

~~Two~~ Two dimensional random variables

Let 'S' be the sample space associated with a random experiment Ω^E . Let $X = X(S)$, $Y = Y(S)$ two functions each assigning a real number to each element $S \in S$. Then we call the pair X, Y as two dimensional random variables.

Ex: Height, weight of randomly selected person.

(X, Y) is two dimensional DRV if the range space of (X, Y) is finite or countably infinite or uncountable set.

Joint probability distribution:

Let (X, Y) be an 2-D DRV. With each possible outcome (x_i, y_j) be associate a number $P(x_i, y_j)$ representing probability at $P(X=x_i, Y=y_j)$ and satisfying the following condition.

$$(i) P(x_i, y_j) \geq 0$$

$$(ii) \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} P(x_i, y_j) = 1.$$

Then $P(x_i, y_j)$ is called joint probability mass function of (X, Y) .

Let (X, Y) be a CRV then its joint p.d.f $f(x, y)$ satisfy the following conditions,

$$(i) f(x_i, y_j) \geq 0$$

$$(ii) \iint_{-\infty}^{\infty} f(x, y) dx dy = 1.$$

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Marginal probability distribution of 'x':

$$\text{DRV} \quad P(x=a_i) = P(a_i) = P(x=a_i, y=y_1) + P(x=a_i, y=y_2) + \dots$$

$$\checkmark \text{Joint PMF} \quad P(a_i) = \sum_{j=1}^{\infty} P(a_i, y_j). \quad (\text{marginal PMF of } a)$$

~~x bno~~ Similarly, continuous probability distribution of

$$P(y=y_i) = q(y_i) = P(x=a_1, y=y_i) + P(x=a_2, y=y_i) + \dots \quad (\text{m})$$

$$q(y_i) = \sum_{i=1}^{\infty} P(a_i, y_i) \quad (\text{marginal PMF of } y).$$

x	y₁ y₂ y₃ y₄	P(a_i)
a₁	P₁₁ P₁₂ P₁₃ P₁₄	P(a₁)
a₂	P₂₁ P₂₂ P₂₃ P₂₄	P(a₂)
a₃	P₃₁ P₃₂ P₃₃ P₃₄	P(a₃)

$$a(y_i) | a(y_1) a(y_2) a(y_3) a(y_4) = 1$$

CRV

Let $f(x_i)$ be the joint PDF of two CRV x & y

then the marginal PDF of 'x' is given by

$$g(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$h(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

Conditioned probability distribution,

DRV

Let (x, y) be a discrete random variable with point PMF $p(x_i, y_j)$ and marginal PMF $P(x_i)$ and $q(y_j)$,

$$\rightarrow P(x_i | y_j) = P(x=x_i | y=y_j) = \frac{P(x_i, y_j)}{q(y_j)}$$

$$\rightarrow P(y_j | x_i) = \frac{P(x_i, y_j)}{P(x_i)}$$

CRV

Let (x, y) be a CRV with joint PDF $f(x, y)$ and marginal PDF $h(y)$ and $g(x)$ then

$$\rightarrow P(x_i | y) = P(x=x_i | y=y) = \frac{f(x_i, y)}{h(y)}$$

$$\rightarrow P(y_j | x) = P(y=y_j | x=x) = \frac{f(x, y_j)}{g(x)}$$

(Q1) A coin is tossed 3 times. Let X denotes 0 or 1 after
dmg as a tail or a head occurs on the first toss.
Let Y denotes the no. of tails. Determine

① Joint Probability distribution of X and Y

② Marginal probability distribution of X and Y

Ans)

$X \setminus Y$	0	1	2	3	$P(X,Y)$
0	0	$\frac{1}{8}$	$\frac{1}{4}$	$\frac{1}{8}$	$\frac{1}{2}$
1	$\frac{1}{8}$	$\frac{1}{4}$	$\frac{1}{8}$	0	$\frac{1}{2}$
	$\frac{1}{8}$	$\frac{3}{8}$	$\frac{3}{8}$	$\frac{1}{8}$	1

(Q2) Suppose that a 2 dimensional CRV (X, Y) has
Joint PDF $f(x,y) = kx(a-y)$, $0 < x < 2$, $-a < y < a$

0 elsewhere.

① Find k

② Find marginal pdf of X .

Ans)

$$= \iint_{-\infty}^{\infty} f(x,y) dx dy = 1$$

$$= \int_{-a}^{a} \int_{0}^{2} kx(a-y) dy dx = 1$$

$$= k \int_{-a}^{a} \int_{0}^{2} (a^2 - ax) dy dx = 1$$

$$= ak \int_{0}^{2} \int_{-a}^{a} x^2 dy dx = 1$$

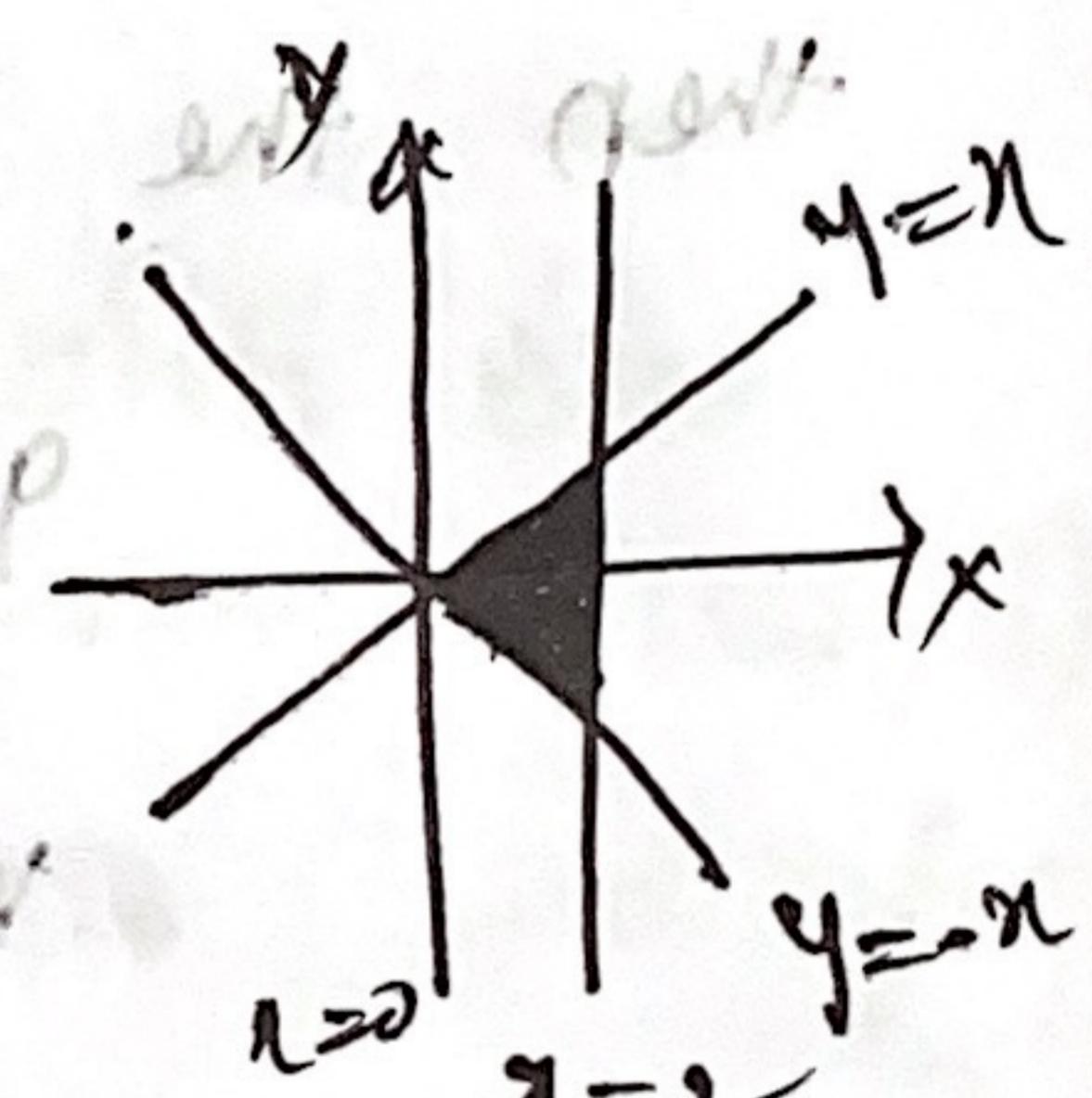
$$= ak \int_{0}^{2} [y]_{-a}^{a} dx = 1$$

