

$$(\Omega, \underline{a}, \underline{P})$$

$$\underline{X}: \Omega \longrightarrow \mathbb{R}$$

X is r.v if $\underline{X^{-1}(B)} \in \underline{a}$

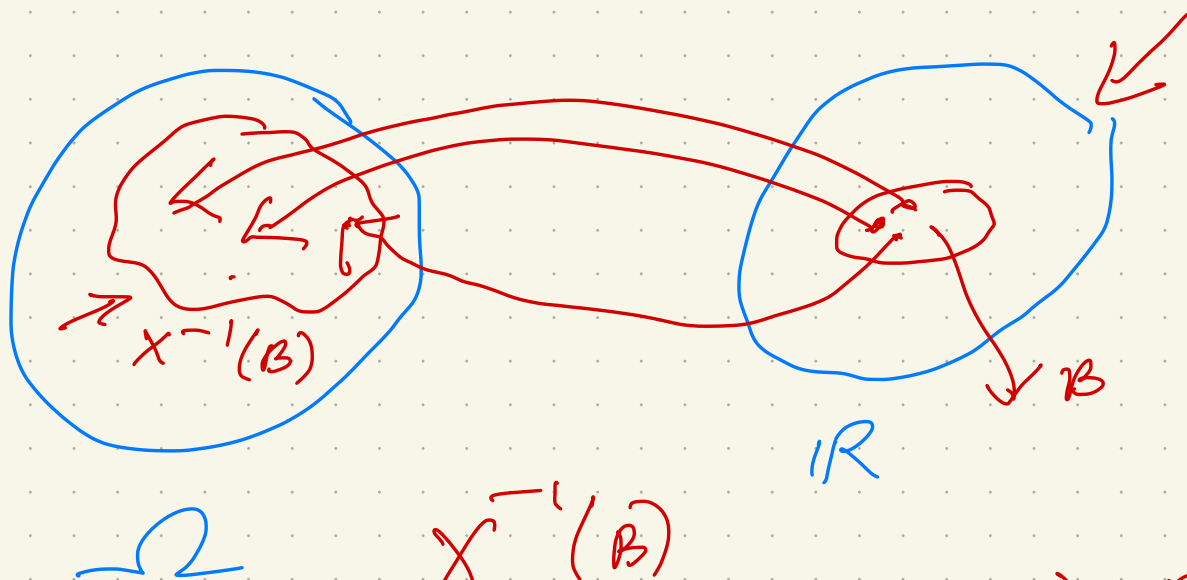
for all $B \in \underline{B_{\mathbb{R}}}$ \rightarrow Borel σ field

Induced probability on $B_{\mathbb{R}}$

Define P_X as:

