Three Essays on Shrinkage Estimation and Model Selection of Linear and Nonlinear Time Series Models

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Overview and Motivation

- Parametric Time Series Models
  - Logistic Smooth Transition Autoregressive Model (LSTAR) (Terasvirta, 1994)
  - Threshold Autoregressive Model (TAR) (Tong, 1990)
  - Autoregressive Moving Average Model (ARMA) (Box and Pierce, 1970)
- Model Differences
  - ARMA = Popularized for Weakly Stationary Time Series
  - LSTAR and TAR = Handling Nonlinear Behavior
    - Changes in Level, Dynamics, and Volatility
    - Alternative Approach to Handling Nonstationary and Cyclical Phenomenon
    - Useful for Clustering and Classification of Realizations

Overview and Motivation

- Purpose
  - Primary Goal is Forecasting
  - Economic Variables: Past Realizations Assist in Prediction (Mann and Wald, 1943)
  - Estimation and Fitting of ARMA (Durbin, 1960)
  - Practical Procedures Developed for Industrial Engineering (Box and Pierce, 1970)
  - Asymmetries Noticed in US Unemployment (Rothman, 1998; Montgomery et al., 1998; Koop and Potter, 1999)

Overview and Motivation

- Model Complexity
  - Determined by Order Parameters
    - Autoregressive (AR)
    - Moving Average (MA)
    - Distributed Lag (DL)
    - Regime-specific Model Orders
  - Frequentist Approach: Select Based on Information Criteria or Forecasting Performance
  - Bayesian Approach: Reversible jump Markov Chain Monte Carlo (RJMCMC) (Green, 1995)
    - AR(P) (Troughton and Godsill, 1997)
    - TAR(P) (Campbell, 2004; Nieto et al., 2013)
    - STAR(P) (Lopes and Salazar, 2006)

Overview and Motivation

- Model Complexity (Cont.)
  - Problem
    - Difficult When Multiple Order Parameters Must Be Chosen
    - Often Leads to Inflexible Representations
    - Overfitting Can Still Occur
  - Solution
    - Intentionally Fix Orders to Be Large
    - Restructure Time Series Model as a High Dimensional Linear Regression
    - Apply Penalized Estimation Methods Aimed at Sparse Models

Introduction: Overview and Motivation

Dissertation Theme

Bayesian Automatic Estimation and Variable Selection Procedures for Flexible Subset Models

Chapter 1: LSTAR Model

- Gaussian LSTAR(P) Model With 2-Regimes
  - Given autoregressive order P, let $\boldsymbol{x}_t' = [y_{t-1}, y_{t-2}, \cdots, y_{t-P}]$, $\boldsymbol{\alpha}' = [\alpha_1, \cdots, \alpha_P]$, and $\boldsymbol{\beta}' = [\beta_1, \cdots, \beta_P]$.

$$y_t = (\mu_\alpha + \boldsymbol{x}_t'\boldsymbol{\alpha})(1 - G(z_t)) + (\mu_\beta + \boldsymbol{x}_t'\boldsymbol{\beta})G(z_t) + \epsilon_t$$

    where $\epsilon_t \sim$ i.i.d. $N(0, \sigma^2)$ and $G(z_t) : \mathbb{R} \to \mathbb{G} \subseteq [0, 1]$.
  - For LSTAR, consider transition function $G(\cdot)$ such that

$$G(z_t, \gamma^*, \delta) = \{1 + \exp[-(\gamma^*/s_Z)(z_t - \delta)]\}^{-1}$$

    - Threshold Parameter δ Determines When Transition Occurs
    - Slope Parameter $\gamma = \gamma^*/s_Z$ Determines the Rate of Transition
    - As $\gamma \to \infty$, $G(z_t, \gamma^*, \delta)$ Becomes a Step Function

Chapter 1: LSTAR Model

- Prior Distributions
  - TAR(Geweke and Terui, 1993; Chen and Lee, 1995; Koop and Potter, 1999), STAR(Lubrano, 2000; Lopes and Salazar, 2006; Livingston Jr. and Nur, 2017)
  - $\mu_\alpha \sim \mathcal{N}(\cdot,\cdot)$ and $\mu_\beta \sim \mathcal{N}(\cdot,\cdot)$
  - $\sigma^2 \sim \mathcal{IG}(\cdot,\cdot)$
  - $\gamma^* \sim \mathcal{LN}(\cdot,\cdot)$
  - $\delta \sim \mathrm{U}[q_Z(0.15), q_Z(0.85)]$ where $q_Z(\cdot)$ is the empirical quantile function
  - Bayesian Global-Local Shrinkage Priors for $\boldsymbol{\alpha}$ and $\boldsymbol{\beta}$

$$\alpha_k, \beta_k | \lambda_k^2, \lambda^2, \sigma^2 \sim N(0, \sigma^2\lambda^2\lambda_k^2)$$

$$\lambda_k^2 \sim \pi_{Local}(.) \text{ and } \lambda^2 \sim \pi_{Global}(.)$$

    - LASSO (Park and Casella, 2008; Hans, 2009; Schmidt and Makalic, 2013)
    - Regime-Specific LASSO (Separate Global Tuning Parameters)
    - Variable Selection LASSO (Lykou and Ntzoufras, 2013)
    - Horseshoe (Carvalho et al., 2010; Makalic and Schmidt, 2016)

Chapter 1: LSTAR Model

- Transition Variable
  - Change-Point Option: $z_t = t$
  - Exogenous Option: $z_t = x_{t-d}$
  - Endogenous Option: $z_t = y_{t-d}$ (Self-Exciting)
  - Let $\boldsymbol{d}_t' = [y_{t-1}, y_{t-2}, \cdots, y_{t-d_{max}}]$ and $\boldsymbol{\phi}_t' = [\phi_1, \phi_2, \cdots, \phi_{d_{max}}]$. Reparameterize transition variable $z_t = \boldsymbol{\phi}' \boldsymbol{d}_t$.

$$\phi \sim Dir\left(\left[\frac{1}{d_{max}}, \frac{1}{d_{max}}, \cdots, \frac{1}{d_{max}}\right]'\right)$$

  Now, z_t is a weighted average of all considered transition variables. Because the weights are constrained to sum to 1, the prior for δ does not require modification.
  - Advantages
    - Allows for a composite transition variable
    - Estimates a more encompassing LSTAR model.

Chapter 1: LSTAR Model

- Simulation Study

Let $\boldsymbol{d}_t = [y_{t-1}, y_{t-2}, y_{t-3}, y_{t-4}]'$ and $\phi' = [\phi_1, \phi_2, \phi_3, \phi_4]$.

$$y_t = (-0.6y_{t-3})[1 - G(y_{t-1})] + (0.02 + 0.75y_{t-3})[G(y_{t-1})] + \epsilon_t$$

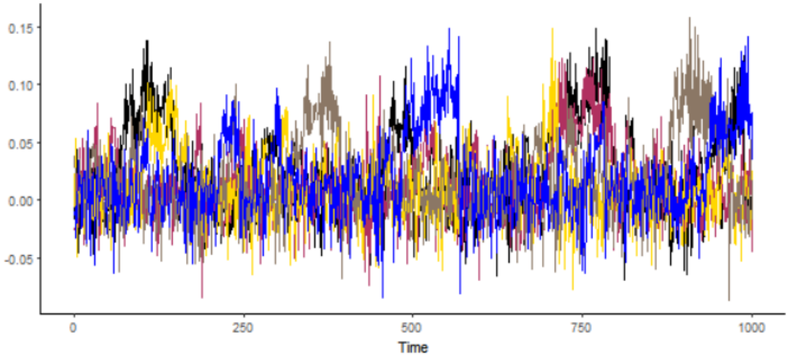
$$\text{where: } G(y_{t-1}) = \left\{1 + \exp\left[-120(\phi'\boldsymbol{d}_t - 0.02)\right]\right\}^{-1}$$

$$\text{and } \epsilon_t \sim \text{i.i.d. } N(0, 0.02^2)$$

Under prior $\phi \sim Dir([0.25, 0.25, 0.25, 0.25]')$, we conduct posterior sampling for three different threshold variables $\{z_{1,t}, z_{2,t}, z_{3,t}\}$ defined through ϕ. BHS priors are used for autoregressive coefficients.

