

NAME: SUCHETA JHUNJHUNWALA

ROLL No.: MDS202151

PASWR FINAL EXAM

1) At first we are defining a function ~~for~~ called 'fun' which takes four arguments as parameters n , μ (default value = 0), σ (default value = 1) and α (default value = 0.95).

Inside the function we calculate the value of z s.t.

$$P(Z > z) = \frac{1 - \alpha}{2} \text{ by defining no variable}$$

$$z = \text{qnorm}((1 - \alpha)/2, \text{lower.tail} = \text{FALSE})$$

qnorm is the quantile function for normal distribution which gives n s.t. $P(X > n) = p$ where p is the parameter given to qnorm . $\text{lower.tail} = \text{FALSE}$ enables $P(X > n)$ instead of $P(X \leq n)$.

$$\text{sd } n = \sigma / \sqrt{\text{length}(n)}$$

This variable calculates the value of $\frac{\sigma}{\sqrt{\text{length}(n)}}$ where

$\text{length}(n)$ is the length of the vector n given to fun .

$$\text{Then, } N = \text{pnorm}((\text{mean}(n) - \mu) / \text{sd } n, \text{lower.tail} = \text{FALSE})$$

calculate the probability of a normal variable $X \sim \text{Normal}(0, 1)$

$$P(X > (\text{mean}(n) - \mu) / \text{sd } n) \text{ where } \text{mean}(n) \text{ is the}$$

mean of the vectors n given to fun . $\text{lower.tail} = \text{FALSE}$

is to calculate $P(X > n)$ instead of $P(X \leq n)$.

At last, the function `fun` returns a vector of length 3 which has three values. The first value is the lower bound of the 95% confidence interval for the expectation of X , second value is the upper bound for the 95% confidence interval and `N` returns the probability.

Then we apply `fun` on the given vector n . Note that, we take $y = \text{fun}(n, 76, 1.5)$ so our $\mu = 76$ and $\sigma = 1.5$, though α is still 0.95. y is a vector which has the lower and upper bound of the 95% confidence interval for expectation corresponding the scores given to it.

$$z \text{ is the value s.t. } P(Z > z) = \frac{1 - \alpha}{2} = 0.025$$

where $Z \sim \text{Normal}(0, 1)$

$$\text{sd } n = \frac{\sigma}{\sqrt{\text{length}(n)}} = \frac{1.5}{\sqrt{9}} \text{ since } n \text{ has 9 elements}$$

`N` gives the probability that given $Z \sim \text{Normal}(0, 1)$,

$$P(Z > \frac{\text{mean}(n) - 76}{1.5/\sqrt{9}})$$

At last we return, $(\text{mean}(n) - z * \frac{1.5}{\sqrt{9}}, \text{mean}(n) + z * \frac{1.5}{\sqrt{9}}, N)$

which has values $(73.9039069, 75.8688709, 0.9868659)$

so the 95% confidence interval for the expected values of scores when the scores (say X) follow Normal distribution s.t. $X \sim \text{Normal}(76, (1.5)^2)$ is

(73.9089069, 75.8688709)

N returns $P\left(Z > \frac{\text{mean}(n) - 76}{1.5/\sqrt{n}}\right)$

that 0.9868659 is the probability of getting a

Score more than $\frac{\text{mean}(n) - 76}{1.5/\sqrt{n}}$.

2) x_1, \dots, x_n iid $x \sim \text{Rayleigh}(\alpha)$

$$f(x) = \begin{cases} \alpha x \exp\left(-\frac{1}{2}\alpha x^2\right) & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

The parameter here is α . The maximum likelihood function is given by,

$$L(\alpha | x_1, x_2, \dots, x_n) = \prod_{i=1}^n \alpha x_i e^{-\frac{1}{2}\alpha x_i^2} \quad \text{when } x_i \geq 0$$

$\forall i=1, 2, \dots, n$

Also note that, if $x_i < 0$ then,

$$L(\alpha | x_1, \dots, x_n) = 0$$

$$\text{Now, } L(\alpha | x_1, \dots, x_n) = \alpha^n \prod_{i=1}^n x_i e^{-\frac{1}{2}\alpha x_i^2}$$

$$= \alpha^n \prod_{i=1}^n x_i \cdot \prod_{i=1}^n e^{-\frac{1}{2}\alpha x_i^2}$$

$$= \alpha^n \left(\prod_{i=1}^n x_i \right) e^{-\frac{1}{2}\alpha \sum_{i=1}^n x_i^2}$$

On taking logarithm on both sides we get,

$$\log L(\alpha) = n \log \alpha + \log \left(\prod_{i=1}^n x_i \right) + \left(-\frac{1}{2}\alpha \sum_{i=1}^n x_i^2 \right) \log e$$

$$= n \log \alpha + \sum_{i=1}^n (\log x_i) - \frac{1}{2}\alpha \sum_{i=1}^n x_i^2$$

To find the maximum likelihood estimate we need to maximize $L(\alpha)$ that is we need to maximize $\log L(\alpha)$.