$$\underline{P_X(B)} = \underline{P[X^{-1}(B)]}$$

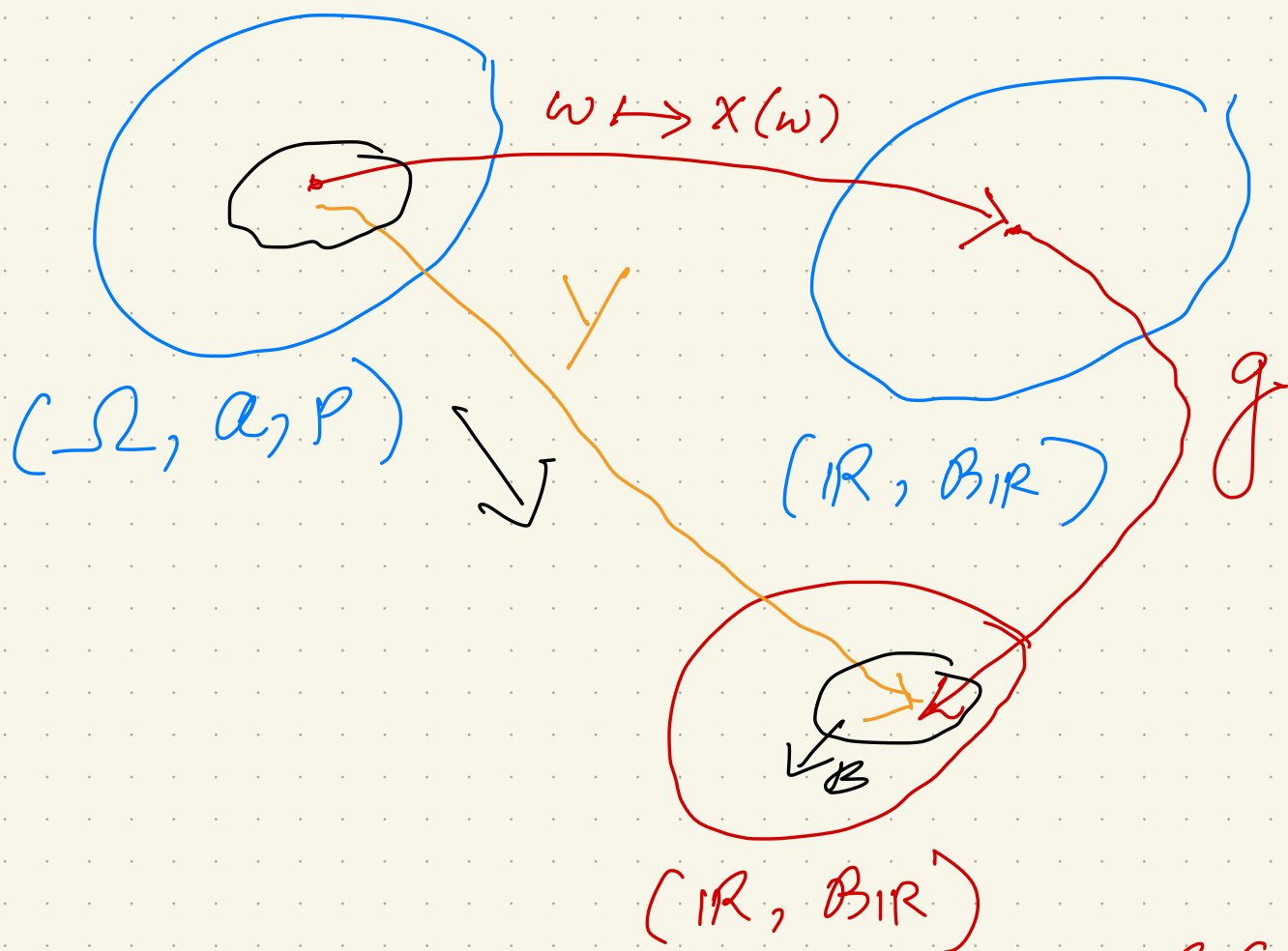
P_X :
distribution
of X



$$X^{-1}(B)$$

$$= \{ \omega \in \Omega : X(\omega) \in B \}$$

$$\underline{P_X(B)} = P(X^{-1}(B))$$



$$X: \Omega \rightarrow \mathbb{R}, \quad X^{-1}(B) \in \mathcal{A} \quad \forall B \in \mathcal{B}_{\mathbb{R}}$$

$$g: \mathbb{R} \rightarrow \mathbb{R}, \quad g \text{ measurable}$$

i.e. $g^{-1}(B) \in \mathcal{B}_{\mathbb{R}}$ for all $B \in \mathcal{B}_{\mathbb{R}}$

$$Y = g(X) \equiv g \circ X$$

$Y = \Omega \rightarrow \mathbb{R}$, Y is also
a proper random variable

P_X and g

Consider P_Y : the distribution of Y .

Let $B \in \mathcal{B}_R$:

$$\begin{aligned} P_Y(B) &= P(Y^{-1}(B)) \\ &= P[(g(X))^{-1}(B)] \\ &= P[X^{-1}(\underbrace{g^{-1}(B)})] \\ &= \underbrace{P_X}_{\text{same}}[g^{-1}(B)] \end{aligned}$$

We can also write:

$$\begin{aligned} P_Y(B) &= P\{\omega : Y(\omega) \in B\} \\ &= P_X[g^{-1}(B)] \\ &= P[\omega : X(\omega) \in g^{-1}(B)] \end{aligned}$$