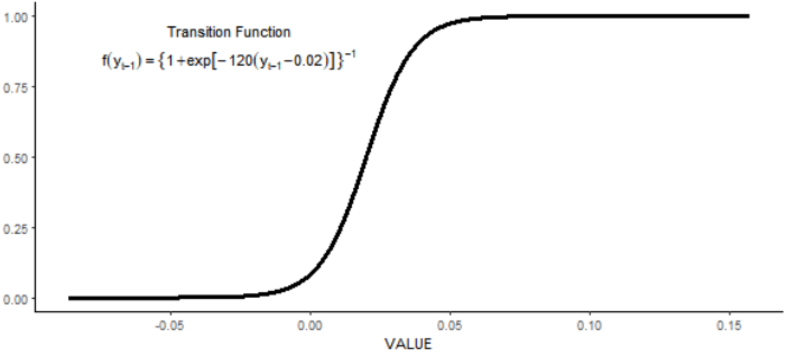
Chapter 1: LSTAR Model

Figure: Ten Random Replications



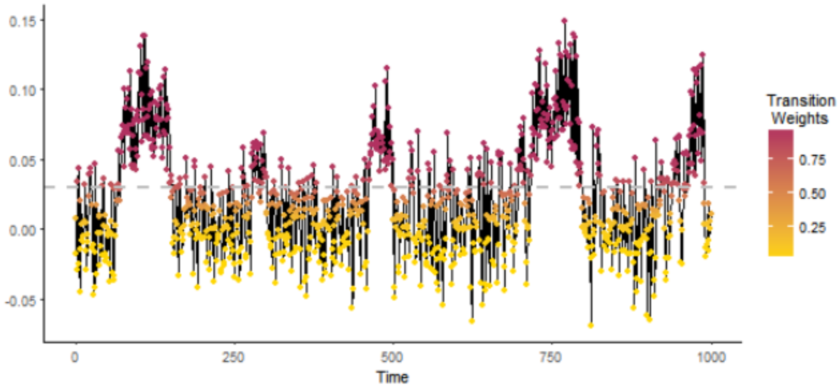
Chapter 1: LSTAR Model

Figure: Transition Function



Chapter 1: LSTAR Model

Figure: Illustration of Regime-switching Behavior

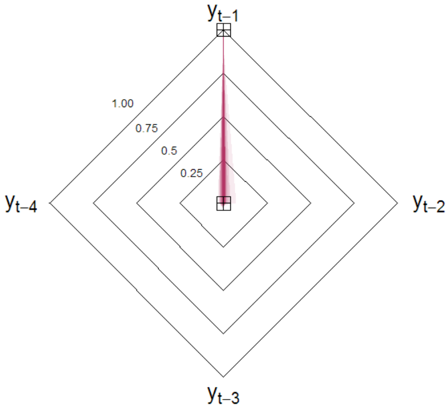


Chapter 1: LSTAR Model

- Bayesian Selection of the Threshold Variable (Scenario 1)

  Consider $z_{1,t} = y_{t-1} = [1, 0, 0, 0]\boldsymbol{d}_t$.

Figure: Posterior Means of ϕ from 100 Replications

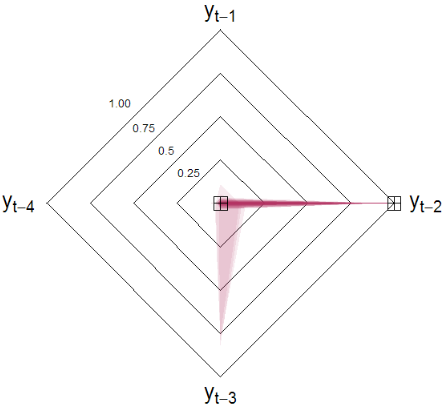


Chapter 1: LSTAR Model

- Bayesian Selection of the Threshold Variable (Scenario 2)

  Consider $z_{2,t} = y_{t-2} = [0, 1, 0, 0]\boldsymbol{d}_t$.

Figure: Posterior Means of ϕ from 100 Replications

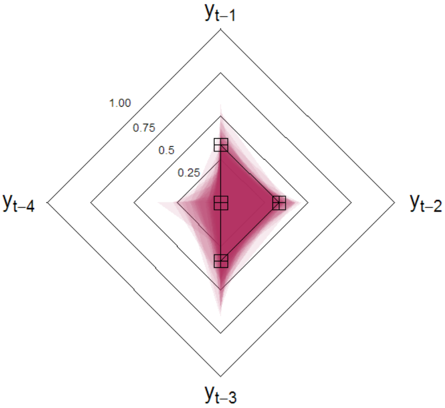


Chapter 1: LSTAR Model

- Bayesian Selection of the Threshold Variable (Scenario 3)

Consider $z_{3,t} = \frac{y_{t-1}+y_{t-2}+y_{t-3}}{3} = [\frac{1}{3}, \frac{1}{3}, \frac{1}{3}, 0]\boldsymbol{d}_t$.

Figure: Posterior Means of ϕ from 100 Replications



Chapter 1: LSTAR Model

- Application to Sunspot Data (Granger, 1957; Teräsvirta et al., 2010)
- Application to Daily Maximum Water Temperatures (Kamarianakis et al., 2016)
  - Data Used From 31 Rivers in Spain
  - Models to Forecast Daily Maximum Water Temperature
  - Inclusion of Exogenous Distributed Lag Terms from Known Air Temperatures
  - Horizon-Specific Models Targeting 3-step and 7-step Ahead Forecasts
  - Nonlinear Models Improved Forecasting Accuracy for Some Rivers

Chapter 1: LSTAR Model

- Contribution and Novelty
  - Efficacy of Bayesian Regularization for Nonlinear Time Series Models
  - Use of the Dirichlet Prior Leads to More Encompassing LSTAR Specification
  - Regime-Specific Tuning Parameters Influences Convergence in MCMC
  - Detailed **R** Code Provided for Reproducibility
- Feedback from *International Journal of Forecasting*
  - Focus on Dirchlet Priors for Estimating Transition Variable
  - Better Forecasting Application
  - Consider Density Forecasts Along with Point Forecasts

