

Macroeconomic Forecasting Lec_01

**Zahid Asghar, Professor, School of Economics,
QAU**

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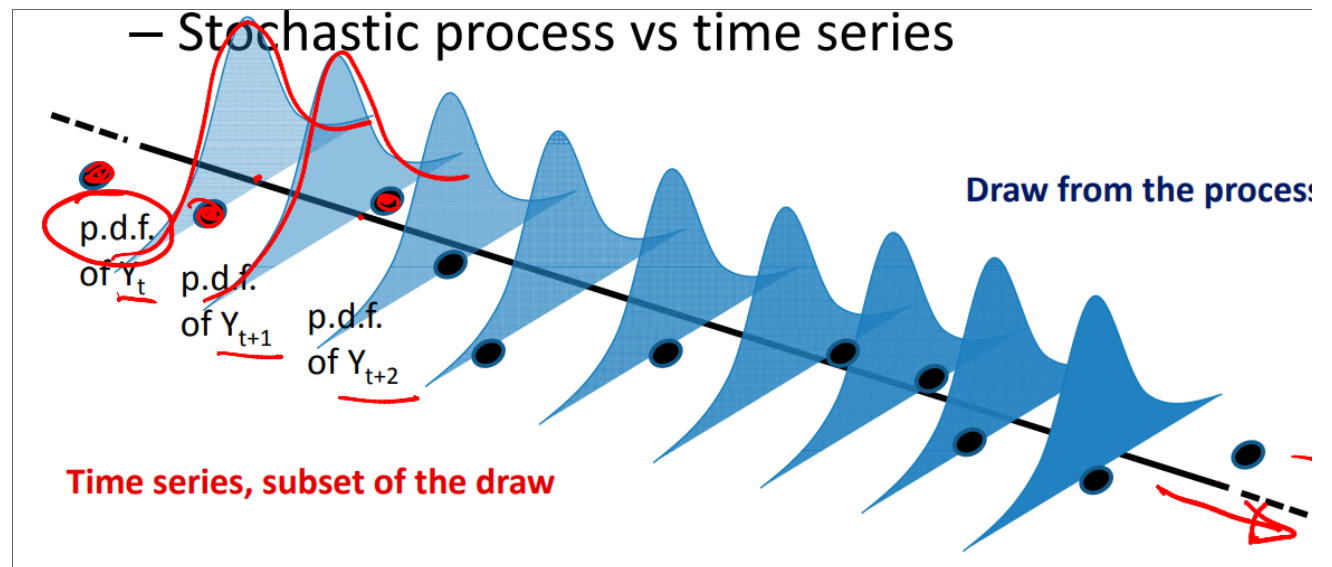


Properties of Time Series

Introduction

- General comments
 - Univariate analysis
 - Two general “classes” of processes
 - Both science and art (judgement):
 - Understanding behavior and forecasting
 - Assessing/testing

Univariate Analysis

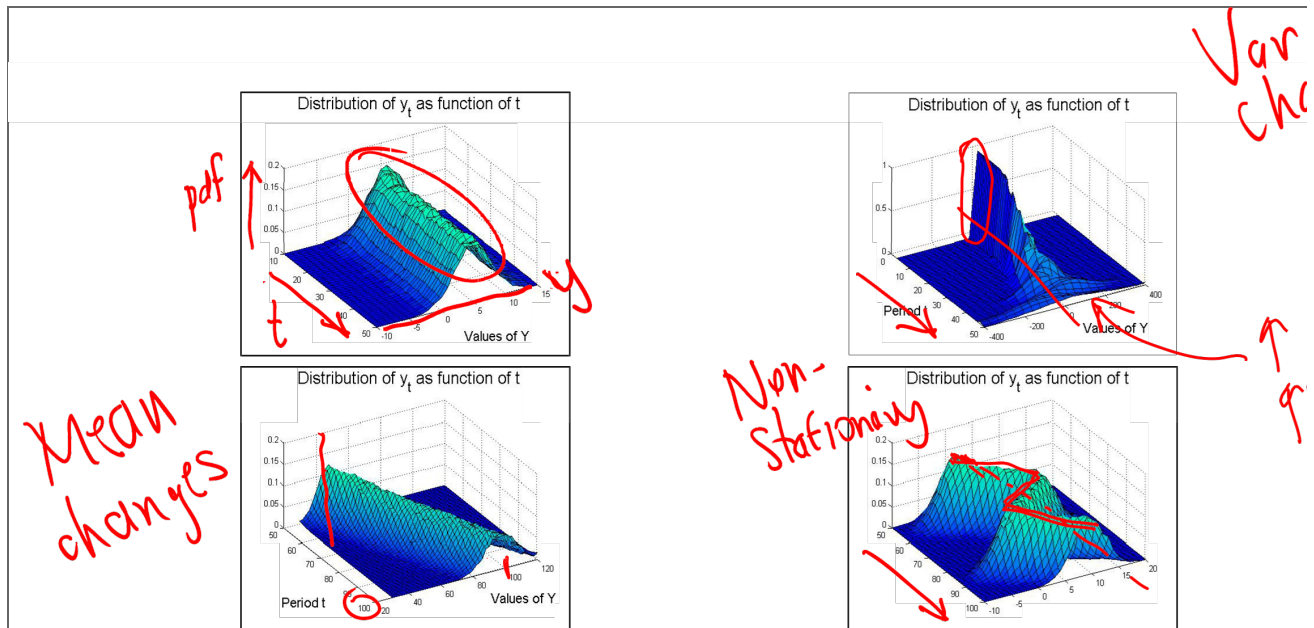


Stochastic Process vs time series

Each distribution is a draw from a **random process**.

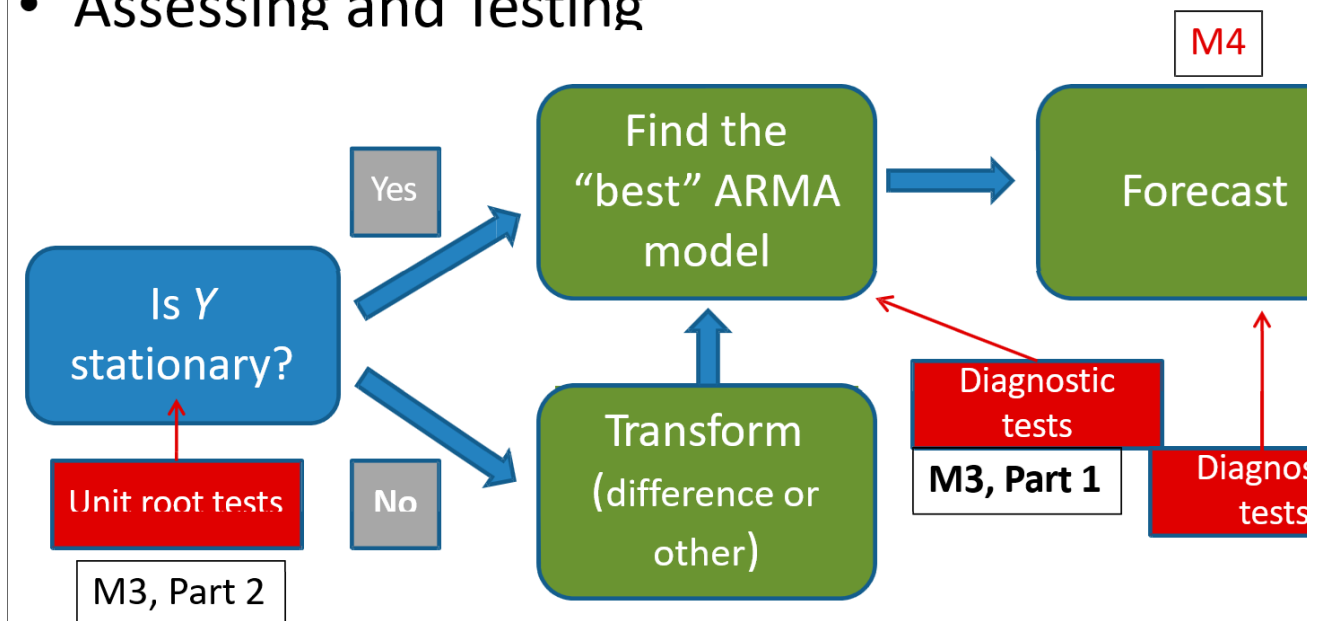
Introduction

- Two general “classes” of processes
 - stationary vs nonstationary
 - Unchanged distribution (pdf) over time?
 - Covariance stationary:
 - i. Unconditional mean $E(Y_t) = E(Y_{t+j}) = \mu$
 - ii. variance constant $Var(Y_t) = Var(Y_{t+j}) = \sigma_y^2$ and
 - iii. Covariance depends on time j that has elapsed between observations, not on reference period $Cov(Y_t, Y_{t+j}) = Cov(Y_s, Y_{s+j}) = \gamma_j$



