

Anticipating stochastic integrals and its large deviations

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CONTENTS

Part 1	Anticipating integrals	1
1.1	Interpretation of exponential processes	2
1.2	Miscellaneous	3
Part 2	Large deviations theory	4
2.1	Friedlin-Wentzell theorem	5
2.2	Friedlin-Wentzell theorem for anticipating initial condition with extension of filtration	6
2.2.1	\tilde{W} is a Wiener process	6
2.2.2	Reformulating the problem	7
	Bibliography	10

PART 1

ANTICIPATING INTEGRALS

1.1 INTERPRETATION OF EXPONENTIAL PROCESSES

In classical Itô theory, there are three equivalent interpretations of exponential processes.

- i. renormalization
- ii. martingale
- iii. SDEs

This is not so in the two-sided stochastic integral theory.

1.2 MISCELLANEOUS

Example 2.1 *The process X given by $X_t = B_{\frac{1}{2}(t+T)} - B_t$ cannot be expressed as a Borel function of $B_T - B_\cdot$.*

PART 2

LARGE DEVIATIONS THEORY

2.1 FRIEDLIN-WENTZELL THEOREM

2.2 FRIEDLIN-WENTZELL THEOREM FOR ANTICIPATING INITIAL CONDITION WITH EXTENSION OF FILTRATION

Notation: In what follows, $T < \infty$ and $t \in [0, T]$.

Our aim is to formulate a large deviations principle for a stochastic differential equation with anticipating initial conditions. Consider a very simple case

$$X_t^\varepsilon = W_T + \sqrt{\varepsilon} \int_0^t \sigma(X_u^\varepsilon) dW_u, \quad t \in [0, T],$$

and σ satisfies the following:

- *bounded:* $|\sigma(x)| \leq M_\sigma$.
- *Lipschitz:* $|\sigma(x) - \sigma(y)| \leq L_\sigma |x - y|$.
- *linear growth:* $|\sigma(x)|^2 \leq G_\sigma (1 + |x|^2)$.

We shall look at the method of enlargement of filtration by [Itô1978]. We denote the enlarged filtration by $\tilde{\mathcal{F}}_t = \mathcal{F}_t \vee \sigma(W_T)$. For $t \in [0, T]$, define the process $A_t = \int_0^t \frac{W_T - W_u}{T - u} du$. Then $W_t = \tilde{W}_t + A_t$, where \tilde{W} is a Wiener process w.r.t. $\tilde{\mathcal{F}}$.

2.2.1 \tilde{W} is a Wiener process

We show that (\tilde{W}_t) is a $(\tilde{\mathcal{F}}_t)$ -martingale with quadratic variation t . Then by Lévy's Characterization of Wiener process, we obtain that \tilde{W} is a Wiener process.

First we prove two lemmas.

Lemma 2.1 *The σ -algebras $\mathcal{F}_s \vee \sigma(W_T)$ and $\mathcal{F}_s \vee \sigma(W_T - W_s)$ are the same.*

Proof. For any Borel set B , the set $\{W_T \in B\} = \{(W_T - W_t) + W_t \in B\}$.

TODO

Lemma 2.2 *For $0 \leq s \leq t \leq T$, we have*

$$\mathbb{E}(W_t - W_s | W_T - W_s) = \frac{t-s}{T-s} (W_T - W_s).$$

Proof. We partition the interval $[0, T]$ into $n = kn_0$ equal parts, where $n_0 = (\min\{s, t-s, T-t\})^{-1}$ and $k \in \mathbb{N}$. Let $n_s = s \frac{n}{T}$ and $n_t = t \frac{n}{T}$. That is, the partition is

$$P = \left\{ 0, \frac{T}{n}, \dots, \frac{n_s T}{n} = s, \dots, \frac{n_t T}{n} = t, \dots, \frac{(n-1)T}{n}, T \right\}.$$

Let $\Delta_i W = W_{\frac{(i+1)T}{n}} - W_{\frac{iT}{n}}$.

