## Classification so far

First solve the inference problem of determining the posterior class probabilities  $p(C_k|\mathbf{x})$ , and then subsequently use decision theory to assign each new  $\mathbf{x}$  to one of the classes. Approaches that model the posterior probabilities directly are called *discriminative models*.

Find a function  $f(\mathbf{x})$ , called a discriminant function, which maps each input  $\mathbf{x}$  directly onto a class label. For instance, in the case of two-class problems,  $f(\cdot)$  might be binary valued and such that f=0 represents class  $\mathcal{C}_1$  and f=1 represents class  $\mathcal{C}_2$ . In this case, probabilities play no role.

#### Generative classifiers

First solve the inference problem of determining the class-conditional densities  $p(\mathbf{x}|\mathcal{C}_k)$  for each class  $\mathcal{C}_k$  individually. Also separately infer the prior class probabilities  $p(\mathcal{C}_k)$ . Then use Bayes' theorem in the form

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{p(\mathbf{x})}$$

Equivalently, we can model the joint distribution  $p(\mathbf{x}, C_k)$  directly and then normalize to obtain the posterior probabilities

Approaches that explicitly or implicitly model the distribution of inputs as well as outputs are known as *generative models* 

Approaches that explicitly or implicitly model the distribution of inputs as well as outputs are known as **generative models**, because by sampling from them it is possible to generate synthetic data points in the input space.

Approaches that model the posterior probabilities directly are called **discriminative models** 

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)} \qquad a = \ln \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_2)p(C_2)}$$
$$= \frac{1}{1 + \exp(-a)} = \sigma(a)$$

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{\sum_j p(\mathbf{x}|C_j)p(C_j)}$$
$$= \frac{\exp(a_k)}{\sum_j \exp(a_j)} \qquad a_k = \ln p(\mathbf{x}|C_k)p(C_k).$$

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)}$$
$$= \frac{1}{1 + \exp(-a)} = \sigma(a) \qquad a = \ln \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_2)p(C_2)}$$

$$p(\mathbf{x}|\mathcal{C}_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^{\mathrm{T}} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)\right\}.$$

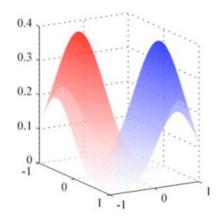
$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)}$$
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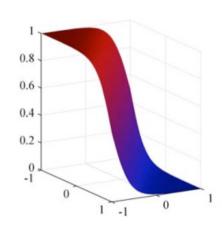
$$p(\mathbf{x}|\mathcal{C}_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^{\mathrm{T}} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)\right\}.$$

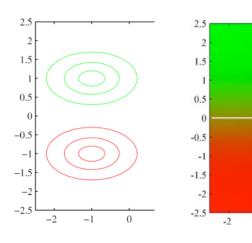
$$\begin{split} p(\mathcal{C}_1|\mathbf{x}) &= \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x} + w_0) \qquad \mathbf{w} &= \Sigma^{-1}(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2) \\ w_0 &= -\frac{1}{2}\boldsymbol{\mu}_1^{\mathrm{T}}\Sigma^{-1}\boldsymbol{\mu}_1 + \frac{1}{2}\boldsymbol{\mu}_2^{\mathrm{T}}\Sigma^{-1}\boldsymbol{\mu}_2 + \ln\frac{p(\mathcal{C}_1)}{p(\mathcal{C}_2)}. \end{split}$$

$$p(\mathbf{x}|\mathcal{C}_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^{\mathrm{T}} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)\right\}.$$

$$p(\mathcal{C}_1|\mathbf{x}) = \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x} + w_0) \qquad \mathbf{w} = \mathbf{\Sigma}^{-1}(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2)$$
$$w_0 = -\frac{1}{2}\boldsymbol{\mu}_1^{\mathrm{T}}\mathbf{\Sigma}^{-1}\boldsymbol{\mu}_1 + \frac{1}{2}\boldsymbol{\mu}_2^{\mathrm{T}}\mathbf{\Sigma}^{-1}\boldsymbol{\mu}_2 + \ln\frac{p(\mathcal{C}_1)}{p(\mathcal{C}_2)}.$$







$$p(\mathbf{x}|\mathcal{C}_k) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\mathbf{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^\mathrm{T} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}_k)\right\}.$$

$$p(\mathcal{C}_1|\mathbf{x}) = \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x} + w_0) \qquad \mathbf{w} = \mathbf{\Sigma}^{-1}(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2)$$