Random variables X and X' have the same distribution if $P_X = P_{X'}$

$$\text{i.e. } \underbrace{P_X(B)}_{=} = \underbrace{P_{X'}(B)}_{=}$$

$$\underbrace{P(X \in B) \quad P(X' \in B)}_{}$$

P_X and $P_{X'}$ are identical iff F_X

and $F_{X'}$ are identical:

$$\text{i.e. } \underbrace{F_X(y)}_{\downarrow} = \underbrace{F_{X'}(y)}_{\downarrow} \text{ for every } y$$

$$P(X \leq y) \quad P(X' \leq y)$$

$$\text{here } B = (-\infty, y]$$

relate p.m.f of Y to that of X in
discrete case

relate p.d.f of Y to that of X in
the continuous case

Start with X discrete

Theorem 2.3

(X) discrete random variable assuming

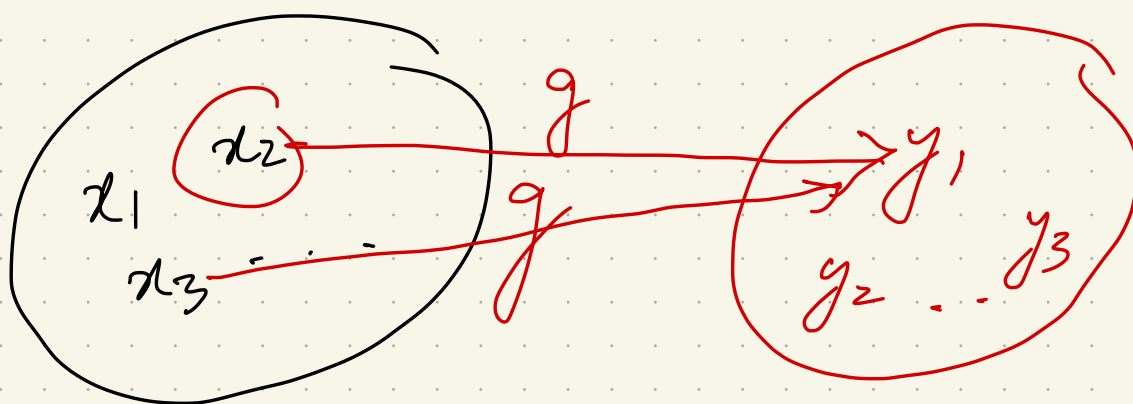
values $\{x_1, x_2, \dots\}$

associated p.m.f $\{p_1, p_2, p_3, \dots\}$

$$p_j = P(X = x_j) \quad p_j > 0$$

$Y = g(X)$ suppose Y assumes

values $\{y_1, y_2, y_3, \dots\}$



p_Y : want to find

what is $p_Y(y_e) = \underline{P(Y = y_e)}$.

$$P(Y = y_e).$$

$$= P[X \in \{\underline{x_j} : \underline{g(x_j)} = y_e\}]$$

$$= \sum_{j: g(x_j) = y_e} \boxed{P(X = x_j)} p_X(x_j)$$

$$= \sum_{\bar{g}^{-1}(\{y_e\})} p_X(x_j) \leftarrow \}$$

Example 2.4: X discrete random

variable.

$$X \in \{\pm m, m = 0, 1, 2, 3, \dots\}$$

$$p_X(0) = \underline{\frac{1}{2}}, \quad p_X(m) = P(X = m) = \left(\frac{2}{3} \cdot \frac{1}{2^{m+1}} \right) \leftarrow$$

What's an intuitive way of generating an X like this?

$$p_X(-m) = \left(\frac{1}{3} \cdot \frac{1}{2^{m+1}} \right)$$

Find p.m.f of $Y = |X| \equiv g(X)$.

$$P(Y=0) = P(X=0) = \frac{1}{2}$$

$$\begin{aligned} P(Y=m) &= P(X \in g^{-1}(m)), \underline{m} > 0 \\ &= P(X \in \{-m, m\}) \\ &= p_X(m) + p_X(-m) = \frac{1}{2^{m+1}} \end{aligned}$$

Continuous random variables with

p.m.f's.

Change of Variable Theorem }:

Theorem 2.4

X : continuous random variable

$$P(X \in I) = 1. \quad I = (\underline{a}, \underline{b})$$

X has a p.d.f on I ~~is~~ denoted f_X .

$$\text{So for any } a < c < d < b \quad P(X \in (c, d)) = \int_c^d f_X(t) dt$$

define $Y = g(X)$

find p.d.f of Y .