Chapter 2: TAR Model

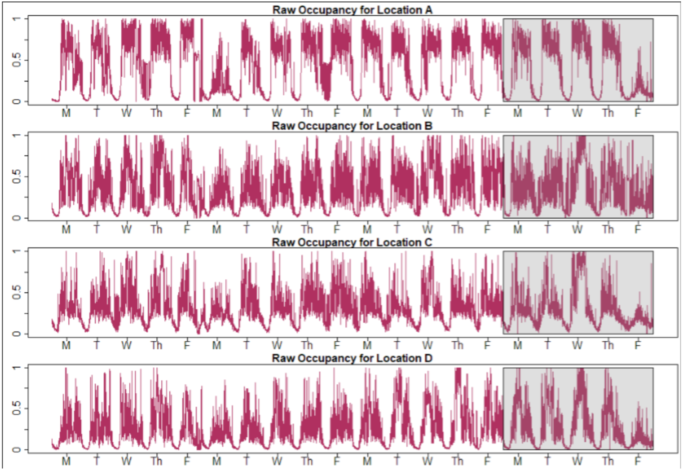
- Need for Traffic Occupancy Models
  - Advanced Traffic Management Systems (ATMS) Monitor Traffic Characteristics in Real Time
  - ATMS Require Fast Short-Term Forecasting to Reduce Congestion
  - Traffic Occupancy is the Percent of Time a Detection Zone is Occupied
  - Different States of Traffic: Free-Flow, Congested, Transitional
  - Factors Influencing Regime Changes : Weekly Work Patterns, Accidents, Weather, etc.
- Traffic Data Considered
  - Major Athens' Arterial: Alexandras Ave.
  - Time Period: April 2000
  - Obtained by National Technical University of Athens
  - Provided for 2013 TRANSportation Data FORecasting Competition (TRANSFOR) Developed by the Traffic Research Board (TRB) for Annual Meeting Workshop (Kamarianakis, 2014)
  - Measured on 90s Interval, but Mean Aggregated to 3min Interval

Chapter 2: TAR Model

Figure: Map of Traffic Network in Athens, Greece

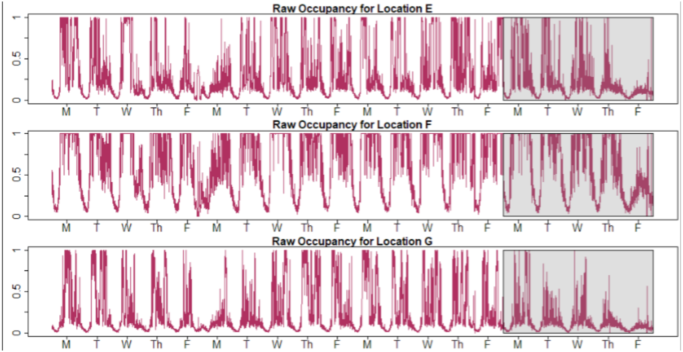
Chapter 2: TAR Model

Figure: Raw Traffic Occupancy From Westbound Detectors



Chapter 2: TAR Model

Figure: Raw Traffic Occupancy From Eastbound Detectors



Chapter 2: TAR Model

- (L, D, h)-Specific Models
  - Location $L \in \{A, B, C, D, E, F, G\}$
  - Work Day $D \in \{M, T, W, Th, F\}$
  - Horizon $h \in \{1, 3, 5\}$
- Data Transformation
  - Let O_t Represent the Traffic Occupancy at Time t
  - $Y_t = \mathrm{logit}(O_t) = \log[O_t/(1 - O_t)]$
  - Raw Data Adjusted at the Boundary so $\mathrm{logit}(\cdot)$ Is Defined
  - Forecasts Evaluated on Original Scale, but

$$\hat{O}_{L,t} \neq \mathrm{logit}^{-1}(\hat{Y}_{L,t})$$

  - Density Forecasts Produced from $\{\mathrm{logit}^{-1}(\hat{Y}_t^{(s)})\}_{s=1}^{S}$ where $\{\hat{Y}_t^{(s)}\}_{s=1}^{S}$ are S posterior samples obtained from the posterior predictive distribution $f(\hat{Y}_t|\mathcal{I}_t^*)$ where $\mathcal{I}_t^* = \{y_k\}_{k=t-h}^{t-h-P+1}$

Chapter 2: TAR Model

- Horizon-Specific Gaussian TAR(P) Model with $(m+1)$-regimes

$$y_t = \phi_0^{(j)} + \sum_{i=1}^{P} \phi_i^{(j)} y_{t-h-i+1} + \sigma\epsilon_t, \text{ for } \delta_{j-1} < y_{t-h} \leq \delta_j$$

where $\sigma > 0$, $j \in \{1, 2, \cdots, m+1\}$, $h \in \mathbb{N}$, and $\epsilon_t \sim \mathcal{N}(0, 1)$.

Vector of Thresholds $\boldsymbol{\delta} = [\delta_1, \cdots, \delta_m]$.

Partitions the Process into $m+1$ regimes such that
$-\infty = \delta_0 < \delta_1 \leq \delta_2 \leq \cdots \leq \delta_m < \delta_{m+1} = \infty$.

Chapter 2: TAR Model

- High Dimensional Linear Representation (Chan et al., 2015, 2017)
  - Let $\boldsymbol{y} = [y_1, \cdots, y_T]'$, $\boldsymbol{\epsilon} = [\epsilon_1, \cdots, \epsilon_T]'$, and define matrix $\boldsymbol{X}$ by

$$\boldsymbol{X} = \begin{bmatrix} 1 & y_{1-h} & y_{1-h-1} & \cdots & y_{1-h-P+1} \\ 1 & y_{2-h} & y_{2-h-1} & \cdots & y_{2-h-P+1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & y_{T-h} & y_{T-h-1} & \cdots & y_{T-h-P+1} \end{bmatrix}.$$

Second Column of $\boldsymbol{X}$ Contains the h-Specific Transition Variable.

Model Matrix $\boldsymbol{X}$ Often Used in Linear AR(P) Regressions.

Chapter 2: TAR Model

- High Dimensional Linear Representation (Cont.)
  - Reorder $\boldsymbol{y}$, $\boldsymbol{\epsilon}$, and $\boldsymbol{X}$ According to Transition Variable
    Sorting function $\pi(i) : \{1, \cdots, T\} \to \{1, \cdots, T\}$ where $\pi(i)$ equates to the time index of the ith smallest element in $[y_{1-h}, y_{2-h}, \cdots, y_{T-h}]'$. Now,

$$\boldsymbol{y}_R = [y_{\pi(1)+h}, \cdots, y_{\pi(T)+h}]',$$

$$\boldsymbol{\epsilon}_R = [\epsilon_{\pi(1)+h}, \cdots, \epsilon_{\pi(T)+h}]',$$

and

$$\boldsymbol{X}_1 = \begin{bmatrix} 1 & y_{\pi(1)} & y_{\pi(1)-1} & \cdots & y_{\pi(1)-P+1} \\ 1 & y_{\pi(2)} & y_{\pi(2)-1} & \cdots & y_{\pi(2)-P+1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & y_{\pi(T)} & y_{\pi(T)-1} & \cdots & y_{\pi(T)-P+1} \end{bmatrix} = \begin{bmatrix} \boldsymbol{y}'_{\pi(1)} \\ \boldsymbol{y}'_{\pi(2)} \\ \vdots \\ \boldsymbol{y}'_{\pi(T)} \end{bmatrix}$$

Chapter 2: TAR Model

- High Dimensional Linear Representation (Cont.)
    - Finite Set of m Thresholds for an $(m+1)$-Regime TAR(P)

      Define the Empirical Quantile Function,

      $q(.) : [0,1] \to [\min\{y_{t-h} : t = 1, 2, \cdots, T\}, \max\{y_{t-h} : t = 1, 2, \cdots, T\}]$

      Select Sequence $\{p_k\}_{k=1}^{m}$ of m Evenly Spaced Percentiles where

      $$p_{min} = p_1 < \cdots < p_m = p_{max}$$

      For a Fully Saturated TAR Model Limited to $(m+1)$ Regimes, Fix *a priori*

      $$\delta = [q(p_1), q(p_2), \cdots, q(p_m)]'$$

Chapter 2: TAR Model

- High Dimensional Linear Representation (Cont.)
  - Finite Set of m Thresholds for an $(m+1)$-Regime TAR(P) (Cont.)