Firstly, note that the $\Delta_i W$ s are independent and identically distributed from the definition of Wiener process. Now, using the linearity of conditional expectation, we have

$$\begin{aligned}
\mathbb{E}(W_t - W_s | W_T - W_s) &= \mathbb{E}\left(\sum_{i=n_s}^{n_t-1} \Delta_i W \mid \sum_{i=n_s}^{n_t-1} \Delta_i W\right) \\
&= \sum_{i=n_s}^{n_t-1} \mathbb{E}\left(\Delta_i W \mid \sum_{i=n_s}^{n-1} \Delta_i W\right) \\
&= \sum_{i=n_s}^{n_t-1} \frac{1}{n - n_s} \sum_{i=n_s}^{n-1} \mathbb{E}\left(\Delta_i W \mid \sum_{i=n_s}^{n-1} \Delta_i W\right) \\
&= \sum_{i=n_s}^{n_t-1} \frac{1}{n - n_s} \mathbb{E}\left(\sum_{i=n_s}^{n-1} \Delta_i W \mid \sum_{i=n_s}^{n-1} \Delta_i W\right) \\
&= \sum_{i=n_s}^{n_t-1} \frac{1}{n - n_s} \sum_{i=n_s}^{n-1} \Delta_i W \\
&= \sum_{i=n_s}^{n_t-1} \frac{1}{n - n_s} (W_T - W_s) \\
&= \frac{n_t - n_s}{n - n_s} (W_T - W_s) = \frac{t - s}{T - s} (W_T - W_s).
\end{aligned}$$

Proposition 2.3 \tilde{W} is a $\tilde{\mathcal{F}}$ -martingale.

Proof. Let $0 \leq s \leq t \leq T$. Then

$$\tilde{W}_t - \tilde{W}_s = (W_t - W_s) - \int_s^t \frac{W_T - W_u}{T - u} du = (W_t - W_s) - \int_s^t \left(\frac{W_T - W_s}{T - u} - \frac{W_u - W_s}{T - u} \right) du.$$

Moreover, since $W_t - W_s$ is independent of \mathcal{F}_s for every $t \geq s$, using lemmas 2.1 and 2.2, we get

$$\mathbb{E}(W_t - W_s | \tilde{\mathcal{F}}_s) = \mathbb{E}(W_t - W_s | \mathcal{F}_s \vee \sigma(W_T - W_s)) = \mathbb{E}(W_t - W_s | W_T - W_s) = \frac{t - s}{T - s} (W_T - W_s).$$

Therefore, using the fact that W_T and W_s are $\tilde{\mathcal{F}}_s$ -measurable with conditional Fubini's theorem, we get

$$\begin{aligned}
\mathbb{E}(\tilde{W}_t - \tilde{W}_s | \tilde{\mathcal{F}}_s) &= \mathbb{E}(W_t - W_s | \tilde{\mathcal{F}}_s) - \int_s^t \left(\frac{W_T - W_s}{T - u} - \frac{\mathbb{E}(W_u - W_s | \tilde{\mathcal{F}}_s)}{T - u} \right) du \\
&= \frac{t - s}{T - s} (W_T - W_s) - \int_s^t \left(\frac{W_T - W_s}{T - u} - \frac{u - s}{T - s} \frac{W_T - W_s}{T - u} \right) du \\
&= \frac{t - s}{T - s} (W_T - W_s) - \int_s^t \frac{W_T - W_s}{T - s} du \\
&= \frac{t - s}{T - s} (W_T - W_s) - \frac{t - s}{T - s} (W_T - W_s) = 0.
\end{aligned}$$

Now, since \tilde{W}_s is $\tilde{\mathcal{F}}_s$ -measurable, $\mathbb{E}(\tilde{W}_t | \tilde{\mathcal{F}}_s) = \mathbb{E}(\tilde{W}_t - \tilde{W}_s | \tilde{\mathcal{F}}_s) + \tilde{W}_s = \tilde{W}_s$.

