

高斯混合模型 (GMM)

1. 模型

$$y_1, y_2, \dots, y_k,$$

$$P(y|\theta) = \sum_{k=1}^K \alpha_k \Phi(y|\theta_k) \quad \forall k \geq 0, \sum_k \alpha_k = 1, \theta_k = (\mu_k, \sigma_k^2), \Phi(y|\theta_k) = \frac{1}{\sqrt{2\pi}\sigma_k} \exp\left(-\frac{(y-\mu_k)^2}{2\sigma_k^2}\right)$$

$y_j \rightarrow$ ① 依据 α_k 选择第 k 个高斯分布 $\Phi(y|\theta_k)$

\rightarrow ② 依据 $\Phi(y|\theta_k)$ 选择 y_j

y_i 是已知的 $j=1, \dots, N$

隐变量 $r_{jk} \begin{cases} 1 & \text{第 } j \text{ 个观测来自 } k \\ 0 & \text{otherwise} \end{cases}$
不可观测的

2. 参数估计

2.1 EM 算法

如果模型有观测数据 (observed), 隐变量 (latent)

例 (三硬币模型)

A B C
 π p q

A $\begin{cases} \text{正} & B \\ \text{反} & C \end{cases}$

Result: 1, 1, 0, 1, 0, 0, 1, 0, 1, 1

$$P(y|\theta) = \sum_z P(y, z|\theta)$$

$$= \sum_z P(z|\theta) P(y|z, \theta) = \pi \cdot p^y (1-p)^{1-y} + (1-\pi) q^y (1-q)^{1-y}$$

$$P(Y|\theta) = \prod_{j=1}^n [\pi p^{y_j} (1-p)^{1-y_j} + (1-\pi) q^{y_j} (1-q)^{1-y_j}]$$

一般用 EM 算法求解

$$y^{(t)} = (\pi^{(t)}, p^{(t)}, q^{(t)})$$

E步: 计算在给定模型参数的情形下, y_j 来自 B 的概率

Expectation

$$P(z_j=1 | y_j, \theta) = \frac{P(z_j, y_j | \theta)}{\sum_j P(z_j, y_j | \theta)}$$

$$\mu^{(i+1)} = E(z_j) = \frac{\pi^{(i)} (p^{(i)})^{y_j} (1-p^{(i)})^{1-y_j}}{\pi^{(i)} (p^{(i)})^{y_j} (1-p^{(i)})^{1-y_j} + (1-\pi^{(i)}) (q^{(i)})^{y_j} (1-q^{(i)})^{1-y_j}}$$

M步: $P(Y, Z | \theta) = \prod_{j=1}^n [\pi p^{y_j} (1-p)^{1-y_j}]^{z_j} [(1-\pi) q^{y_j} (1-q)^{1-y_j}]^{1-z_j}$

Maximize

$$\pi^{(i+1)} = \frac{1}{n} \sum_j \mu_j^{(i+1)}$$

$$p^{(i+1)} = \frac{\sum_j \mu_j^{(i+1)} y_j}{\sum_j \mu_j^{(i+1)}}, \quad q^{(i+1)} = \frac{\sum_j (1-\mu_j^{(i+1)}) y_j}{\sum_j (1-\mu_j^{(i+1)})}$$

· Y : 观测数据 \rightarrow 完全数据 complete data
 Z : 隐变量

EM算法: 输入 $Y, P(Y, Z | \theta), P(Z | \theta)$
 输出 θ

PseudoCode: (1) 选择初始值 $\theta^{(0)}$

(2) E步: $\theta^{(i)}$ 是第 i 步的估计值

求条件概率 $P(Z | \theta, Y) = \mu^{(i)}$

(3) M步: $Q(\theta, \theta^{(i)}) = E_z \{ \log(P(Y, Z | \theta)) | Y, \theta^{(i)} \}$

$$= \sum_z \log P(Y, Z | \theta) P(Z | Y, \theta^{(i)})$$

$$\text{求 } \theta^{(i+1)} = \arg \max_{\theta} Q(\theta, \theta^{(i)})$$

2.2 EM算法的导出

① 为什么EM算法可以实现对数据的极大似然估计?

$$\begin{aligned}\text{似然函数: } L(\theta) &= \log P(Y|\theta) = \log \left\{ \sum_z P(Y, z|\theta) \right\} \\ &= \log \left\{ \sum_z P(Y|z, \theta) P(z|\theta) \right\}\end{aligned}$$