    For $j \in \{2, \cdots, m+1\}$, Let k_j Represent the Number of Elements in $[y_{1-h}, y_{2-h}, \cdots, y_{T-h}]'$ Less than $q(p_{j-1})$ and Define

$$
\boldsymbol{X}_j = \begin{bmatrix} 0 & 0 & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & 0 \\ 1 & y_{\pi(k_j+1)} & y_{\pi(k_j+1)-1} & \cdots & y_{\pi(k_j+1)-P+1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & y_{\pi(T)} & y_{\pi(T)-1} & \cdots & y_{\pi(T)-P+1} \end{bmatrix} = \begin{bmatrix} 0 \\ \vdots \\ 0 \\ \boldsymbol{y}'_{\pi(k_j+1)} \\ \vdots \\ \boldsymbol{y}'_{\pi(T)} \end{bmatrix}.
$$

  - Fully Saturated $(m+1)$-Regime TAR(P) as a Linear Regression

$$
\boldsymbol{y}_R = \boldsymbol{X}_R \boldsymbol{\theta}_R + \boldsymbol{\epsilon}_R
$$

    $\boldsymbol{X}_R = [\boldsymbol{X}_1, \boldsymbol{X}_2, \cdots, \boldsymbol{X}_{m+1}]$ is a $T \times (P+1)(m+1)$ Matrix

    $\boldsymbol{\theta}_R = [\boldsymbol{\theta}'_1, \boldsymbol{\theta}'_2, \cdots, \boldsymbol{\theta}'_{m+1}]'$ is a $(P+1)(m+1) \times 1$ Vector of Grouped Reparameterized Coefficients

Chapter 2: TAR Model

- Baseline (L, D)-Specific Seasonal Model (Cont.)

$$y_t = \mu + \sum_{j=1}^{H} \left[\alpha_j \sin\left(\frac{2\pi tj}{480}\right) + \beta_j \cos\left(\frac{2\pi tj}{480}\right) \right] + \sigma\epsilon_t$$

where $\sigma > 0$, $H \in \mathbb{N}$, and $\epsilon_t \sim \mathcal{N}(0, 1)$.

Representable as a High Dimensional Linear Regession,

$$\boldsymbol{y}_F = \boldsymbol{X}_F \boldsymbol{\theta}_F + \boldsymbol{\epsilon}_F$$

- Considerations for Traffic Occupancy
  - Maximum AR Order $P = 7$
  - Maximum Number of Thresholds $m = 50$
  - Set $p_{min} = 0.15$ and $p_{max} = 0.85$
  - Saturated 51-Regime TAR(7) Model with 408 Parameters in $\boldsymbol{\theta}_R$

  - Maximum Number of Sine/Cosine Pairs $H = 150$
  - Saturated Seasonal Harmonic Regression Model with 301 Parameters in $\boldsymbol{\theta}_F$

Chapter 2: TAR Model

- Three-Step Procedure For Automatic Estimation and Selection
  - Full Model $\boldsymbol{y}_R = \boldsymbol{X}_R\boldsymbol{\theta}_R + \boldsymbol{\epsilon}_R$ Nests 6.61×10^{122} Different $(m^* + 1)$-Regime Subset TAR(P) Models where $0 \leq m^* \leq m$
  - Step 1: Sparse Estimation Using Horseshoe+ Shrinkage Priors
    - Group LASSO Used by Chan et al. (2015)
    - BHS^+ Hierarchy for Each θ_i in $\boldsymbol{\theta}_R$ (Bhadra et al., 2016)

$$\begin{aligned}
\theta_i|\lambda_i, \tau, \sigma^2 &\sim \mathcal{N}(0, \lambda_i^2\tau^2\sigma^2) \\
\lambda_i &\sim \mathcal{C}^+(0, \eta_i) \\
\eta_i &\sim \mathcal{C}^+(0, 1) \\
\tau &\sim \mathcal{C}^+(0, 1)
\end{aligned}$$

    - Modified Hierarchy Required for Gibbs Sampling (Makalic and Schmidt, 2016)

    If $\lambda_i^2|\nu_i \sim \mathcal{IG}(\frac{1}{2}, \frac{1}{\nu_i})$ and $\nu_i \sim \mathcal{IG}(\frac{1}{2}, 1)$, then $\lambda_i^2 \sim \mathcal{C}^+(0, 1)$ (Wand et al., 2011).

Chapter 2: TAR Model

- Three-Step Procedure For Sparse Estimation (Cont.)
  - Step 2: Regime Identification
    - Good Starting Point Under Full Saturated Model $\mathcal{M}_R$
    - Samples $\{\boldsymbol{\theta}_R^{(s)}\}_{s=1}^S$ and $\{\sigma^{(s)}\}_{s=1}^S$ from Joint Posterior Distribution
    - Given Candidate Submodel $\mathcal{M}_\perp$, Posterior Samples $\{\boldsymbol{\theta}_\perp^{(s)}\}_{s=1}^S$ and $\{\sigma_\perp^{(s)}\}_{s=1}^S$ Obtained Via Projection
    - Gaussian Linear Models (Piironen and Vehtari, 2015, 2017)

$$\boldsymbol{\theta}_\perp^{(s)} = (\mathbf{X}_\perp'\mathbf{X}_\perp)^{-1}\mathbf{X}_\perp'\mathbf{X}_R\boldsymbol{\theta}_R^{(s)}$$

$$\sigma_\perp^{(s)} = \sqrt{(\sigma^{(s)})^2 + \frac{(\mathbf{X}_R\theta_R^{(s)} - \mathbf{X}_\perp\boldsymbol{\theta}_\perp^{(s)})'(\mathbf{X}_R\boldsymbol{\theta}_R^{(s)} - \mathbf{X}_\perp\boldsymbol{\theta}_\perp^{(s)})}{T}}$$

Chapter 2: TAR Model

- Three-Step Procedure For Sparse Estimation (Cont.)
  - Step 2: Regime Identification (Cont.)
    - Kullback-Leibler (KL) Divergence (Kullback and Leibler, 1951) Measures the Overall Discrepancy Between the Posterior Predictive Distributions $p(y_{T+1}|\mathcal{M}_R, \mathbf{y}_R, \mathbf{X}_R)$ and $p(y_{T+1}|\mathcal{M}_R, \mathbf{y}_\perp, \mathbf{X}_\perp)$
    - KL Divergence for a Particular Sample

$$d_\perp^{(s)}(\boldsymbol{\theta}_R^{(s)}, \sigma^{(s)}) = \frac{1}{2} \log \left(\frac{\sigma_\perp^{(s)}}{\sigma^{(s)}} \right)^2$$

    - Overall Discrepancy

$$D(\mathcal{M}_R || \mathcal{M}_\perp) = \frac{1}{S} \sum_{s=1}^{S} d_\perp^{(s)}(\boldsymbol{\theta}_R^{(s)}, \sigma^{(s)})$$

Chapter 2: TAR Model

- Three-Step Procedure For Sparse Estimation (Cont.)
  - Step 2: Regime Identification (Cont.)
    - Forward Stepwise Selection Algorithm (Peltola et al., 2014)

      Begin with Linear AR(P) Model, Denoted $\mathcal{M}_{\perp}^{(1)}$, where

      $$\boldsymbol{\theta}_{\perp}^{(1)} = [\boldsymbol{\theta}_1', \mathbf{0}', \mathbf{0}', \cdots, \mathbf{0}']',$$

      with initial discrepancy $D(\mathcal{M}_R || \mathcal{M}_{\perp}^{(1)})$

      For each $j \in \{2, \cdots, m+1\}$, $\boldsymbol{\theta}_j$ is Added to $\boldsymbol{\theta}_{\perp}^{(1)}$ and the Best 2-Regime TAR(P) Model $\mathcal{M}_{\perp}^{(2)}$ Minimizes the Discrepancy $D(\mathcal{M}_R || \mathcal{M}_{\perp}^{(2)})$.