$$= 2ak \int_{0}^{2} x^2 dx = 1$$

$$= 2ak \cdot \frac{24}{9} \Big|_0^2 = 1$$

$$\Rightarrow 8ak^2 = (k+1)x = 1 = (1)^2$$

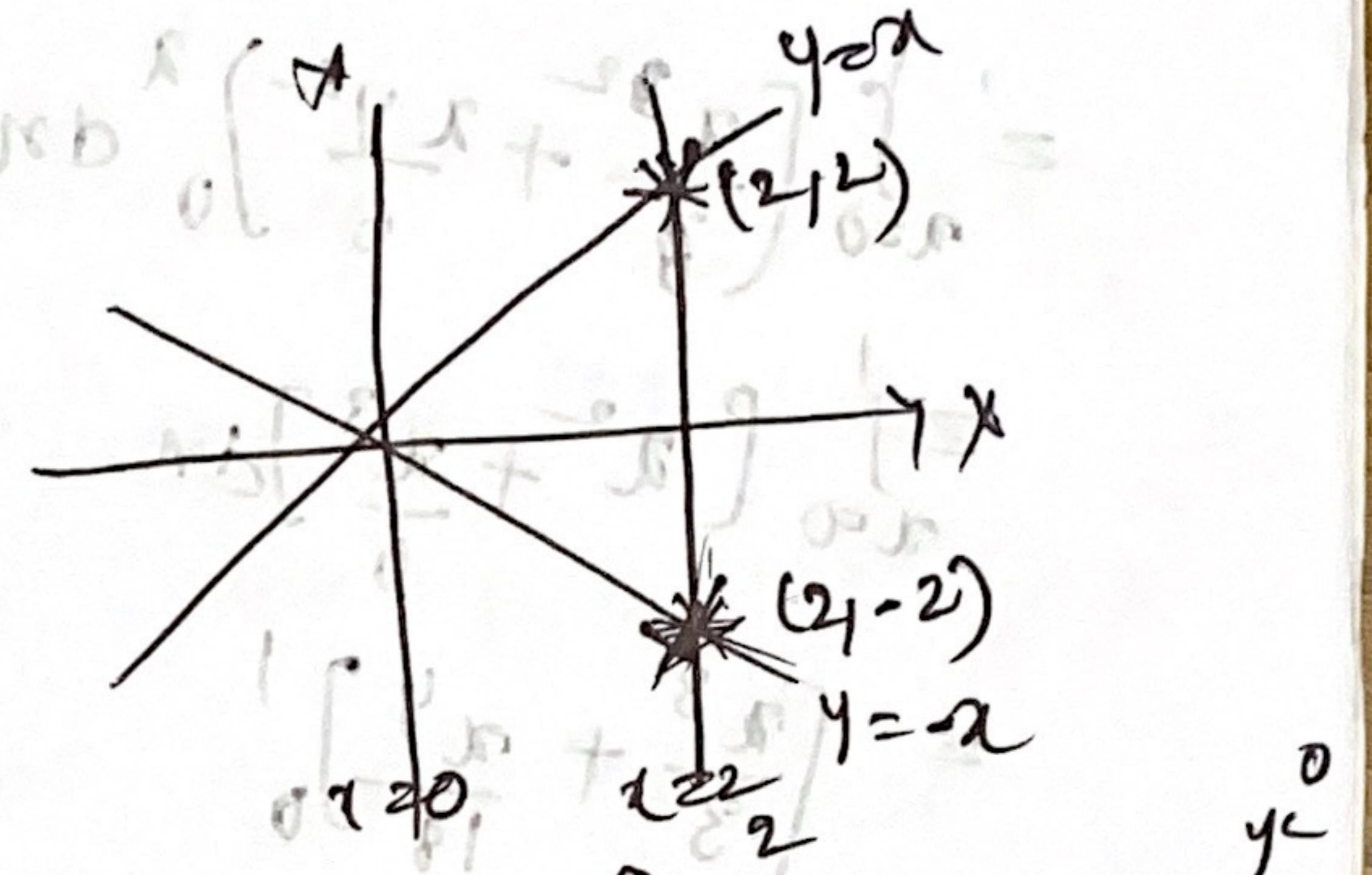
$$\Rightarrow k^2 = 1/8 = (k+1)^2 = 1/9 = (1)^2$$



$$\begin{aligned}
 Q) f(x) &= \int_{-\infty}^{\infty} f(x,y) dy \\
 &= \frac{1}{8} \int_{-x}^{x} (x(y-x)) dy, \quad 0 < x < 2 \\
 &= \frac{1}{8} \int_{-x}^{x} x^2 - xy dy \\
 &= \frac{1}{8} [x^3 - \frac{x^2 y}{2}] \\
 &= x^3 / 4
 \end{aligned}$$

marginal pdt of y

$$h(y) = \int_{-\infty}^{\infty} f(x,y) dx = h(y) = \int_{-\infty}^{\infty} \frac{1}{8} \int_{-y}^{y} (x(x-y)) dx dy$$



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 ① 1st order boundary condition $\Rightarrow y'$
 ② Laplace, own heat
 ③ Paretian total probability, random variable
 $x_1 = 0, x_2 = 0, x_3 = 1$
 $y_1 = ?, y_2 = ?$

h will be given

Q3) Suppose a point pdt of $x+y=0$ random variable is given by $f(x,y) = \begin{cases} x^2 + \frac{xy}{3}, & 0 < x < 1, 0 < y < 2 \\ 0, & \text{elsewhere} \end{cases}$.

Find: (i) $P(X > \frac{1}{2})$ (ii) $P(Y < X)$

(iii) $P(Y < 1/2 | X < 1/2)$.

$$(i) P(X > \frac{1}{2}) = P(\frac{1}{2} < x < \infty)$$

$$= \frac{1}{3} \int_{\frac{1}{2}}^{\infty} \int_0^2 (x^2 + \frac{xy}{3}) dy dx = P(1/2 < x < 1, 0 < y < 2)$$

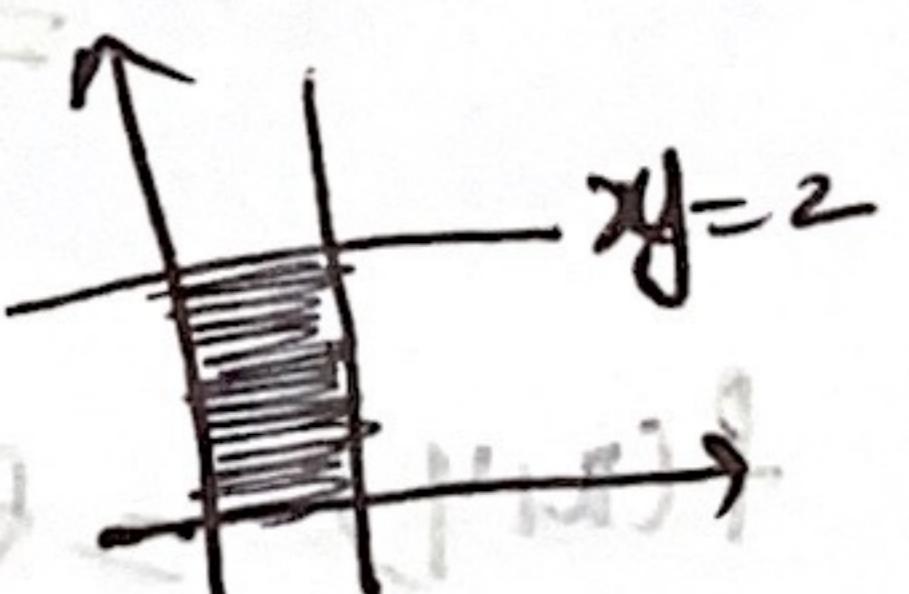
$$= \frac{1}{3} \int_{\frac{1}{2}}^1 (6x^2 + 2x) dx = \int_{\frac{1}{2}}^1 \int_{y=0}^{y=2} (x^2 + \frac{xy}{3}) dy dx$$

$$= \frac{1}{3} \left[\frac{6x^3}{3} + \frac{2x^2}{2} \right]_{\frac{1}{2}}^1 = \int_{\frac{1}{2}}^1 \int_{y=0}^{y=2} \left(x^2 + \frac{xy^2}{6} \right) dy dx$$

$$= \frac{1}{3} \left[x + 1 - \frac{1}{4} - \frac{1}{4} \right]_{\frac{1}{2}}^1 = \int_{\frac{1}{2}}^1 x + \frac{x^3}{6} dx$$

$$= \frac{1}{3} \left[3 - \frac{2}{4} \right]_{\frac{1}{2}}^1 = \int_{\frac{1}{2}}^1 \left(\frac{x^2}{2} + \frac{x^4}{6} \right) dx$$

$$= \frac{5}{6} = \frac{5}{6}$$



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$$(ii) P(Y < X) = P(-\infty < X < \infty)$$

