

Quick proof on univariate OLS estimator

Antoine Mayerowitz

October 1, 2018

We suppose that the data consist of N observations $\{y_i, x_i\}_{i=1}^N$ where each observation i consist of a dependent variable y_i and an explanatory variable x_i . We suppose a linear relationship in our variables such that

$$y_i = b_0 + b_1 x_i + \varepsilon_i \quad (1)$$

Where b_0 and b_1 are scalars, shared by all observations and ε_i is the error term.

Our goal is to find b_0 and b_1 such that the *sum of squared of residuals*(SSR) is minimized. Where the SSR is defined by

$$\begin{aligned} SSR &= \sum_{i=1}^N e^2 \\ &= \sum_{i=1}^N (y_i - \hat{b}_0 - \hat{b}_1 x_i)^2 \end{aligned}$$

Mathematicaly we seek

$$\min_{\hat{b}_0, \hat{b}_1} \sum_{i=1}^N (y_i - \hat{b}_0 - \hat{b}_1 x_i)^2 \quad (2)$$

As we need to minimize a quadratic function, we take the derivative with respect to both argument and equalize to 0.

$$\begin{cases} \frac{\partial SSR}{\partial \hat{b}_0} = -2 \sum_{i=1}^N (y_i - \hat{b}_0 - \hat{b}_1 x_i) \\ \frac{\partial SSR}{\partial \hat{b}_1} = -2 \sum_{i=1}^N x_i (y_i - \hat{b}_0 - \hat{b}_1 x_i) \end{cases}$$

Then we equalize both equations to 0

$$\begin{cases} -2 \sum_{i=1}^N (y_i - \hat{b}_0 - \hat{b}_1 x_i) &= 0 \\ -2 \sum_{i=1}^N x_i (y_i - \hat{b}_0 - \hat{b}_1 x_i) &= 0 \end{cases}$$

We now have our system of 2 equations with 2 unknown that we need to solve for \hat{b}_0 and \hat{b}_1

We simplify

$$\begin{cases} \sum_{i=1}^N (y_i - \hat{b}_0 - \hat{b}_1 x_i) &= 0 \\ \sum_{i=1}^N x_i (y_i - \hat{b}_0 - \hat{b}_1 x_i) &= 0 \end{cases}$$

Rewriting

$$\begin{cases} \sum_{i=1}^N y_i &= \sum_{i=1}^N \hat{b}_0 + \hat{b}_1 \sum_{i=1}^N x_i \\ \sum_{i=1}^N x_i y_i &= \hat{b}_0 \sum_{i=1}^N x_i + \hat{b}_1 \sum_{i=1}^N x_i^2 \end{cases}$$

Using the fact that $\sum_{k=1}^K a = K \times a$ and that \hat{b}_0 do not depends on i we get

$$\begin{cases} \sum_{i=1}^N y_i &= N\hat{b}_0 + \hat{b}_1 \sum_{i=1}^N x_i \\ \sum_{i=1}^N x_i y_i &= \hat{b}_0 \sum_{i=1}^N x_i + \hat{b}_1 \sum_{i=1}^N x_i^2 \end{cases}$$

We multiply the first equation by $\frac{1}{N}$

$$\begin{cases} \frac{1}{N} \sum_{i=1}^N y_i &= \frac{1}{N} N\hat{b}_0 + \hat{b}_1 \frac{1}{N} \sum_{i=1}^N x_i \\ \sum_{i=1}^N x_i y_i &= \hat{b}_0 \sum_{i=1}^N x_i + \hat{b}_1 \sum_{i=1}^N x_i^2 \end{cases}$$

We define the average z by $\bar{z} := \frac{1}{N} \sum_{i=1}^N z_i$.

$$\begin{cases} \bar{y} &= \hat{b}_0 + \hat{b}_1 \bar{x} \\ \sum_{i=1}^N x_i y_i &= \hat{b}_0 \sum_{i=1}^N x_i + \hat{b}_1 \sum_{i=1}^N x_i^2 \end{cases}$$

We rewrite to get an expression of \hat{b}_0

$$\begin{cases} \hat{b}_0 &= \bar{y} - \hat{b}_1 \bar{x} \\ \sum_{i=1}^N x_i y_i &= \hat{b}_0 \sum_{i=1}^N x_i + \hat{b}_1 \sum_{i=1}^N x_i^2 \end{cases}$$