      Continue to Identify the Best 3-Regime TAR, 4-Regime TAR, $\cdots$

      Stopping Rule Based on Relative Explanatory Power ($RelE$) from Dupuis and Robert (2003)

      $$RelE(\mathcal{M}_{\perp}) = 1 - \frac{D(\mathcal{M} || \mathcal{M}_{\perp})}{D(\mathcal{M} || \mathcal{M}^1)}$$

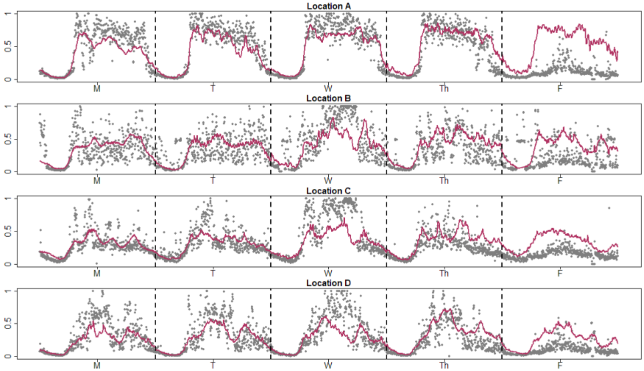
Chapter 2: TAR Model

- Three-Step Procedure For Sparse Estimation (Cont.)
  - Step 3: Final Subset Selection
    - Let $\mathcal{J} = \{j : \boldsymbol{\theta}_j \neq 0\}$ indicate the AR(P) parameter groups in $\boldsymbol{\theta}_R$ Selected Via Forward Algorithm
    - Let $\theta_{i,j}$ Represent the ith Parameter in the jth Vector $\boldsymbol{\theta}_j$ for $i \in \{1, 2, \cdots, P+1\}$ and $j \in \{1, 2, \cdots, m+1\}$.
    - The Set $\mathcal{I} = \{\theta_{i,j} : i = 1, \cdots, P+1 \text{ and } j \in \mathcal{J}\}$ Contains Potentially Relevant Parameters in $\boldsymbol{\theta}_R$
    - Repeat Forward Stepwise Algorithm Across $\mathcal{I}$. The Intercept Only Model Utilized for $\mathcal{M}_{\perp}^{(1)}$.
  - Result: Final Choice $\mathcal{M}_*$ is a $(m^* + 1)$-Regime Subset TAR(P) Model where m^* is the Number of Parameter Groups with at Least 1 Selected Parameter.

Chapter 2: TAR Model

- Results
  - (L, D)-Specific Seasonal Profiles (Quickly Forecasts at All Horizons)

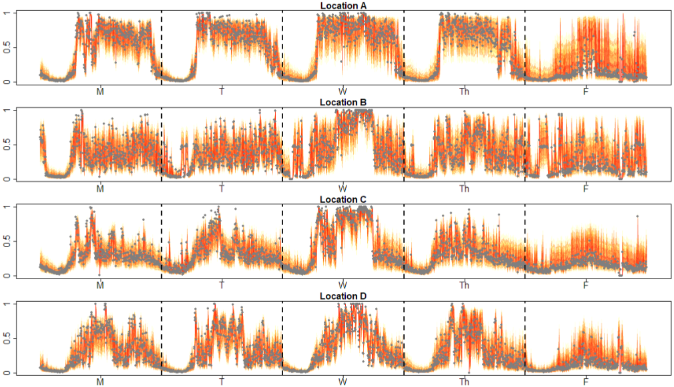
Figure: Forecasts Based on Seasonal Profiles for Westbound Detectors



Chapter 2: TAR Model

- Results (Cont.)
  - $(L, D, 1)$-Specific Final Subset TAR(7) Models

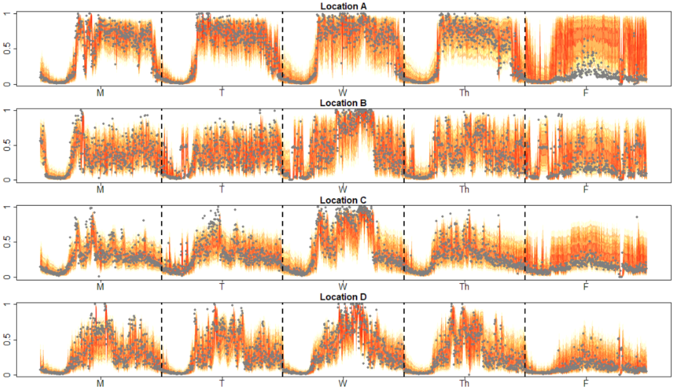
Figure: 1-Step Ahead Density TAR Forecasts for Westbound Detectors



Chapter 2: TAR Model

- Results (Cont.)
  - $(L, D, 3)$-Specific Final Subset TAR(7) Models

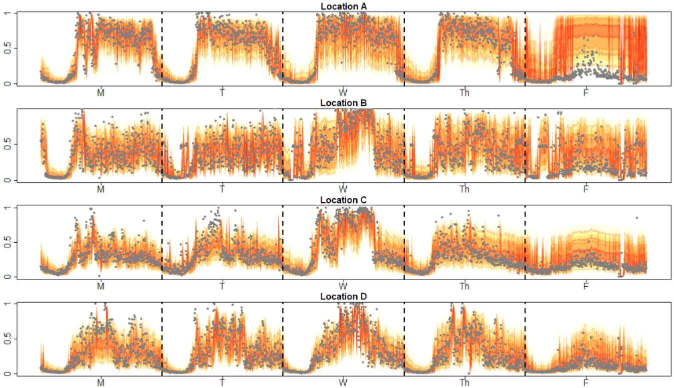
Figure: 3-Step Ahead Density TAR Forecasts for Westbound Detectors



Chapter 2: TAR Model

- Results (Cont.)
  - $(L, D, 5)$-Specific Final Subset TAR(7) Models

Figure: 5-Step Ahead Density TAR Forecasts for Westbound Detectors



Chapter 2: TAR Model

- Results (Cont.)
  - Comparison of Point Forecasts
    - Evaluated on Mean Absolute Scaled Forecast Error (Hyndman and Koehler, 2006)

$$\mathrm{MASFE}(h) = \frac{1}{T_h} \sum_{t=P+h}^{480} \left| \frac{O_t - \widehat{O}_t}{\mathrm{MAE}_{RW}(h)} \right|$$

    - $\mathrm{MAE}_{RW}(h)$ is the Fitted MAE from Naive Random Walk where $\widehat{O}_t = O_{t-h}$

Table: 1-Step Ahead MASFE Forecast Comparison

Day	Model	A	B	C	D	E	F	G
				Location				
M	TAR	1.02	1.10	0.88	1.07	1.87	0.81	1.48
	SEAS	1.80	1.47	1.15	1.43	4.02	1.57	3.66
T	TAR	0.90	1.05	1.04	0.98	1.36	1.03	1.95
	SEAS	1.35	1.36	1.22	1.46	3.36	1.65	3.29
W	TAR	1.04	1.11	0.91	0.97	2.27	1.86	1.55
	SEAS	1.39	2.01	2.18	1.61	4.65	2.90	2.80
Th	TAR	0.93	0.89	0.82	0.92	1.48	1.52	1.83
	SEAS	1.44	1.43	1.51	1.42	3.98	2.74	4.07
F	TAR	1.80	1.08	1.01	0.85	1.45	2.40	1.16
	SEAS	4.77	2.23	1.98	1.83	4.37	6.24	3.78

Chapter 2: TAR Model

- Results (Cont.)
  - Comparison of Point Forecasts (Cont.)