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$$p(\mathcal{C}_k|\mathbf{x}) = \frac{p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)}{\sum_j p(\mathbf{x}|\mathcal{C}_j)p(\mathcal{C}_j)}$$

$$= \frac{\exp(a_k)}{\sum_j \exp(a_j)}$$

$$a_k = \ln p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k). \qquad a_k(\mathbf{x}) = \mathbf{w}_k^T \mathbf{x} + w_{k0}$$

$$\mathbf{w}_k = \mathbf{\Sigma}^{-1} \mu_k$$

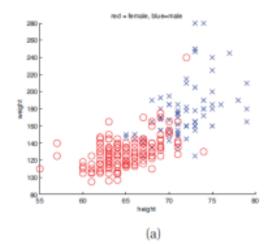
$$w_{k0} = -\frac{1}{2} \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \ln p(\mathcal{C}_k).$$

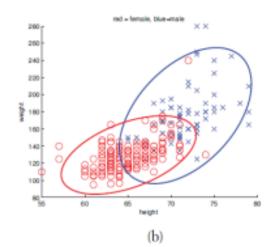
model the class-conditional densities  $p(x/C_k)$ , as well as the class priors  $p(C_k)$ 

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)}$$

class-conditional densities p(x/C)

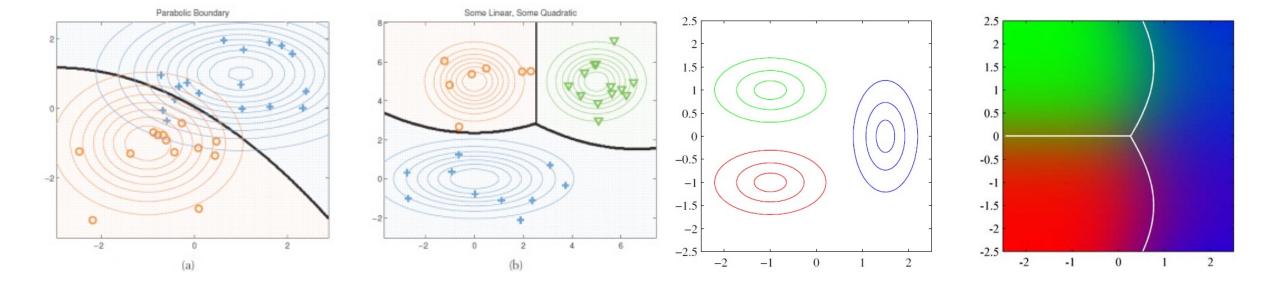
$$p(\mathbf{x}|y=c, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c)$$





## Quadratic discriminant analysis

$$p(y = c | \mathbf{x}, \boldsymbol{\theta}) = \frac{\pi_c |2\pi \boldsymbol{\Sigma}_c|^{-\frac{1}{2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_c)^T \boldsymbol{\Sigma}_c^{-1} (\mathbf{x} - \boldsymbol{\mu}_c)\right]}{\sum_{c'} \pi_{c'} |2\pi \boldsymbol{\Sigma}_{c'}|^{-\frac{1}{2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_c)^T \boldsymbol{\Sigma}_c^{-1} (\mathbf{x} - \boldsymbol{\mu}_c)\right]}$$



## Linear Discriminant Analysis

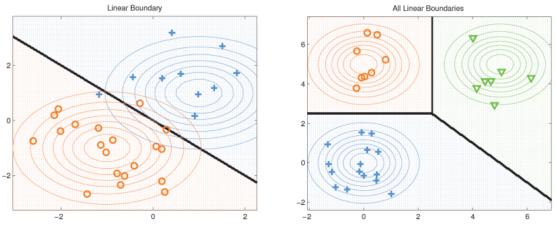
• covariance matrices are **tied** or **shared** across classes,  $\Sigma c = \Sigma$ .