$$= \int_{x=0}^1 \int_{y=0}^x (x^3 + \frac{xy}{3}) dx dy$$

$$= \int_{x=0}^1 \left[\frac{x^4}{4} + \frac{xy^2}{6} \right]_0^x dx$$

$$= \int_{x=0}^1 \left[x^2 + \frac{x^3}{6} \right] dx$$

$$= \left[\frac{x^3}{3} + \frac{x^4}{18} \right]_0^1$$

$$= \left[\frac{1}{3} + \frac{1}{18} \right] = \frac{7}{24}$$

$$(iii) P(Y < 1/2 | X < 1/2) = \frac{P(X < 1/2, Y < 1/2)}{P(X < 1/2)}$$

$$= \frac{\int_0^{1/2} \int_0^{1/2} (x^2 + \frac{xy}{3}) dy dx}{\int_0^{1/2} \int_0^{1/2} (x^2 + \frac{xy}{3}) dy dx}$$

$$= \frac{\int_0^{1/2} \int_0^{1/2} (x^2 + \frac{xy}{3}) dy dx}{\int_0^{1/2} (x^2 + \frac{x}{6}) dx}$$

$$= \frac{\frac{1}{94} + \frac{1}{96}}{\frac{1}{24} + \frac{1}{12}} = \frac{\left[\frac{x^3}{3} + \frac{x^2}{18} \right]_0^{1/2}}{\frac{1}{24} + \frac{1}{12}}$$

$$= \frac{\frac{1}{94} + \frac{1}{96}}{\frac{1}{24} + \frac{1}{12}} = \left[\frac{1}{24} + \frac{1}{12} \right] = \frac{1}{93}$$

$$= \frac{5}{3211}$$

$$(iv) f(x, y) = 6e^{-2x-3y}, x, y \geq 0.$$

$$\text{find (i) } P(0 < x < 2, y > 0)$$

(ii) Marginal pd. f. of x , $f_x(x)$, $f_y(y)$.

$$(i) = 6 \int_{x=0}^2 \int_{y=0}^{\infty} 6e^{-(2x+3y)} dy dx$$

$$= 6 \int_0^2 e^{-2x} \left[-\frac{1}{3} e^{-3y} \right]_0^\infty dx$$

$$= 2 \int_0^2 e^{-2x} dx = 1 - e^{-4}$$

$$\begin{aligned}
 q(x) &= \int_{-\infty}^{\infty} f(x,y) dy \\
 &= \int_0^{\infty} e^{-(2x+3y)} dy \\
 &= \left[\frac{e^{-(2x+3y)}}{-3} \right]_0^{\infty} \\
 &= 2e^{-2x}, x \geq 0.
 \end{aligned}$$

Wert für den Rest kann als σ ist (P(X) ist
 $h(y) = \int_{-\infty}^{\infty} f(x,y) dx$ für die Verteilung von X)

$$Wert für \mu_2 = \int_0^{\infty} e^{-(2x+3y)} dx \text{ ist } (P(X) ist)$$

$$(E(X^2) = \int_0^{\infty} x^2 e^{-(2x+3y)} dx + \text{Abbildung von P(X)}$$

$$= \int_0^{\infty} \left(\frac{e^{-(2x+3y)}}{-2} \right) dx$$

$$= -\frac{1}{2} e^{-3y}, y \geq 0.$$

* Q5) Suppose that CRV (2D) (X,Y) is uniformly distributed over the square with vertices (0,0), (1,0), (0,1), (1,1), giving marginal pdf of a $g(y), g(x), h(y)$.

$$\begin{aligned}
 \text{a) } p_{xy} &= \frac{1}{\text{Area of Region}}, x, y \in \mathbb{R} \\
 f(x,y) &= \frac{1}{2}, x, y \in \text{Square} \\
 q(x) &= \int_{-\infty}^{\infty} f(x,y) dy \\
 &= \int_{-1-x}^{1+x} \frac{1}{2} dy, -1 \leq x \leq 0 \\
 &= 1+x+1-x \\
 &= 2+2x \text{ for } x \in [-1, 0]
 \end{aligned}$$

$$\begin{aligned}
 q(x) &= \int_{-1}^x \frac{1}{2} dy, 0 \leq x \\
 &= 1-x-(-1) \\
 &= 2-2x
 \end{aligned}$$

$$h(y) = \int_{-\infty}^{\infty} f(x,y) dx,$$

$$\begin{aligned}
 &= \int_{-1-y}^{1+y} \frac{1}{2} dx, (-1-y \leq y \leq 0).
 \end{aligned}$$

$$= \int_0^1 \frac{1}{2} dx, 0 \leq y \leq 1.$$

Independent Random Variables:

Let (X_i) be a 2D RV then we say that x_i, y_i are independent if $P(x_i, y_i) = P(x_i) \cdot P(y_i)$

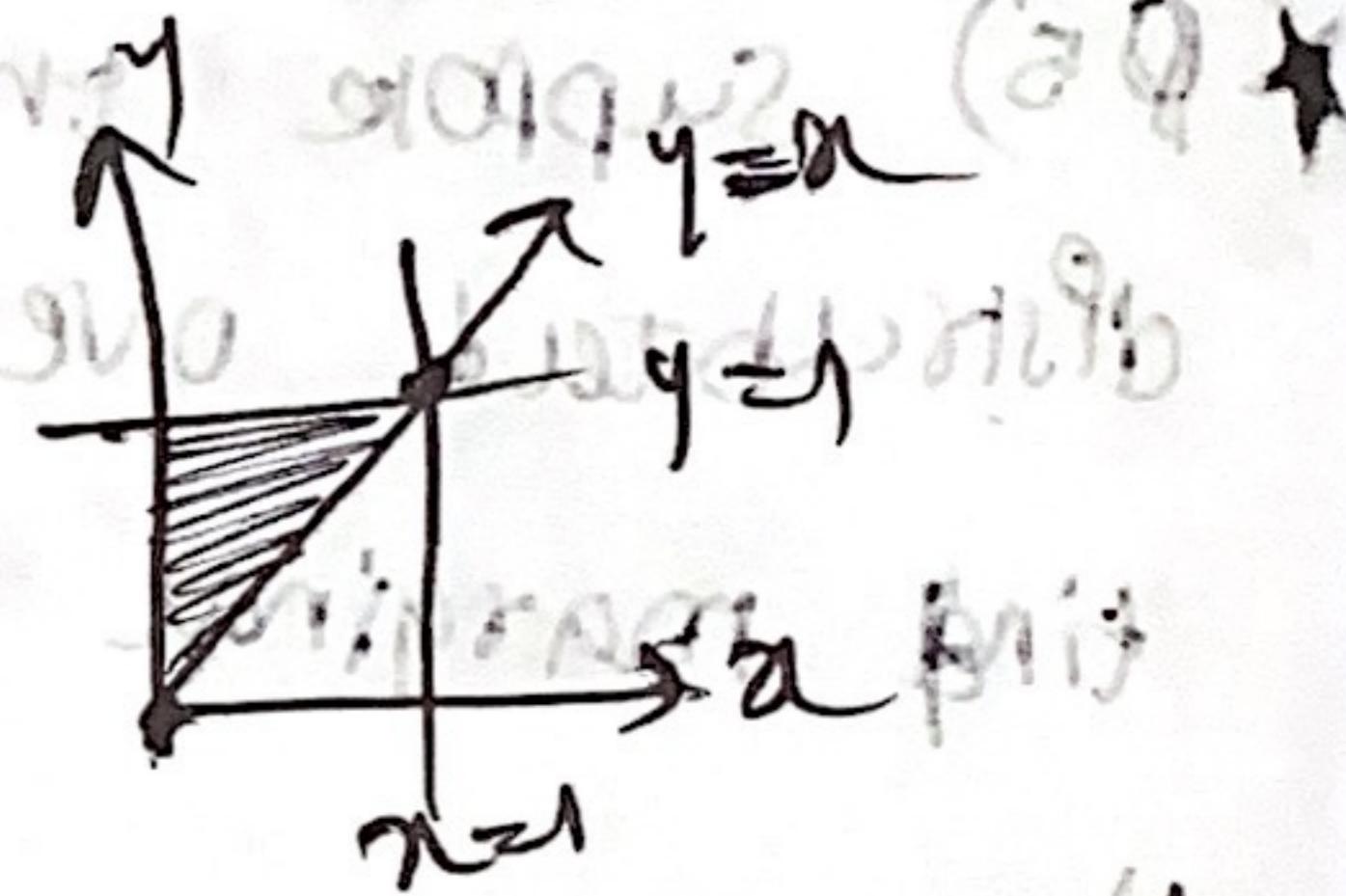
Let (X_i, Y_i) be a 2D CRV then we say that x_i, y_i are independent if $f(x_i, y_i) = g(x_i) h(y_i)$.

