

Econometric Analysis of Cross Section and Panel Data

Lecture 1: Introduction, CEF and Projection

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This Lecture

Hansen (2022): Chapter 1 and 2

- ▶ Basic econometric concepts
- ▶ Conditional expectation function
- ▶ Best linear predictor

Econometrics

Econometrics typically studies the following model:

$$h(X, Y, U) = 0.$$

- ▶ We don't know the joint distribution of (X, Y, U) .
- ▶ We have observed a **sample** of (X, Y) . But we are not able to observe U .
- ▶ We want to learn something about $h(\cdot)$ according to the sample.

Econometrics

In this course we only study parametric models half-parametric and non-parametric

$$h(X, Y, U; \beta) = 0 \quad \text{norm distribution a typical parametric model}$$

where the functional form of $h(\cdot)$ is known up to a parameter β . More specifically, we will be focusing on linear models with various assumptions of U :

linear is the linear parameter, and n dimension of x_n is allowed

Gaussian-Markov assumption:

$$Y = X'\beta + U$$

1.random sample, or observation are iid(为满足大数定理和中心极限定理以便估计 β)

2.linear 3.不完全共线(ols中求(伪)逆) 4. $E(U|X)=0$ 均值独立 无偏性, $E(\beta|X)=\beta$; 大样本下无偏性对应一致性

- ▶ The estimation of β is aimed at finding a statistic $\hat{\beta}$ that is "reasonably" close to β 's true value.
- ▶ The statistical inference on β involves hypothesis testing and constructing confidence intervals for β .

假设 $E(U|X)$ 强于 $E(UX)$, 若 $E(UX)=0$ 不满足, OLS失效, 而使用工具变量方法

Observational Data

- ▶ A common econometric question is to quantify the causal impact of one set of variables on another variable.
 - ▶ the returns to schooling – the change in earnings induced by increasing a worker's education, holding other variables constant
- ▶ **Experimental Data** & **Observational Data**
- ▶ **Experimental Data**: draw data from experiments (Randomized Control Trials are prevailing)
 - ▶ Banerjee and Duflo, *Poor Economics* (2019 Nobel Prize with Michael Kremer)
 - ▶ Money burning, immoral, general equilibrium effects, external validity concern
- ▶ **Observational Data**: survey data, census data (人口普查)
 - ▶ draw balls from a black box, call people to survey their characteristics

实验方法：不道德，贵，一般均衡效应（发钱扰动市场），不具有现实指导意义

Standard Data Structures

1. **Cross-sectional data:** one observation per individual
 - ▶ survey data (one period), census data
 - ▶ It is conventional to assume that cross-sectional observations are mutually independent.
2. **Times series data:** several periods for one variable
 - ▶ macroeconomic aggregates, prices, and interest rates
 - ▶ Serial dependence: $Y_t = \rho Y_{t-1} + e_t$
3. **Panel data:** a set of individuals measured repeatedly over time
 - ▶ independent across individuals, but dependent across periods for a given individual
4. **Clustered sample**
5. **Spatial dependent sample**

Standard Data Structures

- ▶ **Independently Distributed:** i^{th} observation (Y_i, X_i) is independent of j^{th} observation (Y_j, X_j)
- ▶ **Identically Distributed:** Observations are drawn from the same probability distribution.
 - ▶ The variables (Y_i, X_i) are a sample from the distribution $F(Y, X)$ if they are identically distributed with distribution $F(Y, X)$.
 - ▶ $F(Y, X)$ is called as **the population**.
- ▶ Think of drawing balls from one box or two box with different ball-color distribution
- ▶ **Independent and identically distributed, i.i.d., or a random sample:** The variables (Y_i, X_i) are a random sample if they are mutually independent and identically distributed (i.i.d.) across $i = 1, \dots, n$.