Table: 3-Step Ahead MASFE Forecast Comparison

Day	Model	A	B	C	D	E	F	G
					Location			
M	TAR	0.94	1.06	0.88	1.13	1.85	0.88	1.50
	SEAS	1.36	1.17	0.93	1.14	2.91	1.19	2.57
T	TAR	0.87	1.04	0.96	1.03	1.46	1.15	1.21
	SEAS	1.04	1.06	0.91	1.10	2.24	1.22	2.15
W	TAR	1.09	1.15	1.10	0.99	1.75	2.00	1.26
	SEAS	1.15	1.61	1.69	1.21	2.94	2.24	1.71
Th	TAR	0.90	0.96	0.82	0.93	1.88	1.47	1.14
	SEAS	1.15	1.09	1.14	1.03	2.69	1.96	2.68
F	TAR	3.04	1.09	1.00	0.66	1.14	2.47	0.92
	SEAS	3.53	1.57	1.42	1.33	3.12	4.35	2.42

Chapter 2: TAR Model

- Results (Cont.)
  - Comparison of Point Forecasts (Cont.)

Table: 5-Step Ahead MASFE Forecast Comparison

Day	Model	A	B	C	D	E	F	G
					Location			
M	TAR	0.94	0.99	0.89	1.12	1.97	0.86	1.68
	SEAS	1.24	1.06	0.88	1.06	2.46	1.08	2.17
T	TAR	0.81	1.05	0.95	1.00	1.47	1.13	1.15
	SEAS	0.95	0.95	0.85	0.99	1.85	1.07	1.77
W	TAR	1.02	1.12	1.04	0.99	1.67	1.94	1.26
	SEAS	1.01	1.44	1.56	1.12	2.30	1.95	1.44
Th	TAR	0.84	0.98	0.82	0.87	1.51	1.48	1.22
	SEAS	1.05	1.03	1.07	0.94	2.13	1.73	2.24
F	TAR	2.85	1.20	0.88	0.70	1.26	2.52	1.00
	SEAS	3.07	1.48	1.33	1.19	2.59	3.82	2.01

Chapter 2: TAR Model

- Contribution and Novelty
  - Advances Methodology for Estimating TAR(P) Models with Potentially Many Regimes
  - Shows Relevancy in an Industry Needing Short-term Forecasting
  - Easy 3-Step Bayesian Approach Capable of Selecting Regimes and AR Parameters Within Regimes (More Flexible Final Models)
  - Appendix Provides Defense for Horseshoe+ Hierarchy
- Future Developments
  - Modifications for Student t Distributed Errors (Shows Promise)
  - Quality of Forecast Credible Regions
  - Other Techniques for Modeling Heteroskedasticity
  - Look at Using Different Transition Variables (Composite, Difference, etc.)

Chapter 3: ARMA Model

- ARMA(p, q) Model
  - Classic Parametric Form
  $$\phi(B)y_t = \theta(B)\epsilon_t$$
  where
  $$\phi(B) = 1 - \sum_{j=1}^{p} \phi_j B^j \text{ and } \theta(B) = 1 + \sum_{k=1}^{q} \theta_k B^K$$
  and B Represents the Backshift Operator such that $B^k y_t = y_{t-k}$
  - Autoregressive Order p and Moving Average Order q
  - Stationary: Roots of $1 - \phi_1 z - \phi_2 z^2 - \cdots - \phi_p z^p = 0$
  - Invertible: Roots of $1 + \theta_1 z + \theta_2 z^2 + \cdots + \theta_q z^q = 0$
  - Multiplicative Seasonal ARMA (Box and Jenkins, 1976) are Subset ARMA Processes
  - Presence of MA Terms Poses an Estimation Problem (Hamilton, 1994; Cryer and Chan, 2008)

Chapter 3: ARMA Model

- Fast Estimation of ARMA(p, q)
  - Residuals $\{\hat{\epsilon}_t : t = p' + 1, \cdots, T\}$ of a Long AR(p') Process Approximate the Unobserved $\{\epsilon_t\}$ (Hannan and Rissanen, 1982; Brockwell and Davis, 2016)
  - Let $\boldsymbol{y} = [y_m, \cdots, y_T]'$, $\boldsymbol{\epsilon} = [\epsilon_m, \cdots, \epsilon_T]'$, $\boldsymbol{\beta} = [\boldsymbol{\phi}', \boldsymbol{\theta}']' = [\phi_1, \cdots, \phi_p, \theta_1, \cdots, \theta_q]'$, $m = p' + max\{p, q\} + 1$, and

$$\boldsymbol{X} = \begin{bmatrix} \boldsymbol{x}_m' \\ \boldsymbol{x}_{m+1}' \\ \vdots \\ \boldsymbol{x}_T' \end{bmatrix} = \begin{bmatrix} y_{m-1} & \cdots & y_{m-p} & \hat{\epsilon}_{m-1} & \cdots & \hat{\epsilon}_{m-q} \\ y_m & \cdots & y_{m-p+1} & \hat{\epsilon}_m & \cdots & \hat{\epsilon}_{m-q+1} \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ y_{T-1} & \cdots & y_{T-p} & \hat{\epsilon}_{T-1} & \cdots & \hat{\epsilon}_{T-q} \end{bmatrix}.$$

    The ARMA(p, q) Model is Expressed by $\boldsymbol{y} = \boldsymbol{X\beta} + \boldsymbol{\epsilon}$
  - Consider $p' = 10\log_{10}(T)$. Hannan and Kavalieris (1984) and Chen and Chan (2011) Suggest Using Information Criteria to Select p'. Reduces the Loss of Data, but Substantially Effects Results.

Chapter 3: ARMA Model

- Subset ARMA(p, q) Penalized Estimation Methods
  - Classic Techniques
    - Adaptive LASSO (Zou, 2006)

$$\hat{\boldsymbol{\beta}}_{AL}(\lambda) = \underset{\boldsymbol{\beta}}{\operatorname{argmin}} \; ||\mathbf{y} - \mathbf{X}\boldsymbol{\beta}||^2 + \lambda \sum_{i=1}^{p+q} \hat{w}_i |\beta_i|$$

    - Adaptive Elastic Net (Zou and Zhang, 2009)

$$\hat{\boldsymbol{\beta}}_{AE}(\lambda, \alpha) = \underset{\boldsymbol{\beta}}{\operatorname{argmin}} \; ||\mathbf{y} - \mathbf{X}\boldsymbol{\beta}||^2 + \lambda \left[(1-\alpha) \sum_{i=1}^{p+q} \beta_i^2 + \alpha \sum_{i=1}^{p+q} \hat{w}_i |\beta_i| \right]$$

    - Vector of Weights $\hat{\mathbf{w}} = |\hat{\boldsymbol{\beta}}_L + 1/T|^{-\eta}$ where

$$\hat{\boldsymbol{\beta}}_{L}(\lambda) = \underset{\boldsymbol{\beta}}{\operatorname{argmin}} \; ||\mathbf{y} - \mathbf{X}\boldsymbol{\beta}||^2 + \lambda \sum_{i=1}^{p+q} |\beta_i|.$$

$\hat{\boldsymbol{\beta}}_L$ is the Original LASSO Estimate of Tibshirani (1996), and $\eta = 2$, as Recommended by Zou (2006); Chen and Chan (2011)

Chapter 3: ARMA Model

- Subset ARMA(p, q) Penalized Estimation Methods
  - Classic Techniques (Cont.)
    - Optimal Selection of Tuning Parameters λ and α

      For ADLASSO and ADENET, Model Selection Determined by λ and α. Two Stages Require Tuning. Estimate Prediction Error for Grid of λ and α

      Time Series Studies Often Perform Selection Based on Out-of-Sample (OOS) Forecasting

      Penalized Information Criteria Such as AIC or BIC (Chen and Chan, 2011)

      Cross-Validated (CV) Measures of Error (Stone, 1974; Hastie et al., 2009)

      Appropriateness of CV Questioned in Time Series Analysis. Blocked Approaches Used for Dependent Data (Burman et al., 1994; Racine, 2000; Arlot et al., 2010; Bergmeir and Benítez, 2012)

      Regular K-Fold CV Consistently Outperforms OOS in Estimating Prediction Error (Bergmeir et al., 2018).

