

## 2023-2024 Graphical Model Notes

### Chapter 3 Exponential Families and Contingency Tables

- Denote  $X_V \equiv (X_v : v \in V)$ , indexed by  $V = \{1, \dots, p\}$ . Each  $X_v$  takes values in the set  $\mathcal{X}_v$ . For a subset  $A \subseteq V$ , we write  $X_A$  to denote  $(X_v : v \in A)$ .
- Let  $p(\cdot; \theta)$  be a collection of probability densities over  $\mathcal{X}$  indexed by  $\theta \in \Theta$ . We say that  $p$  is an **exponential family** if it can be written as

$$\begin{aligned} p(x; \theta) &= \exp \left\{ \sum_i \theta_i \phi_i(x) - A(\theta) - C(x) \right\} \\ &= \exp \{ \langle \theta, \phi(x) \rangle - A(\theta) - C(x) \} \end{aligned}$$

- The family is said to be **regular** if  $\Theta$  is a non-empty open set.
- The functions  $\phi_i$  are the **sufficient statistics**.
- The components  $\theta_i$  are the **canonical/natural parameters**.
- The function  $A(\theta)$  is the **cumulant function** such that the distribution normalises:

$$A(\theta) = \log \int \exp \{ \langle \theta, \phi(x) \rangle - C(x) \} dx$$

- The function  $Z(\theta) \equiv e^{A(\theta)}$  is the **partition function**.
- We have

$$\nabla A(\theta) = \mathbb{E}_\theta \phi(X) \quad \nabla \nabla^T A(\theta) = \text{Cov}_\theta \phi(X)$$

$A(\theta)$  and  $-\log p(x; \theta)$  are convex in  $\theta$ , and the map  $\mu(\theta) : \theta \mapsto \nabla A(\theta)$  is bijective, named the **mean function**. We care about convexity because it does not have multiple local minima, which in turn facilitates computation.

- Let  $X_V = (X_1, \dots, X_p)^T \in \mathbb{R}^p$  be a random vector. Let  $\mu \in \mathbb{R}^p$  and  $\Sigma \in \mathbb{R}^{p \times p}$  be a positive definite symmetric matrix. We say that  $X_V$  has a **multivariate Gaussian distribution** with  $\mu$  and  $\Sigma$  if the joint density is

$$\begin{aligned} f(x_V) &= \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} (x_V - \mu)^T \Sigma^{-1} (x_V - \mu) \right\} \\ &= \frac{1}{(2\pi)^{\frac{p}{2}}} \exp \left\{ -\frac{1}{2} x_V^T K x_V + \mu^T K x_V - \frac{1}{2} \mu^T K \mu + \frac{1}{2} \log |K| \right\} \quad x_V \in \mathbb{R}^p \end{aligned}$$

Here,  $K \equiv \Sigma^{-1}$  is called the **concentration matrix**. Let

$$\phi(x_V) = \left( x_v, -\frac{1}{2} x_V x_V^T \right) \quad \theta = (K\mu, K)$$

we could easily tell that the multivariate Gaussian distribution is an exponential family<sup>1</sup>.

<sup>1</sup>For two matrices  $A$  and  $B$ , we have  $\langle A, B \rangle = \text{tr}(A, B^T)$ .

- Let  $X_V$  have a multivariate Gaussian distribution with concentration matrix  $K = \Sigma^{-1}$ , then  $X_i \perp\!\!\!\perp X_j \mid X_{V \setminus \{i,j\}}$  iff  $k_{ij} = 0$ .
- Let  $X_V^{(i)} = (X_1^{(i)}, \dots, X_p^{(i)})$  be sampled over individuals  $i = 1, \dots, n$  and define

$$n(x_V) \equiv \sum_{i=1}^n \mathbb{1}\{X_1^{(i)} = x_1, \dots, X_p^{(i)} = x_p\}$$

the number of individuals who have the response pattern  $x_V$ . These counts are the sufficient statistics for the multinomial model with log-likelihood

$$\begin{aligned} l(p; n) &= \sum_{x_V} n(x_V) \log p(x_V) \\ &= \sum_{x_V \neq 0_V} n(x_V) \log \frac{p(x_V)}{p(0_V)} + n \log p(0_V) \end{aligned}$$

where  $0_V$  is the vector of zeros,  $p(x_V) \geq 0$ , and  $\sum_{x_V} p(x_V) = 1$ . The multinomial distribution is also an exponential family with