We substitute the first equation in the second

$$\begin{cases} \hat{b}_0 &= \bar{y} - \hat{b}_1 \bar{x} \\ \sum_{i=1}^N x_i y_i &= (\bar{y} - \hat{b}_1 \bar{x}) \sum_{i=1}^N x_i + \hat{b}_1 \sum_{i=1}^N x_i^2 \end{cases}$$

Developing

$$\begin{cases} \hat{b}_0 &= \bar{y} - \hat{b}_1 \bar{x} \\ \sum_{i=1}^N x_i y_i &= \bar{y} \sum_{i=1}^N x_i - \hat{b}_1 \bar{x} \sum_{i=1}^N x_i + \hat{b}_1 \sum_{i=1}^N x_i^2 \end{cases}$$

We factorize by \hat{b}_1

$$\begin{cases} \hat{b}_0 &= \bar{y} - \hat{b}_1 \bar{x} \\ \sum_{i=1}^N x_i y_i &= \bar{y} \sum_{i=1}^N x_i + \hat{b}_1 \left(\sum_{i=1}^N x_i^2 - \bar{x} \sum_{i=1}^N x_i \right) \end{cases}$$

Isolating \hat{b}_1

$$\begin{cases} \hat{b}_0 &= \bar{y} - \hat{b}_1 \bar{x} \\ \hat{b}_1 &= \frac{\sum_{i=1}^N x_i y_i - \bar{y} \sum_{i=1}^N x_i}{\sum_{i=1}^N x_i^2 - \bar{x} \sum_{i=1}^N x_i} \end{cases}$$

After a bit of algebra (see appendix), one can simplify to

$$\begin{cases} \hat{b}_0 &= \bar{y} - \hat{b}_1 \bar{x} \\ \hat{b}_1 &= \frac{\sum_{i=1}^N (y_i - \bar{y})(x_i - \bar{x})}{\sum_{i=1}^N (x_i - \bar{x})^2} \end{cases}$$

Which is strictly equivalent to

$$\begin{cases} \hat{b}_0 &= \bar{y} - \hat{b}_1 \bar{x} \\ \hat{b}_1 &= \frac{Cov(x, y)}{Var(x)} \end{cases}$$

Appendix A Simplifying \hat{b}_1

A.1 Finding the covariance

$$\sum_{i=1}^N x_i y_i - \bar{y} \sum_{i=1}^N x_i = N \times Cov(x, y)$$

This proof need some tricks and heavily play with the fact that $\sum_{i=1}^N x_i = \bar{x}$ and that $\sum_{i=1}^N \bar{y} = N\bar{y}$. For example, one could write

$$\bar{y} \sum_{i=1}^N x_i = N\bar{y}\bar{x} = \sum_{i=1}^N \bar{x}\bar{y} = \sum_{i=1}^N y_i \bar{x}$$

Proof.

$$\begin{aligned} \sum_{i=1}^N x_i y_i - \bar{y} \sum_{i=1}^N x_i &= \sum_{i=1}^N x_i y_i - N\bar{y}\bar{x} \\ &= \sum_{i=1}^N x_i y_i - N\bar{y}\bar{x} - N\bar{y}\bar{x} + N\bar{y}\bar{x} \\ &= \sum_{i=1}^N x_i y_i - \sum_{i=1}^N x_i \bar{y} - \sum_{i=1}^N \bar{x} y_i + \sum_{i=1}^N \bar{y}\bar{x} \\ &= \sum_{i=1}^N (x_i y_i - x_i \bar{y} - \bar{x} y_i + \bar{y}\bar{x}) \\ &= \sum_{i=1}^N (y_i - \bar{y})(x_i - \bar{x}) \\ &= N \times Cov(x, y) \end{aligned}$$

□

A.2 Finding the variance

$$\sum_{i=1}^N x_i^2 - \bar{x} \sum_{i=1}^N x_i = N \times Var(x)$$

This proof is very similar to the on for covariance.

Proof.

$$\begin{aligned}\sum_{i=1}^N x_i^2 - \bar{x} \sum_{i=1}^N x_i &= \sum_{i=1}^N x_i^2 - N\bar{x}\bar{x} \\ &= \sum_{i=1}^N x_i^2 - 2N\bar{x}\bar{x} + N\bar{x}\bar{x} \\ &= \sum_{i=1}^N x_i^2 - 2\bar{x} \sum_{i=1}^N x_i + N\bar{x}^2 \\ &= \sum_{i=1}^N (x_i^2 - 2\bar{x}x_i + \bar{x}^2) \\ &= \sum_{i=1}^N (x_i - \bar{x})^2 \\ &= N \times \text{Var}(x)\end{aligned}$$

□