$$p(y = c | \mathbf{x}, \boldsymbol{\theta}) \propto \pi_c \exp \left[ \mu_c^T \mathbf{\Sigma}^{-1} \mathbf{x} - \frac{1}{2} \mathbf{x}^T \mathbf{\Sigma}^{-1} \mathbf{x} - \frac{1}{2} \mu_c^T \mathbf{\Sigma}^{-1} \mu_c \right]$$

$$= \exp \left[ \mu_c^T \mathbf{\Sigma}^{-1} \mathbf{x} - \frac{1}{2} \mu_c^T \mathbf{\Sigma}^{-1} \mu_c + \log \pi_c \right] \exp \left[ -\frac{1}{2} \mathbf{x}^T \mathbf{\Sigma}^{-1} \mathbf{x} \right]$$

$$p(y = c | \mathbf{x}, \boldsymbol{\theta}) = \frac{e^{\boldsymbol{\beta}_c^T \mathbf{x} + \gamma_c}}{\sum_{c'} e^{\boldsymbol{\beta}_{c'}^T \mathbf{x} + \gamma_{c'}}} \qquad \gamma_c = -\frac{1}{2} \mu_c^T \mathbf{\Sigma}^{-1} \mu_c + \log \pi_c$$

$$\boldsymbol{\beta}_c = \mathbf{\Sigma}^{-1} \mu_c$$



#### Generative Models: Parameter estimation

Models joint probability of observing input and output

$$p(\mathbf{x}_n, \mathcal{C}_1) = p(\mathcal{C}_1)p(\mathbf{x}_n|\mathcal{C}_1) = \pi \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_1, \boldsymbol{\Sigma}).$$

$$p(\mathbf{x}_n, \mathcal{C}_2) = p(\mathcal{C}_2)p(\mathbf{x}_n|\mathcal{C}_2) = (1 - \pi)\mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_2, \boldsymbol{\Sigma}).$$

$$p(\boldsymbol{t}|\pi, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}) = \prod_{n=1}^{N} \left[\pi \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})\right]^{t_n} \left[(1 - \pi)\mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_2, \boldsymbol{\Sigma})\right]^{1 - t_n}$$

#### Generative Models: Parameter estimation

Models joint probability of observing input and output

$$\log p(\mathcal{D}|\boldsymbol{\theta}) = \left[ \sum_{i=1}^{N} \sum_{c=1}^{C} \mathbb{I}(y_i = c) \log \pi_c \right] + \sum_{c=1}^{C} \left[ \sum_{i:y_i = c} \log \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c) \right]$$

Use Maximum Likelihood Estimation (MLE) to estimate parameters

#### Generative Models: Parameter estimation

Models joint probability of observing input and output

$$\log p(\mathcal{D}|\boldsymbol{\theta}) = \left[ \sum_{i=1}^{N} \sum_{c=1}^{C} \mathbb{I}(y_i = c) \log \pi_c \right] + \sum_{c=1}^{C} \left[ \sum_{i:y_i = c} \log \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c) \right]$$

$$\hat{\boldsymbol{\pi}}_c = \frac{N_c}{N}, \qquad \hat{\boldsymbol{\mu}}_c = \frac{1}{N_c} \sum_{i:y_i=c} \mathbf{x}_i, \quad \hat{\boldsymbol{\Sigma}}_c = \frac{1}{N_c} \sum_{i:y_i=c} (\mathbf{x}_i - \hat{\boldsymbol{\mu}}_c)(\mathbf{x}_i - \hat{\boldsymbol{\mu}}_c)^T$$

- MLE can badly overfit in high dimensions.
- Use a diagonal covariance matrix for each class, which assumes the features are conditionally independent; this is equivalent to using a **naive Bayes classifier**
- Other approaches: Use MAP and Bayesian approaches

## Naïve Bayes Classifier

Features are discrete

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)}$$

- class-conditional density p(x|y)
  - Number of parameters ?