Stochastic Process vs time series

- Assessing and Testing



Diagnostic-ARMA

Outline

Part 1 : Stationary processes

- Identification
- Estimation & Model Selection
- Putting it all together

Part 2: Nonstationary processes

- Characterization
- Testing

Part1 : Stationary Process

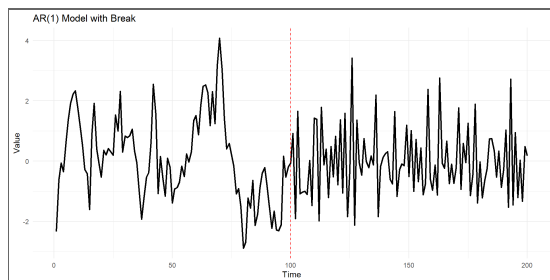
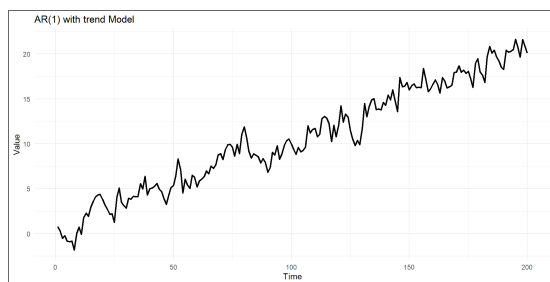
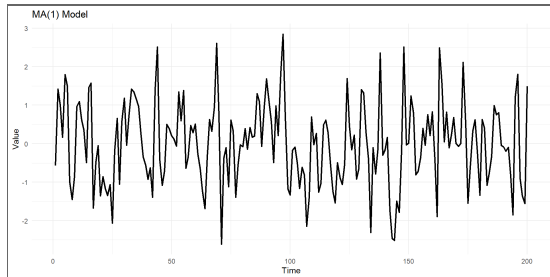
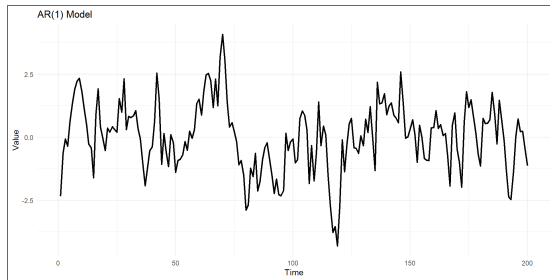
Just to remind you....

- Identification
- Estimation & Model Selection
- Putting it all together

The first step is visual inspection: graph and observe your data.

“You can observe a lot just by watching” Yogi Berra

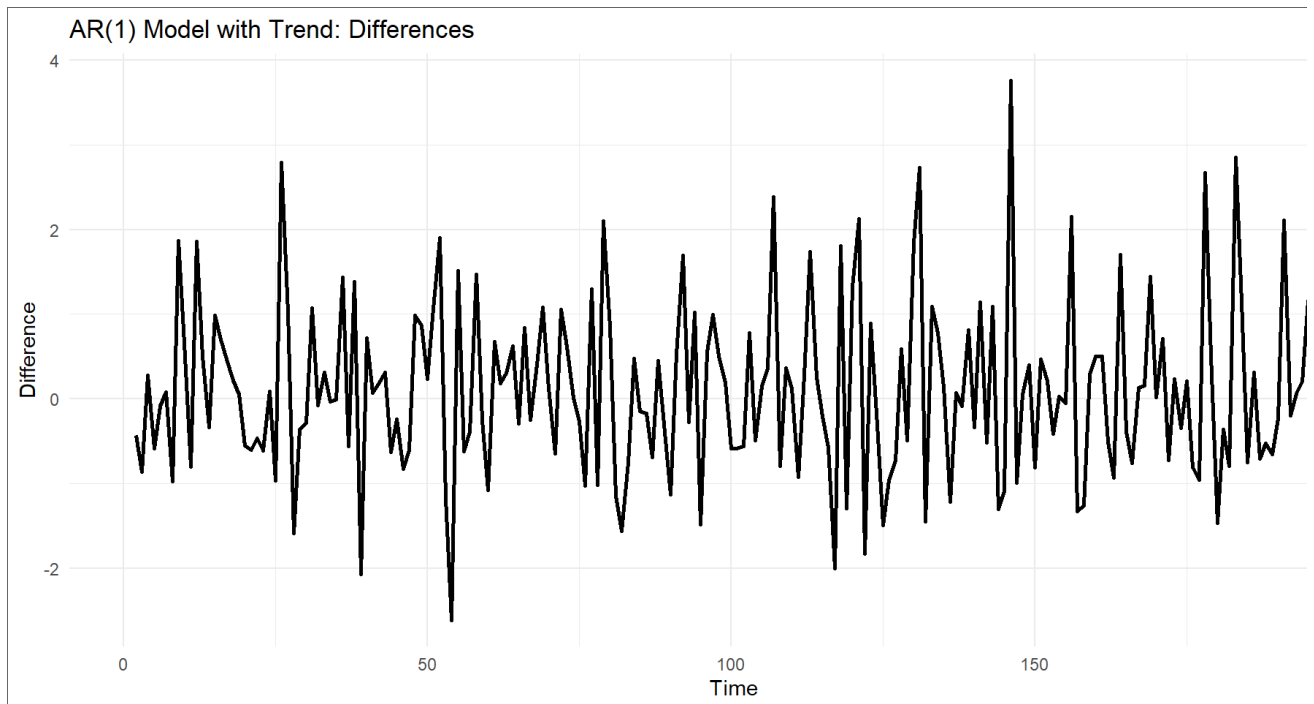
Plot plot and plot your data



Differenced model of trend variable (third case)

Difference can remove the trend.

$$y_t^* = y_t - y_{t-1}$$



Identification

Assuming that the process is stationary, there are three basic types that interest us:

Autoregressive (process) $y_t = a + b_1 y_{t-1} + b_2 y_{t-2} + \dots + b_p y_{t-p} + \epsilon_t$

Moving Average : AR(process) :

$$y_t = \mu + u_t + \phi_1 u_{t-1} + \phi_2 u_{t-2} + \dots + \phi_q u_{t-q} + \epsilon_t$$

Combined ARMA-process $y_t = a + b_1 y_{t-1} + b_2 y_{t-2} + \dots + b_p y_{t-p} + u_t + \phi_1 u_{t-1} + \phi_2 u_{t-2} + \dots + \phi_q u_{t-q} + \epsilon_t$

Some notation: AR(p), MA(q), ARMA(p,q), where p,q refer to the order (maximum lag) of the process

ϵ_t is a white noise disturbance:

$$E(\epsilon_t) = 0, Var(\epsilon_t) = \sigma^2, Cov(\epsilon_t, \epsilon_s) = 0, if t \neq s$$