$\theta^{(i)}$ 希望 θ 增加 $L(\theta)$

$$\begin{aligned}L(\theta) - L(\theta^{(i)}) &= \log \left(\sum_z P(Y|z, \theta) P(z|\theta) \right) - \log P(Y|\theta^{(i)}) \\ &= \log \left(\sum_z P(z|Y, \theta^{(i)}) \frac{P(Y|z, \theta) P(z|\theta)}{P(z|Y, \theta^{(i)})} \right) - \log P(Y|\theta^{(i)}) \\ &\geq \sum_z P(z|Y, \theta^{(i)}) \log \frac{P(Y|z, \theta) P(z|\theta)}{P(z|Y, \theta^{(i)})} - \log P(Y|\theta^{(i)}) \\ &= \sum_z P(z|Y, \theta^{(i)}) \log \frac{P(Y|z, \theta) P(z|\theta)}{P(z|Y, \theta^{(i)}) P(Y|\theta^{(i)})}\end{aligned}$$

$\log \sum \lambda_j y_j \geq \sum \lambda_j \log y_j$
if $\lambda_j \geq 0, \sum \lambda_j = 1$

$$\therefore B(\theta, \theta^{(i)}) \triangleq L(\theta^{(i)}) + \sum_z P(z|Y, \theta^{(i)}) \log \frac{P(Y|z, \theta) P(z|\theta)}{P(z|Y, \theta^{(i)}) P(Y|\theta^{(i)})}$$

$$\therefore L(\theta) \geq \underbrace{B(\theta, \theta^{(i)})}_{\text{下界}}, \text{ 选择 } \theta^{(i+1)} = \arg \max_{\theta} B(\theta, \theta^{(i)})$$

$$\begin{aligned}\therefore \theta^{(i+1)} &= \arg \max_{\theta} \left\{ L(\theta^{(i)}) + \sum_z P(z|Y, \theta^{(i)}) \log \frac{P(Y|z, \theta) P(z|\theta)}{P(z|Y, \theta^{(i)}) P(Y|\theta^{(i)})} \right\} \\ &= \arg \max_{\theta} \left\{ \sum_z P(z|Y, \theta^{(i)}) \log P(Y|z, \theta) P(z|\theta) \right\} \\ &= \arg \max_{\theta} \left\{ \sum_z P(z|Y, \theta^{(i)}) \log P(Y, z|\theta) \right\} \\ &= \arg \max_{\theta} \{ Q(\theta, \theta^{(i)}) \}\end{aligned}$$

(1) EM对初值敏感

(2) 在迭代中使 $L(\theta) \uparrow$

(3) 在一定条件下可以保证收敛, 但不一定收敛到极大值点

2.3 通过 EM 算法求解 GMM

(1) 明确隐变量, 写出完全数据的对数似然 $(y_j, r_{j1}, r_{j2}, \dots, r_{jk}) \quad j=1, 2, \dots, N$

$$\begin{aligned} P(y, r | \theta) &= \prod_{j=1}^N P(y_j, r_{j1}, r_{j2}, \dots, r_{jk} | \theta) \\ &= \prod_{j=1}^N \prod_{k=1}^K [\alpha_k \Phi(y_k | \theta_k)]^{r_{jk}} \\ &= \prod_{k=1}^K \alpha_k^{r_k} \prod_{j=1}^N [\Phi(y_k | \theta_k)]^{r_{jk}}, \quad r_k = \sum_j r_{jk} \\ &= \prod_{k=1}^K \alpha_k^{r_k} \prod_{j=1}^N \left[\frac{1}{\sqrt{2\pi} \sigma_k} \exp\left(-\frac{(y_j - \mu_k)^2}{2\sigma_k^2}\right) \right]^{r_{jk}} \end{aligned}$$

$$\log P(y, r | \theta) = \sum_{k=1}^K \left\{ r_k \log \alpha_k + \sum_{j=1}^N r_{jk} \left[\log\left(\frac{1}{\sqrt{2\pi}}\right) - \log \sigma_k - \frac{1}{2\sigma_k^2} (y_j - \mu_k)^2 \right] \right\}$$

$$\hat{r}_{jk} = E(r_{jk} | y, \theta) = P(r_{jk}=1 | y_j, \theta) \quad \leftarrow \text{子模型 } k \text{ 对于 } y_j \text{ 的响应度}$$

$$\begin{aligned} &= \frac{P(r_{jk}=1, y_j | \theta)}{\sum_{k=1}^K P(r_{jk}, y_j | \theta)} = \frac{P(y_j | r_{jk}=1, \theta) P(r_{jk}=1 | \theta)}{\sum_k P(y_j | r_{jk}=1, \theta) P(r_{jk}=1 | \theta)} \\ &= \frac{\alpha_k \cdot \Phi(y_j | \theta_k)}{\sum_k \alpha_k \cdot \Phi(y_j | \theta_k)} \end{aligned}$$