Ex. 2D RV (x, y) has joint pdf $f(x, y) = 8xy, 0 < x < 1$

i) Find marginal pdf of (x, y) $f(x, y) = \text{say}$

ii) Are they independent?

$$g(x) = \int_y 8xy dy, 0 < x < 1.$$



$$= \int_0^x 8xy dy = 4x^2, 0 < x < 1.$$

$$= 4x^2 - 4x^3, 0 < x < 1.$$

$$h(y) = \int_x 8xy dx, 0 < y < 1 = 4y - 4y^2, 0 < y < 1.$$

$$= (4y)^2 - (4y^2)^2, 0 < y < 1.$$

$$= 16y^2 - 16y^4, 0 < y < 1.$$

ii) $g(x) \cdot h(y) \neq f(x, y)$

x and y are not independent

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Expectation (Mean value): $E(X) = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} x_i y_j P(x_i, y_j)$

If x_i, y_j are DRV,

$$E(X) = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} x_i y_j P(x_i, y_j)$$

$$= \sum_{i=1}^{\infty} x_i P(x_i) \quad [\because \text{marginal Pmf}]$$

$$E(Y) = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} y_j P(x_i, y_j) = \sum_{j=1}^{\infty} y_j P(y_j)$$

If x_i, y_j are CRV,

$$E(X) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f(x, y) dy dx$$

$$E(Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f(x, y) dy dx$$

$$= \int_{-\infty}^{\infty} y h(y) dy$$

Properties:

$$(i) E(X+Y) = E(X) + E(Y)$$

(ii) If x_i, y_j are 2D independent RV with point Pmf $f(x_i, y_j)$

then, $E(XY) = E(X) \cdot E(Y)$.

Variance:

If x_i, y_j are 2D-RV & x_i, y_j are independent then,

$$V(X+Y) = V(X) + V(Y).$$

$$V(X+Y) = E(X+Y)^2 - [E(X+Y)]^2$$

$$= E(X^2 + 2XY + Y^2) - [E(X) + E(Y)]^2$$

$$= E(X^2) + 2E(XY) + E(Y^2) - [E(X) + E(Y)]^2$$

$$= V(X) + V(Y)$$

$E(XY) = E(X) \cdot E(Y)$ [since x_i, y_j are independent].

$$\frac{1}{\lambda_1} = \frac{k_1}{\lambda_1} + \frac{l_1}{\lambda_2} = \left(\frac{k_1}{\lambda_1} + p_1 k_1 \right) = p_1 \left(\frac{k_1}{\lambda_1} + \frac{l_1}{\lambda_1} \right) =$$

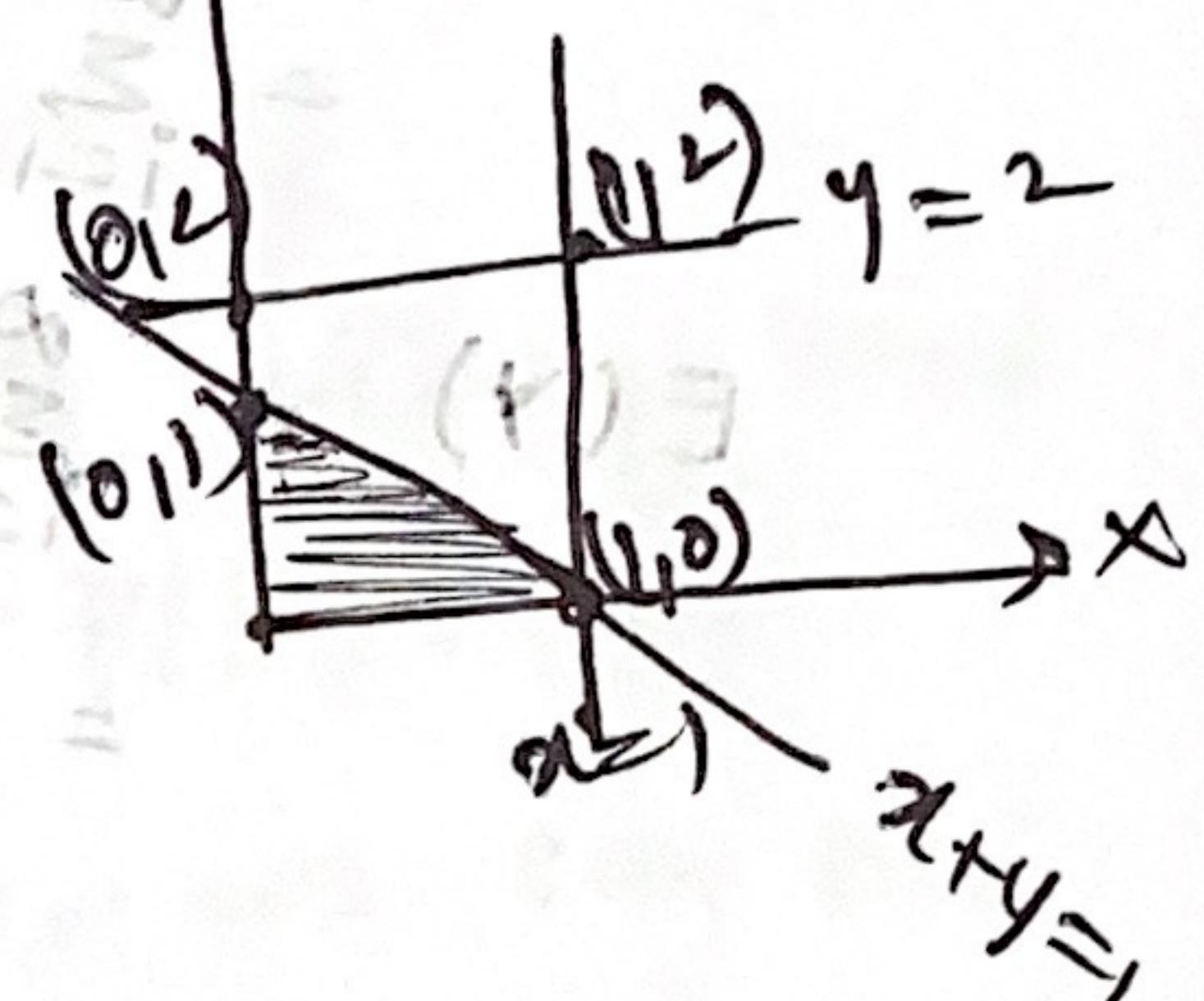
(i)

A 2-D RV x, y has joint PDF $f(x, y) = \begin{cases} x^2 + \frac{2y}{3}, & 0 \leq x \leq 1, 0 \leq y \leq 2 \\ 0, & \text{elsewhere} \end{cases}$

(i) $P(X+Y \geq 1)$ (ii) $E(X)$

(iii) $V(Y)$.

$$(i) P(X+Y \geq 1) = 1 - P(X+Y \leq 1)$$



$$= \int_{x=1}^{x=2} \int_{y=1-x}^{y=2} x^2 + \frac{2y}{3} dy dx$$

$$= \int_{x=1}^{x=2} \left[x^2 y + \frac{2y^2}{6} \right]_{1-x}^{2} dx$$

$$= 1 - \int_{x=0}^{x=2} \int_{y=0}^{y=2-x} x^2 + \frac{2y}{3} dy dx = \int_{x=0}^{x=2} \int_{y=0}^{y=2-x} (x^2 + \frac{2y}{3}) dy dx$$

$$= 1 - \int_{x=0}^{x=2} \left[x^2 y + \frac{2y^2}{6} \right]_{0}^{2-x} dx = (i)$$

$$= 1 - \int_{x=0}^{x=2} x^2(1-x) + x \left(\frac{1}{6}(1-x)^2 - \frac{2x^2}{6} \right) dx$$

$$= 1 - \int_{x=0}^{x=2} \left[-\frac{1}{6}x^3 + \frac{1}{3}x^2 + x + \frac{1}{6}x^3 - \frac{2x^2}{6} \right] dx = 1 - 5x^3 + 4x^2 + x$$

$$= 1 - \frac{1}{6} \int_{x=0}^{x=2} \left[-\frac{5x^3}{6} + \frac{1}{3}x^2 + \frac{x^2}{2} + \frac{x^4}{4} - \frac{2x^3}{3} \right] dx = (K+X)V$$

$$= 1 - \frac{1}{6} \left[-\frac{3}{2} + \frac{2}{3} + \frac{1}{8} + \frac{1}{4} - \frac{2}{3} \right] = (K+X)V$$

$$= 1 - \frac{1}{6} \left[-\frac{5}{4} + \frac{1}{3} + \frac{1}{8} + \frac{1}{4} - \frac{2}{3} \right] = (K+X)V$$

$$= 1 - \frac{1}{6} \left[-\frac{5}{4} + \frac{1}{3} + \frac{1}{8} + \frac{1}{4} - \frac{2}{3} \right] = (K+X)V$$

$$= 1 - \frac{7}{72}$$

$$= (K+X)V$$

$$= \frac{65}{72}$$

$$(ii) E(X) = \int_{x=0}^{x=2} \int_{y=0}^{y=2} x f(x, y) dy dx =$$

$$= \int_{x=0}^{x=2} \int_{y=0}^{y=2} x (x^2 + \frac{2y}{3}) dy dx$$

$$= \int_0^2 \left(\frac{x^4}{4} + \frac{x^2 y}{9} \right)_0^2 dy$$

$$= \int_0^2 \left(\frac{1}{4} + \frac{4}{9} \right) dy = \left(\frac{y}{4} + \frac{y^2}{18} \right)_0^2 = \frac{1}{2} + \frac{4}{18} = \frac{13}{18}$$