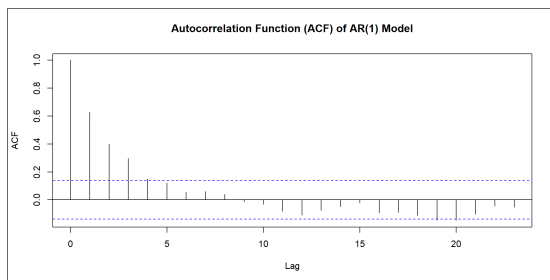
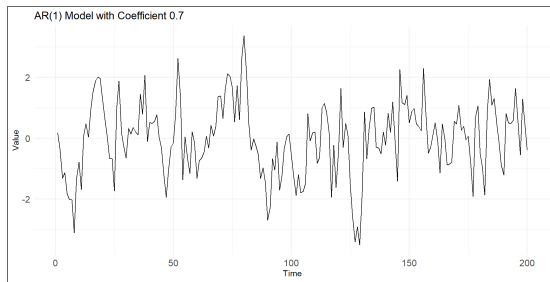
Tools for Identification

- Where are we? Where are we going?
- Stationary process (visual inspection) y
- Learned about possible processes for y
- Need to identify which one in order to understand, then eventually forecast y

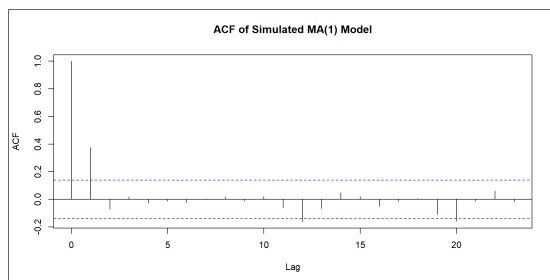
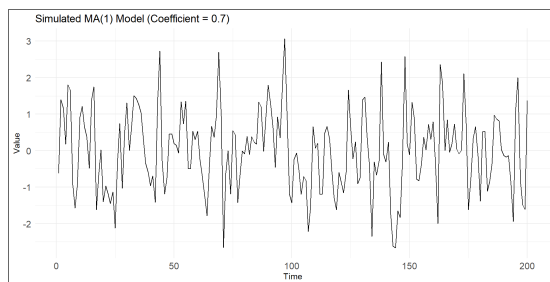
tools to help identify

Autocovariance and autocorrelation Relations between observations at different lags:

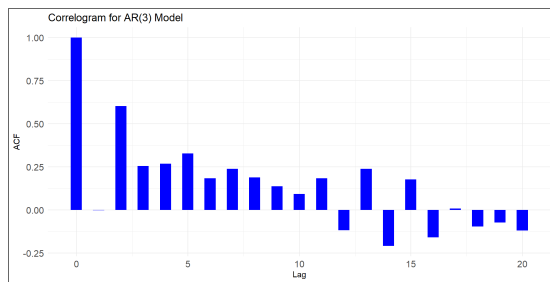
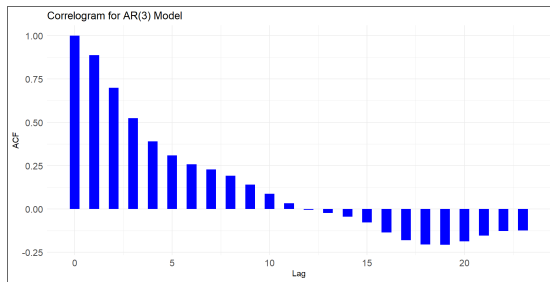
- Autocovariance $\gamma_j = E[(y_t - \mu)(y_{t-j} - \mu)]$
- Autocorrelation $\rho_j = \gamma_j / \gamma_0$
- ACF or Correlogram : graph of autocorrelations at each lag



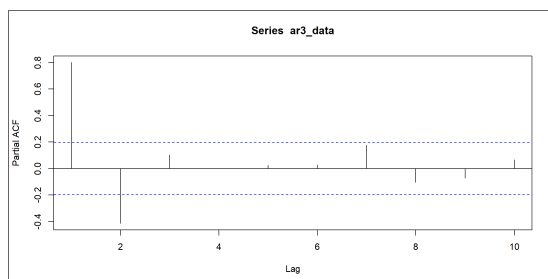
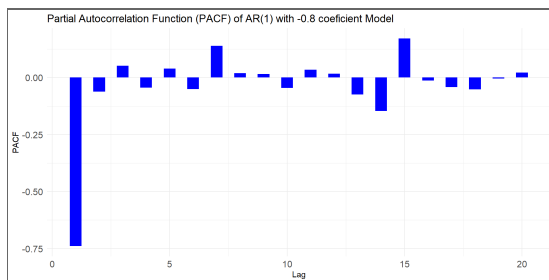
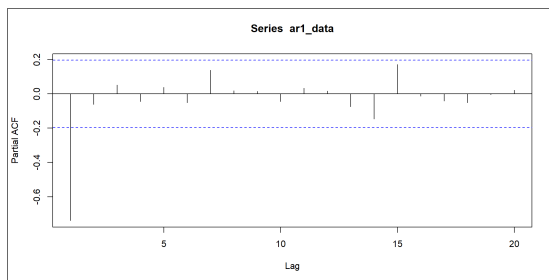
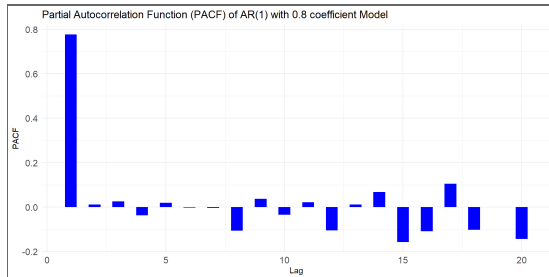
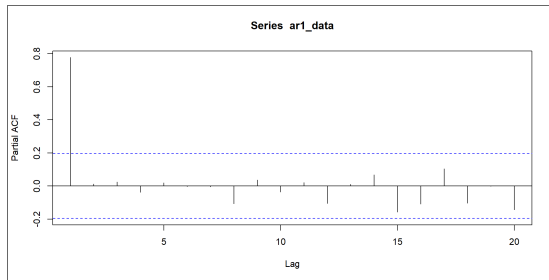
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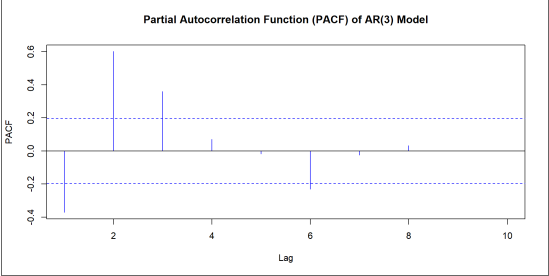
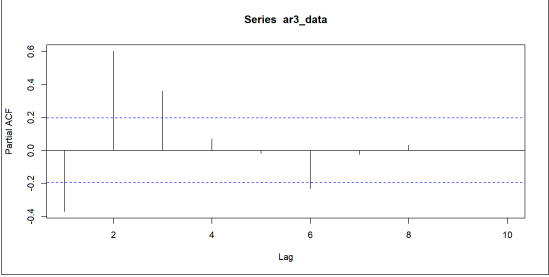
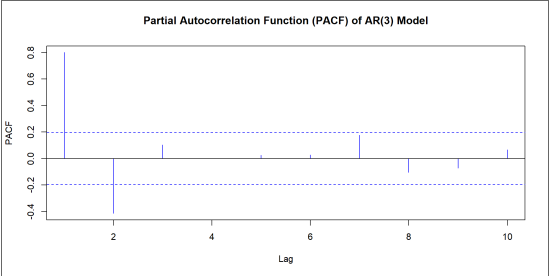


