

Linear Regression

1. Linear Regression Model

1.1 Model & Notations

$$x_i = (x_{i1}, x_{i2}, \dots, x_{ip})^T \in \mathbb{R}^p$$

$$y = (y_1, y_2, \dots, y_N)^T, X = (x_1, x_2, \dots, x_N)^T \in \mathbb{R}^{N \times p}, \varepsilon = (\varepsilon_1, \varepsilon_2, \dots, \varepsilon_N)^T \in \mathbb{R}^N$$

$$y = X\beta + \varepsilon$$

remarks:

1. Quantitative inputs & its transformations (log, squares) & basis expansions ($X_2 = X_1^2, X_3 = X_1^3$)
2. Qualitative inputs: dummy variable coding
3. Interaction between variables ($X_3 = X_1 \cdot X_2$)

1.2 Model Assumptions

(A1) The relationship between response y and covariates X is linear

(A2) X is non-stochastic matrix and $\text{rank}(X) = p$

(A3) $E(\varepsilon) = 0$. This implies $E(y) = X\beta$

(A4) $\text{cov}(\varepsilon) = E(\varepsilon\varepsilon^T) = \sigma^2 I_N$

(A5) ε follows multivariate normal distribution $N(0, \sigma^2 I_N)$

or

(A2*) X is a full rank matrix with probability 1 ($\lambda_{\min}(X^T X) \rightarrow \infty$ a.s.)

(A3*) $E(\varepsilon|X) = 0$

(A4*) $E(\varepsilon\varepsilon^T|X) = \sigma^2 I_N$

(A5*) $\varepsilon|X \sim N(0, \sigma^2 I_N)$

2. Model Estimation

• OLS estimation:
$$RSS(\beta) = \sum_{i=1}^N \{y_i - f(x_i)\}^2 = \sum_{i=1}^N \{y_i - \beta_0 - \sum_j x_{ij} \beta_j\}^2$$
$$= (y - X\beta)^T (y - X\beta)$$

This criterion is valid if y_i 's are conditionally independent given inputs x_i

$$\frac{\partial RSS(\beta)}{\partial \beta} = -2(y - X\beta)^T X = 0, \quad \hat{\beta} = (X^T X)^{-1} X^T y, \quad \hat{y} = X(X^T X)^{-1} X^T y = Hy$$

1. Assume X is full rank, hence $X^T X$ is positive definite

2. Fitted Values: $\hat{y} = Hy$. Residual vector $y - \hat{y}$ is orthogonal to the column space of X

3. The residual sum of squares $RSS(\beta)$ can be used as a goodness-of-fit measure.

3. Statistical Inference

3.1 Mean and Variance of the OLS Estimator

$$E(\hat{\beta}) = \beta, \quad \text{Cov}(\hat{\beta}) = (X^T X)^{-1} \sigma^2$$

$$\hat{\sigma}^2 = \frac{1}{N-p} \sum_{i=1}^N (y_i - \hat{y}_i)^2$$

Theorem 1: (Gauss Markov Theorem) Assume $(A_1) \sim (A_4)$. Then $\hat{\beta}$ is the best linear unbiased estimator (BLUE), provided it exists.

It implies, $\hat{\beta}$ has the smallest variance over all linear unbiased estimator $\tilde{\beta}$, i.e. $\tilde{\beta} = \sum_{i=1}^N w_i y_i$ and $E(\hat{\beta}) = \beta, \forall \eta \in \mathbb{R}^p, \|\eta\|=1, \text{Var}(\eta^T \hat{\beta}) \leq \text{Var}(\eta^T \tilde{\beta})$

proof: $\eta^T \hat{\beta} = \underbrace{\eta^T (X^T X)^{-1} X^T}_{\alpha'} y \quad E(\eta^T \hat{\beta}) = \eta^T (X^T X)^{-1} X^T E(y) = \eta^T \beta \quad \text{Cov}(\eta^T \hat{\beta}) = \alpha' \alpha \cdot \sigma^2$