$$(iii) V(X) = E(X^2) - [E(X)]^2 = \frac{26}{81}$$

$$\begin{aligned} E(X^2) &= \int_0^1 \int_0^x y(x^2 + 2y/3) dy dx = \frac{1}{3} \int_0^1 3x^2 y + \frac{2y^2}{3} dx \\ &= \int_0^1 \frac{4}{2} \left[\frac{x^3}{3} + \frac{2y^3}{9} \right]_0^x dx = \frac{1}{3} \int_0^1 \left[\frac{3x^2}{2} + \frac{2y^3}{3} \right]_0^x dx \\ &= \int_0^1 \left[x^2 + \frac{2y^2}{6} \right] dx = \frac{1}{3} \int_0^1 \left[\frac{3x^2}{2} + \frac{8x^2}{3} \right] dx \\ &= \left[\frac{2x^3}{3} + \frac{8x^2}{21} \right]_0^1 = \frac{10}{9} = \frac{1}{3} \left[\frac{8}{3} + \frac{8}{21} \right] \\ &= \frac{1}{3} \left[\frac{10}{3} \right] = \frac{10}{9}, \end{aligned}$$

$$E(X^2) = \int_0^1 \int_0^x y^2 (x^2 + 2y/3) dy dx.$$

$$\begin{aligned} &= \frac{1}{3} \int_0^1 \int_0^x 3y^2 x^2 + \frac{2y^3}{3} dy dx \\ &= \frac{1}{3} \int_0^1 \left[\frac{3y^3 x^2}{8} + \frac{2y^4}{12} \right]_0^x dx \\ &= \frac{1}{3} \int_0^1 \left[\frac{8x^2}{3} + \frac{18x^4}{41} \right] dx \\ &= \frac{1}{3} \left[\frac{8x^3}{3} + \frac{4x^5}{21} \right]_0^1 \\ &= \frac{1}{3} \left[\frac{8}{3} + \frac{2}{21} \right] = \frac{1}{3} \left[\frac{8+6}{3} \right] = \frac{14}{9}. \end{aligned}$$

$$E(Y^2) = \frac{14}{9}$$

$$V(Y) = \frac{14}{9} - \frac{100}{81} = \frac{26}{81}.$$

$$+ (Q_2), \text{ joint PDF } f(x,y) = \begin{cases} e^{-(x+y)}, & x, y \geq 0 \\ 0, & \text{elsewhere.} \end{cases}$$

$$(i) P(X < Y)$$

$$(ii) P(X+Y \leq 1)$$

$$(iii) P(1 < X+Y < 2).$$

$$\begin{aligned} &\text{for part (i) } 0 < y < 1 \text{ and } 0 < x < y \\ &= \int_0^y \int_0^y e^{-(x+y)} dx dy \\ &= \int_0^y y e^{-2y} dy = y \left[e^{-2y} \right]_0^y = y [1 - e^{-2y}] \end{aligned}$$

$$A) P(x+y < 2) \quad B) P(x+y \leq 1)$$

$$P(X+Y) =$$

$$e^{-y} e^{-x}$$

$$= \int_0^{\infty} \int_0^{\infty} e^{-(x+y)} dy dx = \int_0^{\infty} \int_0^{\infty} e^{-(x+y)} dy dx$$

$$= \int_0^{\infty} e^{-(x+2-x)} dx = \int_0^{\infty} e^{-2} dx$$

$$= \left[e^{-2} \right]_0^1$$

$$= [e^{-2}] = 0.13533.$$

$$\left[\frac{e^{-2}}{e^{-2}} + \frac{e^{-2}}{e^{-2}} \right] =$$

$$\left[\frac{1}{e^{-2}} + \frac{1}{e^{-2}} \right] = \frac{2}{e^{-2}}$$

$$\frac{2}{e^{-2}} = \left[\frac{2}{e^{-2}} \right] =$$

$$= 2e^2 = 2 \cdot 7.35 = 14.7$$

$$(standard deviation = \sqrt{V(X)})$$

The correlation coefficient

Let (X, Y) be RV we define correlation coefficient
b/w X & Y

$$\rho_{xy} = \frac{E\{(X - E(X))(Y - E(Y))\}}{\sqrt{V(X)V(Y)}}.$$

$E\{(X - E(X))(Y - E(Y))\} \rightarrow$ covariance of X and Y
 σ_{xy} or $Cov(X, Y)$.

Theorem:

$$\rho = \frac{E(XY) - E(X)E(Y)}{\sqrt{V(X)V(Y)}}.$$

$$E(XY) = E(X)E(Y) + E(X)E(Y) + E(Y)E(X) + E(X)E(Y)$$

$$E(XY) = E(X)E(Y) + E(X)E(Y) + E(Y)E(X) + E(X)E(Y)$$

$$E(XY) = E(X)E(Y).$$

Note: If X and Y are independent $E(XY) = E(X)E(Y)$

$$\Rightarrow \sigma_{xy} = 0, \rho_{xy} = 0$$

Note: The converse is not true i.e. if $\rho = 0$ then X & Y need not be independent

Eg: consider RV X define RV $Y = X^2$

$$f(x) = \frac{1}{2}, -1 \leq x \leq 1$$

$$\sigma_{xy} = E(X^3) - E(X)E(X^2)$$

$$\sigma_{xy} = 0 - 0 = 0.$$

$\boxed{\sigma_{xy} = 0 \therefore \rho_{xy} = 0}$ but X & Y are not independent (note 2)

Note: $\rho = 0$

(if X and Y are uncorrelated then $\rho = 0$)

Theorem ②:

$$-1 \leq \rho \leq 1. \quad \text{if } V(X+Y) \geq 0$$

$$E\left[\frac{X-E(X)}{\sqrt{V(X)}} + \frac{Y-E(Y)}{\sqrt{V(Y)}}\right]^2 \geq 0.$$

$$E\left[\frac{X-E(X)}{\sqrt{V(X)}} + \frac{Y-E(Y)}{\sqrt{V(Y)}} + \frac{2(X-E(X))(Y-E(Y))}{\sqrt{V(X)V(Y)}}\right]^2 \geq 0.$$

$$= \left[\frac{E(X)-E(X)}{\sqrt{V(X)}} + \frac{E(Y)-E(Y)}{\sqrt{V(Y)}} + \frac{2(X-E(X))(Y-E(Y))}{\sqrt{V(X)V(Y)}}\right]^2 \geq 0$$

$$= \left[\frac{V(X)}{\sqrt{V(X)}} + \frac{V(Y)}{\sqrt{V(Y)}} + \frac{2(X-E(X))(Y-E(Y))}{\sqrt{V(X)V(Y)}}\right]^2 \geq 0$$

$$= [2 + 2\rho \geq 0]$$

$$= 1 + \rho \geq 0, \quad [-1 \leq \rho \leq 1]$$

Theorem ③: If (X, Y) are linearly related then

$$(\rho = \pm 1 \text{ if } Y = a + bX)$$

Since (X, Y) are linearly related

$$Y = a + bX$$

$$\rho = \frac{E(XY) - E(X)E(Y)}{\sqrt{V(X)V(Y)}}$$

$$= \frac{E(aX + bX^2) - E(a+bX)E(X)}{\sqrt{V(aX+bX) V(X)}}$$

$$= \frac{aE(X) + bE(X^2) - aE(X) - bE(X)^2}{\sqrt{V(X) \sigma^2(V(X))}}$$

$$= \pm \frac{E(X^2) - [E(X)]^2}{V(X)}$$

$$ab(a+b) = \pm \frac{v(x)}{v(y)} = \pm 1.$$

(Ex) $\exists x \exists y \exists z$ - ($\forall u \exists v$ $\neg \mu u v$)

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(9) If $u = a + bx$ and $v = c + dy$ then PT

$$C_{\mu\nu} = \pm R_{\mu\nu} + \text{top term} + \text{other terms} = \mu^2 X$$

-Any)

$$\rho_{uv} = \frac{E\{uv\} - E(u)E(v)}{\sqrt{V(u)V(v)}}.$$

$$\frac{E[(at+bx)(ct+dy)] - E(at+bx)E(ct+dy)}{\sqrt{V(at+bx)V(ct+dy)}} = 1$$

$$= -E[a\epsilon + adY + b\alpha x + b\beta dY] - \cancel{a(E(\epsilon))} - \cancel{bE(\alpha)} + (E(x) - bE(\beta))$$

~~$E(\alpha) = E(\beta)$~~

$$[a + bE(\alpha)][c + bE(\beta)].$$

05 [REDACTED] (U) ~~SECRET~~

$$= \frac{a(x) + a(E(x)) + b(E(x)) + b^2 E(x) - (a+bE(x)) C_C f_d E(x)}{\sqrt{b^2 a(x) V(x) V(x)}}$$

$$= \frac{b^2 [E(UV) - E(U)E(V)]}{\pm b/2 \sqrt{r(X)V(Y)}} = \pm \frac{\text{Pay}}{\text{Var}} \quad \text{group}$$

$$Q_2) \text{ PT } v(ax+bx)=a^2v(x)+b^2xv(x)+2abv_{xy}.$$

$$V(ax+by) = E[(ax+by) - E(ax+by)]^2$$

$$= E[(ax+bx - aE(x) - bE(x))^2]$$

$$= E[(a(x - E(x)) + b(u - E(u))]^2$$

$$= E[a^2(x-Ea)^2 + b^2(x-E(X))^2]$$

$$= a^2 F(x - E(a))^2 + b^2 F(x - E(y))^2 + da_b \sigma_{xy}$$

$$= a^2(v(x)) + b^2(v(y)) + 2ab \sigma_{xy}$$

When proved.

