durbin-	===================================	=======	2.044			
Prob(Omnibus)):	0.001	Jarque-Bera (JB):		42.703				
Skew:		0.293	Prob(JB):		5.34e-10				
Kurtosis:		5.703	Cond. No	ο.		878.			

Notes:

[1] Standard Errors are heteroscedasticity robust (HCO)

EKC turning point:66626.62091065194

White test for heteroskedasticity: 62.623501547779185

P_value (of t-stat): 0.17172887541964668

F Test stat: 1.3243300262956288

P-value (of F-test): 0.12649741419268612

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

 ${\tt heteroskedasticity}.$

BP Test statistics: 15.008299877096468

```
p_value (Chi-sq distribution): 0.13175955060780753
F stat: 1.551386258853481
p_value: 0.12926593356847557
Fail to rejet the null hypothesis of homoskedasticity. Evidence of no heteroskedasticity.
harvey_test (LM stat): 20.846861140784867
p-value: 0.022187086598483008
Reject the null hypothesis of heteroskedasticity. Evidence of heteroskedasticity.
LR statistics: 86.85371224185042
p_value (chi-sq): 1.3802444842278738e-19
Reject the null hypothesis of restricted model provides a good fit. The unrestricted model fits the model better.
```

1.5 Table 3 : Restricted model: CO2 emission 1971-2010 \sim IPAT Model (Column 5)

By applying restrictions to the general model shown above, the IPAT model is estimated. Equation (6) shows the regression equation:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \sum_{j=4}^k \beta_j X_{j,i} + \epsilon_i$$

The restrictions on the general model are: $\beta_1 = \beta_2 = \beta_3 = 0$

The variables in the above regression equation have the same meaning as mentioned above. This regression has considered all the explanatory variables (X_i) , but the results are not specified in the table.

```
# White test for heteroskedasticity:
reg5_exog = reg5.model.exog
white_test = het_white(reg5.resid, reg5_exog)
print(f"White test statistics: {white test[0]}")
print(f"P_value: {white_test[1]}")
print(f"F stat: {white test[2]}")
print(f"p_value: {white_test[3]}")
if white_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.")
# BP test for heteroskedasticity.
bp_test = het_breuschpagan(reg5.resid, reg5_exog)
print(f"bp_test_stat: {bp_test[0]}")
print(f"p_value: {bp_test[1]}")
if bp_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of,
→heteroskedasticity.")
else:
    print("Fail to accept the null hypothesis of homoskedasticity. Evidence of_{\sqcup}
→no heteroskedasticity.")
# Harvey-test - Has to be manually calculated
# calculate the log of squared residuals and store it in a column
df['log sq resid'] = np.log(reg5.resid ** 2)
# Regress the log od squared residuals on the independent variables
harvey_reg = smf.ols("log_sq_resid ~ GRGDPPC + CPE + legorfr + legorge + ∪
→legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971",data=df).
→fit(cov_type='HCO')
# Obtain the R^2 from the harvey regression
rsquared_harvey = harvey_reg.rsquared
# Obtain the no of obs. of the data
n=len(df)
# Harvey test stat
harvey_stat = n * rsquared_harvey
# Calculate the p_value of the harvey test - it follows a chi-sq distribution
```

```
p_value_harvey = chi2.sf(harvey_stat, df = 9) # Degree of freedom is the no of_
→ independent variables in the harvey regression.
print(f"Harvey test: {harvey stat}")
print(f"p_value: {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
    print('Reject the null hypothesis of homoskedasticity. Evidence of \Box
⇔heteroskedasticity.')
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.")
# LR test : Comparing the restricted model with the unrestricted model
# Unrestricted model (General Model)
reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE,
\hookrightarrow+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +_{\sqcup}
→DMLNPOPD1971", data=df).fit(cov_type='HCO')
# collect the llf stat from both the models
llf unrestricted = reg1.llf
llf_restricted = reg5.llf
# Calculate the lr stat:
lr_stat= 2 * (llf_unrestricted - llf_restricted)
# Calculate the p_value:
df_diff = llf_unrestricted - llf_restricted
p_value_lr = chi2.sf(lr_stat, df_diff)
print(f"LR statistics: {lr_stat}")
print(f"p_value: {p_value_lr}")
if p value lr < 0.05:
    print("Reject the null hypothesis of restricted model provides a good fit:⊔
→The unrestricted model fits significantly better.")
    print("Fail to reject the null hypothesis of restricted model provides a
→good fit: No significan timprovement in fit for the unrestricted model.")
```

 Dep. Variable:
 GREPC
 R-squared:
 0.431

 Model:
 OLS
 Adj. R-squared:
 0.389

 Method:
 Least Squares
 F-statistic:
 12.37

 Date:
 Fri, 21 Jun 2024
 Prob (F-statistic):
 7.94e-14

No. Observation Df Residuals: Df Model: Covariance Typ		134 124 9 HCC	BIC:			-662.1 -633.1
	coef	std err	z	P> z	[0.025	0.975]
Intercept	-0.0053	0.003	-1.648	0.099	-0.012	0.001
GRGDPPC	0.9890	0.118	8.377	0.000	0.758	1.220
CPE	-0.0114	0.008	-1.410	0.158	-0.027	0.004
legorfr	0.0049	0.004	1.250	0.211	-0.003	0.013
legorge	-0.0034	0.007	-0.453	0.651	-0.018	0.011
legorsc	-0.0088	0.009	-0.933	0.351	-0.027	0.010
DMSUMT	0.0005	0.000	1.355	0.176	-0.000	0.001
DMWINT	0.0001	0.000	0.396	0.692	-0.000	0.001
DMLNFFPC1971	-0.0004	0.000	-0.874	0.382	-0.001	0.000

21:25:25 Log-Likelihood:

341.04

0.002

Omnibus:	16.819	Durbin-Watson:	2.023					
Prob(Omnibus):	0.000	Jarque-Bera (JB):	51.697					
Skew:	0.319	Prob(JB):	5.94e-12					
Kurtosis:	5.975	Cond. No.	719.					

-0.869

0.385

-0.004

0.001

Notes:

DMLNPOPD1971

Time:

[1] Standard Errors are heteroscedasticity robust (HCO)

White test statistics: 59.56368515878082

-0.0012

P_value: 0.058762025963511044 F stat: 1.6185794620785032 p_value: 0.02790738794517342

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

bp_test_stat: 15.885667741819455
p value: 0.06930787287573381

Fail to accept the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

Harvey test: 14.714180180131681 p_value: 0.09909212354157193

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

LR statistics: 98.3834528217285 p_value: 3.9388383719571115e-05

Reject the null hypothesis of restricted model provides a good fit: The

unrestricted model fits significantly better.

1.6 Table 5 : General model period 1: CO2 emission 1971-1990

In this regression, the authors have estimated the general model for CO2 emission as in equation (2) in the paper but only for a truncated period 1971-1990.

Rewriting equation (2) for convenience:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_1 \hat{G}_i G_{i,0} + \beta_2 E_{i,0} + \beta_3 G_{i,0} + \sum_{j=4}^k \beta_j X_{j,i} + \epsilon_i$$

where hats indicate long-run growth rates, E is the log of emissions per-capita, and G is the log of GDP per capita. X_i is a vector of explanatory and control variables. The sample mean has been deducted from each continuous level variable in this vector.

```
[3]: # TABLE 5 CO2 - 1971-1990
     import numpy as np
     import matplotlib.pyplot as plt
     import pandas as pd
     import statsmodels.api as sm
     import statsmodels.formula.api as smf
     from stargazer.stargazer import Stargazer
     from IPython.core.display import HTML
     from statsmodels.stats.diagnostic import het_breuschpagan, het_white
     from sklearn.linear_model import LinearRegression
     from scipy.stats import chi2
     df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx',sheet_name='CO2 1971-1990')
     reg6 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE_
      \hookrightarrow+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +_{\sqcup}
     →DMLNPOPD1971", data=df).fit(cov_type='HCO')
     print(reg6.summary())
     # Turning point - EKC
     alpha1 = reg6.params['GRGDPPC']
     beta1 = reg6.params['DMMULGDP']
     EKC_turning_point = np.exp( - alpha1 / beta1 + 8.099206859)
     print(f"EKC Turning point:{EKC_turning_point}")
     # White test
     reg6_exog = reg6.model.exog
     white_test=het_white(reg6.resid, reg6_exog)
     print(f"White test stat: {white_test[0]} ")
     print(f"P_value: {white_test[1]}")
```

```
if white_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of ⊔
 ⇔heteroskedasticity.")
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
 →no heteroskedasticity.")
# Breusch-Pagan test
bp_test = het_breuschpagan(reg6.resid, reg6_exog)
print(f"BP test stat: {bp_test[0]}")
print(f"p_value: {bp_test[1]}")
if bp_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of \Box
 ⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of ⊔
 →no heteroskedasticity.")
# Harvey Test:
# Calculate the log of squared residuals from the original regression.
df['log_sq_resid'] = np.log(reg6.resid ** 2)
# Run the harvey regression
harvey_test = smf.ols("log_sq_resid ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + L
\hookrightarrow DMLNEPC1971 + CPE + legorfr + legorge + legorsc + DMSUMT + DMWINT +
→DMLNFFPC1971 + DMLNPOPD1971", data=df).fit(cov_type='HCO')
rsquared = harvey_test.rsquared
n=len(df)
harvey_stat = n * rsquared # Harvey test: n * R^2 from harvey test
p_value_harvey = chi2.sf(harvey_stat, df=12) # 12 nos of independent variables
print(f"Harvey-Godfrey test for heteroskedasticity: {harvey_stat}")
print(f"p_value: {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of \Box
⇔heteroskedasticity.")
else:
    print('Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
 →no heteroskedasticity.')
```