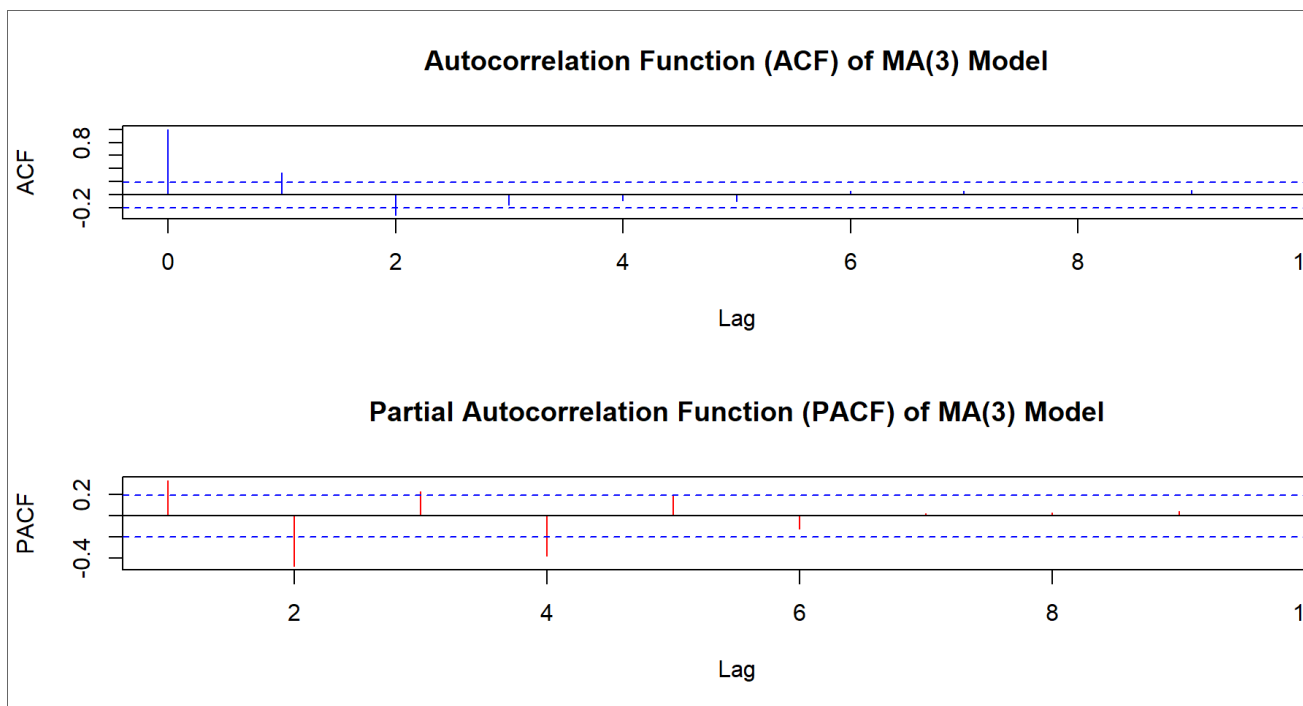
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Patterns for PACF







Stationary Time Series

We now have a tool (ACF, PACF) to help us identify the stochastic process underlying time series we are observing. Now we will:

- Summarize the basic patterns to look for
- Observe an actual data series and make an initial guess
- Observe an actual data series and make an initial guess
- Next step: estimate (several alternatives) based on this guess

Summary Table Pattern

	ACF	PACF
White noise	All ρ 's = 0	All b 's = 0
AR(1)	Geometric decay (oscillating if $b < 0$)	Cutoff after lag 1; $\rho_1 = b_1$
AR(p)	Decays toward zero, may oscillate	Cutoff after lag p.
MA(1)	Cutoff after lag 1.	Geometric decay (oscillating if $\phi < 0$)
MA(q)	Cutoff after lag q.	Decay (oscillating if $\phi < 0$)
ARMA(1,1)	Geometric decay after lag 1 (oscillating if $b < 0$)	Geometric decay after lag 1 (oscillating if $b < 0$)
ARMA(p,q)	Decay (direct or oscillatory) after lag q	Decay (direct or oscillatory) after lag p

Pattern for model identification

Tips

- ACF's that do not go to zero could be sign of nonstationarity
- ACF of both AR, ARMA decay gradually, drops to 0 for MA
- PACF decays gradually for ARMA, MA, drops to 0 for AR

Possible approach: begin with parsimonious low order AR, check residuals to decide on possible MA terms.

When looking at ACF, PACF

- Box-Jenkins provide sampling variance of the observed ACF and PACFs (r_s and b_s)
- Permits one to construct confidence intervals around each \rightarrow assess whether significantly $\neq 0$
- Computer packages provide this automatically

Estimation and Model Selection

- Decide on plausible alternative specifications (ARMA)
- Estimate each specification
- Choose “best” model, based on:
 - Significance of coefficients
 - Fit vs parsimony (criteria)
 - White noise residuals
 - Ability to forecast
- Account for possible structural breaks

Fit vs parsimony (information criteria):

- Additional parameters (lags) automatically improve fit but reduce forecast quality
- Tradeoff between fit and parsimony; widely used criteria:
- Akaike Information Criterion (AIC) $AIC = T\ln(SST) + 2(p + q + 1)$
- Schwartz Bayesian Criterion (BIC) $SBC = T\ln(SST) + (p + q + 1)\ln(T)$
- SBC is considered to be preferable for having more parsimonious models than AIC

White noise errors :

- Aim to eliminate autocorrelation in the residuals (could indicate that model does not reflect the lag structure well)
- Plot “standardized residuals” (ϵ_{it}) No more than 5% of them should lie outside $[-2, +2]$ over all periods
- Look at r_s , b_s (and significance) at different lags Box-Pierce Statistic: joint significance test up to lag s : $\chi^2 =$

$$Q = T \sum_{k=1}^s r_k^2$$

$$H_0 : \text{all } r_k = 0, H_1 : \text{at least one } r_k \neq 0$$

Forecastability

Can assess how well the model forecasts "out of sample":

- Estimate the model for a sub-sample (for example, the first 250 out of 300 observations).
- Use estimated parameters to forecast for the rest of the sample (last 50)
- Compute the "forecast errors" and assess:
 - Mean Squared Prediction Error
 - Granger-Newbold Test
 - Diebold-Mariano Test

Account for possible structural breaks:

- Does the same model apply equally well to the entire sample, or do parameters change (significantly) within the sample?

How to approach:

- Own priors/suspicion : Chow test for parameter change
- If priors not strong, recursive estimation, tests for parameter stability over the sample, for example, CUSUM

Learning all above with simulated data

Lets simulate MA(1), AR(1) series in Excel, R and one can use STATA/EVIEWS as well