Standard Data Structures

- ▶ **The distribution F is unknown, and the goal of statistical inference is to learn about features of F from the sample.**
- ▶ The assumption of random sampling provides the mathematical foundation for treating economic statistics with the tools of mathematical statistics.
- ▶ Law of Large Number, Central Limit Theorem

The Distribution of Wages

- ▶ Remember, we treat variables as random!
- ▶ For example, we view the wage of an individual worker as a random variable *wage* with the probability distribution

$$F(y) = \mathbb{P}[wage \leq y]$$

Condition expectation(CEF): $E(e|x)=0$
BLP: $E(xe)=0$

- ▶ Probability density function:

$$f(y) = \frac{d}{dy} F(y)$$

The Distribution of Wages

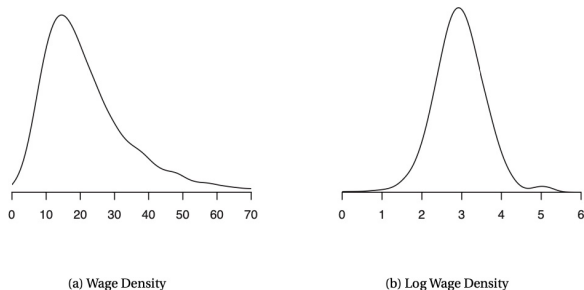


Figure 2.1: Density of Wages and Log Wages

- ▶ Median m : $F(m) = \frac{1}{2}$
- ▶ Mean or expectation of a random variable Y :
 - ▶ Discrete: $\mu = \mathbb{E}[Y] = \sum_{j=1}^{\infty} \tau_j \mathbb{P}[Y = \tau_j]$
 - ▶ Continuous: $\mu = \mathbb{E}[Y] = \int_{-\infty}^{\infty} y f(y) dy$
- ▶ The expectation is not robust, so usually we take log of wage.

Conditional Expectation

- ▶ Conditional expectation is defined as $\mathbb{E}[Y|X = x]$ or $\mathbb{E}[Y|X]$.
- ▶ Conditional expectation of log wages given gender:

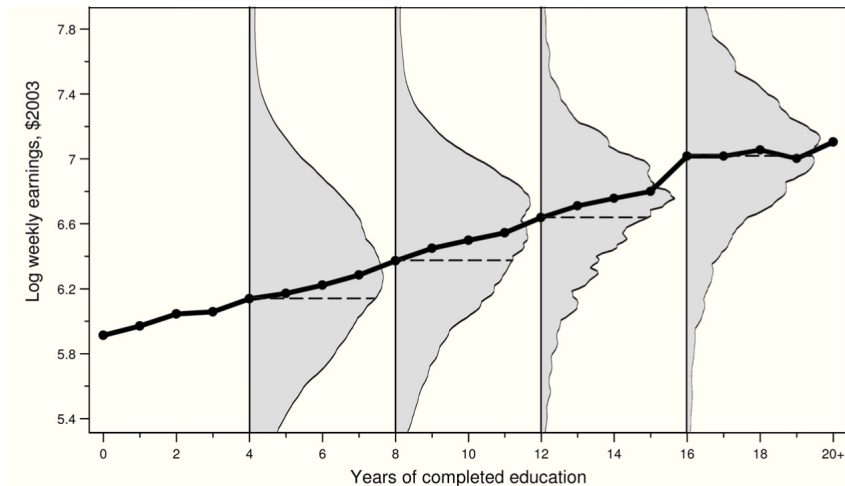
$$\mathbb{E}[\log(wage) \mid gender = man] = 3.05$$

$$\mathbb{E}[\log(wage) \mid gender = woman] = 2.81$$

$$E(Y|X) = \int Y f(Y|x) dY = \int \frac{f(x,y)}{f(x)} dY$$

Conditional Expectation

It reminds me of EM algorithm.



Conditional Expectation Function

- ▶ More generally, we can define the conditional expectation as a function of many random variables

$$\mathbb{E}[Y \mid X_1 = x_1, X_2 = x_2, \dots, X_k = x_k] = m(x_1, x_2, \dots, x_k)$$

- ▶ We call this the **conditional expectation function** (CEF).
- ▶ We typically write the conditioning variables as a vector in \mathbb{R}^k :

$$X = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_k \end{pmatrix}$$

- ▶ Given this notation, the CEF can be compactly written as

$$\mathbb{E}[Y \mid X = x] = m(x)$$

- ▶ When X takes the value x then the average value of Y is $m(x)$.