Not all g 's produce a continuous Y .

Note that g needs to be sufficiently rich / complex for $g(X)$ to be continuous and have a pdf.

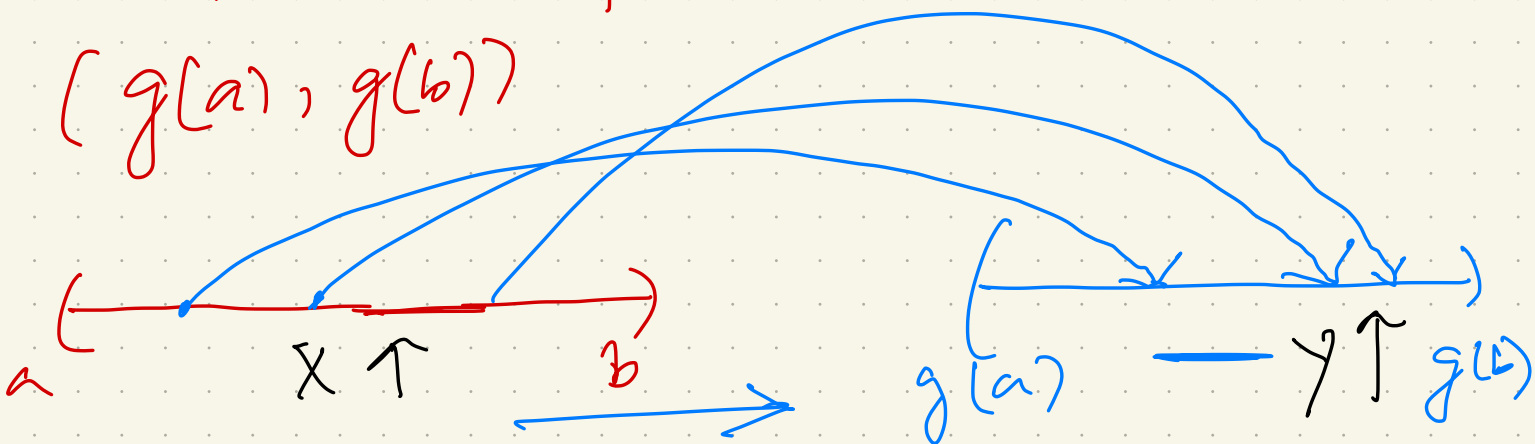
Regularity conditions on g :

Assume g is continuously differentiable and that $g'(x) \neq 0$ for $x \in \underline{I}$.

$\Rightarrow g$ is either \uparrow on \underline{I} or \downarrow on \underline{I} .