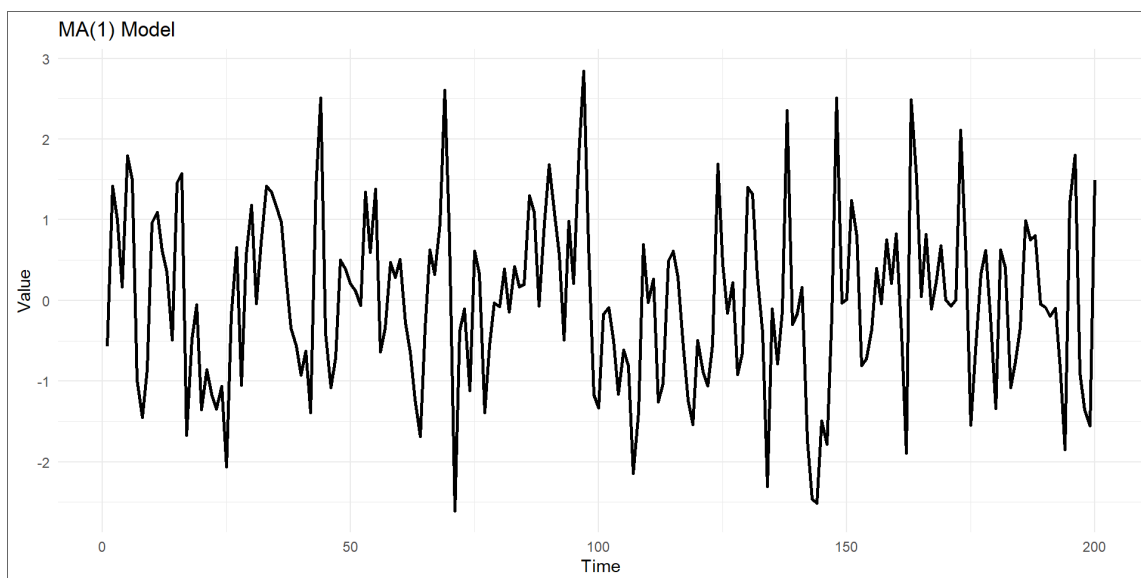
MA-1 code

MA1-output

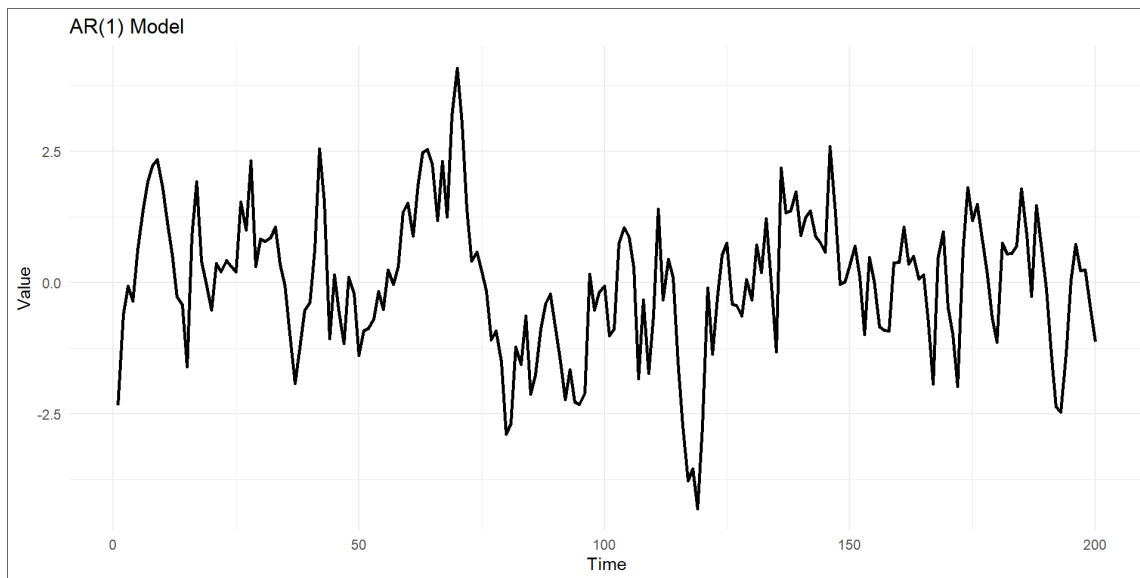
AR-1 code

AR-1 output

```
1 # Set the parameters
2 mu <- 0      # Mean of the series
3 theta <- 0.6  # Moving average coefficient
4 n <- 200     # Number of time points
5
6 # Simulate data from the MA(1) model
7 set.seed(123) # For reproducibility
8 ma1_data <- arima.sim(model = list(ma = theta), n = n, me
9
10 # Create a data frame with time series data
11 time_series_data <- data.frame(Time = 1:n, Value = ma1_da
12
13 # Create a ggplot2 line plot
14 ggplot(time_series_data, aes(x = Time, y = Value)) +
15   geom_line(linewidth=1) + labs(title = "MA(1) Model ", x :
16   theme_minimal()
```



```
1 phi <- 0.8 # Autoregressive coefficient
2 n <- 200   # Number of time points
3
4 # Simulate data from the AR(1) model
5 set.seed(123) # For reproducibility
6 ar1_data <- arima.sim(model = list(ar = phi), n = n)
7 # Create a data frame with time series data
8 time_series_data <- data.frame(Time = 1:n, Value = ar1_data)
9 # Create a ggplot2 line plot
10 ggplot(time_series_data, aes(x = Time, y = Value)) +
11   geom_line(linewidth=1) + labs(title = "AR(1) Model ", x = "Time", y = "Value") +
12   theme_minimal()
```