CEF with Continuous Variables

- ▶ Assume that the variables (Y, X) are continuously distributed with a joint density function $f(y, x)$.
- ▶ The marginal density of x is

$$f_X(x) = \int_{-\infty}^{\infty} f(y, x) dy$$

- ▶ For any x such that $f_X(x) > 0$ the conditional density of Y given X is defined as

$$f_{Y|X}(y | x) = \frac{f(y, x)}{f_X(x)}$$

- ▶ The CEF of Y given $X = x$ is the expectation of the conditional density (Note: $\mathbb{E}[Y] = \int_{-\infty}^{\infty} y f_Y(y) dy$)

$$m(x) = \mathbb{E}[Y | X = x] = \int_{-\infty}^{\infty} y f_{Y|X}(y | x) dy$$

Law of Iterated Expectations

Theorem 2.1 Simple Law of Iterated Expectations

- ▶ If $\mathbb{E}|Y| < \infty$ then for any random vector X , $\mathbb{E}[\mathbb{E}[Y | X]] = \mathbb{E}[Y]$.

Theorem 2.2 Law of Iterated Expectations

- ▶ If $\mathbb{E}|Y| < \infty$ then for any random vector X_1 and X_2 ,
 $\mathbb{E}[\mathbb{E}[Y | X_1, X_2] | X_1] = \mathbb{E}[Y | X_1]$.

Theorem 2.3 Conditioning Theorem

- ▶ If $\mathbb{E}|Y| < \infty$ then $\mathbb{E}[g(X)Y | X] = g(X)\mathbb{E}[Y | X]$
- ▶ If in addition $\mathbb{E}|g(X)| < \infty$ then $\mathbb{E}[g(X)Y] = \mathbb{E}[g(X)\mathbb{E}[Y | X]]$.

CEF Error

Define **CEF error**:

$$e = Y - m(X) \quad X \setminus \beta = m(x)$$

Theorem 2.4 Properties of the CEF error

If $\mathbb{E}|Y| < \infty$ then

1. $\mathbb{E}[e | X] = 0$ (**mean independence**)
 2. $\mathbb{E}[e] = 0$ 无条件期望，迭代法则求解
 3. For any function $h(x)$ such that $\mathbb{E}|h(X)e| < \infty$ then $\mathbb{E}[h(X)e] = 0$.
- The equations that $Y = m(X) + e$ and $\mathbb{E}[e | X] = 0$ together imply that $m(X)$ is the CEF of Y given X .
 $g(x) = \mathbb{E}[Y|x] = m(x)$

Regression Variance

Unconditional variance of the CEF error:

$$\sigma^2 = \text{var}[e] = \mathbb{E} [(e - \mathbb{E}[e])^2] = \mathbb{E} [e^2]$$

- ▶ The magnitude of σ^2 measures the amount of variation in Y which is not “explained” or accounted for in the conditional expectation $E[Y|X]$.
- ▶ The variance of this unexplained portion e decreases when we condition on more variables.

Theorem 2.6 If $\mathbb{E} [Y^2] < \infty$ then

$$\text{var}[Y] \geq \text{var} [Y - \mathbb{E} [Y | X_1]] \geq \text{var} [Y - \mathbb{E} [Y | X_1, X_2]]$$

- ▶ See the proof in Section 2.33 in Hansen (2022) using Jensen’s inequality.

Best Predictor

Suppose that given a random vector X , we want to predict or forecast Y .

- ▶ We can write any predictor as a function $g(X)$ of X . The (ex-post) prediction error is the realized difference $Y - g(X)$.
- ▶ A non-stochastic measure of the magnitude of the prediction error is the expectation of its square

$$\mathbb{E} [(Y - g(X))^2]$$

- ▶ We can define the **best predictor** as the function $g(X)$ which **minimizes** the above equation.