Q3) If x_1, x_2, x_3 are uncorrelated random variables having the same standard deviation find the correlation coefficient between x_1+x_2 and x_2+x_3 .

Sol)

$$\rho_{x_1 x_2} = \rho_{x_2 x_3} = \rho_{x_1 x_3} = 0$$

$$V(x_1) = V(x_2) = V(x_3) = \sigma^2$$

$$\text{Let } U = x_1 + x_2, V = x_2 + x_3$$

$$\rho_{UV} = \frac{E(UV) - E(U)E(V)}{\sqrt{V(U)V(V)}}$$

$$= \frac{E((x_1+x_2)(x_2+x_3)) - E(x_1+x_2)E(x_2+x_3)}{\sqrt{V(x_1+x_2)V(x_2+x_3)}}$$

$$= \frac{E(x_1 x_2 + x_1 x_3 + x_2^2 + x_2 x_3) - E(x_1)E(x_2) - E(x_1)E(x_3) - [E(x_2)]^2 - E(x_2)E(x_3)}{\sqrt{V(x_1) + V(x_2)V(x_2) + V(x_3)}}$$

$$= \frac{E(x_2^2) - E(x_2)^2}{\sqrt{4\sigma^4}} = \frac{E(x_2^2) - E(x_2)^2}{\pm 2\sigma^2} = \frac{V(x_2)}{2\sigma^2}$$

$$= \pm \frac{\sigma^2}{2\sigma^2} = \pm \frac{1}{2}$$

Q4) If x, x_1, x_2 are uncorrelated RV with standard deviation 5, 12, 9 respectively. If $U = x+y, V = y+z$, find ρ_{UV} (correlation coefficient).

Ans)

$$V(x) = 25, V(y) = 144, V(z) = 81$$

$$\rho_{xx} = \rho_{yy} = \rho_{zz} = 0$$

$$\text{Let } U = x+y, V = y+z$$

$$\rho_{UV} = \frac{E(UV) - E(U)E(V)}{\sqrt{V(U)V(V)}}$$

$$\begin{aligned}
 & \text{Let } p(x) = E[(x+y)(x+z)] = E(x+y)E(x+z) \\
 & = \sqrt{V(x+y)V(x+z)} \\
 & = E[(xx+xy+xz+y^2+yz)] - E(x)E(y) - E(x)E(z) - E(y)E(z).
 \end{aligned}$$

$$\begin{aligned}
 & = E(xy) + E(xz) + E(y^2) + E(yz) - E(x)E(y) - E(x)E(z) - E(y)E(z).
 \end{aligned}$$

$$\begin{aligned}
 & = \sqrt{V(x)V(y) + V(x) + V(z)}
 \end{aligned}$$

$$= \frac{E(y) - E(x)}{\sqrt{V(x) + V(y) + V(z)}}$$

$$\begin{aligned}
 & = \frac{75 - 70}{\sqrt{25 \times 144 + 25 \times 8}} = 1.92
 \end{aligned}$$

18/4/22
conditional expectation

If (X, Y) DRV, we define conditional expectation of X for given Y .

$$P(X|Y) = \sum_{i=1}^{\infty} x_i P(X|y_i) \quad E(Y|X) = \sum_{j=1}^{\infty} y_j P(Y|x_j)$$

$$E(X|Y) = \sum_{i=1}^{\infty} x_i \frac{P(X|y_i)}{P(Y|y_i)}$$

If (X, Y) is CRV

$$E(X|Y) = \int_{-\infty}^{\infty} x \frac{f(x|y)}{h(y)} dy$$

$$E(Y|X) = \int_{-\infty}^{\infty} y \frac{f(x|y)}{g(x)} dy$$

(Q1)

$$\text{Let } f(x,y) = \begin{cases} 8xy & , 0 < x < y < 1 \\ 0 & , \text{ otherwise} \end{cases}$$

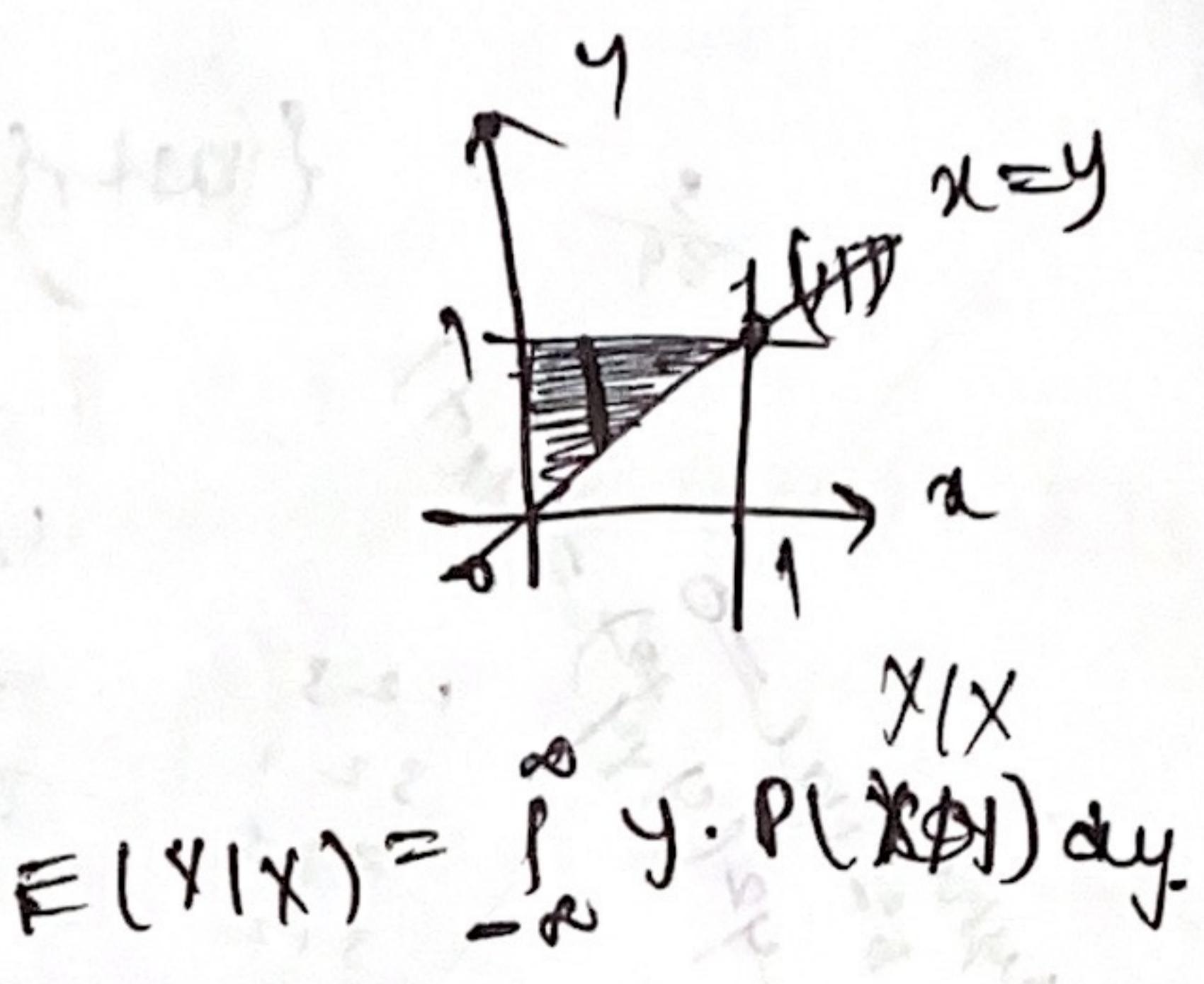
Find,
Ans) (i) $E(Y|X)$ (ii) $V(Y|X)$

(i) $E(Y|X) = ?$

$$g(x) = \int_{-\infty}^{\infty} f(x,y) dy.$$

$$\begin{aligned} &= \int_{y=x}^{\infty} 8xy dy \\ &= \left[8 \frac{y^2}{2} \right]_x^1 \\ &= 4x - 4x^3 \end{aligned}$$