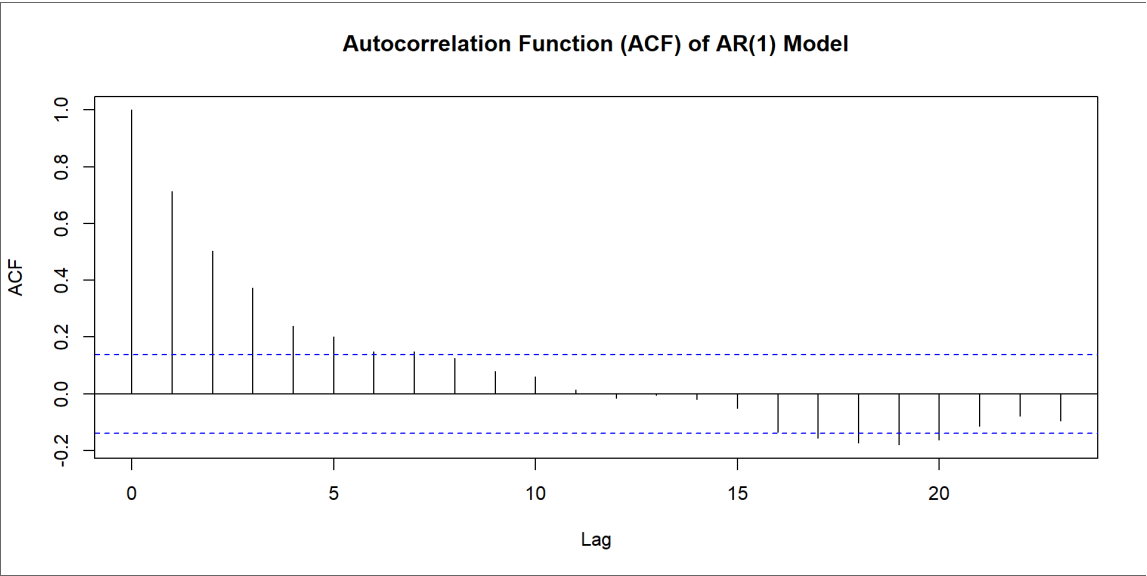
ACF and PACF of AR-1 and MA-`

ACF of AR-1

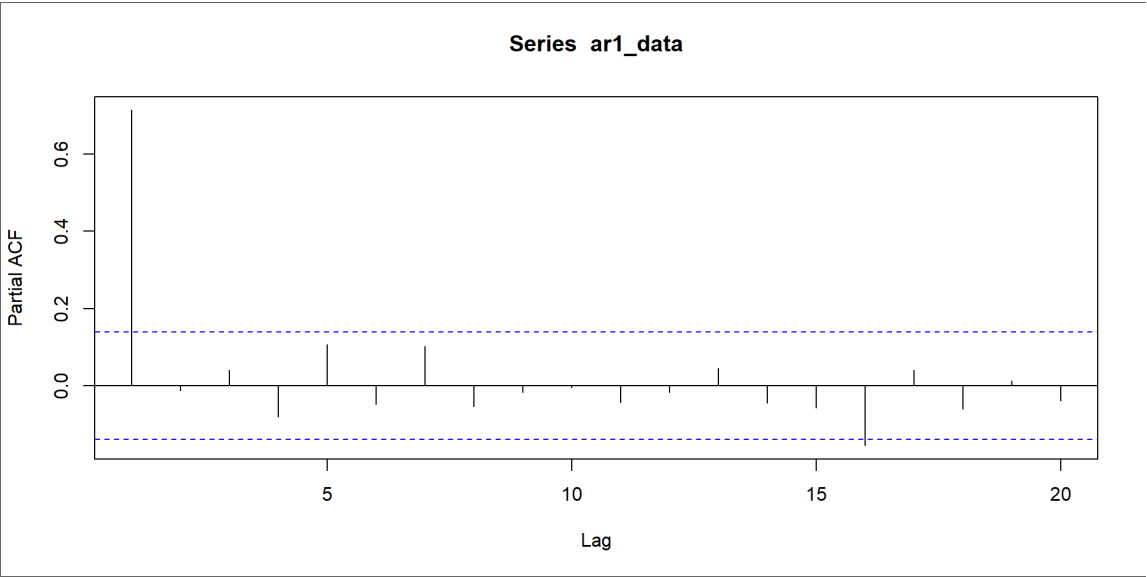
PACF of AR-1

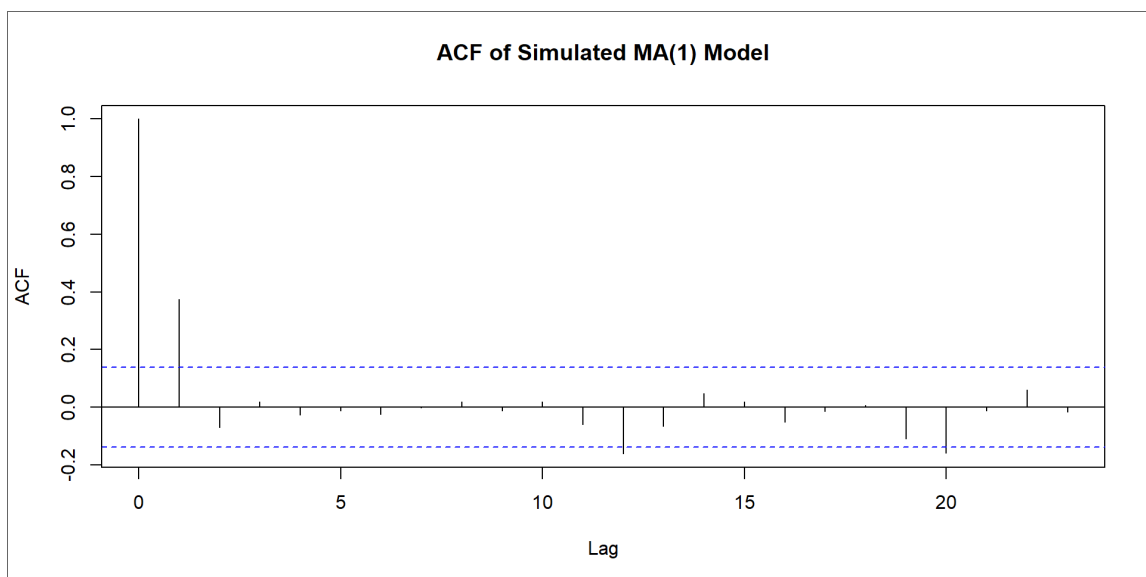
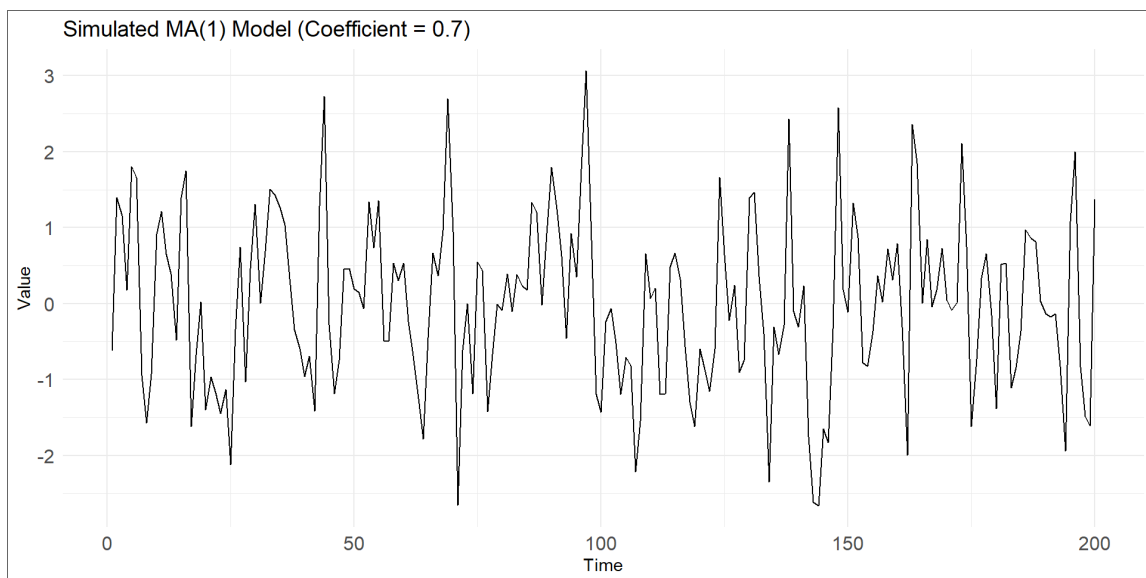
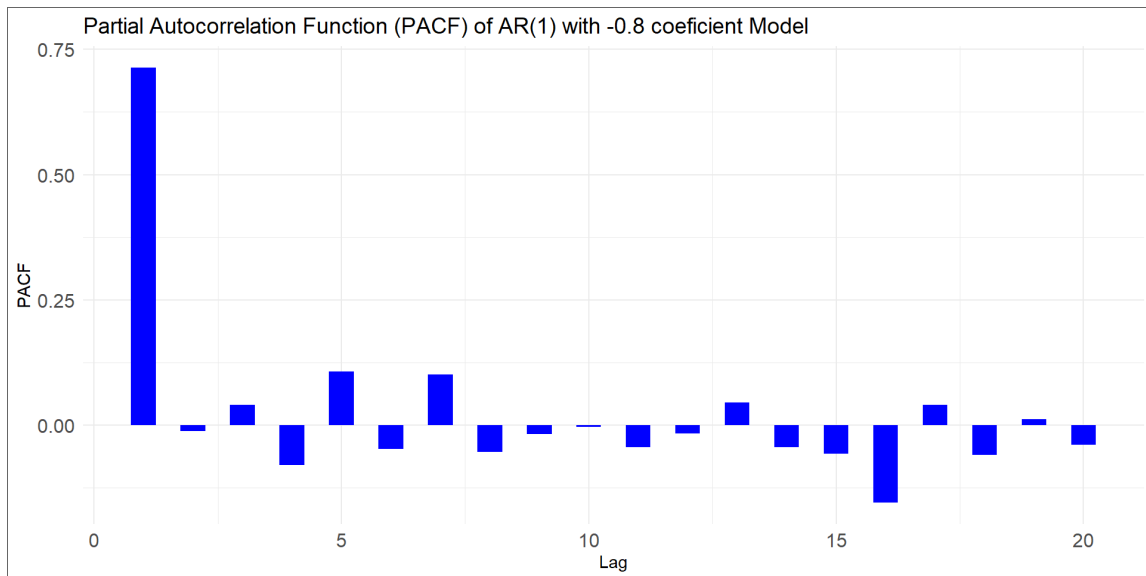
ACF-MA1

