

RUCStatBeamer—Typst Template

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Outline

Section A

Section B

Appendix



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Section A

Section B

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How to choose a threshold?

- Control Per-Comparison Type I Error (**PCER**)
 1. a.k.a. “*uncorrected testing*”, many type I errors
 2. Gives $\mathbb{P}\{\text{FD}_i > 0\} \leq \alpha$ marginally for all $1 \leq i \leq m$
- Control Familywise Type I Error (**FWER**)
 1. e.g. Bonferroni method, or using per-comparison significance level $\frac{\alpha}{m}$
 2. Guarantees $\mathbb{P}\{\text{FD} > 0\} \leq \alpha$
- Control False Discovery Rate (**FDR**)
 1. First defined by Benjamini & Hochberg [1]
 2. Guarantees $\text{FDR} \equiv \mathbb{E}\left(\frac{\text{FD}}{D}\right) \leq \alpha$

BH Procedure

Theorem 1

The Benjamini–Hochberg procedure (BH step-up procedure) controls the FDR at level α for independent multiple tests.

Visualization and Algorithm

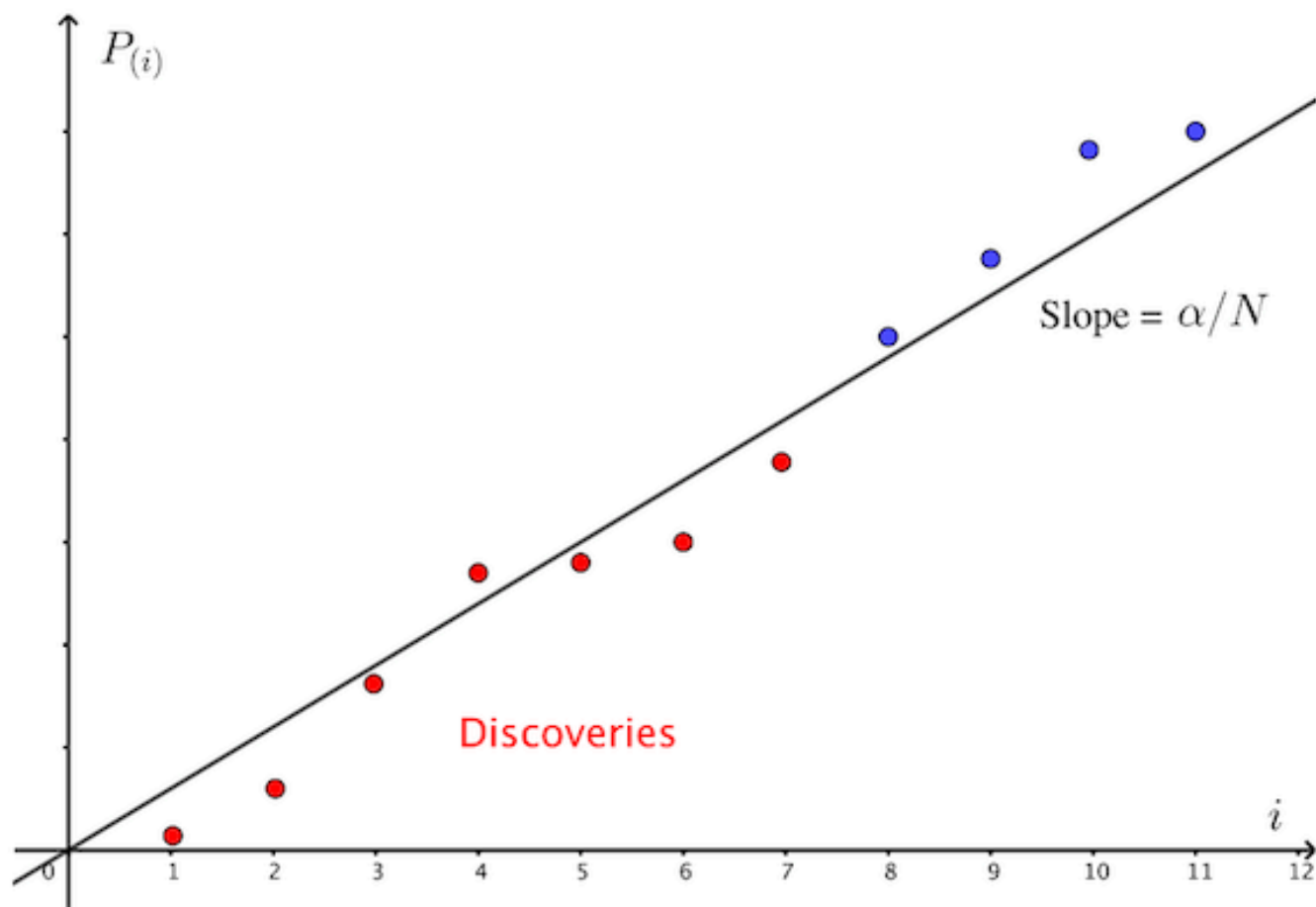


Figure 1: BH Procedure

Visualization and Algorithm

1. For a given α , find the largest k such that $P_{(k)} \leq \frac{k}{m} \frac{k}{m} \alpha$.
2. Reject the null hypothesis (i.e., declare discoveries) for all $H_{(i)}$ for $i = 1, \dots, k$.