Chapter 3: ARMA Model

- Subset ARMA(p, q) Penalized Estimation Methods (Cont.)
  - Bayesian Techniques
    - Bayesian Horseshoe (Carvalho et al., 2009, 2010) and Bayesian Horseshoe+ (Bhadra et al., 2016) Hierarchies Considered for Initial Estimation
    - Projection Method with Forward Selection Algorithm Used to Identify the Best Model at Every Level of Flexibility from Intercept-Only to Fully Saturated ARMA(p, q)
    - Final Model Selected Based on Relative Explanatory Power ($RelE$) or Out-of-Sample Forecasting Results (Piironen and Vehtari, 2017)

Chapter 3: ARMA Model

- Overview of Methods

ADLASSO and ADENET Variants Denoted AL_m and AE_m where $m \in \{1, 2, \cdots, 11\}$

Table: Summary of ADLASSO and ADENET Variants

Method (m)	Initial Weights (Stage 1)	Final Model (Stage 2)
1	AIC	AIC
2	AIC	BIC
3	BIC	BIC
4	OOS	
5	depOOS	
6	CV-5	
7	CV-10	
8	LOOCV	
9	BCV-5	
10	BCV-10	
11	LOBOCV	

Chapter 3: ARMA Model

- Overview of Methods (Cont.)

  Bayesian Variants Denoted BHS_m and BHS_m^+ where $m \in \{1, 2, \cdots, 4\}$

Table: Summary of BHS and BHS^+ Variants

Method (m)	Final Model Selection
1	$RelE(\cdot) > 0.90$
2	$RelE(\cdot) > 0.95$
3	$RelE(\cdot) > 0.98$
4	OOS

Chapter 3: ARMA Model

- Simulation Study
  - Gaussian ARMA Processes (Chen and Chan, 2011)

$$y_{1,t} = 0.8y_{1,t-1} + 0.7y_{1,t-6} - 0.56y_{1,t-7} + \epsilon_{1,t}$$

$$\begin{aligned} y_{2,t} &= 0.8y_{2,t-1} + 0.7y_{2,t-6} - 0.56y_{2,t-7} \\ &\quad + 0.8\epsilon_{2,t-1} + 0.7\epsilon_{2,t-6} + 0.56\epsilon_{2,t-7} + \epsilon_{2,t} \end{aligned}$$

$$y_{3,t} = 0.8\epsilon_{3,t-1} + 0.7\epsilon_{3,t-6} + 0.56\epsilon_{3,t-7} + \epsilon_{3,t}$$

Standard Gaussian Errors $\{\epsilon_{1,t}\}$, $\{\epsilon_{2,t}\}$, and $\{\epsilon_{3,t}\}$. Abbreviated Models I, II, and III, Respectively. Samples of Length $T \in \{120, 240, 360\}$.

Data Generating Processes are Subset ARMA$(7, 7)$. Maximum ARMA Orders $P = Q = 14$.

  - Evaluating Subset ARMA Selection
    - C: Relative Frequency of Selecting All Relevant Parameters
    - I: Relative Frequency of Identifying the True Model
    - $-$: False Negative Rate (Probability of Missing a Relevant Parameter)
    - $+$: False Positive Rate (Probability of Selecting an Irrelevant Parameter)

Chapter 3: ARMA Model

- Simulation Study (Cont.)
  - Sensitivity: Order Selection for Long AR(p') Process

Table: Effect of Using AIC to Select p' on ADLASSO Subset ARMA(14, 14) Estimation of Model I Based on 500 Replications

	T	Long AR(p') C	Long AR(p') I	Long AR(p') $-$	Long AR(p') $+$	Short AR(p') C	Short AR(p') I	Short AR(p') $-$	Short AR(p') $+$
AL_1	120	0.19	0.01	0.36	0.28	0.02	0.00	0.49	0.28
	240	0.40	0.05	0.24	0.27	0.02	0.00	0.43	0.34
	360	0.46	0.07	0.21	0.26	0.03	0.00	0.38	0.36
AL_2	120	0.16	0.04	0.40	0.17	0.02	0.00	0.53	0.18
	240	0.36	0.13	0.27	0.18	0.01	0.00	0.47	0.24
	360	0.45	0.17	0.22	0.18	0.03	0.01	0.41	0.27
AL_3	120	0.05	0.01	0.44	0.12	0.01	0.00	0.48	0.14
	240	0.15	0.05	0.35	0.17	0.01	0.00	0.44	0.20
	360	0.21	0.09	0.31	0.20	0.01	0.01	0.39	0.24

Chapter 3: ARMA Model

- Simulation Study (Cont.)
  - Results for Model I

Table: Subset ARMA(14, 14) Results from 200 Replications of Model I

m	T	AL_m C	I	$-$	$+$	AE_m C	I	$-$	$+$
1	360	0.50	0.08	0.20	0.24	0.54	0.08	0.17	0.25
2	360	0.42	0.16	0.23	0.19	0.50	0.20	0.19	0.15
3	360	0.20	0.10	0.32	0.19	0.20	0.10	0.30	0.19
4	360	0.28	0.12	0.30	0.14	0.70	0.00	0.10	0.67
5	360	0.24	0.12	0.32	0.15	0.66	0.00	0.12	0.66
6	360	0.36	0.16	0.27	0.16	0.52	0.18	0.19	0.17
7	360	0.44	0.16	0.23	0.15	0.54	0.16	0.18	0.17
8	360	0.42	0.21	0.24	0.15	0.53	0.17	0.19	0.16
9	360	0.44	0.20	0.23	0.12	0.60	0.04	0.15	0.27
10	360	0.36	0.16	0.26	0.13	0.54	0.04	0.17	0.26
11	360	0.46	0.24	0.23	0.11	0.62	0.05	0.14	0.29
m	T	BHS_m C	I	$-$	$+$	BHS^+_m C	I	$-$	$+$
1	360	0.70	0.60	0.18	0.04	0.66	0.57	0.21	0.04
2	360	0.88	0.62	0.08	0.04	0.88	0.60	0.07	0.04
3	360	0.92	0.42	0.05	0.06	0.90	0.49	0.06	0.05
4	360	0.92	0.30	0.04	0.09	0.92	0.32	0.05	0.08

Chapter 3: ARMA Model

- Simulation Study (Cont.)
  - Results for Model II

Table: Subset ARMA(14, 14) Results from 200 Replications of Model II

		AL_m				AE_m			
m	T	C	I	$-$	$+$	C	I	$-$	$+$
1	360	0.26	0.01	0.15	0.38	0.42	0.00	0.13	0.38
2	360	0.26	0.02	0.16	0.32	0.36	0.02	0.14	0.33
3	360	0.20	0.01	0.18	0.30	0.23	0.02	0.17	0.31
4	360	0.06	0.03	0.28	0.08	0.70	0.00	0.05	0.78
5	360	0.06	0.04	0.27	0.08	0.67	0.00	0.06	0.78
6	360	0.18	0.05	0.18	0.25	0.30	0.00	0.14	0.32
7	360	0.16	0.04	0.19	0.24	0.26	0.01	0.16	0.30
8	360	0.16	0.02	0.20	0.26	0.26	0.01	0.16	0.31
9	360	0.06	0.04	0.29	0.06	0.34	0.00	0.14	0.36
10	360	0.08	0.06	0.26	0.07	0.26	0.00	0.16	0.37
11	360	0.04	0.03	0.27	0.06	0.31	0.00	0.14	0.38
		BHS_m				BHS_m^+			
m	T	C	I	$-$	$+$	C	I	$-$	$+$
1	360	0.13	0.02	0.24	0.09	0.12	0.03	0.25	0.09
2	360	0.57	0.18	0.10	0.09	0.51	0.16	0.12	0.10
3	360	0.89	0.05	0.03	0.15	0.88	0.08	0.04	0.14
4	360	0.84	0.09	0.05	0.17	0.87	0.10	0.04	0.15