- Sufficient statistics given by  $n(x_V)$ .
- Canonical parameters given by  $\log \frac{p(x_V)}{p(0_V)}$ .
- Convex cumulant function given by

$$-\log p(0_V) = \log \left( 1 + \sum_{x_V \neq 0_V} e^{\theta(x_V)} \right)$$

Each possibility  $x_V$  is called a **cell** of the table. Given  $A \subseteq V$ ,

$$n(x_A) \equiv \sum_{x_B} n(x_A, x_B)$$

where  $B = V \setminus A$  is called the **marginal table**.

- The **log-linear** parameters for  $p(x_V) > 0$  are defined by the relation

$$\begin{aligned} \log p(x_V) &= \sum_{A \subseteq V} \lambda_A(x_A) \\ &= \lambda_\emptyset + \lambda_1(x_1) + \dots + \lambda_V(x_V) \end{aligned}$$

and the identifiability constraint  $\lambda_A(x_A) = 0$  whenever  $x_a = 1$  for some  $a \in A$ .

- Consider a  $2 \times 2$  contingency table with probabilities  $\pi_{ij}$ . The log-linear parametrisation has

$$\begin{aligned} \log \pi_{11} &= \lambda_\emptyset & \log \pi_{21} &= \lambda_\emptyset + \lambda_X \\ \log \pi_{12} &= \lambda_\emptyset + \lambda_Y & \log \pi_{22} &= \lambda_\emptyset + \lambda_X + \lambda_Y + \lambda_{XY} \end{aligned}$$

We can deduce that

$$\lambda_{XY} = \log \frac{\pi_{11}\pi_{22}}{\pi_{21}\pi_{12}}$$

and  $e^{\lambda_{XY}}$  is called the **odds ratio** between X and Y.

- Let  $X_i \sim P(\mu_i)$  independently, and let  $N = \sum_{i=1}^k X_i$ . Then,

$$N \sim P\left(\sum_i \mu_i\right)$$

$$(X_1, \dots, X_k)^T \mid N = n \sim \text{Multinom}(n, (\pi_1, \dots, \pi_k)^T)$$

where  $\pi_i = \frac{\mu_i}{\sum_j \mu_j}$ .

- Let  $p > 0$  be a discrete distribution on  $X_V$  with associated log-linear parameters  $\lambda_C, C \subseteq V$ . The conditional independence  $X_a \perp\!\!\!\perp X_b \mid X_{V \setminus \{a,b\}}$  holds if and only if  $\lambda_C = 0$  for all  $\{a,b\} \subseteq C \subseteq V$ .

## Chapter 4 Undirected Graphical Models

- Let  $V$  be a finite set. An **undirected graph**  $\mathcal{G}$  is a pair  $(V, E)$  where
  - $V$  are the **vertices/nodes**.
  - $E \subseteq \{\{i, j\} : i, j \in V, i \neq j\}$  is a set of unordered distinct pairs of  $V$  called **edges**.

We represent graphs by drawing the vertices and then joining pairs of vertices by a line if there is an edge between them.

- We write  $i \sim j$  if  $\{i, j\} \in E$ , and say they are **adjacent** in the graph. The vertices adjacent to  $i$  are called the **neighbours** of  $i$ , and the set of neighbours is often called the **boundary** of  $i$  and denoted by  $\text{bd}_{\mathcal{G}}(i)$ .
- A **path** in a graph is a sequence of adjacent vertices without repetition. The **length** of a path is the number of edges in it.
- Given a subset of vertices  $W \subseteq V$ , we define the **induced subgraph**  $\mathcal{G}_W$  of  $\mathcal{G}$  to be the graph with vertices  $W$ , and all edges from  $\mathcal{G}$  whose endpoints are contained in  $W$ .
- We say  $C \subseteq V$  is **complete** if  $i \sim j$  for every  $i, j \in C$ . A maximal<sup>2</sup> complete set is called a **clique**. The set of cliques in a graph is denoted by  $\mathcal{C}(\mathcal{G})$ .
- Let  $A, B, S \subseteq V$ . We say that  $A$  and  $B$  are **separated** by  $S$  in  $\mathcal{G}$  ( $A \perp_s B \mid S[\mathcal{G}]$ ) if every path from any  $a \in A$  to any  $b \in B$  contains at least one vertex in  $S$ .  $A$  and  $B$  are separated by  $S$  (where  $S \cap A = S \cap B = \emptyset$ ) iff  $A$  and  $B$  are separated by  $\emptyset$  in  $\mathcal{G}_{V \setminus S}$ .