Outline

Section A

Section B

Appendix

R Code

```

bh <- function() {
  UseMethod("bh")
}

bh.func <- function(pv, alpha =
0.05) {
  m <- length(pv)
  i <- 1:m
  sorted_pv <- sort(pv)
  if (sorted_pv[1] > alpha / m) {
    return(rep(0, m))
  }
  k <- max(i[sorted_pv <= i / m
* alpha])
  criterion <- sorted_pv[k]
  return(1 * (pv <= criterion))
}

```

Lorem ipsum dolor sit amet, consectetur adipiscing elit, sed do eiusmod tempor incididunt ut labore et dolore magnam aliquam quaerat voluptatem. Ut enim aequaleamur animo, cum corpore dolemus, fieri.

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P P-value, or realization of a P-variable.

E E-value, or realization of a E-variable.

Q Q-value, or realization of a Q-variable.

Subsection B.2

You can use `#pause` to pause display some proof. 🤔

and this completes the proof.



Subsection B.2

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Proof: Let $\alpha_r = \alpha r/K$ for $r \in \mathcal{K}$. We can write

$$\mathbb{E} \left[\frac{F_{\mathcal{D}}}{R_{\mathcal{D}}} \right] = \mathbb{E} \left[\frac{\sum_{k \in \mathcal{N}} \mathbb{1}_{\{P_k \leq \alpha_{R_{\mathcal{D}}}\}}}{R_{\mathcal{D}}} \right] = \sum_{k \in \mathcal{N}} \sum_{r=1}^K \frac{1}{r} \mathbb{E} \left[\mathbb{1}_{\{P_k \leq \alpha_r\}} \mathbb{1}_{\{R_{\mathcal{D}}=r\}} \right]. \quad (1)$$

and this completes the proof. □

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For $k \in \mathcal{N}$, let R_k be the number of rejection from the BH procedure if it is applied to \mathbf{P} with P_k replaced by 0. Note that $\{P_k \leq \alpha_r, R_{\mathcal{D}} = r\} = \{P_k \leq \alpha_r, R_k = r\}$ for each k, r . Hence, we have

$$\mathbb{E} \left[\mathbb{1}_{\{P_k \leq \alpha_r\}} \mathbb{1}_{\{R_{\mathcal{D}}=r\}} \right] = \mathbb{E} \left[\mathbb{1}_{\{P_k \leq \alpha_r\}} \mathbb{1}_{\{R_k=r\}} \right]. \quad (2)$$

and this completes the proof. □

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Proof: Let $\alpha_r = \alpha r / K$ for $r \in \mathcal{K}$. We can write

$$\mathbb{E} \left[\frac{F_{\mathcal{D}}}{R_{\mathcal{D}}} \right] = \mathbb{E} \left[\frac{\sum_{k \in \mathcal{N}} \mathbb{1}_{\{P_k \leq \alpha_{R_{\mathcal{D}}}\}}}{R_{\mathcal{D}}} \right] = \sum_{k \in \mathcal{N}} \sum_{r=1}^K \frac{1}{r} \mathbb{E} \left[\mathbb{1}_{\{P_k \leq \alpha_r\}} \mathbb{1}_{\{R_{\mathcal{D}}=r\}} \right]. \quad (1)$$

For $k \in \mathcal{N}$, let R_k be the number of rejection from the BH procedure if it is applied to \mathbf{P} with P_k replaced by 0. Note that $\{P_k \leq \alpha_r, R_{\mathcal{D}} = r\} = \{P_k \leq \alpha_r, R_k = r\}$ for each k, r . Hence, we have

$$\mathbb{E} \left[\mathbb{1}_{\{P_k \leq \alpha_r\}} \mathbb{1}_{\{R_{\mathcal{D}}=r\}} \right] = \mathbb{E} \left[\mathbb{1}_{\{P_k \leq \alpha_r\}} \mathbb{1}_{\{R_k=r\}} \right]. \quad (2)$$

Putting this into Equation 1, we get

$$\mathbb{E} \left[\frac{F_{\mathcal{D}}}{R_{\mathcal{D}}} \right] = \frac{\alpha}{K} \sum_{k \in \mathcal{N}} \sum_{r=1}^K \mathbb{P}(R_k = r) = \frac{K_0 \alpha}{K},$$

and this completes the proof. □



Outline

Section A

Section B

Appendix

Appendix

- [1] Benjamini Y, Hochberg Y. Controlling the false discovery rate: a practical and powerful approach to multiple testing. *Journal of the Royal statistical society: series B (Methodological)*. 1995;57(1):289–300.
- [2] Bonferroni C. *Teoria statistica delle classi e calcolo delle probabilita*. Pubblicazioni del R Istituto Superiore di Scienze Economiche e Commerciali di Firenze. 1936;8:3–62.
- [3] Dwork C, McSherry F, Nissim K, et al. Calibrating noise to sensitivity in private data analysis. In *Theory of Cryptography: Third Theory of Cryptography Conference, TCC 2006, New York, NY, USA, March 4-7, 2006 Proceedings* 3. 2006. pp. 265–284.
- [4] Holm S. A simple sequentially rejective multiple test procedure. *Scandinavian Journal of Statistics*. 1979;6(2):65–70.
- [5] Hochberg Y. A sharper Bonferroni procedure for multiple tests of significance. *Biometrika*. 1988;75(4):800–802.