OLS Regression Results

Dep. Variable: Model: Method: Date: Time: No. Observations Df Residuals: Df Model: Covariance Type:	Thu,	GREPC OLS east Squares 27 Jun 2024 10:48:24 134 121 12 HC0	OLS Adj. R-squared: quares F-statistic: 1 2024 Prob (F-statistic): 148:24 Log-Likelihood: 134 AIC: 121 BIC: 12			0.586 0.545 16.09 5.90e-20 311.99 -598.0 -560.3
0.975]	coef	std err	z	P> z	[0.025	
- Intercept 0.007	-0.0014	0.004	-0.339	0.734	-0.010	
GRGDPPC 1.012	0.8177	0.099	8.259	0.000	0.624	
DMMULGDP 0.083	-0.0448	0.065	-0.688	0.492	-0.173	
DMLNGDPPC1971	0.0245	0.006	4.408	0.000	0.014	
DMLNEPC1971 -0.016	-0.0234	0.004	-6.589	0.000	-0.030	
CPE 0.014	-0.0093	0.012	-0.796	0.426	-0.032	
legorfr 0.014	0.0049	0.005	1.051	0.293	-0.004	
legorge 0.028	0.0075	0.010	0.715	0.475	-0.013	
legorsc 0.008	-0.0101	0.009	-1.109	0.268	-0.028	
DMSUMT	0.0017	0.001	2.504	0.012	0.000	
DMWINT -0.000	-0.0010	0.000	-3.060	0.002	-0.002	
DMLNFFPC1971 0.003	0.0020	0.001	2.838	0.005	0.001	
DMLNPOPD1971 0.005	0.0017	0.002	0.954	0.340	-0.002	
Omnibus: Prob(Omnibus): Skew: Kurtosis:		2.242 0.326 0.003 3.603	Durbin-Wa Jarque-Be	atson: era (JB): :		2.295 2.028 0.363 550.

[1] Standard Errors are heteroscedasticity robust (HCO)

EKC Turning point:274175751573.85257 White test stat: 106.15309225202549

P value: 0.004365902246960979

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

BP test stat: 23.802254612861898 p_value: 0.02163769164636923

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

Harvey-Godfrey test for heteroskedasticity: 11.358018236214125

p_value: 0.4985110168932617

Fail to reject the null hypothesis of homoskedasticity. Evidence of no heteroskedasticity.

1.7 Table 6: General model period 2: CO2 emission 1990-2010

In this regression, the authors have estimated the general model for CO2 emission as in equation (2) in the paper but only for a truncated period 1990-2010.

Rewriting equation (2) for convenience:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_1 \hat{G}_i G_{i,0} + \beta_2 E_{i,0} + \beta_3 G_{i,0} + \sum_{j=4}^k \beta_j X_{j,i} + \epsilon_i$$

where hats indicate long-run growth rates, \$ E\$ is the log of emissions per-capita, and G is the log of GDP per capita. X_i is a vector of explanatory and control variables. The sample mean has been deducted from each continuous level variable in this vector.

```
reg7 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1990 + DMLNEPC1990 + CPE_
\hookrightarrow+ LEGORFR + LEGORGE + LEGORSC + DMSUMT + DMWINT + DMLNFFPC1971 +_{\sqcup}
→DMLNPOPD1990 ", data=df).fit(cov_type='HCO')
print(reg7.summary())
# EKC turning point:
alpha1 = reg7.params['GRGDPPC']
beta1 = reg7.params['DMMULGDP']
EKC_turning_point = np.exp(- alpha1 / beta1 + 8.341672) # Here \mu_G is just_\to
→ the average of the column LNGDPPC1990
print(f"EKC turning point: {EKC_turning_point}")
# White test for heteroskedasticity
white_test = het_white(reg7.resid, reg7.model.exog)
print(f"White test for heteroskedasticity : {white_test[0]}")
print(f"p-value: {white_test[1]}")
if white_test[1] < 0.05:</pre>
    print('Reject the null hypothesis of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.')
    print('Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
→no heteroskedasticity.')
# Breusch-Pagan test for heteroskedasticity
bp_test = het_breuschpagan(reg7.resid, reg7.model.exog)
print(f"Breusch-Pagan Test: {bp_test[0]}")
print(f"p_value: {bp_test[1]}")
if bp_test[1] < 0.05:</pre>
    print('Reject the null hypothesis of homoskedasticity. Evidence of \Box
⇔heteroskedasticity.')
else:
    print('Fail to reject the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.')
# Harvey-Godfrey test:
# Calulate the log of squared residuals from the original regression: reg?
df['log_sq_resid'] = np.log(reg7.resid ** 2)
#print(df.head())
# Regress the log of squared residuals with the independent variables of the
→model (Auxiliary regression)
harvey_reg = smf.ols("log_sq_resid ~ GRGDPPC + DMMULGDP + DMLNGDPPC1990 + L
→DMLNEPC1990 + CPE + LEGORFR + LEGORGE + LEGORSC + DMSUMT + DMWINT + L
 →DMLNFFPC1971 + DMLNPOPD1990 ",data=df).fit(cov_type = 'HCO')
```

```
# Store the r-square value from the harvey_reg
harvey_rsquared = harvey_reg.rsquared
# Calculate the number of independent variable in the model
n = len(df)
# Find the Harvey test stat and the p-value (follows a chi-sq distribution)
harvey_test = n * harvey_rsquared
p_value_harvey = chi2.sf(harvey_test, df=12) # the degree of freedom is equal_
→ to the number of independent variables in the model.
print(f"Harvey Test Stat: {harvey_test}")
print(f"p value : {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of_{\sqcup}
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of
→no heteroskedasticity.")
```

=======================================	=======	=========				:=====
Dep. Variable: Model: Method: Date: Time: No. Observations Df Residuals: Df Model: Covariance Type:	Wed,	ast Squares 26 Jun 2024		quared: :ic: :tatistic):	2.	0.580 0.538 38.10 17e-35 335.95 -645.9 -608.2
=======================================				========	========	:======
0.975]	coef	std err	z	P> z	[0.025	
- Intercept 0.002	-0.0036	0.003	-1.191	0.234	-0.009	
GRGDPPC	0.9279	0.136	6.848	0.000	0.662	
1.193 DMMULGDP 0.007	-0.1095	0.060	-1.836	0.066	-0.227	
DMLNGDPPC1990	0.0137	0.004	3.338	0.001	0.006	
0.022 DMLNEPC1990 -0.006	-0.0133	0.003	-3.814	0.000	-0.020	
CPE	-0.0043	0.010	-0.455	0.649	-0.023	

0.014 LEGORFR	-0.0027	0.004	-0.754	0.451	-0.010	
0.004 LEGORGE 0.010	-0.0031	0.007	-0.473	0.636	-0.016	
LEGORSC 0.012	0.0007	0.006	0.125	0.900	-0.011	
DMSUMT	0.0009	0.001	1.591	0.112	-0.000	
DMWINT 0.001	-6.769e-05	0.000	-0.188	0.851	-0.001	
DMLNFFPC1971 0.002	0.0010	0.001	1.953	0.051	-3.82e-06	
DMLNPOPD1990 0.002	-0.0013	0.002	-0.765	0.444	-0.005	
======================================	=======	======================================	======= Durbin-V		=======	2.010
Prob(Omnibus)	:	0.000				91.734
Skew:		-1.065	Prob(JB)			1.20e-20
Kurtosis:	========	6.448	Cond. No). =======	========	891.

