

# Econometrics I, ECO341A, Summer 2023

## Homework II (100 points)

Instructor: M.A. Rahman

Deadline: 4:00 pm, June 28, 2023.

**Please read the instructions carefully and follow them while writing answers.**

- *Solutions to homework should be written in A4 size loose sheets. If you are not comfortable writing on white sheets, please ask for biology paper in Tarun Book Store.*
  - *Questions should be answered in order as they appear in the homework. Every new question should begin in a new page. Please number all the pages of your homework solution.*
  - *Please leave a margin of one inch from top and one inch from left. Staple the sheets on the top-left.*
  - *For Matlab assignments (if any), please answer the questions. Please do not dump the codes as an answer, they should be in the appendix.*
  - *Please write your name and names of your group members on the first page of your answer script.*
- 

**1. (20 points).** Suppose  $Y \sim \text{Beta}(\alpha, \delta)$  such that the *pdf* is given by,

$$f(y|\alpha, \delta) = \frac{\Gamma(\alpha + \delta)}{\Gamma(\alpha)\Gamma(\delta)} y^{\alpha-1}(1-y)^{\delta-1}, \quad \alpha > 0, \delta > 0, 0 < y < 1.$$

Construct the log-likelihood for a sample of  $n$  observations and find the Fisher's Information matrix.

Hint: To simplify the derivations, please make use of the following notations:

$$\begin{aligned} \Gamma'(\alpha) &= \frac{\partial \Gamma(\alpha)}{\partial \alpha}, & \psi(\alpha) &= \frac{\Gamma'(\alpha)}{\Gamma(\alpha)}, & \psi'(\alpha) &= \frac{\partial \psi(\alpha)}{\partial \alpha} \\ \Gamma'(\delta) &= \frac{\partial \Gamma(\delta)}{\partial \delta}, & \psi(\delta) &= \frac{\Gamma'(\delta)}{\Gamma(\delta)}, & \psi'(\delta) &= \frac{\partial \psi(\delta)}{\partial \delta}, \\ \Gamma'_\alpha(\alpha + \delta) &= \frac{\partial \Gamma(\alpha + \delta)}{\partial \alpha}, & \psi_\alpha(\alpha + \delta) &= \frac{\Gamma'_\alpha(\alpha + \delta)}{\Gamma(\alpha + \delta)}, & \psi'_\alpha(\alpha + \delta) &= \frac{\partial \psi_\alpha(\alpha + \delta)}{\partial \alpha}, \\ \Gamma'_\delta(\alpha + \delta) &= \frac{\partial \Gamma(\alpha + \delta)}{\partial \delta}, & \psi_\delta(\alpha + \delta) &= \frac{\Gamma'_\delta(\alpha + \delta)}{\Gamma(\alpha + \delta)}, & \psi'_\delta(\alpha + \delta) &= \frac{\partial \psi_\delta(\alpha + \delta)}{\partial \delta}. \end{aligned}$$

**2. (10 points)** Consider the simple linear regression model  $y_i = \beta_1 + x_{2i}\beta_2 + \epsilon_i$ . Show that the matrix formulation of the OLS estimator  $\hat{\beta} = (X'X)^{-1}(X'y)$  yields the following,

$$\begin{aligned}\hat{\beta}_1 &= \bar{y} - \hat{\beta}_2\bar{x}, \\ \hat{\beta}_2 &= \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}.\end{aligned}$$

**3. (10+10 = 20 points).** Let  $X_1, X_2, \dots$  be i.i.d. random variable with exponential pdf  $f(x) = \frac{1}{\lambda} \exp(-x/\lambda)$  if  $x > 0$  (and  $= 0$  otherwise),  $\lambda > 0$ .

(a) Show that  $\frac{\sqrt{n}(\bar{X}_n - \lambda)}{\bar{X}_n} \xrightarrow{d} N(0, 1)$  as  $n \rightarrow \infty$ .

(b) Let  $\beta = 1/\lambda$ . Show that  $\sqrt{n}(\beta\bar{X}_n - 1) \xrightarrow{d} N(0, 1)$  as  $n \rightarrow \infty$ .

**4. (20 points)** Consider the linear regression model  $y = X\beta + \epsilon$ . We know the least squares estimator  $\hat{\beta} = (X'X)^{-1}(X'y)$  is BLUE, so let's prove it. Consider another linear and unbiased estimator,

$$\bar{\beta} = [(X'X)^{-1}X' + c]y,$$

where  $c$  is matrix of dimension  $k \times n$ . Show that  $\Sigma_{\bar{\beta}} - \Sigma_{\hat{\beta}}$  is a positive definite matrix, where  $\Sigma_{\eta}$  denotes the covariance matrix of the estimator  $\eta$ .

**5. (5+5+5+5+10=30 points)** Consider the data presented in the file 'ProdFuncData.xlsx'. The data are statewide observations on SIC33, the primary metals industry in the US. There are 3 variables in the data, **value added** (which is the output,  $Y$ ), **labor**, and **capital**. They were originally constructed by Hilderbrand and Liu (1957) and subsequently used by many authors.

(a) Estimate a Cobb-Douglas production function by regressing  $\log Y$  on an intercept,  $\log L$ , and  $\log K$ , where  $Y$ ,  $L$ , and  $K$  denote output, labor, and capital, respectively. Report the regression coefficients, standard errors, and t-statistics in a table. Also, report the covariance matrix of  $\hat{\beta}$  and the R-square.

(b) The hypothesis of constant returns to scale is often tested in studies of production. This hypothesis is equivalent to the restriction that the two coefficients of Cobb-Douglas production function sum to 1 i.e.,  $H_0 : \beta_L + \beta_K = 1$ , versus  $H_1 : \beta_L + \beta_K \neq 1$ . Test the hypothesis and report your inference at 95% confidence level.

(Hint:  $\text{Var}(\hat{\beta}_L + \hat{\beta}_K) = \text{Var}(\hat{\beta}_L) + \text{Var}(\hat{\beta}_K) + 2 * \text{cov}(\hat{\beta}_L, \hat{\beta}_K)$ .)

(c) A generalization of the Cobb-Douglas model is the *translog* model, which is

$$\log Y = \beta_1 + \beta_2 \log L + \beta_3 \log K + \beta_4 \left( \frac{1}{2} \log^2 L \right) + \beta_5 \left( \frac{1}{2} \log^2 K \right) + \beta_6 (\log L * \log K) + \epsilon, \quad (1)$$

Estimate the model and report the regression coefficients, standard errors, and t-statistics in a table. Also, report the covariance matrix of  $\hat{\beta}$  and the R-square.

- (d) Test the hypothesis that the Cobb-Douglas model is the appropriate model i.e., test  $H_0 : \beta_4 = \beta_5 = \beta_6 = 0$ , versus  $H_1 : H_0$  **is not true**. Report the F-statistic and draw your inference at 95% confidence level.
- (e) The hypothesis of constant returns to scale can also be tested from the translog model. In this case, the null hypothesis is  $H_0 : \beta_2 + \beta_3 = 1$  and  $\beta_4 + \beta_5 + 2\beta_6 = 0$ , versus  $H_1 : H_0$  **is not true**. Report your F-statistics and draw your inference at 95% confidence level.