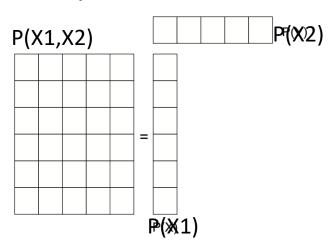
## Naïve Bayes Classifier

Features are discrete

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)}$$

- "naive" since we do not expect the features to be independent, but results in classifiers that work well
  - model is quite simple and hence it is relatively immune to overfitting, as a lower number of parameters need to be estimated due to independence assumption.
- Class-conditional density p(x|y)

$$p(\mathbf{x}|y=c, \boldsymbol{\theta}) = \prod_{j=1}^{D} p(x_j|y=c, \boldsymbol{\theta}_{jc})$$



# Naïve Bayes Classifier

• Features are conditionally independent given the class label.

$$p(\mathbf{x}|y=c,\boldsymbol{\theta}) = \prod_{j=1}^{D} p(x_j|y=c,\boldsymbol{\theta}_{jc})$$

- class-conditional density p(x|y)
  - In the case of real-valued features, we can use the Gaussian distribution:  $p(\mathbf{x}|y=c,\boldsymbol{\theta}) = \prod_{j=1}^{D} \mathcal{N}(x_j|\mu_{jc},\sigma_{jc}^2)$ , where  $\mu_{jc}$  is the mean of feature j in objects of class c, and  $\sigma_{jc}^2$  is its variance.
  - In the case of binary features,  $x_j \in \{0, 1\}$ , we can use the Bernoulli distribution:  $p(\mathbf{x}|y = c, \boldsymbol{\theta}) = \prod_{i=1}^{D} \text{Ber}(x_j|\mu_{jc})$ , where  $\mu_{jc}$  is the probability that feature j occurs in class c.
  - In the case of categorical features,  $x_j \in \{1, ..., K\}$ , we can model use the multinoulli distribution:  $p(\mathbf{x}|y=c, \boldsymbol{\theta}) = \prod_{j=1}^{D} \operatorname{Cat}(x_j|\boldsymbol{\mu}_{jc})$ , where  $\boldsymbol{\mu}_{jc}$  is a histogram over the K possible values for  $x_j$  in class c.

## Training NBC: ML Estimation

 $p(\mathbf{x}_i, y_i | \boldsymbol{\theta}) = p(y_i | \boldsymbol{\pi}) \prod_j p(x_{ij} | \boldsymbol{\theta}_j) = \prod_c \pi_c^{\mathbb{I}(y_i = c)} \prod_j \prod_c p(x_{ij} | \boldsymbol{\theta}_{jc})^{\mathbb{I}(y_i = c)}$ 

$$\log p(\mathcal{D}|\boldsymbol{\theta}) = \sum_{c=1}^{C} N_c \log \pi_c + \sum_{j=1}^{D} \sum_{c=1}^{C} \sum_{i:y_i=c} \log p(x_{ij}|\boldsymbol{\theta}_{jc})$$

Bernoulli class condiitonal likelihood

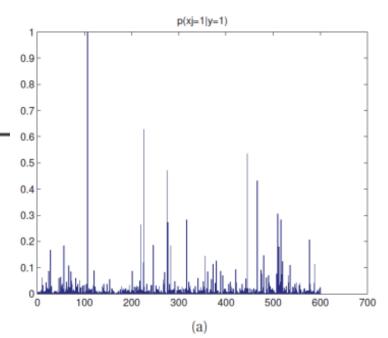
$$\hat{\pi}_c = \frac{N_c}{N} \qquad \hat{\theta}_{jc} = \frac{N_{jc}}{N_c}$$

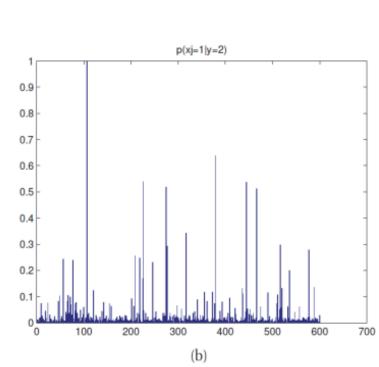
The method is easily generalized to handle features of mixed type.