In order to do so, we need to find its critical points. On taking derivative w.r.t. α on both sides we get,

$$(\log L)'(\alpha) = \frac{n}{\alpha} + 0 - \frac{1}{\alpha} \cdot \frac{1}{2} \sum_{i=1}^n x_i^2$$

(Note that $\log L$ is a function of α, x_1, \dots, x_n so here by derivative we mean partial derivative w.r.t. α)

Now any pt. $\hat{\alpha}$ is a critical point if $(\log L)'(\hat{\alpha}) = 0$

$$\Rightarrow \frac{n}{\hat{\alpha}} - \frac{1}{2} \sum_{i=1}^n x_i^2 = 0$$

$$\Rightarrow \frac{n}{\hat{\alpha}} = \frac{1}{2} \sum_{i=1}^n x_i^2 \Rightarrow \hat{\alpha} = \frac{2n}{\sum_{i=1}^n x_i^2}$$

We shall perform the second derivative test to check if $\hat{\alpha}$ is a point of maximum or not.

$$(\log L)''(\alpha) = -\frac{n}{\alpha^2} + 0$$

$$(\log L)''(\hat{\alpha}) = -\frac{n}{\left(\frac{2n}{\sum_{i=1}^n x_i^2}\right)^2} = -n \cdot \left(\frac{\sum_{i=1}^n x_i^2}{2n}\right)^2$$

$$\left(\frac{\sum_{i=1}^n x_i^2}{2n}\right)^2 \geq 0 \quad \forall n \text{ and } n > 0$$

$$\therefore (\text{Jeg } D''(\hat{\alpha})) < 0.$$

$\Rightarrow \hat{\alpha}$ is the point of maximum for L .

\Rightarrow Maximum likelihood estimate of d is $\frac{2n}{\sum_{i=1}^n X_i^0 + 2}$

4) Brown wrinkled seeds $\rightarrow 219$

Brown round seeds $\rightarrow 81$

White wrinkled seeds $\rightarrow 69$

White round seeds $\rightarrow 31$

Ratios should be $9:3:3:1$.

i.e. $\left(\frac{9}{16} \times 100\right)\%$ should be brown wrinkled, $\left(\frac{3}{16} \times 100\right)\%$

should be brown round, $\left(\frac{3}{16} \times 100\right)\%$ should be white wrinkled and $\left(\frac{1}{16} \times 100\right)\%$ should be white round.

The total number of samples is $n = 400$.

This model is for ~~checking~~ ^{is} the χ^2 -test ^{for} goodness of fit.

which is given by
$$\chi^2 = \sum_{j=1}^k \frac{(y_j^0 - np_j^0)^2}{np_j^0}$$

where y_j^0 is the number of data points and p_j^0 is

$P(X = y_j^0) \cdot y_j^0 = \# \{ \omega : X(\omega) = y_j^0 \}$.

> n = c(219, 81, 69, 31)

Here n represents our y_i^0 which is the number of data points corresponding to the range of X i.e.

{ brown wrinkled, brown round, white wrinkled, white round }

n is the number of samples taken so

> n = 400

This is χ^2 distribution with $(k-1)$ degrees of freedom.

Here since the range of X has 4 ~~rows~~ elements,

$k = 4$ so we get,

> T = pchisq(Xsquare, df = 4 - 1, lower.tail = FALSE)

∴ the final code becomes,

> x = c(219, 81, 69, 31)

+ n = 400

+ prob = c(9/16, 3/16, 3/16, 1/16)

+ Xsquare = sum((x - n * prob)^2) / n * prob))

> Xsquare

> T = pchisq(Xsquare, df = 4 - 1, lower.tail = FALSE)

> T

As output we get two values, Xsquare and T.

Xsquare is the value of the test statistic and it

comes out as 2.733333.

whereas T is the value of $P(W > \chi^2)$ where
 $W \sim \chi^2(4-1) = \chi^2(3)$ and χ^2 is the test statistic.

$$\chi^2 = \sum_{i=1}^n \frac{(y_i - np_i)^2}{np_i}$$

This probability comes out as 0.2549554

i.e. Such type of samples are observed for more than
approximately 25.49% times.

The null hypothesis is ~~that~~ the theory of predicting such
seeds in the ratio 9:3:3:1.

The level of significance is 0.05.

$$0.2549554 > 0.05$$

\therefore We can't reject the null hypothesis. Since the value
is much larger than the significance level, we can
say that there is no evidence to reject the null
hypothesis.