[1] Standard Errors are heteroscedasticity robust (HCO)

EKC turning point: 20005157.212090995

White test for heteroskedasticity: 107.9285530174746

p-value: 0.0031082454974498777

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

Breusch-Pagan Test: 17.140654481707116

p_value: 0.14438252588514103

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

Harvey Test Stat: 26.704368966986404

p value : 0.008520708310142702

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

1.8 Table 6: General model period 2: SO2 emission 1988-2005

In this regression, the authors have estimated the general model for SO2 emission as in equation (2) in the paper but only for a truncated period 1988-2005.

Rewriting equation (2) for convenience:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_1 \hat{G}_i G_{i,0} + \beta_2 E_{i,0} + \beta_3 G_{i,0} + \sum_{i=4}^k \beta_i X_{j,i} + \epsilon_i$$

where hats indicate long-run growth rates, \$ E\$ is the log of emissions per-capita, and G is the log of GDP per capita. X_i is a vector of explanatory and control variables. The sample mean has been deducted from each continuous level variable in this vector.

```
[14]: # Table 6 - SO2 1988-2005
      import numpy as np
      import matplotlib.pyplot as plt
      import pandas as pd
      import statsmodels.api as sm
      import statsmodels.formula.api as smf
      from stargazer.stargazer import Stargazer
      from IPython.core.display import HTML
      from statsmodels.stats.diagnostic import het_breuschpagan, het_white
      from sklearn.linear model import LinearRegression
      from scipy.stats import chi2
      df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx',sheet_name='S02 1988-2005')
      # print(df.head())
      # Note in the regression the change of initial date to 1990 from 1971. But the
      →log of fossial fuel consumption is calculated in 1971 not in 1990.
      reg7 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1988 + DMLNEPC1988 + CPE_
      →+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + U
      →DMLNPOPD1988 ", data=df).fit(cov_type='HCO')
      print(reg7.summary())
      # EKC turning point:
      alpha1 = reg7.params['GRGDPPC']
      beta1 = reg7.params['DMMULGDP']
      EKC_turning_point = np.exp(- alpha1 / beta1 + 8.5637) # Here \mu_G is just the_
      →average of the column LNGDPPC1988
      print(f"EKC turning point: {EKC_turning_point}")
      # White test for heteroskedasticity
      white_test = het_white(reg7.resid, reg7.model.exog)
      print(f"White test for heteroskedasticity : {white_test[0]}")
      print(f"p-value: {white_test[1]}")
      if white_test[1] < 0.05:</pre>
          print('Reject the null hypothesis of homoskedasticity. Evidence of \Box
      ⇔heteroskedasticity.')
      else:
          print('Fail to reject the null hypothesis of homoskedasticity. Evidence of ⊔
      →no heteroskedasticity.')
      # Breusch-Pagan test for heteroskedasticity
      bp_test = het_breuschpagan(reg7.resid, reg7.model.exog)
      print(f"Breusch-Pagan Test: {bp_test[0]}")
      print(f"p_value: {bp_test[1]}")
```

```
if bp_test[1] < 0.05:</pre>
    print('Reject the null hypothesis of homoskedasticity. Evidence of ⊔
→heteroskedasticity.')
else:
    print('Fail to reject the null hypothesis of homoskedasticity. Evidence of,
→no heteroskedasticity.')
# Harvey-Godfrey test:
# Calulate the log of squared residuals from the original regression: reg7
df['log_sq_resid'] = np.log(reg7.resid ** 2)
#print(df.head())
# Regress the log of squared residuals with the independent variables of the
→model (Auxiliary regression)
harvey_reg = smf.ols("log_sq_resid ~ GRGDPPC + DMMULGDP + DMLNGDPPC1988 +__
\hookrightarrow DMLNEPC1988 + CPE + legorfr + legorge + legorsc + DMSUMT + DMWINT +
→DMLNFFPC1971 + DMLNPOPD1988 ",data=df).fit(cov_type = 'HCO')
# Store the r-square value from the harvey req
harvey_rsquared = harvey_reg.rsquared
# Calculate the number of independent variable in the model
n = len(df)
# Find the Harvey test stat and the p-value (follows a chi-sq distribution)
harvey_test = n * harvey_rsquared
p_value_harvey = chi2.sf(harvey_test, df=12) # the degree of freedom is equal_
→ to the number of independent variables in the model.
print(f"Harvey Test Stat: {harvey_test}")
print(f"p value : {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
→no heteroskedasticity.")
```

GREPC Dep. Variable: R-squared: 0.624 Model: OLS Adj. R-squared: 0.562 Method: Least Squares F-statistic: 12.29 Wed, 26 Jun 2024 Prob (F-statistic): 3.41e-13 Date: Time: 11:05:47 Log-Likelihood: 173.99 No. Observations: AIC: 86 -322.0BIC: Df Residuals: 73 -290.1Df Model: 12

HC0

Covariance Type:

		========				=======
= 0.975]	coef	std err	z	P> z	[0.025	
_						
Intercept	-0.0223	0.008	-2.893	0.004	-0.037	
-0.007						
GRGDPPC	0.7678	0.190	4.041	0.000	0.395	
1.140						
DMMULGDP	-0.5485	0.189	-2.903	0.004	-0.919	
-0.178						
DMLNGDPPC1988	0.0210	0.005	3.891	0.000	0.010	
0.032						
DMLNEPC1988	-0.0135	0.004	-3.250	0.001	-0.022	
-0.005						
CPE	-0.0062	0.023	-0.272	0.786	-0.051	
0.039						
legorfr	-0.0099	0.008	-1.261	0.207	-0.025	
0.005						
legorge	-0.0264	0.017	-1.583	0.113	-0.059	
0.006						
legorsc	-0.0733	0.017	-4.306	0.000	-0.107	
-0.040						
DMSUMT	0.0045	0.001	4.473	0.000	0.003	
0.007						
DMWINT	-0.0007	0.001	-1.171	0.242	-0.002	
0.000						
DMLNFFPC1971	-0.0007	0.001	-0.513	0.608	-0.003	
0.002						
DMLNPOPD1988	-0.0094	0.003	-3.233	0.001	-0.015	
-0.004						
==========	=======	========				======
Omnibus:		0.373	Durbin-Watson:		1.817	
<pre>Prob(Omnibus):</pre>		0.830				0.268
Skew:			Prob(JB):			0.875
Kurtosis:		2.965	Cond. No			917.

[1] Standard Errors are heteroscedasticity robust (HCO)

EKC turning point: 21238.08513049237

White test for heteroskedasticity : 76.44044721348092

p-value: 0.2258895650380592

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

Breusch-Pagan Test: 9.977120443058821

p_value: 0.6179679539055802

```
Fail to reject the null hypothesis of homoskedasticity. Evidence of no heteroskedasticity.

Harvey Test Stat: 10.424952369779191

p value: 0.578732888079803

Fail to reject the null hypothesis of homoskedasticity. Evidence of no heteroskedasticity.
```

1.9 Table 5: General model period 1: SO2 emission 1971-1988

In this regression, the authors have estimated the general model for SO2 emission as in equation (2) in the paper but only for a truncated period 1971-1988.