Best Predictor

The CEF is the best predictor.

Prove:

$$\begin{aligned}\mathbb{E}[(Y - g(X))^2] &= \mathbb{E}[(e + m(X) - g(X))^2] && \text{CEF 3: } \mathbb{E}[h(x)e] = 0 \\ &= \mathbb{E}[e^2] + 2\mathbb{E}[e(m(X) - g(X))] + \mathbb{E}[(m(X) - g(X))^2] \\ &= \mathbb{E}[e^2] + \mathbb{E}[(m(X) - g(X))^2] \\ &\geq \mathbb{E}[e^2] \\ &= \mathbb{E}[(Y - m(X))^2]\end{aligned}$$

- It means that the CEF minimizes the squared error among all the functions $g(X)$.

Conditional Variance

If $\mathbb{E}[Y^2] < \infty$, the **conditional variance** of Y given $X = x$ is

$$\sigma^2(x) = \text{var}[Y \mid X = x] = \mathbb{E}[(Y - \mathbb{E}[Y \mid X = x])^2 \mid X = x]$$

$$\text{var}(Y) = \mathbb{E}[Y - \mathbb{E}(Y)]^2, \text{ and } \mathbb{E}(Y) = \mathbb{E}[Y|X=x]$$

- ▶ The conditional variance is a function of the conditioning variables.
- ▶ It can be easily seen that the conditional variance of Y is exactly the conditional variance of the CEF error.

$$\sigma^2(x) = \mathbb{E}[e^2 \mid X = x] = \text{var}[e \mid X = x]$$

- ▶ Relate to unconditional error variance: 需自行推导

$$\sigma^2 = \mathbb{E}[e^2] = \mathbb{E}[\mathbb{E}[e^2 \mid X]] = \mathbb{E}[\sigma^2(X)]$$

Homoskedasticity and Heteroskedasticity

最佳线性无偏估计量：同方差、不完全共线、随机抽样等

- ▶ The error is homoskedastic if $\sigma^2(x) = \sigma^2$ does not depend on x .
- ▶ The error is heteroskedasticity if $\sigma^2(x)$ depend on x . 异方差
- ▶ **Heteroskedasticity is generic and “standard”, while homoskedasticity is unusual and exceptional.**

Linear CEF

The CEF can be both linear or non-linear in x .

- ▶ An important special case is when the CEF is linear in x :

$$m(x) = x_1\beta_1 + x_2\beta_2 + \cdots + \beta_k = x'\beta$$

where

$$\beta = \begin{pmatrix} \beta_1 \\ \vdots \\ \beta_k \end{pmatrix}$$

- ▶ It is called the **linear CEF model**, **linear regression model**, or the regression of Y on X .

Linear CEF with Nonlinear Effects

$$m(x_1, x_2) = x_1\beta_1 + x_2\beta_2 + x_1^2\beta_3 + x_2^2\beta_4 + x_1x_2\beta_5 + \beta_6$$

- ▶ This equation is quadratic in the regressors (x_1, x_2) , but we still call it a linear CEF because it is a linear function of the coefficients.
- ▶ Define $x_3 = x_1^2$, $x_4 = x_2^2$, $x_5 = x_1x_2$, and $x_6 = 1$, and transform the equation into a common and linear one.

$$\frac{\partial}{\partial x_1} m(x_1, x_2) = \beta_1 + 2x_1\beta_3 + x_2\beta_5$$

- ▶ We typically call β_5 the interaction effect.

Best Linear Predictor

While the conditional expectation $m(X) = \mathbb{E}[Y | X]$ is the best predictor of Y among all functions of X , its functional form is typically unknown. (It is rarely linear.)