PACF-MA1



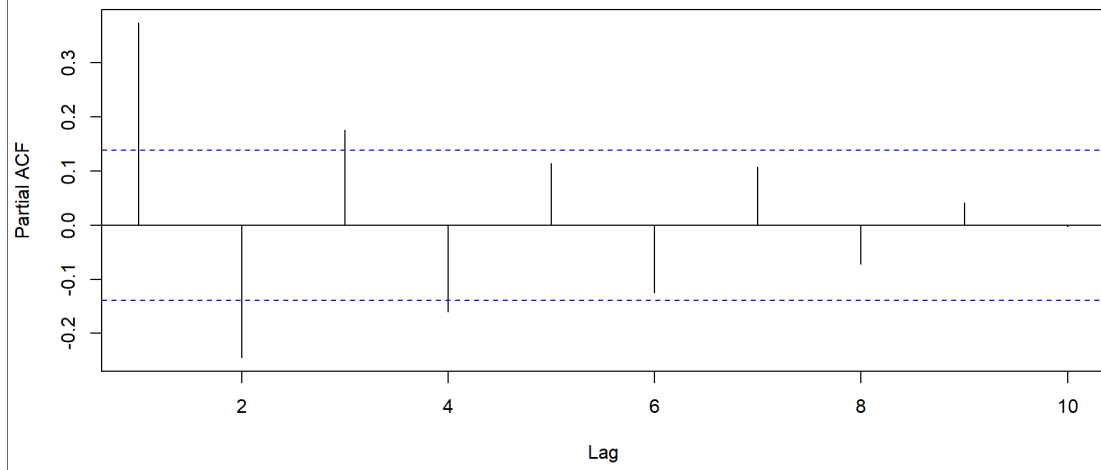
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NULL

PACF of MA(1) Model



Now let's work with real world data

Rows: 733

Columns: 8

```
$ dateid01 <date> 1954-02-01, 1954-03-01, 1954-04-01, 1954-05-01, 1954-06-01, ...
```

```
$ dateid    <dtm> 1954-03-01 00:00:00, 1954-04-01 00:00:00,
1954-05-01 00:00:0...
```

```
$ date <date> 1954-02-26, 1954-03-31, 1954-04-30, 1954-05-31, 1954-06-30, ...
```

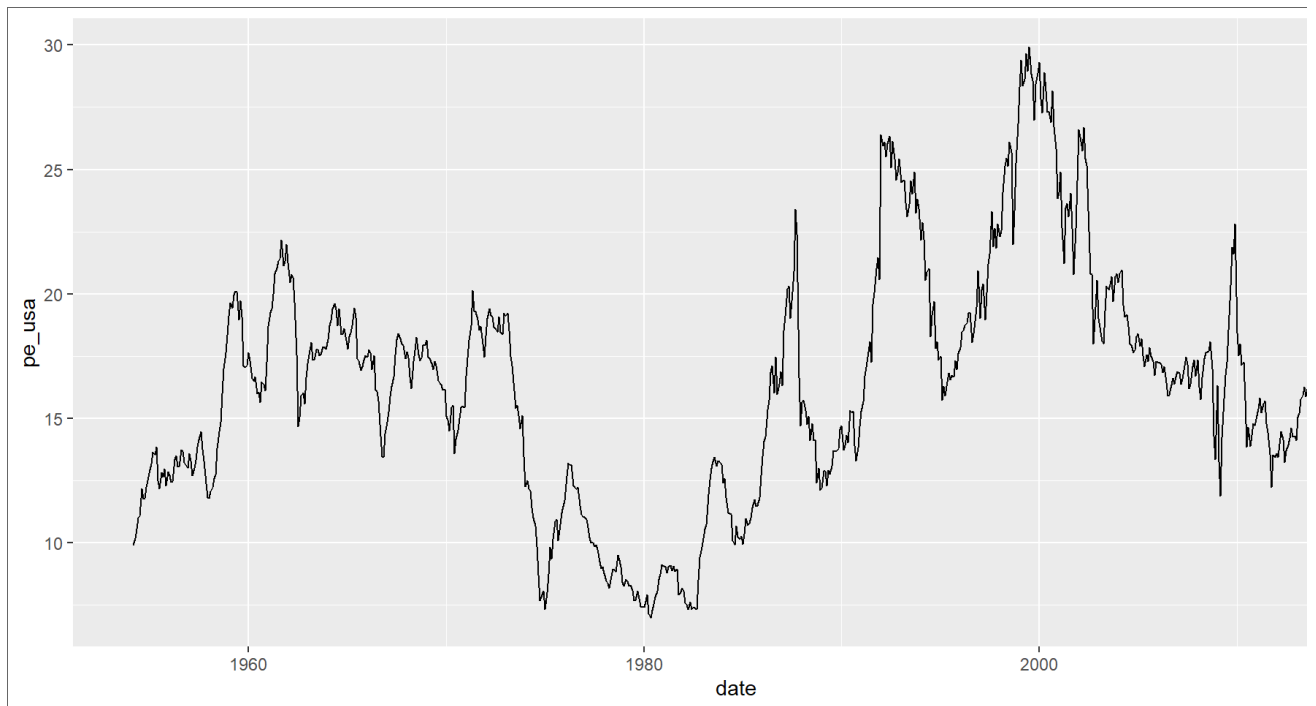
```
$ pe_aus <dbl> NA, NA, NA, NA, NA, NA, NA, NA, NA, NA,  
NA, NA, NA, NA, NA, N...
```

```
$ pe_ind      <dbl> NA, NA, NA, NA, NA, NA, NA, NA, NA, NA,
NA, NA, NA, NA, NA, N...
```

```
$ pe_ndo <dbl> NA, NA, NA, NA, NA, NA, NA, NA, NA, NA,  
NA, NA, NA, NA, NA, N...
```

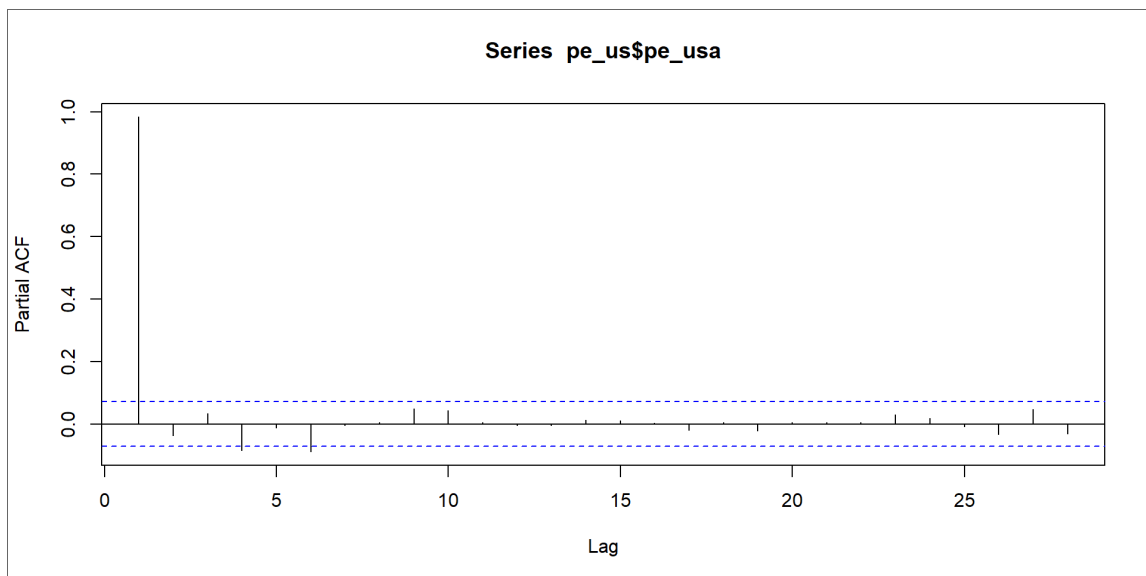
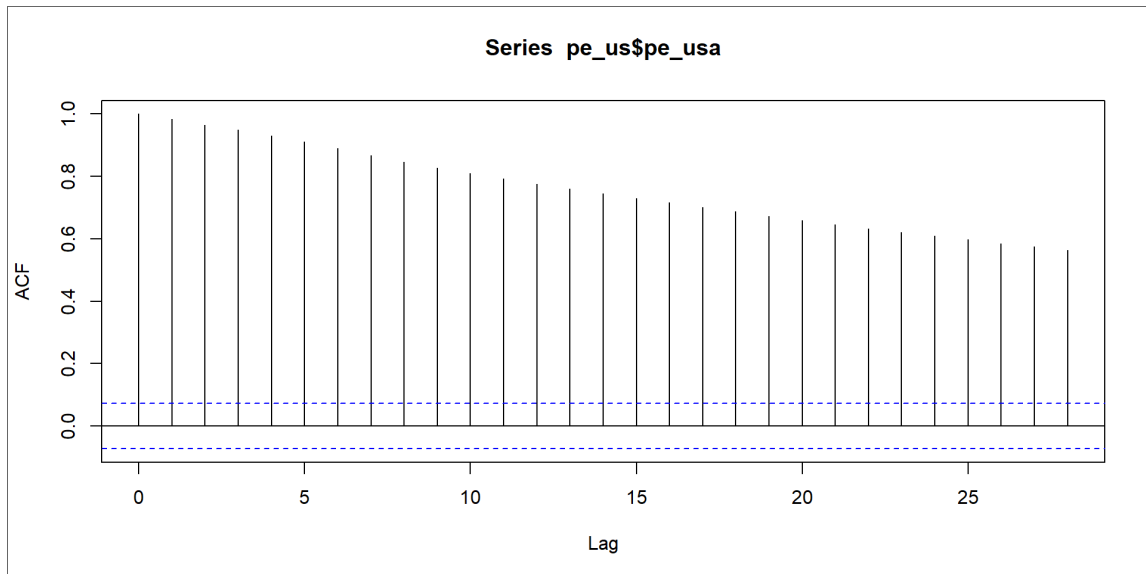
```
$ pe_saf      <dbl> NA, NA, NA, NA, NA, NA, NA, NA, NA, NA,
```