$$g(x) = 4(x - x^3)$$



$$E(Y|X) = \int_0^\infty y \cdot P(X > y) dy.$$

$$E(Y|X) = \int_x^1 y \cdot \frac{8xy}{4x-4x^3} dy$$

$$\begin{aligned} &= \frac{2}{3(1-x^2)} (1-x^3) \left| \frac{y^2}{2} \cdot \frac{8xy^2}{4x-4x^3} \right|_x^1 \\ &\geq \frac{2}{3(1+x)} \cdot \frac{(1-x)}{(1+x)(1-x)} \left| \frac{y^2}{2} \cdot \frac{8xy^2}{4x-4x^3} \right|_x^1 \\ &\geq \frac{2(1+x+x^2)}{3(1+x)} \left| \frac{y^2}{2} \cdot \frac{8xy^2}{4x-4x^3} \right|_x^1 \end{aligned}$$

$$\begin{aligned} &\stackrel{\text{num}}{=} \frac{2(1+x+x^2)}{3(1+x)} \left| \frac{2x^5 - x^5}{4x-4x^3} \right|_x^1 = \frac{x(2-x^4)}{4(1-x^3)} = \frac{2-x^4}{4(1-x^3)} \end{aligned}$$

$$E(Y|X) = \frac{2(1+x+x^2)}{3(1+x)}$$

(ii) $V(Y|X) = E(Y^2|X) - [E(Y|X)]^2$

$$E(Y^2|X) = \int_x^1 y^2 \cdot \frac{8xy}{4x-4x^3} dy$$

$$= \frac{2}{1-x^2} \int_x^1 y^3 dy$$

$$= \frac{2}{4(1-x)} (1-x^4)$$

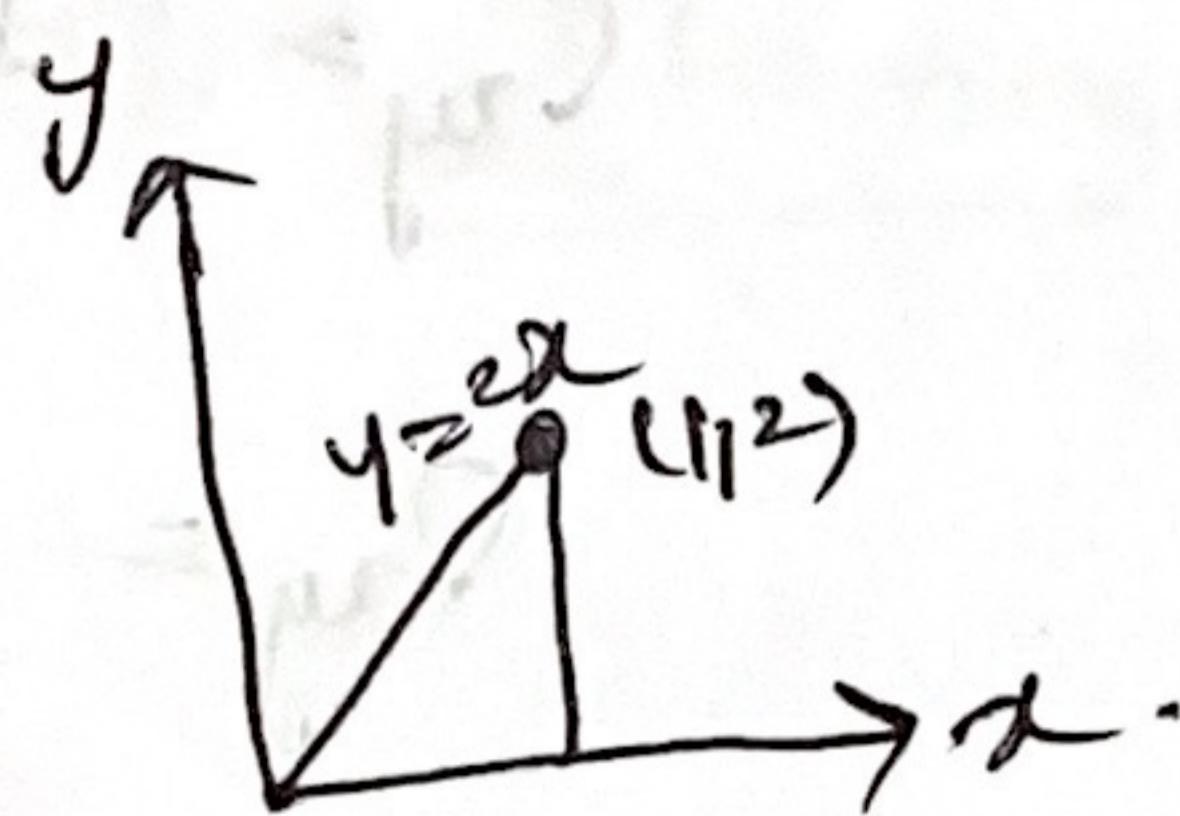
$$= \frac{1}{4(1-x^2)} (1+x^2)(1-x^2)$$

$$= \frac{1+x^2}{2}$$

$$V(Y|X) = \frac{1+x^2}{2} - \left[\frac{2(1+x+x^2)}{3(1+x)} \right]^2$$

(Q2) suppose that (X,Y) is uniformly distributed over triangular region. Then find marginal and conditional P.d.f's of X & Y .

$$\text{Ans) } f(x|y) = \begin{cases} 1, & (x,y) \in D^{\text{tri}} \\ 0, & \text{elsewhere} \end{cases}$$



$$\rightarrow \text{P.d.f} = \frac{1}{\text{area of given region}} = \frac{1}{\frac{1}{2} \cdot 1^2 \cdot 2} = 1 \quad \text{①}$$

$$g(x) = ? \quad h(y) = ? \quad P(X|Y) = ? \quad P(Y|X) = ?$$

$$\begin{aligned} g(x) &= \int_{-\infty}^{\infty} f(x|y) dy \\ &= \int_0^{2x} 1 \cdot dy \\ &= 2x \quad (0 \leq x \leq 1). \end{aligned}$$

$$P(X|Y) = \frac{f(x|y)}{g(x)} = \frac{1}{2x} \quad (0 < x < 1)$$

$$\begin{aligned} h(y) &= \int_{-\infty}^{\infty} f(x|y) dx \\ &= \int_{y/2}^1 1 dx \\ &= 1 - y/2 = \frac{2-y}{2}, \quad (0 \leq y \leq 2) \end{aligned}$$

$$P(Y|X) = \frac{f(x|y)}{h(y)} = \frac{1}{\frac{2-y}{2}} = \frac{2}{2-y} \cdot (0 < y < 2)$$

$$\begin{aligned} ① E(X|Y) &= \int_{-\infty}^{\infty} x \cdot P(X|Y) dy \\ &= \int_{y/2}^1 \frac{2x}{2-y} dx, \quad 0 < y < 2. \end{aligned}$$

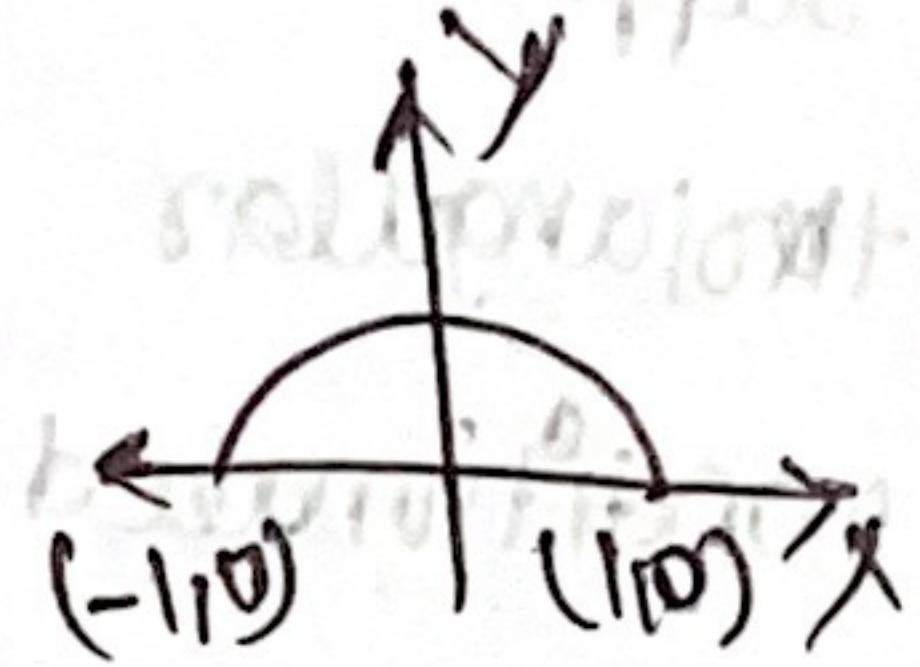
$$\begin{aligned} ② E(Y|X) &= \int_{-\infty}^{\infty} y \cdot P(Y|X) dx \\ &= \int_0^{2x} \frac{y}{2x} dy, \quad 0 < x < 1. \end{aligned}$$

$$\frac{1}{2} \left(\frac{y^2}{2x} \right) \Big|_0^{2x} = \frac{1}{2} \frac{2x \cdot 4x^2}{2x} = x$$

(Q3) suppose a 2D RV (X, Y) uniformly distributed over the region R defined by $R: \{(x, y) / x^2 + y^2 \leq 1, y \geq 0\}$. find correlation coefficient.