Alternative estimator: $d^T Y, E(d^T Y) = (X^T d)' \beta \quad \therefore X^T d = \eta$

$$\begin{aligned} \text{Cov}(d^T Y) &= d^T d \cdot \sigma^2 = \sigma^2 (d - a + a)^T (d - a + a) \quad a'(d - a) = \eta^T (X^T X)^{-1} X^T d - \eta^T (X^T X)^{-1} \eta \\ &= \sigma^2 (d^T a + (d - a)^T (d - a) + 2a^T (d - a)) \quad = \eta^T (X^T X)^{-1} \eta - \eta^T (X^T X)^{-1} \eta \\ &\geq \sigma^2 a^T a \quad = 0 \end{aligned}$$

property 1: $(N-p)\hat{\sigma}^2 \sim \sigma^2 \chi_{N-p}^2$

proof $RSS = (y - \hat{y})'(y - \hat{y}) = y'(I - H)y$

$$\text{tr}(I - H) = \text{tr}(I) - \text{tr}(X(X'X)^{-1}X') = N - p$$

$$\exists U, \text{ s.t. } U'(I - H)U = \begin{pmatrix} 1 & & \\ & \ddots & \\ & & 0 \end{pmatrix}$$

$$RSS = (U'y)' D (U'y) \sim \chi_{N-p}^2$$

property 2: $\sum_{j=1}^n (y_j - \bar{y})^2 = \sum_{j=1}^n (\hat{y}_j - \bar{y})^2 + \sum_{j=1}^n (y_j - \hat{y}_j)^2$

proof $\sum_{j=1}^n (y_j - \bar{y})^2 = \sum_{j=1}^n (y_j - \hat{y}_j + \hat{y}_j - \bar{y})^2$

$$= \sum_{j=1}^n (y_j - \hat{y}_j)^2 + \sum_{j=1}^n (\hat{y}_j - \bar{y})^2 + 2 \sum_{j=1}^n (\hat{y}_j - \bar{y})(y_j - \hat{y}_j)$$

$$\sum_{j=1}^n (\hat{y}_j - \bar{y})^T (\hat{y}_j - y_j) = y^T (H - \frac{1}{n} 11^T) (I - H) y$$

$$= y^T (\frac{1}{n} 11^T - \frac{1}{n} 11^T) y = 0$$

3.2 Sampling property

• $\hat{\beta} \sim N(\beta, (X^T X)^{-1} \sigma^2)$, $(N-p)\hat{\sigma}^2 \sim \sigma^2 \chi_{N-p}^2$

• $z_j = \frac{\hat{\beta}_j}{\hat{\sigma} \sqrt{v_j}} \sim t_{N-p}$

• $F = \frac{(RSS_0 - RSS_1)/(p_1 - p_0)}{RSS_1/(N - p_1)} \sim F(p_1 - p_0, N - p_1)$

4. Goodness-of-fit

• $R^2 = 1 - \frac{RSS}{TSS}$, Adjusted $R^2 = 1 - \frac{RSS/(n-p)}{TSS/(n-1)}$

5. Model Selection

1. Subset Selection

1.1 Best-subset Selection: time-consuming

1.2 Forward-stepwise selection (greedy algorithm): Starts with the intercept, and then sequentially adds into the model the predictor that most improves the fit

1.3 Backward-stepwise selection: starts with the full model, and sequentially deletes the predictor that has the least impact on the fit. ($N > p$)

1.4 Stepwise selection: consider both forward and backward moves at each step, and select the "best"

$$AIC = -\frac{1}{N} \ln(\beta) + 2 \frac{d}{N}$$

$$BIC = -2 \ln(\beta) + (\log N) d$$

Best model has smallest AIC

comment:

1. BIC can consistently select the true model

2. other criterion including C_p

2. Shrinkage Methods

2.1 Ridge Regression

$$\hat{\beta}^{\text{ridge}} = \arg \min_{\beta} \left\{ \sum_{i=1}^N (y_i - \beta_0 - \sum_j x_{ij} \beta_j)^2 + \lambda \sum_j \beta_j^2 \right\}$$

$$= (X^T X + \lambda I)^{-1} X^T y$$

$$RSS(\lambda) = (y - X\beta)^T (y - X\beta) + \lambda \beta^T \beta$$

2.2 Lasso Regression

$$\hat{\beta}^{\text{Lasso}} = \arg \min_{\beta} \left\{ \sum_{i=1}^N (y_i - \beta_0 - \sum_j x_{ij} \beta_j)^2 + \lambda \sum_j |\beta_j| \right\}$$