<sup>2</sup>**Maximal** means if one is to add another vertex into the graph, the graph will no longer be complete. However, graphs in the cliques do not necessarily need to have the same number of vertices.

- Let  $\mathcal{G}$  be a graph with vertices  $V$ , and let  $p$  be a probability distribution over the random variables  $X_V$ . We say that  $p$  satisfies the **pairwise Markov property** for  $\mathcal{G}$  if

$$i \not\sim j \in \mathcal{G} \implies X_i \perp\!\!\!\perp X_j \mid X_{V \setminus \{i,j\}}[p]$$

We say that  $p$  satisfies the **global Markov property** for  $\mathcal{G}$  if for any disjoint sets  $A, B, S$

$$A \perp_s B \mid S \subseteq \mathcal{G} \implies X_A \perp\!\!\!\perp X_B \mid X_S[p]$$

- A distribution  $p$  is said to **factorises** according to graph  $\mathcal{G}$  if

$$p(x_V) = \prod_{C \in \mathcal{C}(\mathcal{G})} \psi_C(x_C)$$

The functions  $\psi_C : \mathbb{R}^{|C|} \rightarrow \mathbb{R}$  are called **potentials**.

- If  $p(x_V)$  factorises according to  $\mathcal{G}$ , then  $p$  is globally Markov with respect to  $\mathcal{G}$ .
- (**Hammersley-Clifford Theorem**). If  $p(x_V) > 0$  obeys the pairwise Markov property with respect to  $\mathcal{G}$ , then  $p$  factorises according to  $\mathcal{G}$ .

– The followings always hold:

$$\text{factorisation} \implies \text{global Markov} \implies \text{pairwise Markov}$$

The following holds if  $p$  is strictly positive:

$$\text{pairwise Markov} \implies \text{factorisation}$$

- Given a graph  $\mathcal{G}$  with vertices  $V = A \cup B \cup S$  where  $A, B, S$  are disjoint sets. We say that  $(A, S, B)$  constitutes a **decomposition** of  $\mathcal{G}$  if:

- $\mathcal{G}_S$  is complete;
- $A$  and  $B$  are separated by  $S$  in  $\mathcal{G}$

If  $A$  and  $B$  are both non-empty, we say the decomposition is **proper**. If not, we say the decomposition is a **prime**.

- A graph is decomposable if either it is complete or there is a proper decomposition  $(A, S, B)$  and  $\mathcal{G}_{A \cup S}, \mathcal{G}_{B \cup S}$  are decomposable.
- Let  $C_1, C_2, \dots, C_k$  be a collection of subsets. We say that the sequence satisfies the **running intersection property** if  $\forall j \geq 2$ ,

$$C_j \cap \bigcup_{i=1}^{j-1} C_i = C_j \cap C_{\sigma(j)} \quad \sigma(j) < j$$

- If  $C_1, \dots, C_k$  satisfy the running intersection property, then there is a graph whose cliques are  $\mathcal{C} = \{C_1, \dots, C_k\}$ .
- Let  $\mathcal{G}$  be an undirected graph. A **cycle** is a sequence of vertices  $\langle v_1, \dots, v_k \rangle$  for  $k \geq 3$  such that there is a path  $v_1 - \dots - v_k$  and an edge  $v_k - v_1$ . A **chord** on a cycle is any edge between two vertices not adjacent on the cycle. A graph is **chordal** or **triangulated** if whenever there is a cycle of length greater or equal to 4, it contains a chord.
- Let  $\mathcal{G}$  be an undirected graph. The followings are equivalent:
  - $\mathcal{G}$  is decomposable;
  - $\mathcal{G}$  is triangulated;
  - Every minimal separator of  $a \not\sim b$  is complete;
  - The cliques of  $\mathcal{G}$  satisfy the running intersection property, starting with  $C$ .