and then g is a 1-1 continuously differentiable fn. from (a, b) to

$(g(a), g(b))$

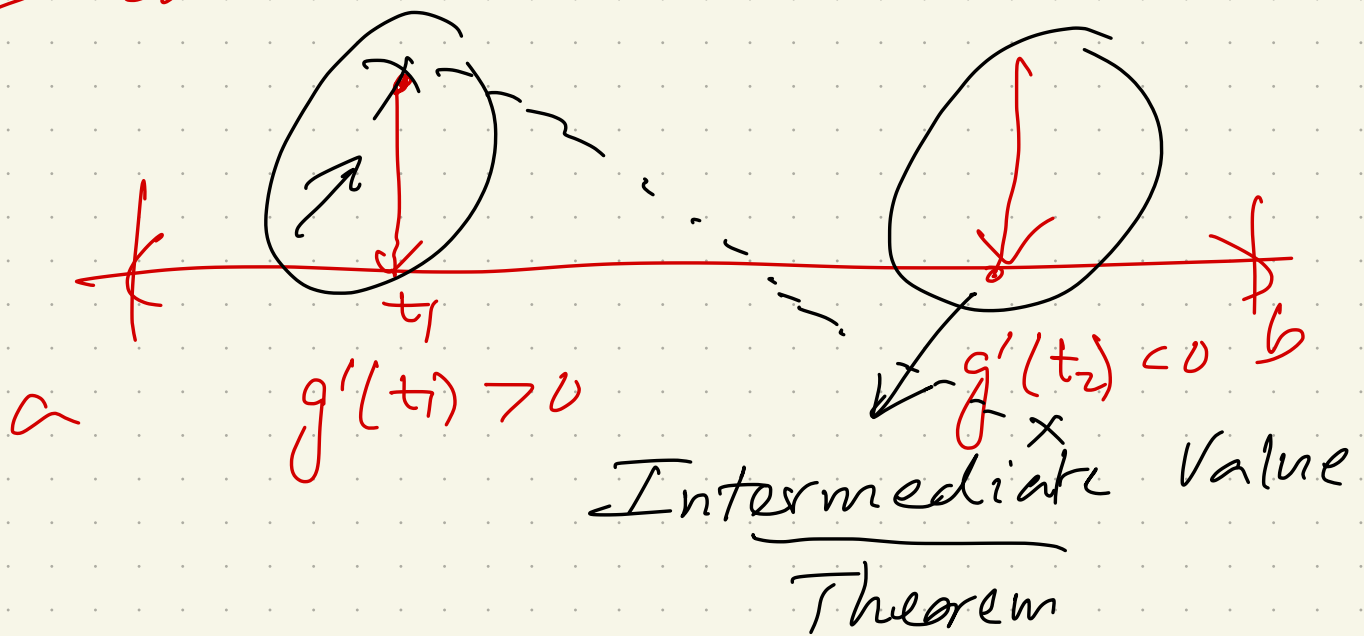


$$\underline{g'(x) \neq 0 \quad \forall x \in I_2}$$

you know $x_1 \neq x_2 \Rightarrow g(x_1) \neq g(x_2)$

because were $g(x_1) = g(x_2)$, the mean value theorem gives you a point between x_1, x_2 , say x^* s.t. $g'(x^*) = 0$.

You also know that g' cannot be both +ve and -ve. \rightarrow because g' is continuous.



either $g'(t) > 0$ for all $t \Rightarrow g \uparrow$
 or $g'(t) < 0$ for all $t \Rightarrow g \downarrow$

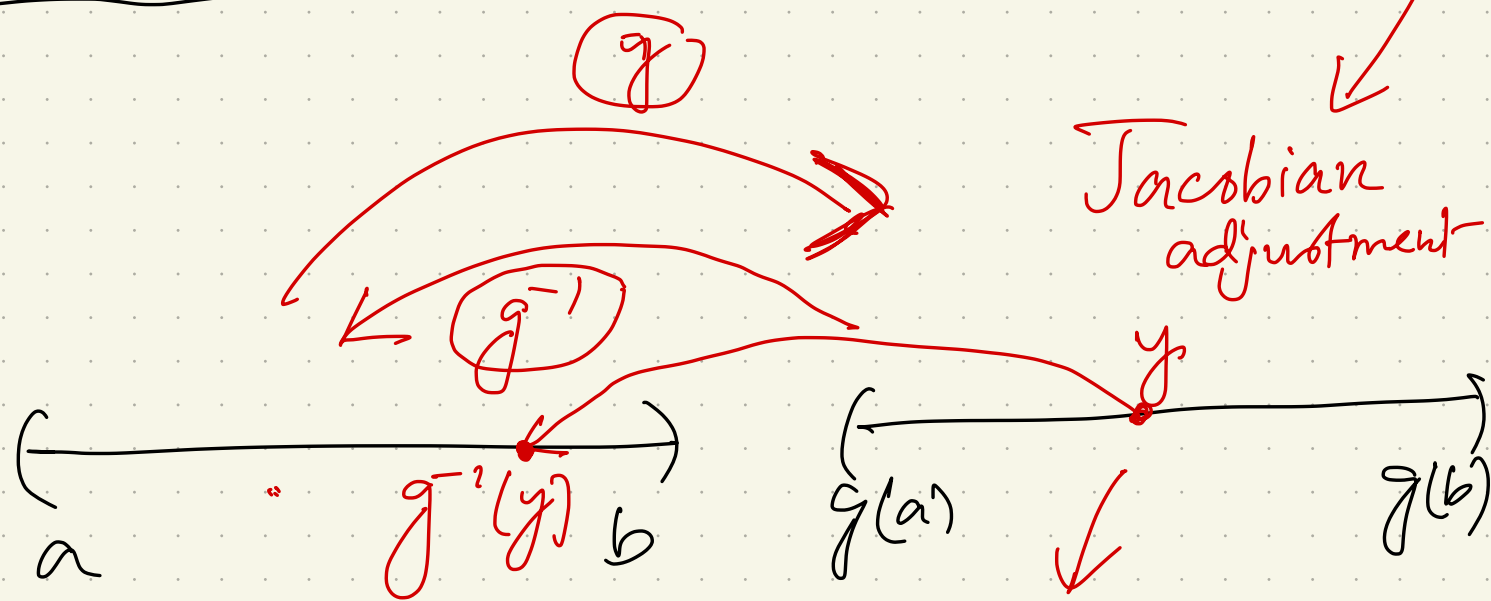
Define $y = g(x)$.