Rewriting equation (2) for convenience:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_1 \hat{G}_i G_{i,0} + \beta_2 E_{i,0} + \beta_3 G_{i,0} + \sum_{j=4}^k \beta_j X_{j,i} + \epsilon_i$$

where hats indicate long-run growth rates. E is the log of emissions per-capita, and G is the log of GDP per capita. X_i is a vector of explanatory and control variables. The sample mean has been deducted from each countinuous levels variables in this vector.

```
[15]: # Table 5 - SO2 1971-1988
      import numpy as np
      import matplotlib.pyplot as plt
      import pandas as pd
      import statsmodels.api as sm
      import statsmodels.formula.api as smf
      from stargazer.stargazer import Stargazer
      from IPython.core.display import HTML
      from statsmodels.stats.diagnostic import het breuschpagan, het white
      from sklearn.linear_model import LinearRegression
      from scipy.stats import chi2
      df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx',sheet_name='S02 1971-1988')
      # print(df.head())
      # Note in the regression the change of initial date to 1990 from 1971. But the
       →log of fossial fuel consumption is calculated in 1971 not in 1990.
      reg7 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE,
       \hookrightarrow+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +_{\sqcup}
       →DMLNPOPD1971 ", data=df).fit(cov_type='HCO')
      print(reg7.summary())
      # EKC turning point:
      alpha1 = reg7.params['GRGDPPC']
      beta1 = reg7.params['DMMULGDP']
```

```
EKC_turning_point = np.exp(- alpha1 / beta1 + 8.318) # Here \mu_G is just the_\u00cd
→average of the column LNGDPPC1971
print(f"EKC turning point: {EKC_turning_point}")
# White test for heteroskedasticity
white test = het white(reg7.resid, reg7.model.exog)
print(f"White test for heteroskedasticity : {white_test[0]}")
print(f"p-value: {white_test[1]}")
if white_test[1] < 0.05:</pre>
    print('Reject\ the\ null\ hypothesis\ of\ homoskedasticity.\ Evidence\ of_{\sqcup}
⇔heteroskedasticity.')
    print('Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
\hookrightarrowno heteroskedasticity.')
# Breusch-Pagan test for heteroskedasticity
bp_test = het_breuschpagan(reg7.resid, reg7.model.exog)
print(f"Breusch-Pagan Test: {bp_test[0]}")
print(f"p_value: {bp_test[1]}")
if bp test[1] < 0.05:
    print('Reject the null hypothesis of homoskedasticity. Evidence of \Box
→heteroskedasticity.')
else:
    print('Fail to reject the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.')
# Harvey-Godfrey test:
# Calulate the log of squared residuals from the original regression: reg?
df['log_sq_resid'] = np.log(reg7.resid ** 2)
#print(df.head())
# Regress the log of squared residuals with the independent variables of the
→ model (Auxiliary regression)
harvey_reg = smf.ols("log_sq_resid ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + L
→DMLNEPC1971 + CPE + legorfr + legorge + legorsc + DMSUMT + DMWINT + L
→DMLNFFPC1971 + DMLNPOPD1971 ",data=df).fit(cov_type = 'HCO')
# Store the r-square value from the harvey_reg
harvey_rsquared = harvey_reg.rsquared
# Calculate the number of independent variable in the model
n = len(df)
# Find the Harvey test stat and the p-value (follows a chi-sq distribution)
harvey_test = n * harvey_rsquared
p_value_harvey = chi2.sf(harvey_test, df=12 ) # the degree of freedom is equal_
→ to the number of independent variables in the model.
```

```
print(f"Harvey Test Stat: {harvey_test}")
print(f"p value : {p_value_harvey}")

if p_value_harvey < 0.05:
    print("Reject the null hypothesis of homoskedasticity. Evidence of 
    →heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of 
    →no heteroskedasticity.")</pre>
```

Dep. Variable: Model: Method: Date: Time: No. Observations Df Residuals: Df Model: Covariance Type:	Wed,	ast Squares 26 Jun 2024 11:11:58 100 87 12 HC0	Prob (F-s Log-Likel AIC: BIC:	uared: ic: tatistic): ihood:	0.635 0.585 8.892 7.62e-11 172.87 -319.7 -285.9	
0.975]	coef	std err	z	P> z	[0.025	
Intercept 0.026	0.0076	0.009	0.812	0.417	-0.011	
GRGDPPC 1.208	0.8340	0.191	4.373	0.000	0.460	
DMMULGDP 0.127	-0.1498	0.141	-1.060	0.289	-0.427	
DMLNGDPPC1971	0.0246	0.012	2.101	0.036	0.002	
DMLNEPC1971 -0.012	-0.0304	0.009	-3.248	0.001	-0.049	
CPE 0.121	0.0728	0.024	2.985	0.003	0.025	
legorfr -0.006	-0.0241	0.009	-2.602	0.009	-0.042	
legorge -0.015	-0.0573	0.022	-2.647	0.008	-0.100	
legorsc -0.027	-0.0714	0.023	-3.139	0.002	-0.116	
DMSUMT 0.006	0.0031	0.001	2.495	0.013	0.001	

==========	=========				===============
Kurtosis:		9.295	Cond. No.		547.
Skew:		1.523	Prob(JB):	:	5.65e-45
<pre>Prob(Omnibus):</pre>		0.000	Jarque-Be	era (JB):	203.768
Omnibus:		48.133	Durbin-Wa	atson:	1.853
-0.004 =========	========			.=======	
DMLNPOPD1971	-0.0120	0.004	-2.942	0.003	-0.020
0.001					
DMLNFFPC1971	-0.0017	0.001	-1.444	0.149	-0.004
-0.001					
DMWINT	-0.0022	0.001	-3.667	0.000	-0.003

[1] Standard Errors are heteroscedasticity robust (HCO)

EKC turning point: 1072271.276296573

White test for heteroskedasticity: 99.16947725009595

p-value: 0.010104134059967036

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

Breusch-Pagan Test: 26.56389227610282

p value: 0.008924297940185963

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

Harvey Test Stat: 12.60396061309577

p value : 0.39847104634769565

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

1.10 Table 4 : Restricted model: SO2 emission 1971-2005 \sim Ordas Criado et al. (2011) (Column 2)

By applying restrictions to the general model shown above, Ordas Criado et al. (2011) is estimated. Equation (3) depicts the regression equation:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_2 E_{i,0} + \beta_3 G_{i,0} + \sum_{i=4}^k \beta_j X_{j,i} + \epsilon_i$$

The restrictions on the general model is: $\beta_1 = 0$

```
[16]: import numpy as np
  import matplotlib.pyplot as plt
  import pandas as pd
  import statsmodels.api as sm
  import statsmodels.formula.api as smf
```

```
from stargazer.stargazer import Stargazer
from IPython.core.display import HTML
from statsmodels.stats.diagnostic import het_breuschpagan, het_white
from sklearn.linear_model import LinearRegression
from scipy.stats import chi2
df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='S02 1971-2005')
reg2= smf.ols("GREPC ~ GRGDPPC + DMLNGDPPC1971 + DMLNEPC1971 + CPE + legorfr +
→legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).
→fit(cov type='HCO')
print(reg2.summary())
# For conducting different statistical tests we have to follow the following ...
# 1. Create a matrix of exogenous variables
exog_het = reg2.model.exog
# 2. Run the white test: HO: Homoskedasticity
white_test = het_white(reg2.resid, exog_het)
white_pvalue = white_test[1]
# 3. Print the results
print(f"White test for heteroskedasticity:")
print(f"T-stats for white test: {white_test[0]}")
print(f"P-value: {white_test[1]}")
print(f" F statistics: {white_test[2]}")
print(f" p_value: {white_test[3]}")
if white_pvalue < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of,
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. No evidence⊔
→of heteroskedasticity")
# 3. Run the Breusch-Pagan test: HO: Homoskedasticity
# The het breuschpagan function returns a tuple with four elements:
# a) Lagrange Multiplier (LM) Statistic: The first element, accessible with \square
→ bp_test[0], is the test statistic for the Breusch-Pagan test.
# b) p-value: The second element, accessible with bp_test[1], is the p-value__
\hookrightarrowassociated with the LM statistic.
```

```
# c) f-value: The third element, accessible with bp test[2], is the f-value of
\rightarrow the test.
# d) f p-value: The fourth element, accessible with bp_test[3], is the p-value_
\rightarrow associated with the f-statistic.
bp_test = het_breuschpagan(reg2.resid, exog_het)
bp_pvalue= bp_test[1]
print(f"BP test:")
print(f"LM Statistic (Chi-square): {bp_test[0]}")
print(f"LM Test p-value: {bp test[1]}")
print(f" F statistics: {bp_test[2]}")
print(f" p_value: {bp_test[3]}")
if bp_pvalue < 0.05:</pre>
      print("Reject the null hypothesis of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
else:
     print("Fail to reject the null hypothesis of homoskedasticity. Evidence⊔
# 4. Run the Harvey test (Harvey test doesn't exist in any statistical packagesu
→ in python. Thus it has to do it manually)
    # Step 1: Calculate the log of the squared residuals
df['log sq residuals'] = np.log(reg2.resid**2)
    # Step 2: Regress log_squared_residuals on the independent variables
harvey_model = smf.ols("log_sq_residuals ~ GRGDPPC + DMLNGDPPC1971 +__
\hookrightarrow DMLNEPC1971 + CPE + legorfr + legorge + legorsc + DMSUMT + DMWINT +
→DMLNFFPC1971 + DMLNPOPD1971", data=df).fit()
    # Step 3: Obtain the R-squared value from this regression
r_sq = harvey_model.rsquared
    # Step 4: Calculate the LM statistic
n = len(df) # this command helps in identifying the number of obs in the
→dataset. len() ~ determines the length of various objects.
lm_stat_harvey = n * r_sq # The LM statistics is calulated as N * r_squared
    # Step 5: Calculate the p-value
```

```
p_value_harvey = chi2.sf(lm_stat_harvey, df=11) # degree of freedom (df) = 11_L
→ for eleven independent variables
    # Step 6: Print the results
print("Harvey-Godfrey test for heteroskedasticity:") # HO: Homoskedasticity
print(f" LM Statistics: {lm_stat_harvey}")
print(f"LM test p-value: {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
→no hoteroskedasticity.")
# 5. Performing the Likelihood test (LR test) against the general model. \Box
\hookrightarrow (follows a chi-sq distribution)
# The Likelihood Ratio (LR) test is used to compare the goodness-of-fit of two.
→nested models. The HO: The restricted model provides a good fit.
reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE_
\hookrightarrow+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +_{\sqcup}
→DMLNPOPD1971", data=df).fit(cov_type='HCO')
    # Calculating the log-likelihood values of both the models
llf_restricted = reg2.llf
llf_unrestricted = reg1.llf
    # Calculating the LR statistics
lr_statistic = -2 * (llf_restricted - llf_unrestricted)
    # Degrees of freedom (df): number of restrictions (difference in number of \Box
→parameters). In this case it should be 1
df_diff = reg1.df_model - reg2.df_model # Note: reg1 is the unrestricted model_
→and reg2 is a restricted model
    # Calculating the p_value
p_value = chi2.sf(lr_statistic, df_diff)
print("LR test against the general model:")
print(f"LR Statistic: {lr_statistic}")
print(f"p-value: {p_value}")
```