Proposition 2.4 The quadratic variation of \tilde{W}_t is t .

Proof. TODO

2.2.2 Reformulating the problem

For now, we shall bound $t \in [0, T_b]$, where $T_b \in [0, T)$. What happens when $t \rightarrow T$?

Using this, we write our original stochastic differential equation as

$$X_t^\varepsilon = W_T + \sqrt{\varepsilon} \int_0^t \sigma(X_s^\varepsilon) d\tilde{W}_s + \sqrt{\varepsilon} \int_0^t \sigma(X_s^\varepsilon) \frac{W_T - W_s}{T-s} ds.$$

Let $\tilde{X}_t^\varepsilon = X_t^\varepsilon - W_T$ and $Y_t^\varepsilon = \sqrt{\varepsilon}(W_T - W_t)$. Then we have

$$\begin{aligned} \tilde{X}_t^\varepsilon &= \sqrt{\varepsilon} \int_0^t \sigma\left(\tilde{X}_s^\varepsilon + \frac{Y_0^\varepsilon}{\sqrt{\varepsilon}}\right) d\tilde{W}_s + \int_0^t \sigma\left(\tilde{X}_s^\varepsilon + \frac{Y_0^\varepsilon}{\sqrt{\varepsilon}}\right) \frac{Y_s^\varepsilon}{T-s} ds, \text{ and} \\ Y_t^\varepsilon &= \sqrt{\varepsilon} W_T - \sqrt{\varepsilon} \int_0^t d\tilde{W}_s - \int_0^t \frac{Y_s^\varepsilon}{T-s} ds. \end{aligned}$$

So together we have the joint process

$$Z_t^\varepsilon := \begin{pmatrix} \tilde{X}_t^\varepsilon \\ Y_t^\varepsilon \end{pmatrix} = \begin{pmatrix} 0 \\ \sqrt{\varepsilon} W_T \end{pmatrix} + \sqrt{\varepsilon} \int_0^t \begin{pmatrix} \sigma\left(\tilde{X}_s^\varepsilon + \frac{Y_0^\varepsilon}{\sqrt{\varepsilon}}\right) \\ -1 \end{pmatrix} d\tilde{W}_s + \int_0^t \begin{pmatrix} \sigma\left(\tilde{X}_s^\varepsilon + \frac{Y_0^\varepsilon}{\sqrt{\varepsilon}}\right) \\ -1 \end{pmatrix} \frac{Y_s^\varepsilon}{T-s} ds.$$

Note that we expect $Z_t^\varepsilon \rightarrow 0$ as $\varepsilon \searrow 0$. We first obtain a large deviation principle for Z_t^ε .

Fix $R > 0$, and define the exit time from the R -ball centered at the origin as

$$\tau_R = \inf\{t : Y_t^\varepsilon > R\} \wedge T_b.$$

Clearly $\tau_R \nearrow T_b$ as $R \nearrow \infty$.

We now show that there is a unique solution of Z_t^ε on $t \in [0, \tau_R]$. For convenience, let

$$\tilde{\sigma}_{t,v}^\varepsilon(x,y) = \left(\sigma\left(x + \frac{v}{\sqrt{\varepsilon}}\right), -1 \right) \quad \text{and} \quad \tilde{b}_{t,v}^\varepsilon(x,y) = \frac{y}{T-s} \left(\sigma\left(x + \frac{v}{\sqrt{\varepsilon}}\right), -1 \right).$$

Lemma 2.5 *The process Z_t^ε exists and is unique for $t \in [0, \tau_R]$.*

Proof. Let $v = Y_0^\varepsilon$. We first show that $\tilde{\sigma}$ and \tilde{b} satisfy the linear growth and Lipschitz conditions locally.