Pick pe_usa

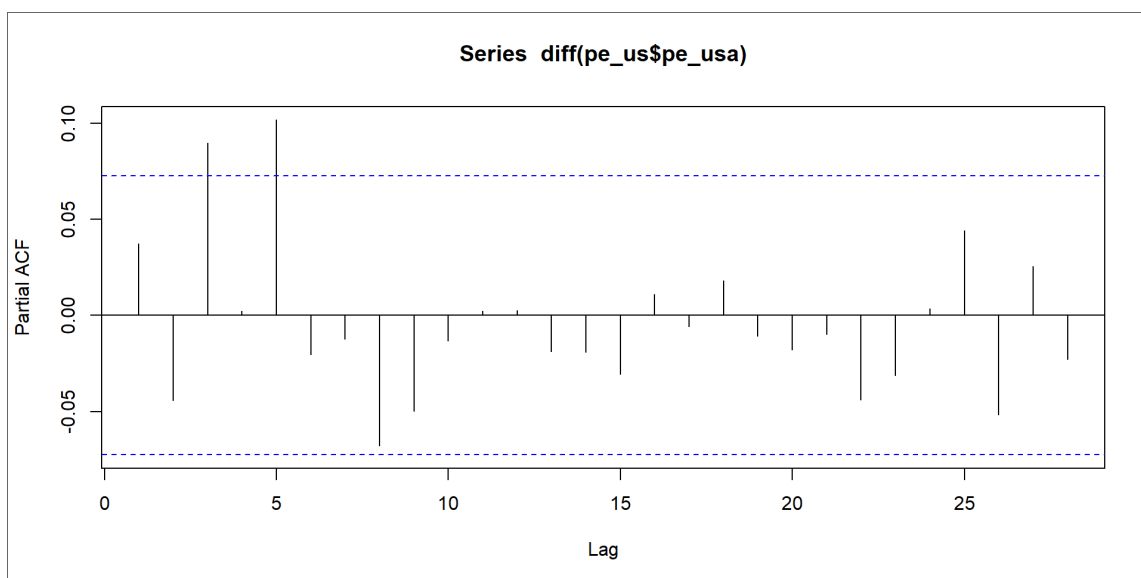
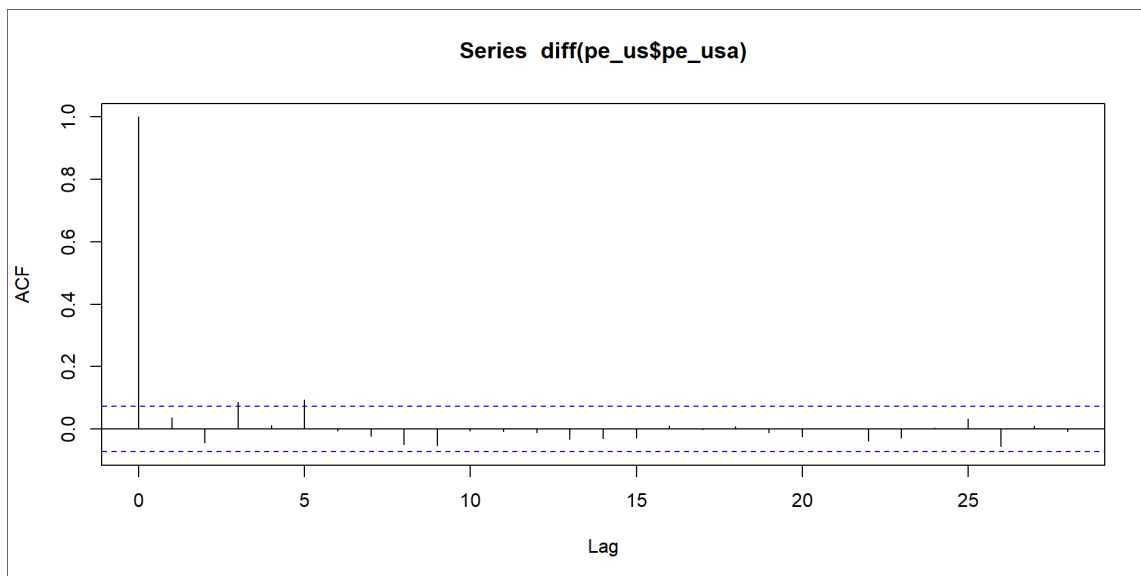


Identify model

Use SACF and SPACF to choose model(s) and also use auto.arima and see which of these two wins: your judged model or auto.arima



acf and pacf patterns indicate series is non-stationary. So here we run acf and pacf of difference of the series.



From these two graphs it seems model is ARIMA(0,1,1)

Series: pe_us\$pe_usa
ARIMA(1,1,1)

Coefficients:

	ar1	ma1
	-0.7554	0.8268
s.e.	0.1106	0.0944

sigma² = 0.7151: log likelihood = -914.95
AIC=1835.91 AICc=1835.94 BIC=1849.69

So our model was ARIMA(0,1,1) while auto.ARIMA is ARIMA(1,1,1) not bad.

Nonstationary Series

Nonstationary Series

- Introduction:
- Key Questions:
- What is nonstationarity?
- Why is it important?
- How do we determine whether a time series is nonstationary?

What is nonstationarity?

Recall from earlier part on stationarity:

- Covariance stationarity of y implies that, over time, y has:
- Constant mean
- Constant variance
- Co-variance between different observations that do not depend on that time (t), only on the “distance” or “lag” between them (j):

$$Cov(Y_t, Y_{tj}) = Cov(Y_s, Y_{s+j}) = \gamma_j$$

What is nonstationarity?

Thus, if any of these conditions does not hold, we say that y is nonstationary:

There is no long-run mean to which the series returns (economic concept of long-term **equilibrium**)

The variance is time-dependent. For example, could go to infinity as the number of observations goes to infinity

Theoretical autocorrelations do not decay, sample autocorrelations do so very slowly.

Nonstationary series can have a trend:

- Deterministic: nonrandom function of time:

$$y_t = \mu + \beta t + u_t, \text{ where } u_t \text{ is "iid"}$$

- Example $\beta = 0.45$