if p_value < 0.05:</pre>

print("Reject the null hypothesis of restricted model provides a good fit: $_\sqcup$ $_\to The$ unrestricted model fits significantly better.")

else:

print("Fail to reject the null hypothesis of restricted model provides a_{\sqcup} \hookrightarrow good fit: No significant improvement in fit for the unrestricted model.")

OLS Regression Results

		:=======			
Dep. Variable: Model: Method: Date: Time: No. Observations Df Residuals: Df Model: Covariance Type:	Wed,	GREPC OLS east Squares 26 Jun 2024 11:31:24 100 88 11 HC0		quared: cic: statistic): lihood:	0.735 0.702 17.98 3.83e-18 225.08 -426.2 -394.9
0.975]	coef	std err	z	P> z	[0.025
- Intercept	-0.0129	0.006	-2.212	0.027	-0.024
-0.001 GRGDPPC 1.287	0.9723	0.161	6.056	0.000	0.658
DMLNGDPPC1971 0.028	0.0179	0.005	3.482	0.000	0.008
DMLNEPC1971 -0.015	-0.0220	0.004	-5.988	0.000	-0.029
CPE 0.059	0.0336	0.013	2.559	0.010	0.008
legorfr -0.003	-0.0141	0.006	-2.386	0.017	-0.026
legorge -0.013	-0.0329	0.010	-3.214	0.001	-0.053
legorsc -0.012	-0.0431	0.016	-2.712	0.007	-0.074
DMSUMT 0.005	0.0033	0.001	4.582	0.000	0.002
DMWINT -0.000	-0.0011	0.000	-2.892	0.004	-0.002
DMLNFFPC1971 0.001	-0.0010	0.001	-1.169	0.242	-0.003
DMLNPOPD1971	-0.0097	0.002	-4.368	0.000	-0.014

-0.005

1.728
5.166
0.0756
774.

Notes:

[1] Standard Errors are heteroscedasticity robust (HCO)

White test for heteroskedasticity:

T-stats for white test: 80.39682298732005

P-value: 0.04055479364221336

F statistics: 2.6657890661271875 p_value: 0.0007695731167671187

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

BP test:

LM Statistic (Chi-square): 19.756826591929567

LM Test p-value: 0.048789231219805405 F statistics: 1.9696954397810522

p_value: 0.0410117574185589

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

Harvey-Godfrey test for heteroskedasticity:

LM Statistics: 16.358572378400943 LM test p-value: 0.1283323721092265

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

hoteroskedasticity.

LR test against the general model: LR Statistic: 3.3711394882227523 p-value: 0.06634785046250796

Fail to reject the null hypothesis of restricted model provides a good fit: No

significant improvement in fit for the unrestricted model.

1.11 Table 4 : Restricted model: SO2 emission 1971-2005 \sim Green Solow Model (Column 3)

By applying restrictions to the general model shown above, Green Solow model (GSM) is estimated. Equation (4) shows the regression equation:

$$\hat{E}_{i} = \alpha_{0} + \beta_{2} E_{i,0} + \sum_{j=4}^{k} \beta_{j} X_{j,i} + \epsilon_{i}$$

The restrictions on the general model are: $\alpha_1 = \beta_1 = \beta_3 = 0$

```
[17]: # Table 4 SO2 1971-2005 Regression Results
      # Green Solow model
      import numpy as np
      import matplotlib.pyplot as plt
      import pandas as pd
      import statsmodels.api as sm
      import statsmodels.formula.api as smf
      from stargazer.stargazer import Stargazer
      from IPython.core.display import HTML
      from statsmodels.stats.diagnostic import het_breuschpagan, het_white
      from sklearn.linear_model import LinearRegression
      from scipy.stats import chi2
      df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='S02_1971-2005')
      reg3= smf.ols("GREPC ~ DMLNEPC1971 + CPE + legorfr + legorge + legorsc + DMSUMT_
      →+ DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).fit(cov_type = 'HCO')
      print(reg3.summary())
      # White test for heteroskedasticity
       # Step 1: Create a matrix of exogenous variables
      exog_het = reg3.model.exog
       # Follow the above steps for the White test.
      white_test = het_white(reg3.resid, exog_het)
      print(f"White test for heteroskedasticity: {white_test[0]}")
      print(f"p-value: {white_test[1]}")
      print(f" F stat: {white_test[2]}")
      print(f"p-value : {white test[3]}")
      if white_test[1] < 0.05:</pre>
          print("Reject the null hypothesis of homoskedasticity. Evidence of \Box
      ⇔heteroskedasticity.")
      else:
          print("Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
      →no heteroskedasticity.")
      # Breusch-Pagan Test for heteroskedasticity
      bp_test = het_breuschpagan(reg3.resid,exog_het)
      print(f"Breusch-Pagan test - LM (chi-sq) stat: {bp_test[0]}")
      print(f"p-value: {bp_test[1]}")
      print(f"F stat: {bp_test[2]}")
      print(f"p-value of F stat: {bp_test[3]}")
```

```
if bp_test[1] < 0.05:</pre>
    print("Rejecting the null of homoskedasticity. Evidence of
⇔heteroskedasticity.")
else:
    print("Fail to reject the null of homoskedasticity. Evidence of no,,
→heteroskedasticity.")
# Harvey-Godfrey test. Note the python packages donot have the harvey-test_{\sqcup}
→ module, so have to do it manually.
# Step 1: Calculate the log of squared residuals
df['log_sq_resid'] = np.log(reg3.resid ** 2)
# Step 2: Regress the log of residuals on the independent variables
harvey_reg = smf.ols("log_sq_resid ~ DMLNEPC1971 + CPE + legorfr + legorge + ⊔
→legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).
→fit(cov_type='HCO')
# Step 3: Calculate the r-sq value from the regression
r_sq = harvey_reg.rsquared
# Step 4: Calculate the LM statistics:
          # Explain the number of obs
n=len(df)
lm_stat = n * r_sq # the harvey test statistic is calculated as N * R^2
# Step 5: Calculate the p_value
p_value_harvey = chi2.sf(lm_stat,df=9)
print("Harvey test:")
print(f"LM (chi-sq) test statistics: {lm_stat}")
print(f"p_value: {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of_{\sqcup}
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
→no heteroskedasticity.")
```

```
\# LR statistics: It is used to find the goodness of fit of restricted model \sqcup
\rightarrow against the unrestricted model.
# The unrestricted model:
reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE,
→+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +
→DMLNPOPD1971", data=df).fit(cov_type='HCO')
# Calculate the log likelihood ratio of both the models
llf unrestricted = reg1.llf
llf_restricted = reg3.llf
# Caculating the LR Statistics
lr_stat = - 2 * (llf_restricted - llf_unrestricted)
# Degrees of freedom (difference in the number of parameters between the
→restricted and unrestricted model)
df_diff = reg1.df_model - reg3.df_model
# LR stat follows a chi-sq distribution
p_value = chi2.sf(lr_stat, df_diff)
print (f"LR test statistics against the general model: {lr_stat}")
print(f"p_value: {p_value}")
if p_value < 0.05:</pre>
    print("Reject the null hypothesis of restricted model provides a good fit:⊔
→The unrestricted model fits significantly better.")
else:
    print("Fail to reject the null hypothesis of restricted model provides and
→good fit: No significan timprovement in fit for the unrestricted model.")
```

