

## Question-1

\* Linear Regression \*

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Ans: Considering linear model

$$y(x, w) = w_0 + \sum w_i x_i \quad \text{for } i=1 \text{ to } D$$

$D$  - features

Error:  $E_D(w) = \frac{1}{2} \sum_{n=1}^N \{y_n(x_n, w) - t_n\}^2$

Note: This Soln is formulated in terms of matrix / vector calculations.

→ Denotions:

Data →

$$\bar{X} = N \times (D+1)$$

$$\begin{bmatrix} \bar{x}^1 \\ \vdots \\ \bar{x}^N \end{bmatrix}$$

where  $\bar{x}_n \rightarrow$  row vector

$$\bar{x}^n = [x_0^n \quad x_1^n \quad \dots \quad x_{D+1}^n]$$

Here,  $\boxed{x_0^n = 1 \quad \forall n=1 \text{ to } N}$

$$\therefore \bar{x}^n = [1, x_1^n, \dots, x_{D+1}^n]$$

Here, for every data point  $\bar{x}^n$ ,  $\bar{\varepsilon}^n$  is added,  $\rightarrow$  which is Gaussian distributed

$$\bar{x}^n = \bar{x}^n + \bar{\varepsilon}^n$$

Such that

$$\bar{\varepsilon}^n = [\varepsilon_0^n \ \varepsilon_1^n \ \dots \ \varepsilon_D^n]$$

Here,  ~~$\varepsilon_0^n$~~   $\boxed{E(\varepsilon_i^n) = 0}$

$$E(\varepsilon_i^n \varepsilon_j^n) = \sigma^2 \delta_{ij}$$

$\Rightarrow$  ~~clear~~ NOTE:

Since ~~the~~ here  $\bar{\varepsilon}^n$  is a row vector, (Dx1)

where, it has (Dx1) elements, since

no error is added to  $\bar{x}_0^n \Rightarrow \boxed{\varepsilon_0^n = 0}$

$$\boxed{\varepsilon_0^n = 0}$$

$\rightarrow$  Each element among other 'D'  $\Rightarrow i \in [1, D]$

$\varepsilon_i$  are generated from a

Gaussian distribution of mean '0'

and variance:  $\sigma^2 =$

$$\therefore E(\varepsilon_i) = 0 \quad \forall i$$

$$\Rightarrow E(\varepsilon_i \varepsilon_j) = \sigma^2 \delta(i, j) ; i, j \in \{1, D\}$$

$$E(\varepsilon_0 \varepsilon_j) = 0 \Rightarrow \boxed{\varepsilon_0 = 0}$$

$$\Rightarrow \text{cov}(\varepsilon) = \sigma^2 \begin{bmatrix} 0 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ 0 & \dots & 1 & \dots \\ 0 & \dots & \dots & 1 \\ 0 & \dots & \dots & \dots & 1 \end{bmatrix}_{(D+1) \times (D+1)}$$

Note: First & row of  $\text{cov}(\varepsilon) = \overline{0}$

(a) Minimising  $E_D$  averaged over the noise distribution.

$$\Rightarrow J(w) = E_{\varepsilon_i} \left( \left( \frac{1}{2} \right) \sum_{n=1}^N \{ y_n \{ \bar{x}_n, w \} - t_n \}^2 \right)$$

$$\Rightarrow y_n \{ \bar{x}_n, w \} = \underbrace{\bar{x}_n^T}_{1 \times (D+1)} \cdot \underbrace{\bar{w}}_{(D+1) \times 1}$$

$t_n \rightarrow$  ground truth.

Matrix form

$$W = \begin{bmatrix} w_0 \\ \vdots \\ w_D \end{bmatrix}_{(D+1, 1)} \rightarrow \text{weight vector}$$

$$J(W) = E_{\epsilon} \left( \frac{1}{2} \| \hat{X}W - t \|^2 \right)$$

Here  $t = \begin{bmatrix} t_1 \\ \vdots \\ t_N \end{bmatrix}_{N \times 1} \rightarrow$  ground truth.