So $P(Y \in (g(a), g(b))) = 1$

(g is being assumed \uparrow)

Then the density function of Y is given by:

$$\underline{f_Y(y)} = \underline{f_X(g^{-1}(y))} \left| \frac{d}{dy} g^{-1}(y) \right| \cdot \underline{1(y \in g(I))}$$



$$\left| \frac{d}{dy} g^{-1}(y) \right| = \frac{1}{g'(g^{-1}(y))}$$

chain rule

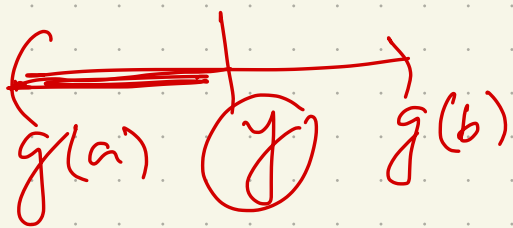
Proof:

$$I = (a, b), \quad g \uparrow.$$

F_Y : distribution function of Y .

$$\underline{F_Y(y)} = P(Y \leq y)$$

$$= \begin{cases} 0 & \text{if } y \leq g(a) \\ (*) & \text{if } y \in (g(a), g(b)) \\ 1 & \text{if } y \geq g(b) \end{cases}$$



$$\underline{P(Y \leq y)} = \underbrace{P(Y \leq g(a))}_{=0} + P(g(a) < Y \leq y)$$

$$= P(g(a) < Y \leq y)$$

$$= P(g(a) < g(x) \leq y)$$

$$= P(a < x \leq g^{-1}(y))$$

applying g^{-1}

$$P(a < x \leq g^{-1}(y))$$

$$= P(x \leq g^{-1}(y)) \text{ , since } P(x > a) = 1$$

$$\text{So: } \underline{F_Y(y)} = P(Y \leq y)$$

$$= P(x \leq g^{-1}(y))$$

$$= \underline{F_X(g^{-1}(y))}$$

$$\text{so } \underline{f_Y(y)} = \frac{d}{dy} F_Y(y)$$

$$= \frac{d}{dy} (F_X(g^{-1}(y)))$$

Try
the
same
derivation
with g

↓ chain rule

$$= f_X(g^{-1}(y)) \frac{d}{dy} (g^{-1}(y))$$

gives me the
desired expression.

where did I lose 1?

Exercise 2.5 $U \sim \text{Uniform}(0, 1)$.

define $w = -\frac{1}{\lambda} \log u$.

$\lambda > 0$.

Find density of w .

Transfer to the notation of our theorem:

$U \equiv X$, $(a, b) = (0, 1)$

$f_U(u) = 1 \cdot \mathbb{1}(0 < u < 1)$

$w = g(u) = -\frac{1}{\lambda} \log u$.

$(g(a), g(b)) = (0, \infty)$.