OLS Regression Results

______ Dep. Variable: GREPC R-squared: 0.605 Model: OLS Adj. R-squared: 0.565 Method: Least Squares F-statistic: 10.22 Date: Wed, 26 Jun 2024 Prob (F-statistic): 1.07e-10 Time: 11:38:07 Log-Likelihood: 204.97 No. Observations: 100 AIC: -389.9Df Residuals: 90 BIC: -363.99

Df Model:

Covariance	Type:	HC	0

==========		========	=======	========		=======
	coef	std err	Z	P> z	[0.025	0.975]
Intercept	0.0082	0.006	1.381	0.167	-0.003	0.020
DMLNEPC1971	-0.0176	0.003	-6.188	0.000	-0.023	-0.012
CPE	0.0219	0.010	2.236	0.025	0.003	0.041
legorfr	-0.0213	0.007	-2.982	0.003	-0.035	-0.007
legorge	-0.0342	0.010	-3.329	0.001	-0.054	-0.014
legorsc	-0.0341	0.021	-1.645	0.100	-0.075	0.007
DMSUMT	0.0035	0.001	4.711	0.000	0.002	0.005
DMWINT	-0.0019	0.000	-4.098	0.000	-0.003	-0.001
DMLNFFPC1971 -	-3.701e-05	0.001	-0.041	0.968	-0.002	0.002
DMLNPOPD1971	-0.0050	0.003	-1.947	0.051	-0.010	3.26e-05
==========		=========	======	========	=======	======
Omnibus:		6.475	Durbin-	Watson:		1.966
Prob(Omnibus)	:	0.039	Jarque-	Bera (JB):		7.419
Skew:		0.353	Prob(JB	:		0.0245
Kurtosis:		4.132	Cond. N	o.		74.6

[1] Standard Errors are heteroscedasticity robust (HCO) White test for heteroskedasticity: 69.74169865974989

p-value: 0.0060800548846500395
F stat: 3.001701821699393
p-value : 6.566761639620607e-05

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

Breusch-Pagan test - LM (chi-sq) stat: 22.711000979026164

p-value: 0.006879045105691347
F stat: 2.9384519487518665

p-value of F stat: 0.004208768486539984

Rejecting the null of homoskedasticity. Evidence of heteroskedasticity.

Harvey test:

LM (chi-sq) test statistics: 5.700805440724463

p_value: 0.7694491335840247

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

LR test statistics against the general model: 43.59137824180459

p_value: 1.843034714071239e-09

Reject the null hypothesis of restricted model provides a good fit: The unrestricted model fits significantly better.

1.12 Table 4 : Restricted model: SO2 emission 1971-2005 \sim Basic EKC Model (Column 4)

By applying restrictions to the general model shown above, the basic EKC model is estimated. Equation (5) shows the regression equation:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_1 \hat{G}_i G_{i,0} + \sum_{j=4}^k \beta_j X_{j,i} + \epsilon_i$$

The restrictions on the general model are: $\beta_2=\beta_3=0$

```
[8]: # Table 4 SO2 1971-2005 Regression Results
     # EKC model
     import numpy as np
     import matplotlib.pyplot as plt
     import pandas as pd
     import statsmodels.api as sm
     import statsmodels.formula.api as smf
     from stargazer.stargazer import Stargazer
     from IPython.core.display import HTML
     from statsmodels.stats.diagnostic import het breuschpagan, het white
     from sklearn.linear_model import LinearRegression
     from scipy.stats import chi2
     df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='S02 1971-2005')
     reg4 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + CPE + legorfr + legorge + legorsc⊔
     \hookrightarrow+ DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).

→fit(cov_type='HC0')
     print(reg4.summary())
     # Finding the EKC income per capita turning point:
     # Step 1: Storing the estimates from the regression
     reg4.alpha1 = reg4.params['GRGDPPC']
     reg4.beta1 = reg4.params['DMMULGDP']
     EKC_turning_point = np.exp( - reg4.alpha1 / reg4.beta1 + 8.3187336) # Log of_
      →LNGDPPC1971
     #print(reg4.alpha1)
     #print(req4.beta1)
     print(f"EKC turning point:{EKC_turning_point}")
     # White test for heteroskedasticity
```

```
reg4_exog = reg4.model.exog
white_test = het_white(reg4.resid, reg4_exog)
print(f"White test for heteroskedasticity: {white_test[0]}")
print(f"P_value (of t-stat): {white_test[1]}")
print(f"F Test stat: {white_test[2]}")
print(f"P-value (of F-test): {white_test[3]}")
if white test[1] < 0.05:
    print("Reject the null hypothesis of homoskedsaticity. Evidence of,
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of \Box
→no heteroskedasticity.")
# Breusch-Pagan test for heteroskedasticity
bp_test = het_breuschpagan(reg4.resid, reg4.model.exog)
print(f"BP Test statistics: {bp_test[0]}")
print(f"p_value (Chi-sq distribution): {bp_test[1]}")
print(f"F stat: {bp_test[2]}")
print(f"p_value: {bp_test[3]}")
if bp_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedsaticity. Evidence of,
⇔heteroskedasticity.")
else:
    print("Fail to rejet the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.")
# Harvey-Godfrey test for heteroskedasticity (n * R^2)
# Step 1: Calculate the log of squared residuals from reg4
df['log sq resid'] = np.log(reg4.resid ** 2)
# Step 2: Regress the log of squared residuals on the independent variables in \Box
\rightarrow the model
harvey_reg = smf.ols(" log_sq_resid ~ GRGDPPC + DMMULGDP + CPE + legorfr + L
⇒legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971", data=df).
→fit(cov_type='HCO')
# Step 3: Store the r-sq value from the regression
r_sq_harvey = harvey_reg.rsquared
# calculate the number of obs
n = len(df)
# Calculate the Harvey test statistics: n * R^2
harvey_test_stat = n * r_sq_harvey
```

```
# Calculate the chi_sq stat:
p_value = chi2.sf(harvey_test_stat , df = 10) # 10 independent variables in the_
→model (in harvey_req)
print(f"harvey_test (LM stat): {harvey_test_stat}")
print(f"p-value: {p value}")
if p_value < 0.05:</pre>
    print("Reject the null hypothesis of heteroskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of heteroskedasticity. Evidence,
→of no heteroskedasticity.")
# LR test stat against the general model:
# Estimating the unrestricted model here again to obtain the number of \Box
→ paprameters from the unrestricted model
reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE,
→+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +
→DMLNPOPD1971", data=df).fit(cov_type='HCO')
llf_unrestricted = reg1.llf
llf_restricted = reg4.llf # Here llf stands for log-likelihood function.
lr_stat = 2 * (llf_unrestricted - llf_restricted)
df diff = reg1.df model - reg4.df model
p_value_lr = chi2.sf(lr_stat, df_diff)
print(f"LR statistics: {lr stat}")
print(f"p_value (chi-sq): {p_value_lr}")
if p_value_lr < 0.05:</pre>
    print("Reject the null hypothesis of restricted model provides a good fit. ⊔
→ The unrestricted model fits the model better.")
else:
    print("Fail to reject the null hypothesis of restricted model provides a_{\sqcup}
⇒good fit. No significant improvement in fit for the unrestricted model.")
```

OLS Regression Results

```
Dep. Variable: GREPC R-squared: 0.529
Model: OLS Adj. R-squared: 0.476
Method: Least Squares F-statistic: 14.23
```

Date:	Thu, 27 Jun 2024	Prob (F-statistic):	1.08e-14
Time:	11:32:28	Log-Likelihood:	196.25
No. Observations:	100	AIC:	-370.5
Df Residuals:	89	BIC:	-341.8
Df Model:	10		

HC0

=====
0.975]
0.016
1.313
-0.251
0.032
-0.006
-0.018
-0.011
0.005
0.000
-0.001
-0.004
2.031 2.431 2.431 2.431 881.
- 2. e

Notes:

[1] Standard Errors are heteroscedasticity robust (HCO)

EKC turning point:19838.54727680577

White test for heteroskedasticity: 64.09531979328712

P_value (of t-stat): 0.1030160275013332

F Test stat: 1.6801432714889248

P-value (of F-test): 0.03610453324727679

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

Covariance Type:

BP Test statistics: 16.296637382261036

p_value (Chi-sq distribution): 0.09144986354240889

F stat: 1.732786690595694 p_value: 0.08554107801825975

Fail to rejet the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

harvey_test (LM stat): 11.194721737508557

p-value: 0.3425500693706504

Fail to reject the null hypothesis of heteroskedasticity. Evidence of no

heteroskedasticity.