- ▶ Let's find a linear approximation.
- ▶ Assumptions:
 1. $\mathbb{E}[Y^2] < \infty$
 2. $\mathbb{E}\|X\|^2 < \infty$
 3. $\mathbf{Q}_{XX} = \mathbb{E}[XX']$ is positive definite

Best Linear Predictor

Define the **Best Linear Predictor or Projection** of Y given X as

$$\mathcal{P}[Y | X] = X'\beta$$

where β minimizes the mean squared prediction error

$$S(\beta) = \mathbb{E} \left[(Y - X'\beta)^2 \right]$$

The minimizer

not vector

$$\beta = \underset{b \in \mathbb{R}^k}{\operatorname{argmin}} S(b)$$

is called the **Linear Projection Coefficient**.

Best Linear Predictor

We now calculate an explicit expression for its value.

$$S(\beta) = \mathbb{E}[Y^2] - 2\beta' \mathbb{E}[XY] + \beta' \mathbb{E}[XX'] \beta$$

The first-order condition for minimization is dim = 1

$$0 = \frac{\partial}{\partial \beta} S(\beta) = -2\mathbb{E}[XY] + 2\mathbb{E}[XX'] \beta$$

$[\mathbb{E}(XX')]^T$, 不管对左右 β

Then we have

$$\beta = \mathbb{E}[XX']^{-1} \mathbb{E}[XY]$$

Sometimes we define $\mathbf{Q}_{XY} = \mathbb{E}[XY]$ and $\mathbf{Q}_{XX} = \mathbb{E}[XX']$. So

$$\beta = \mathbf{Q}_{XX}^{-1} \mathbf{Q}_{XY}$$

Best Linear Predictor

We now have an explicit expression for the best linear predictor:

$$\mathcal{P}[Y | X] = X'\beta = X'\mathbb{E}[XX']^{-1}\mathbb{E}[XY]$$

Define the **projection error** as

$$e = Y - X'\beta$$

An important property of the projection error e is

$$\mathbb{E}[Xe] = 0$$

proof:

$$\begin{aligned} &= \mathbb{E}[X(Y - X'\beta)] \\ &= \mathbb{E}[XY] - \mathbb{E}[XX']^{-1}\mathbb{E}[XY]\mathbb{E}[XY] \\ &= 0 \end{aligned}$$

- ▶ It is equivalent to $\mathbb{E}[X_j e] = 0$ for $j = 1, \dots, k$.
- ▶ If the X contains a constant, then $\mathbb{E}[e] = 0$.

$Y = \beta_0 + \sum_i \beta_i x_i + e$, where β_0 can be regarded as $\beta_0 x_0$, $x_0 = 1$. If $\mathbb{E}[x_0 e] = 0$, $\mathbb{E}[e] = 0$

不具有常数项的回归方程穿过0，且 R^2 可能为负？

Linear Predictor Error Variance

As in the CEF model, we define the error variance as

$$\sigma^2 = \mathbb{E} [e^2]$$

Setting $Q_{YY} = \mathbb{E} [Y^2]$ and $\mathbf{Q}_{YX} = \mathbb{E} [YX']$ we have

$$\begin{aligned}\sigma^2 &= \mathbb{E} \left[(Y - X'\beta)^2 \right] \\ &= \mathbb{E} [Y^2] - 2\mathbb{E} [YX'] \beta + \beta' \mathbb{E} [XX'] \beta \\ &= Q_{YY} - 2\mathbf{Q}_{YX} \mathbf{Q}_{XX}^{-1} \mathbf{Q}_{XY} + \mathbf{Q}_{YX} \mathbf{Q}_{XX}^{-1} \mathbf{Q}_{XX} \mathbf{Q}_{XX}^{-1} \mathbf{Q}_{XY} \\ &= Q_{YY} - \mathbf{Q}_{YX} \mathbf{Q}_{XX}^{-1} \mathbf{Q}_{XY} \\ &\stackrel{\text{def}}{=} Q_{YY \cdot X}.\end{aligned}$$

Regression Coefficients with Intercept

Sometimes it is useful to separate the constant from the other regressors

$$Y = X'\beta + \alpha + e$$

where α is the intercept and X does not contain a constant.