$u \mapsto -\frac{1}{\lambda} \log(u) \equiv g(u)$

is a strictly \downarrow function from
 $(0, 1)$ to $(0, \infty)$

so W has a density on $(0, \infty)$

$$g(u) = -\frac{1}{\lambda} \log u$$

$$g'(u) = -\frac{1}{\lambda u} < 0!$$

$$g(u) = -\frac{1}{\lambda} \log u =$$

Finding inverse means express u in terms of $-g(u)$.

$$\checkmark \text{ so } + \log u = -\lambda g(u)$$

$$\Rightarrow u = \exp[-\lambda g(u)]$$

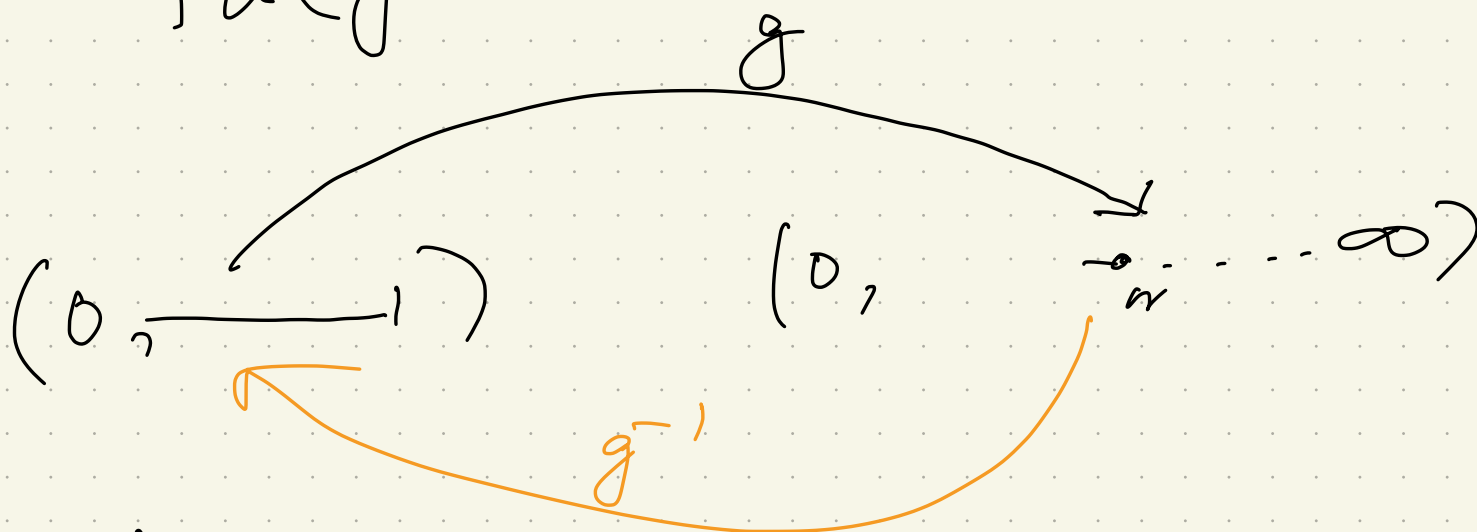
so $g^{-1}: (0, \infty) \rightarrow (0, 1)$

$$\text{with } g^{-1}(w) = \exp[-\lambda w]$$

$$(w \equiv g(u))$$

$$\underline{f_W(w)} = f_U(g^{-1}(w)) \left| \frac{d}{dw} g^{-1}(w) \right| \mathbb{1}_{\{w \in (0, \infty)\}}.$$

$$f_u(g^{-1}(w)) = 1.$$



$$\frac{d}{dw} g^{-1}(w)$$

$$= \frac{d}{dw} [e^{-\lambda w}] = -\lambda e^{-\lambda w}$$

$$f_w(w) = \lambda e^{-\lambda w} \mathbb{1}(w \in (0, \infty))$$

✓
This is the exponential λ density.

$$\widetilde{w} = -\frac{1}{\lambda} \log(1-u).$$

Find density of \widetilde{w} }
 $u = 1 - e^{-\lambda w}$

use COVT! $\widetilde{u} \sim \text{Unif}(0, 1)$

$$\widetilde{w} = -\frac{1}{\lambda} \log \widetilde{u}$$

$$\widetilde{u} \sim \text{Unif}(0, 1).$$

$$\text{so } \widetilde{w} \sim \exp(\lambda).$$

$g \uparrow$ or \downarrow and we considered
a mapping from an interval to
an interval.

Would like generalizations.

What if g is a many to one
function?

What if g is a many to many
function?

$$g(x) = x^2$$

Next Class.