## NBC: Algorithm

#### Algorithm 3.1: Fitting a naive Bayes classifier to binary features

```
1 N_c = 0, N_{jc} = 0;
2 for i = 1 : N do
      c=y_i // Class label of i'th example;
     N_c := N_c + 1;
     for j = 1 : D do
        6
8 \hat{\pi}_c = \frac{N_c}{N}, \hat{\theta}_{jc} = \frac{N_{jc}}{N}
```





# Spam Classification example

#### **HAM** examples

- d1: {good}
- d2: {very good}

```
P(y=ham) = 2/5
```

#### **SPAM** examples

- d3: {bad}
- d4: {very bad}
- d5: {very bad very bad}

$$p(y=spam)=3/5$$

- Test data: d6: {good bad very bad} SPAM OR HAM??
- Vocabulary : V={good, bad, very}
- Use Naive Bayes classifier:
- P(ham | d6)=  $\Pi$  P(word | y=ham) P(y=ham) = P(good | ham) \* P(bad | ham) \* P(very | ham)

Word frequencies wrt class	P(good   y)	P(bad   y)	P(very   y)
Class 0: y = spam	0/3	3/3	2/3
Class 1: y = ham	2/2	0/2	1/2

- Estimate parameters using ML:
- P(ham | d6)= P(good | ham) \* P(bad | ham) \* P(very | ham)\* P(ham)
- P(spam | d6)= P(good | spam) \* P(bad | spam) \* P(very | spam)\*
   P(spam)
- What is the problem with this?

Word frequencies wrt class	P(good   y)	P(bad   y)	P(very   y)
Class 0: y = spam	0/3	3/3	2/3
Class 1: y = ham	2/2	0/2	1/2

Estimate parameters using ML:

P(ham | d6)= P(good | ham) \* P(bad | ham) \* P(very | ham)\* P(ham)

Soln : Bayesian Naïve Bayes

Classifiers!

P(spam | d6)= P(good | spam) \* P(bad | spam) \* P(very | spam)\*
 P(spam)

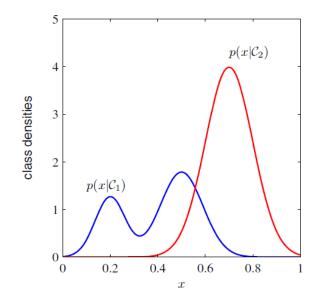
What is the problem with this?

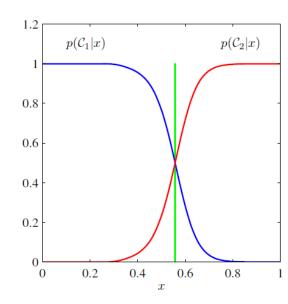
#### Discriminative vs Generative

• **Generative Models:** model the joint distribution p(x, Ck) directly and then normalize to obtain the posterior probabilities.

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{p(\mathbf{x})}$$

 Most demanding because it involves finding the joint distribution over both x and Ck. For many applications, x will have high dimensionality, and consequently we may need a large training set. if we only wish to make classification decisions, then it can be wasteful of computational resources but can be useful for detecting outliers or novel classes.





#### Discriminative vs Generative

• **Generative Models**: model the joint distribution p(x, Ck) directly and then normalize to obtain the posterior probabilities.

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{p(\mathbf{x})}$$

- Easy to adapt to new data: Use posterior probabilities as prior
- Multi-Modal Data: combining multiple modalities and domains

$$p(\mathbf{x}_{\mathrm{I}}, \mathbf{x}_{\mathrm{B}} | \mathcal{C}_k) = p(\mathbf{x}_{\mathrm{I}} | \mathcal{C}_k) p(\mathbf{x}_{\mathrm{B}} | \mathcal{C}_k).$$

- Easy to fit? very easy to fit generative classifiers. we can fit a naive Bayes model and an LDA model by simple counting and averaging. logistic regression requires solving a convex optimization problem which is much slower.
- Fit classes separately? In a generative classifier, we estimate the parameters of each class conditional density independently, so we do not have to retrain the model when we add more classes.
- Well-calibrated probabilities? Some generative models, such as naive Bayes, make strong independence assumptions which are often not valid
- Generative models can easily handle unlabelled data and missing features