$$Q(\theta, \theta^{(i)}) = \sum_{k=1}^K \left\{ \sum_{j=1}^N \hat{r}_{jk} \log \alpha_k + \sum_{j=1}^N \hat{r}_{jk} \left[\log \frac{1}{\sqrt{2\pi}} - \log \sigma_k - \frac{1}{2\sigma_k^2} (y_j - \mu_k)^2 \right] \right\}$$

$$\theta^{(i+1)} = \underset{\theta}{\operatorname{argmax}} Q(\theta, \theta^{(i)}), \quad \text{s.t.} \sum_{k=1}^K \alpha_k = 1$$

$$\Rightarrow \hat{\mu}_k = \frac{\sum_{j=1}^N \hat{r}_{jk} y_j}{\sum_{j=1}^N \hat{r}_{jk}}, \quad \hat{\sigma}_k^2 = \frac{\sum_j \hat{r}_{jk} (y_j - \mu_k)^2}{\sum_j \hat{r}_{jk}}, \quad \hat{\alpha}_k = \frac{\sum_{j=1}^N \hat{r}_{jk}}{N}$$

GMM 的 EM 算法:

(1) 选择初值开始迭代

$$(2) \text{求响应度 } \hat{r}_{jk} = \frac{\alpha_k^{(i)} \Phi(y_j | \theta_k^{(i)})}{\sum_{k=1}^K \alpha_k^{(i)} \Phi(y_j | \theta_k^{(i)})}$$

$$(3) M \text{ 步: } \theta^{(i+1)} = \underset{\theta}{\operatorname{argmax}} Q(\theta, \theta^{(i)})$$

2.4 EM算法的收敛性

定理 9.1: 设 $P(Y|\theta)$ 为观测数据的似然函数, $\theta^{(i)} (i=1, 2, \dots)$ 为 EM 算法得到的参数估计序列, $P(Y|\theta^{(i)}) (i=1, 2, \dots)$ 为对应的似然函数序列, 则 $P(Y|\theta^{(i)})$ 是单调递增的, 即 $P(Y|\theta^{(i+1)}) \geq P(Y|\theta^{(i)})$

$$\text{证明: } \because P(Y|\theta) = \frac{P(Y, Z|\theta)}{P(Z|Y, \theta)}$$
$$\log P(Y|\theta) = \log P(Y, Z|\theta) - \log P(Z|Y, \theta)$$

$$Q(\theta, \theta^{(i)}) = \sum_Z \log P(Y, Z|\theta) \cdot P(Z|Y, \theta^{(i)})$$

$$\text{令 } H(\theta, \theta^{(i)}) = \sum_Z \log P(Z|Y, \theta) \cdot P(Z|Y, \theta^{(i)})$$

$$\therefore \log P(Y|\theta) = Q(\theta, \theta^{(i)}) - H(\theta, \theta^{(i)})$$

$$\therefore \log P(Y|\theta^{(i+1)}) - \log P(Y|\theta^{(i)}) = [Q(\theta^{(i+1)}, \theta^{(i)}) - Q(\theta^{(i)}, \theta^{(i)})]$$
$$- [H(\theta^{(i+1)}, \theta^{(i)}) - H(\theta^{(i)}, \theta^{(i)})]$$

$$\text{又: } H(\theta^{(i+1)}, \theta^{(i)}) - H(\theta^{(i)}, \theta^{(i)}) = \sum_Z \left(\log \frac{P(Z|Y, \theta^{(i+1)})}{P(Z|Y, \theta^{(i)})} \cdot P(Z|Y, \theta^{(i)}) \right)$$
$$\leq \log \left(\sum_Z \frac{P(Z|Y, \theta^{(i+1)})}{P(Z|Y, \theta^{(i)})} \cdot P(Z|Y, \theta^{(i)}) \right)$$
$$= 0$$

$$Q(\theta^{(i+1)}, \theta^{(i)}) = \max_{\theta} Q(\theta, \theta^{(i)}) \geq Q(\theta^{(i)}, \theta^{(i)})$$

$$\therefore \log P(Y|\theta^{(i+1)}) - \log P(Y|\theta^{(i)}) \geq 0$$

$P(Y|\theta^{(i+1)}) \geq P(Y|\theta^{(i)})$ 才保证了收敛性