A)

$$C_{xy} = \frac{E[(x - E(x))(y - E(y))]}{\sqrt{V(x)V(y)}}$$



$$\rho_{xy} = \frac{E(XY) - E(X)E(Y)}{\sqrt{V(X)V(Y)}}$$

$$\text{Pdt} = \frac{1}{\frac{\pi}{2}} = \frac{2}{\pi}$$

$$\begin{aligned} E(X) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x \cdot f(x, y) dy dx \\ &= \int_{-1}^{1} \int_{0}^{\sqrt{1-x^2}} x \cdot \frac{2}{\pi} dy dx \end{aligned}$$

$$f(x, y) = \begin{cases} \frac{2}{\pi}, & (x, y) \in R \\ 0, & \text{otherwise} \end{cases}$$

$$E(X) = 0$$

$$E(X^2) = \int_0^{\pi/2} \int_0^{\sqrt{1-y^2}} x^2 dy dx$$

$$\begin{aligned} &= \int_0^{\pi/2} \int_0^{\sqrt{1-y^2}} y^2 dy dx \\ &= \frac{4}{\pi} \int_0^{\pi/2} \int_0^{\sqrt{1-y^2}} y^2 dy dx \end{aligned}$$

$$\begin{aligned} &= \frac{4}{\pi} \int_0^{\pi/2} \left(\frac{1}{3} y^3 \right) \Big|_0^{\sqrt{1-y^2}} dy \\ &= \frac{4}{3\pi} \int_0^{\pi/2} \cos^3 \theta d\theta \end{aligned}$$

$$\begin{aligned} E(Y) &=? \\ E(Y^2) &=? \\ E(XY) &=? \end{aligned}$$

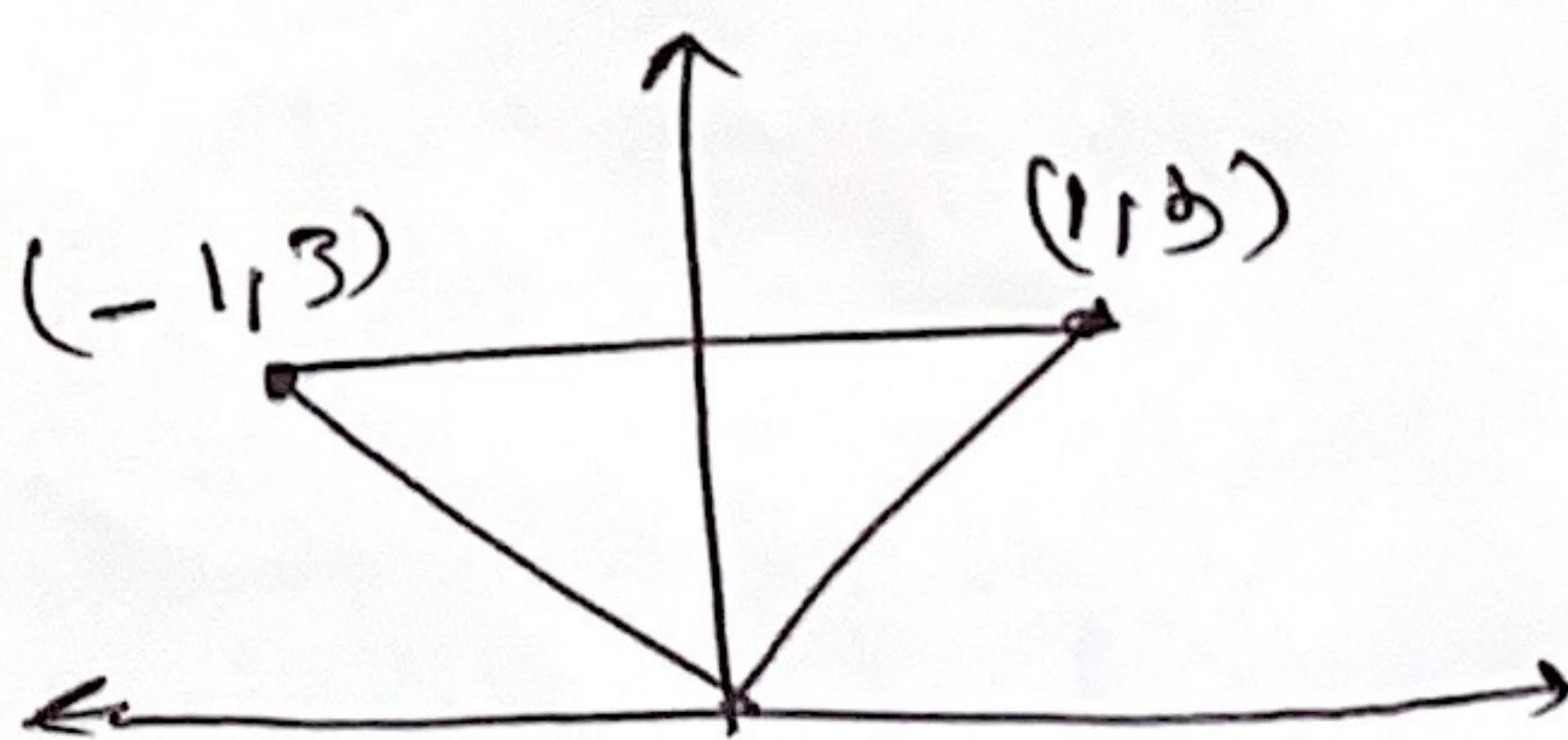
$$= \frac{4}{3\pi} \cdot \frac{8}{9} \cdot \frac{1}{2} \cdot \frac{\pi}{12}$$

$$= \frac{1}{4}$$

$$\begin{aligned} \text{The sum of cosines} &= \int_0^{\pi/2} \cos \theta d\theta = \left[\frac{\sin \theta}{\theta} \right]_0^{\pi/2} \\ &= 1 \\ \text{d}y &= \cos \theta d\theta \\ \text{d}x &= \sin \theta d\theta \end{aligned}$$

(Q4) Suppose that X, Y is uniformly distributed over the triangle. (1) obtain the marginal pdf of X & Y

(2) evaluate $V(X), V(Y)$.



(5) If x & y are independent random variables then find $E(X|Y)$ and $E(Y|X)$.

(6) Suppose that x, y has pdf $f(x,y) = \begin{cases} k e^{-y}, & 0 < x < y \\ 0, & \text{elsewhere} \end{cases}$
Find the correlation coefficient. $K = ? \in \left[\int \int f(x,y) dx dy = 1 \right]$
($P_{xy} = ?$).

Difference Equation:

A difference equation is a relation b/w the difference of an unknown function at one or more general values of the argument.

$$\Delta y_{n+1} + y_n = 0 \rightarrow ①$$

$$\Delta y_{n+1} + \Delta^2 y_{n-1} = 1 \rightarrow ②$$

equation's ④ can be rewritten as

$$① \Rightarrow \Delta y_{n+2} - y_{n+1} + y_n = 2 \rightarrow ③$$

$$② \Rightarrow y_{n+2} - y_{n+1} + y_{n+1} - y_n - y_{n-1} = 1.$$

$$③ \Rightarrow y_{n+2} + y_{n-1} - 2y_n = 1 \rightarrow ④.$$

Formation of Difference equation:

$$y = ax + bx^2$$

$$\begin{aligned} \Delta y &= a\Delta x + b\Delta x^2 \\ &= a(x+1-x) + b[(x+1)^2 - x^2] \\ &= a + b(2x+1). \end{aligned}$$

$$\Delta^2 y = 2b.$$

$$\begin{vmatrix} y & x & x^2 \\ \Delta y & 1 & (x+1) \\ \Delta^2 y & 0 & 2 \end{vmatrix} = 0.$$

$$\begin{aligned} &= q(2-0) - a(2\Delta y - \Delta^2 y(2x+1) + x^2(-\Delta^2 y)) = 0 \\ &\Rightarrow 2q - 2x\Delta y + (2x^2 + a - a^2)\Delta^2 y = 0 \\ &(a^2 + a)\Delta^2 y - 2x\Delta y + 2q = 0 \end{aligned}$$