- *Lipschitz condition for $\tilde{\sigma}$:* Since σ is Lipschitz, we have

$$\begin{aligned} \left\| \tilde{\sigma}_{t,v_2}^\varepsilon(x_2, y_2) - \tilde{\sigma}_{t,v_1}^\varepsilon(x_1, y_1) \right\| &= \left| \sigma\left(x_2 + \frac{v_2}{\sqrt{\varepsilon}}\right) - \sigma\left(x_1 + \frac{v_1}{\sqrt{\varepsilon}}\right) \right| \\ &\leq L_\sigma \left(|x_2 - x_1| + \frac{1}{\sqrt{\varepsilon}} |v_2 - v_1| \right) \\ &\leq L_\sigma \left(1 \vee \frac{1}{\sqrt{\varepsilon}} \right) (|x_2 - x_1| + |v_2 - v_1|). \end{aligned}$$

- *Lipschitz condition for \tilde{b} :* Using the boundedness of σ , we get

$$\begin{aligned} &\left\| \tilde{b}_{t,v_2}^\varepsilon(x_2, y_2) - \tilde{b}_{t,v_1}^\varepsilon(x_1, y_1) \right\| \\ &\leq \frac{1}{T-t} \left(\left| \sigma\left(x_2 + \frac{v_2}{\sqrt{\varepsilon}}\right) y_2 - \sigma\left(x_1 + \frac{v_1}{\sqrt{\varepsilon}}\right) y_1 \right| + |y_2 - y_1| \right) \\ &\leq \frac{1}{T-t} \left(\left| \sigma\left(x_2 + \frac{v_2}{\sqrt{\varepsilon}}\right) \right| |y_2 - y_1| + \left| \sigma\left(x_2 + \frac{v_2}{\sqrt{\varepsilon}}\right) - \sigma\left(x_1 + \frac{v_1}{\sqrt{\varepsilon}}\right) \right| |y_1| + |y_2 - y_1| \right) \\ &\leq \frac{1}{T-t} \left(M_\sigma |y_2 - y_1| + L_\sigma \left(1 \vee \frac{1}{\sqrt{\varepsilon}} \right) |y_1| (|x_2 - x_1| + |v_2 - v_1|) + |y_2 - y_1| \right) \\ &\leq \frac{1}{T-t} \left((M_\sigma + 1) \vee \left(L_\sigma \left(1 \vee \frac{1}{\sqrt{\varepsilon}} \right) R \right) \right) (|x_2 - x_1| + |y_2 - y_1| + |v_2 - v_1|), \end{aligned}$$

where $|y_1| \leq R$ since $t \in [0, \tau_R]$.

- *Linear growth condition for $\tilde{\sigma}$:*

$$\begin{aligned}
\|\tilde{\sigma}_{t,v}^\varepsilon(x,y)\|^2 &= 1 + \left| \sigma \left(x + \frac{v}{\sqrt{\varepsilon}} \right) \right|^2 \\
&\leq 1 + G_\sigma \left(1 + \left| x + \frac{v}{\sqrt{\varepsilon}} \right|^2 \right) \\
&\leq 1 + G_\sigma \left(1 + 2|x|^2 + 2\frac{|v|^2}{\varepsilon} \right) \\
&\leq 2G_\sigma (1 \vee \varepsilon^{-1}) (1 + |x|^2 + |v|^2).
\end{aligned}$$

◦ *Linear growth condition for \tilde{b} :*

$$\begin{aligned}
\|\tilde{b}_{t,v}^\varepsilon(x,y)\|^2 &= 1 + \left| \sigma \left(x + \frac{v}{\sqrt{\varepsilon}} \right) \frac{y}{T-t} \right|^2 \\
&\leq \frac{1}{T-t} 2G_\sigma (1 \vee \varepsilon^{-1}) R^2 (1 + |x|^2 + |v|^2).
\end{aligned}$$

The above implies that Z_t^ε exists and is unique for $t \in [0, \tau_R]$.

BIBLIOGRAPHY