LR statistics: 61.03891614240899

p_value (chi-sq): 5.566314864214063e-14

Reject the null hypothesis of restricted model provides a good fit. The unrestricted model fits the model better.

1.13 Table 4 : Restricted model: SO2 emission 1971-2005 \sim IPAT Model (Column 5)

By applying restrictions to the general model shown above, the IPAT model is estimated. Equation (6) shows the regression equation:

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \sum_{j=4}^k \beta_j X_{j,i} + \epsilon_i$$

The restrictions on the general model are: $\beta_1 = \beta_2 = \beta_3 = 0$

```
[20]: # Table 4 SO2 1971-2005 Regression Results
      # IPAT
      import numpy as np
      import matplotlib.pyplot as plt
      import pandas as pd
      import statsmodels.api as sm
      import statsmodels.formula.api as smf
      from stargazer.stargazer import Stargazer
      from IPython.core.display import HTML
      from statsmodels.stats.diagnostic import het_breuschpagan, het_white
      from sklearn.linear_model import LinearRegression
      from scipy.stats import chi2
      df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx',sheet_name='SO2 1971-2005')
      reg5 = smf.ols("GREPC ~ GRGDPPC + CPE + legorfr + legorge + legorsc + DMSUMT +_
      →DMWINT + DMLNFFPC1971 + DMLNPOPD1971",data=df).fit(cov type='HCO')
      print(reg5.summary())
      # White test for heteroskedasticity:
      reg5_exog = reg5.model.exog
      white_test = het_white(reg5.resid, reg5_exog)
      print(f"White test statistics: {white_test[0]}")
      print(f"P_value: {white_test[1]}")
      print(f"F stat: {white_test[2]}")
```

```
print(f"p_value: {white_test[3]}")
if white_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of \Box
⇔heteroskedasticity.")
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of_{\sqcup}
→no heteroskedasticity.")
# BP test for heteroskedasticity.
bp test = het breuschpagan(reg5.resid, reg5 exog)
print(f"bp_test_stat: {bp_test[0]}")
print(f"p_value: {bp_test[1]}")
if bp_test[1] < 0.05:</pre>
    print("Reject the null hypothesis of homoskedasticity. Evidence of \Box
→heteroskedasticity.")
else:
    print("Fail to accept the null hypothesis of homoskedasticity. Evidence of ⊔
→no heteroskedasticity.")
# Harvey-test - Has to be manually calculated
# calculate the log of squared residuals and store it in a column
df['log_sq_resid'] = np.log(reg5.resid ** 2)
# Regress the log od squared residuals on the independent variables
harvey_reg = smf.ols("log_sq_resid ~ GRGDPPC + CPE + legorfr + legorge + L
→legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + DMLNPOPD1971",data=df).

→fit(cov_type='HCO')
# Obtain the R^2 from the harvey regression
rsquared_harvey = harvey_reg.rsquared
# Obtain the no of obs. of the data
n=len(df)
# Harvey test stat
harvey_stat = n * rsquared_harvey
# Calculate the p_value of the harvey test - it follows a chi-sq distribution
p_value_harvey = chi2.sf(harvey_stat, df = 9) # Degree of freedom is the no of
→ independent variables in the harvey regression.
print(f"Harvey test: {harvey stat}")
print(f"p_value: {p_value_harvey}")
if p_value_harvey < 0.05:</pre>
    print('Reject the null hypothesis of homoskedasticity. Evidence of ⊔
→heteroskedasticity.')
```

```
else:
    print("Fail to reject the null hypothesis of homoskedasticity. Evidence of
 →no heteroskedasticity.")
# LR test : Comparing the restricted model with the unrestricted model
# Unrestricted model (General Model)
reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE,
 →+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 + L
 →DMLNPOPD1971", data=df).fit(cov_type='HCO')
# collect the llf stat from both the models
llf unrestricted = reg1.llf
llf_restricted = reg5.llf
# Calculate the lr stat:
lr_stat= 2 * (llf_unrestricted - llf_restricted)
# Calculate the p_value:
df_diff = llf_unrestricted - llf_restricted
p_value_lr = chi2.sf(lr_stat, df_diff)
print(f"LR statistics: {lr_stat}")
print(f"p_value: {p_value_lr}")
if p_value_lr < 0.05:</pre>
    print("Reject the null hypothesis of restricted model provides a good fit:
 →The unrestricted model fits significantly better.")
else:
    print("Fail to reject the null hypothesis of restricted model provides a
 →good fit: No significan timprovement in fit for the unrestricted model.")
                           OLS Regression Results
Dep. Variable:
                               GREPC
                                       R-squared:
                                                                       0.479
Model:
                                 OLS Adj. R-squared:
                                                                       0.426
                      Least Squares F-statistic:
Method:
                                                                      10.38
                                                                7.70e-11
Date:
                  Wed, 26 Jun 2024 Prob (F-statistic):
Time:
                            12:22:36 Log-Likelihood:
                                                                      191.15
No. Observations:
                                 100 AIC:
                                                                      -362.3
Df Residuals:
                                  90
                                      BIC:
                                                                      -336.2
Df Model:
                                   9
```

coef std err z P>|z| [0.025]

HC0

Covariance Type:

Intercept

-0.0106

GRGDPPC	1.2176	0.262	4.644	0.000	0.704	1.731
CPE	0.0211	0.013	1.634	0.102	-0.004	0.046
legorfr	-0.0231	0.010	-2.319	0.020	-0.043	-0.004
legorge	-0.0361	0.013	-2.685	0.007	-0.062	-0.010
legorsc	-0.0555	0.022	-2.518	0.012	-0.099	-0.012
DMSUMT	0.0045	0.001	4.782	0.000	0.003	0.006
DMWINT	-0.0005	0.001	-0.956	0.339	-0.002	0.001
DMLNFFPC1971	-0.0026	0.001	-2.113	0.035	-0.005	-0.000
DMLNPOPD1971	-0.0153	0.004	-3.556	0.000	-0.024	-0.007
===========		=========	=======			======
Omnibus:		41.410	Durbin-	Watson:		2.036
<pre>Prob(Omnibus):</pre>		0.000	Jarque-	Bera (JB):		137.759
Skew:		1.374	Prob(JE	3):		1.22e-30
Kurtosis:		8.050	Cond. N	o.		748.
==========			=======			=======

[1] Standard Errors are heteroscedasticity robust (HCO)

White test statistics: 61.3182673061752

P_value: 0.03448818771142843 F stat: 2.0644459944892377 p_value: 0.005537988848957214

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

bp_test_stat: 18.950183512066154
p_value: 0.025619821158774553

Reject the null hypothesis of homoskedasticity. Evidence of heteroskedasticity.

Harvey test: 6.63112288718729 p_value: 0.67546299669092

Fail to reject the null hypothesis of homoskedasticity. Evidence of no

heteroskedasticity.

LR statistics: 71.24415196763232 p_value: 0.0003622576264703981

Reject the null hypothesis of restricted model provides a good fit: The

unrestricted model fits significantly better.

1.14 Table 2 : General model (SO2 emission)

The regression shown below is of the general model for SO2 emission as in regression equation 2 in the original paper.

$$\hat{E}_i = \alpha_0 + \alpha_1 \hat{G}_i + \beta_1 \hat{G}_i G_{i,0} + \beta_2 E_{i,0} + \beta_3 G_{i,0} + \sum_{j=4}^k \beta_j X_{j,i} + \epsilon_i$$

where hats indicate long-run growth rates, \$ E\$ is the log of emissions per-capita, and G is the log of GDP per capita. X_i is a vector of explanatory and control variables. The sample mean has been deducted from each continuous level variable in this vector.