Here  $\hat{X}$  is the noised data

$$\hat{X} = \bar{X} + \epsilon$$

(N, D+1)

$$\Rightarrow J(W) = E_{\epsilon} \left( \frac{1}{2} \| (X + \epsilon)W - t \|^2 \right)$$

We need to minimise  $J(W)$

Gradient for  $k$ th feature

$$\Rightarrow \frac{\partial J(W)}{\partial W_k} = \frac{\partial}{\partial W_k} E_{\epsilon} \left( \frac{1}{2} ((X + \epsilon)W - t)^T ((X + \epsilon)W - t) \right)$$

$$\frac{\partial J(W)}{\partial W} = E_{\epsilon} \left( (X + \epsilon)^T ((X + \epsilon)W - t) \right)$$

$$\left( \because \frac{\partial (\|A\|^2)}{\partial A} = \frac{\partial A^T A}{\partial A} = 2A \right)$$

$$\Rightarrow \frac{\partial J}{\partial W} = E_{\epsilon}((X + \epsilon)^T (X + \epsilon)W - (X + \epsilon)^T t)$$

$$= E_{\epsilon}((X^T X + \epsilon^T X + X^T \epsilon + \epsilon^T \epsilon)W - X^T t - \epsilon^T t)$$

$$\Rightarrow = (X^T X + E_{\epsilon}(\epsilon^T)X + X^T E_{\epsilon}(\epsilon) + E_{\epsilon}(\epsilon^T \epsilon))W - X^T t - E(\epsilon^T t)$$

$\because X, t$  are independent of  $\epsilon$

$$\Rightarrow E(X\epsilon) = X E(\epsilon)$$

$$\frac{\partial J}{\partial W} = (X^T X + 0 + 0 + \text{Cov}(\epsilon))W - X^T t - 0$$

$$\boxed{\frac{\partial J}{\partial W} = (X^T X + \text{Cov}(\epsilon))W - X^T t}$$

To minimise  $J$ ,  $\frac{\partial J}{\partial W} = 0$

$$\Rightarrow (X^T X + \text{Cov}(\epsilon))W = X^T t$$

$$\boxed{W = (X^T X + \text{Cov}(\epsilon))^{-1} X^T t} \rightarrow \textcircled{1}$$

$\hookrightarrow$  Result -1.

Part 3:

Minimising sum of squares error for  
noise free variables with weight decay  
regularisation

$$\text{ie; } J(w) = \frac{1}{2} \sum_{n=1}^N \{y_n(x_n, w) - t_n\}^2 + \lambda w^T w$$

$$J(w) = \frac{1}{2} \|XW - t\|^2 + \frac{\lambda}{2} \|\tilde{W}\|^2$$

$\downarrow$  noise-free data       $\downarrow$  regularisation

$$\text{Here, } w = \begin{bmatrix} w_0 \\ \vdots \\ w_D \end{bmatrix}_{(D+1 \times 1)}$$

Note The regularisation has a omitted  $w_0$

$$\text{ie, } \tilde{w} = \begin{bmatrix} 0 \\ w_1 \\ \vdots \\ w_D \end{bmatrix}$$



$$J(W) = \frac{1}{2} \|XW - t\|^2 + \frac{\lambda}{2} \|\tilde{W}\|^2$$

$$\frac{\partial J(\cdot)}{\partial W} = \left(\frac{1}{2}\right)(2) ((XW - t)^T(X)) + \frac{\lambda}{2} (2(W^T) \tilde{I}) = 0$$

Note:  $\frac{\partial \| \tilde{W} \|^2}{\partial W} = 2W^T \tilde{I}$ ,  $\tilde{I} = \begin{bmatrix} 0 & 0 & 0 & \dots & 1 \\ 0 & 1 & 0 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & 0 \end{bmatrix}$

→ First row of  $\tilde{I} = \bar{0}$  (Not (0+1X0+1))

$$\Rightarrow \frac{\partial J}{\partial W} = (XW - t)^T X + \lambda (W^T \tilde{I})$$

$$\Rightarrow \text{To minimise, } \frac{\partial J}{\partial W} = 0$$

$$\Rightarrow (XW - t)^T X + \lambda (W^T \tilde{I}) = 0$$

$$\Rightarrow X^T (XW - t) + \lambda (\tilde{I} W) = 0$$

$$\Rightarrow (X^T X + \lambda \tilde{I}) W = X^T t$$

$$\Rightarrow \underline{W = (X^T X + \lambda \tilde{I})^{-1} (X^T t)} \quad \text{--- (2)}$$

Conclusion:

From (1):

$$W = (X^T X + \text{Cov}(\epsilon))^T (X^T t)$$

From (2):

$$W = (X^T X + \lambda \tilde{I})^T (X^T t)$$

We know that  $\text{Cov}(\epsilon) = \begin{bmatrix} 0 & 0 & \dots & 0 \\ 0 & \sigma^2 & \dots & 0 \\ 0 & 0 & \sigma^2 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & 0 & \sigma^2 \end{bmatrix}$   
( $(D+1) \times (D+1)$ )

$$\text{Cov}(\epsilon) = \sigma^2 \begin{bmatrix} 0 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & 1 \end{bmatrix} = \sigma^2 \tilde{I}$$

Both results are the same

where  $\boxed{\lambda = \sigma^2}$

— X —