Chapter 3: ARMA Model

- Simulation Study (Cont.)
  - Results for Model III

Table: Subset ARMA(14, 14) Results from 200 Replications of Model III

m	T	AL_m C	AL_m I	AL_m $-$	AL_m $+$	AE_m C	AE_m I	AE_m $-$	AE_m $+$
1	360	0.45	0.03	0.26	0.33	0.47	0.02	0.21	0.34
2	360	0.36	0.04	0.33	0.21	0.40	0.07	0.27	0.18
3	360	0.78	0.19	0.10	0.11	0.78	0.18	0.10	0.11
4	360	0.49	0.18	0.24	0.21	0.86	0.00	0.05	0.61
5	360	0.52	0.20	0.23	0.20	0.90	0.00	0.03	0.61
6	360	0.46	0.22	0.25	0.12	0.48	0.12	0.22	0.18
7	360	0.44	0.23	0.28	0.12	0.50	0.14	0.23	0.16
8	360	0.48	0.22	0.24	0.10	0.48	0.12	0.23	0.17
9	360	0.55	0.14	0.20	0.15	0.60	0.04	0.16	0.29
10	360	0.56	0.21	0.18	0.14	0.64	0.02	0.14	0.25
11	360	0.42	0.08	0.27	0.19	0.48	0.01	0.22	0.34
m	T	BHS_m C	BHS_m I	BHS_m $-$	BHS_m $+$	BHS_m^+ C	BHS_m^+ I	BHS_m^+ $-$	BHS_m^+ $+$
1	360	0.26	0.03	0.46	0.14	0.26	0.04	0.46	0.14
2	360	0.40	0.00	0.34	0.22	0.38	0.00	0.36	0.21
3	360	0.62	0.00	0.18	0.39	0.59	0.00	0.20	0.34
4	360	0.40	0.02	0.35	0.25	0.35	0.02	0.36	0.24

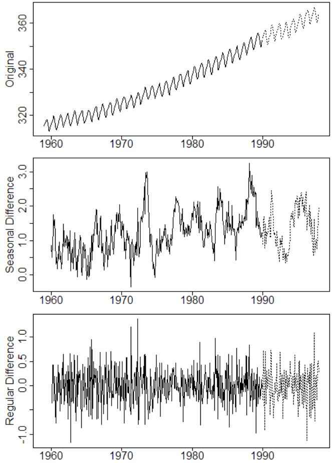
Chapter 3: ARMA Model

- Application
  - Monthly $\mathrm{CO_2}$ Levels Measured in Mauna Loa, Hawaii
  - Popular Textbook Example for Seasonal ARMA Processes with $\textit{Period} = 12$
  - Fitting Period: Jan 1994 - Dec 1989 (468 Time Points)
  - Forecasting Period: Jan 1990 - Dec 1997 (96 Time Points)
  - Apply Seasonal and Regular Differencing

Chapter 3: ARMA Model

- Application (Cont.)

Figure: Plots of Mauna Loa CO_2 Levels



Chapter 3: ARMA Model

- Application (Cont.)
  - Forecasting Metrics
    - Root Mean Squared Error

$$RMSE = \frac{1}{96}\sum_{j=1}^{96}(y_j - \hat{y}_j)^2$$

    - Mean Absolute Scaled Error
    - Mean Bias

$$MB = \frac{1}{96}\sum_{j=1}^{96}(y_j - \hat{y}_j)$$

    - Mean Directional Bias

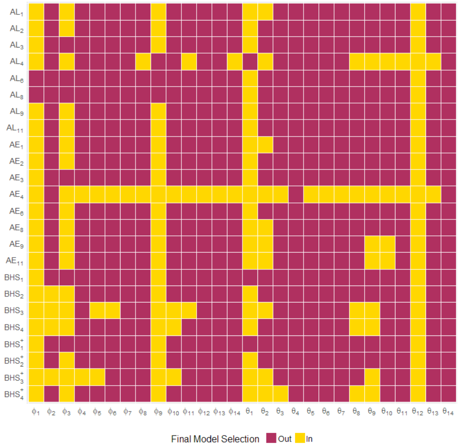
$$MDB = \frac{1}{96}\sum_{j=1}^{96}\operatorname{sgn}(y_j - \hat{y}_j)$$

where $\operatorname{sgn}(x) = 1$ if $x > 0$ and $\operatorname{sgn}(x) = -1$ if $x < 0$

Chapter 3: ARMA Model

- Application (Cont.)
  - Final Model Selection

Figure: Subset ARMA Models for Mauna Loa CO_2



Chapter 3: ARMA Model

- Application (Cont.)
  - Out-of-Sample Prediction

Table: One-Step Ahead Forecasting Results for Mauna Loa CO_2

	RMSE		*MASE*		*MB*		*MDB*	
m	AL_m	AE_m	AL_m	AE_m	AL_m	AE_m	AL_m	AE_m
1	0.34	0.34	0.53	0.53	-0.13	-0.13	-0.31	-0.29
2	0.33	0.33	0.52	0.52	-0.10	-0.10	-0.21	-0.21
3	0.34	0.34	0.52	0.52	-0.10	-0.10	-0.21	-0.21
4	Not Invertible (NI)							
6	0.34	0.33	0.53	0.52	-0.10	-0.09	-0.17	-0.23
8	0.34	0.34	0.53	0.54	-0.10	-0.14	-0.19	-0.31
9	0.34	0.36	0.52	0.57	-0.10	-0.18	-0.23	-0.44
11	0.33	0.36	0.52	0.56	-0.10	-0.17	-0.21	-0.44
m	BHS_m	BHS_m^+	BHS_m	BHS_m^+	BHS_m	BHS_m^+	BHS_m	BHS_m^+
1	0.31	0.31	0.49	0.49	-0.01	-0.01	0.04	0.06
2	0.32	0.32	0.50	0.50	-0.02	-0.02	0.00	-0.02
3	0.32	0.32	0.51	0.51	-0.02	-0.02	0.00	0.00
4	0.32	0.32	0.50	0.50	-0.01	-0.01	0.02	0.02
RW	0.64		1.03		0.00		-0.02	
ARMA	0.37		0.60		0.01		0.13	
SARMA	0.30		0.49		-0.04		-0.02	

Chapter 3: ARMA Model

- Contribution and Novelty
  - Highlight Issues with Previous Usage of Adaptive LASSO in Subset ARMA Selection
  - Extends Subset ARMA Estimation to Adaptive Elastic Net
  - Demonstrate the Appropriateness of Multiple Cross Validation Techniques with Penalized Regression for ARMA Models
  - Multi-step Bayesian Approach that Outperforms in Estimation, Selection, and Forecasting
  - Provides Detailed **R** Code for Reproducibility
- Future Developments
  - Extend Theoretical Asymptotic Results for Adaptive Elastic Net
  - More Focus on Bayesian Predictive Posterior Projection Method
  - Consider Alternative Bayesian Handling of Moving Average Terms

Questions?

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