Taking expectations of this equation, we find

$$\mathbb{E}[Y] = \mathbb{E}[X'\beta] + \mathbb{E}[\alpha] + \mathbb{E}[e] \quad \mathbb{E}[e]=0$$

Let $\mu_Y = \mathbb{E}[Y]$ and $\mu_X = \mathbb{E}[X]$, then we have

同上所证

$$Y - \mu_Y = (X - \mu_X)'\beta + e$$

$$\begin{aligned}\beta &= (\mathbb{E}[(X - \mu_X)(X - \mu_X)'])^{-1} \mathbb{E}[(X - \mu_X)(Y - \mu_Y)] \\ &= \text{var}[X]^{-1} \text{cov}(X, Y)\end{aligned}$$

$$\alpha = \mu_Y - \mu_X'\beta$$

Omitted Variable Bias

Consider two projections of Y :

1. $Y = X_1'\beta_1 + X_2'\beta_2 + e$, $\mathbb{E}[Xe] = 0$, called long regression
2. $Y = X_1'\gamma_1 + u$, $\mathbb{E}[X_1u] = 0$, called short regression

The relation between γ_1 and β_1 is

$$\begin{aligned}\gamma_1 &= (\mathbb{E}[X_1X_1'])^{-1} \mathbb{E}[X_1Y] \\ &= (\mathbb{E}[X_1X_1'])^{-1} \mathbb{E}[X_1(X_1'\beta_1 + X_2'\beta_2 + e)] \\ &= \beta_1 + (\mathbb{E}[X_1X_1'])^{-1} \mathbb{E}[X_1X_2'] \beta_2 \\ &= \beta_1 + \Gamma_{12}\beta_2\end{aligned}$$

where $\Gamma_{12} = \mathbf{Q}_{11}^{-1}\mathbf{Q}_{12}$ is the coefficient matrix from a projection of X_2 on X_1 .

Omitted Variable Bias

- ▶ Observe that $\gamma_1 = \beta_1 + \Gamma_{12}\beta_2 \neq \beta_1$ unless $\Gamma_{12} = 0$ (X_1 and X_2 are uncorrelated) or $\beta_2 = 0$.
- ▶ The difference $\Gamma_{12}\beta_2$ between γ_1 and β_1 is known as **omitted variable bias**.

Best Linear Approximation

前文被跳过了

The best linear predictor $X'\beta$ is also the best linear approximation of the CEF $m(X)$.

Define the mean-square approximation error

$$d(\beta) = \mathbb{E} \left[(m(X) - X'\beta)^2 \right]$$

Select β to minimize $d(\beta)$. Then we have

$$\begin{aligned} \beta &= (\mathbb{E} [XX'])^{-1} \mathbb{E}[Xm(X)] \\ &= (\mathbb{E} [XX'])^{-1} \mathbb{E}[X(m(X) + e)] \\ &= (\mathbb{E} [XX'])^{-1} \mathbb{E}[XY] \end{aligned}$$

区别：
 $\mathbb{E}[(Y - X'\beta)^2]$

Limitations of the Best Linear Projection

- ▶ The linear CEF means $\mathbb{E}[e|X] = 0$, so apparently $\mathbb{E}[Xe] = 0$.
- ▶ The best linear predictor means $\mathbb{E}[Xe] = 0$, but it may not be true that $\mathbb{E}[e|X] = 0$.

Furthermore, the linear projection may be a poor approximation to the CEF.

- ▶ Consider the true process is $Y = X + X^2$ with $X \sim N(0, 1)$. Then the CEF is $m(X) = X + X^2$. when have a constant, $\beta = \text{var}(X)^{-1} \text{cov}(X, Y)$
- ▶ Now consider project Y on X and a constant: $Y = \beta X + \alpha + e$. Then we have $\beta = 1$ and $\alpha = 1$. So $\mathcal{P}[Y | X] = X + 1$.
- ▶ The projection error is $e = X^2 - 1$, which increases infinitely as X increases.

can CEF $\mathbb{E}[e|x]=0$ demonstrate BLP $\mathbb{E}[xe]=0$?
consider $e=x^2-1, x \sim N(0,1), \mathbb{E}[xe]=0$ while $\mathbb{E}[e|x]=x^2-1$