The third term, $\hat{G}_I G_{i,0}$, is the interaction between the rate of economic growth and the initial level of log income per capita. The EKC turning point is calculated with the assumption that $\alpha_1 > 0$ and $\beta_1 < 0$ when $\frac{\partial \hat{E}}{\partial \hat{G}} = 0$ and $\tau = exp\left(-\frac{\alpha_1}{\beta_1} + \mu_G\right)$, where μ_G is the cross-country mean of initial GDP per capita variable prior to the estimation

The following variables are contained in $X_{j,i}$: Dummy variable for countries that are centrally planned or market economy, dummies for countries that have legal origins in the UK, France, Germany or any Scandinavian countries, average summer and winter temperatures, de-meaned log of fossil fuel consumption in 1971 and de-meaned log of population density in 1971.

```
[4]: import numpy as np
     import matplotlib.pyplot as plt
     import pandas as pd
     import statsmodels.api as sm
     import statsmodels.formula.api as smf
     from stargazer.stargazer import Stargazer
     from IPython.core.display import HTML
     from statsmodels.stats.diagnostic import het_breuschpagan, het_white
     from sklearn.linear_model import LinearRegression
     from scipy.stats import chi2
     df = pd.read_excel(r'D:\Stern_etal2017EDE.xlsx', sheet_name='S02 1971-2005')
     # Run OLS regression
     reg1 = smf.ols("GREPC ~ GRGDPPC + DMMULGDP + DMLNGDPPC1971 + DMLNEPC1971 + CPE,
      \hookrightarrow+ legorfr + legorge + legorsc + DMSUMT + DMWINT + DMLNFFPC1971 +_{\sqcup}
      →DMLNPOPD1971", data=df).fit(cov_type='HC3')
     print(reg1.summary())
     #Different Types of Robust Standard Errors
     #'HCO': The original White (1980) heteroskedasticity-consistent standard errors.
     #'HC1': A small sample correction to HCO.
     #'HC2': Another small sample correction, which is more reliable when the sample
      \rightarrow size is small.
     #'HC3': MacKinnon and White's (1985) heteroskedasticity-consistent standard
      →errors, which is more robust in small samples.
     #print(reg1.resid())
     #Create a matrix of independent varibales which can be used for the residual_
      \rightarrow tests
     exog_het = reg1.model.exog
     # Perform the Breusch-Pagan Test for Heteroskedasticity. Note exog_het is the_
     → matrix of explanatory (independent) variables
     bp_test = het_breuschpagan(reg1.resid,exog_het)
```

```
print("Breusch-Pagan Test for Heteroskedasticity:") # HO: Variance of the
→ residuals is constant (homoskedastic)
print (f"LM Statistics: {bp_test[0]}") # The Lagrange Multiplier (LM) test
→ follows a chi-sq distribution and the degree of freedom is the number of ⊔
→independent variables. The LM statistic is calculated based on the residuals ⊔
→ from the regression model and the independent variables
print (f"LM Test p-value: {bp_test[1]}") # If p-value is less than 0.05 (5%
→ level) then we reject the null and vice-versa
#print (f"F-Statistics: {bp_test[2]}") # F-stat is the ratio of the mean_
→ squared error of the model and the mean squared error of the residual.
#print (f"F-Statistics p-value: {bp_test[3]}")
# Performing the White test for heteroskedasticity
white_test = het_white(reg1.resid, exog_het) # HO: Homoskedasticity
# Extracting the test statistic and p-value
white_stat = white_test[0]
white_pvalue = white_test[1]
print(f"White test statistic: {white stat}")
print(f"White test p-value: {white_pvalue}")
#print("White Test for Heteroskedasticity:") # Another way of writing the above∟
→command of white test and p value
#print(f"T Statistics: {white_test[0]}")
#print(f"P-value: {white_test[1]}")
#print(f"F-statistics: {white test[2]}")
#print(f"F-test p-value: {white_test[3]}")
# Interpretation
if white_pvalue < 0.05: # Typically, we use a significance level of 0.05
   print("Reject the null hypothesis of homoskedasticity. Evidence of ⊔
⇔heteroskedasticity.")
else:
   print("Fail to reject the null hypothesis of homoskedasticity. No evidence⊔
→of heteroskedasticity.")
r1s= reg1.resid ** 2 # This command enables python to store the squared of □
\rightarrow residuals in r1s
# Assuming df is your DataFrame and PWTPOP1971 is the column you want to invert
df['invpop'] = 1 / df['PWTPOP1971']
df['lpop1971'] = np.log (df['PWTPOP1971']) # To create a new varible with log
df['rls']=reg1.resid**2 # stored the residual square in the data file.
```

Dep. Variable: Model: Method: Date: Time: No. Observations Df Residuals: Df Model: Covariance Type:	Thu,	OLS Least Squares Thu, 27 Jun 2024 11:04:57 100 87 12 HC3		R-squared: Adj. R-squared: F-statistic: Prob (F-statistic): Log-Likelihood: AIC: BIC:		0.744 0.709 16.45 1.28e-17 226.77 -427.5 -393.7	
0.975]	coef	std err	Z	P> z	[0.025		
Intercept 0.003	-0.0099	0.007	-1.500	0.134	-0.023		
GRGDPPC 1.221	0.8543	0.187	4.563	0.000	0.487		
DMMULGDP 0.038	-0.2668	0.156	-1.715	0.086	-0.572		
DMLNGDPPC1971 0.032	0.0197	0.007	3.026	0.002	0.007		
DMLNEPC1971 -0.011	-0.0209	0.005	-3.963	0.000	-0.031		
CPE 0.065	0.0292	0.018	1.610	0.107	-0.006		
legorfr -0.002	-0.0151	0.007	-2.318	0.020	-0.028		

==========	========			========	========	======
Omnibus: Prob(Omnibus): Skew: Kurtosis:		6.532 0.038 0.531 3.548	Durbin-W Jarque-B Prob(JB) Cond. No	era (JB): :		1.745 5.950 0.0511 893.
==========						======
0.001 DMLNPOPD1971 -0.004	-0.0093	0.003	-3.293	0.001	-0.015	
-0.000 DMLNFFPC1971	-0.0015	0.001	-1.358	0.174	-0.004	
0.005 DMWINT	-0.0010	0.000	-2.188	0.029	-0.002	
-0.004 DMSUMT	0.0028	0.001	3.035	0.002	0.001	
-0.010 legorsc	-0.0446	0.021	-2.177	0.029	-0.085	
legorge	-0.0356	0.013	-2.744	0.006	-0.061	

_ _ . .

Notes:

[1] Standard Errors are heteroscedasticity robust (HC3)

Breusch-Pagan Test for Heteroskedasticity:

LM Statistics: 21.556851039733115 LM Test p-value: 0.042797192474353325 White test statistic: 86.7873337082127 White test p-value: 0.07264959660445151

Fail to reject the null hypothesis of homoskedasticity. No evidence of

heteroskedasticity.

EKC Truning point:100732.05913166571

1.15 Breusch Pagan, Harvey test and LR test as in Stern's code:

The method adopted by Stern et al. to estimate the Breusch-Pagan and Harvey tests for heteroskedasticity for all the estimations is not reproducible using their codes. These are the following codes used by the authors:

 ${\rm set~invpop} = 1/{\rm PWTPOP1971}$

set lpop1971 = log(PWTPOP1971)

source regwhitetest.src

@regwhitetest(type=full)

set r1s = r1**2

LINREG(noprint) r1s

#Constant INVPOP

cdf(title="Breusch-Pagan Test for INVPOP") chisqr %trsquared 1

legorge -0.010	-0.0356	0.013	-2.744	0.006	-0.061	
legorsc	-0.0446	0.021	-2.177	0.029	-0.085	
-0.004						
DMSUMT	0.0028	0.001	3.035	0.002	0.001	
0.005						
DMWINT	-0.0010	0.000	-2.188	0.029	-0.002	
-0.000						
DMLNFFPC1971	-0.0015	0.001	-1.358	0.174	-0.004	
0.001						
DMLNPOPD1971	-0.0093	0.003	-3.293	0.001	-0.015	
-0.004						
===========		.=======		.=======		
Omnibus:		6.532	Durbin-Wats	son:		1.745
Prob(Omnibus):		0.038	Jarque-Bera	(JB):		5.950
Skew:		0.531	Prob(JB):		C	.0511
Kurtosis:		3.548	Cond. No.			893.

[1] Standard Errors are heteroscedasticity robust (HC3)

Breusch-Pagan Test for Heteroskedasticity:

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