线性回归模型的矩阵表示法

- 最小二乘法(最小平方法)是基于误差为高斯分布(正态)的假设的来的。
- 在用线性回归做预测时,或多或少存在误差,因此引入误差变量epsilon,将误差看作是随机变量。
- 根据大数定理,样本数增多,误差会慢慢服从正态分布。

假设epsilon服从正态分布,针对每个样本得出条件概率:

线性回归的误差模型

$$y = W^T x + b + \epsilon \qquad \epsilon \sim N(0, \sigma^2)$$

针对于样本(メル)メル)

$$p(y_i|x_i) \sim N(w^T x_i + b. 6^2) = \frac{1}{\sqrt{2\pi6}} e^{-\frac{(w^T x_i + b - y_i)^2}{26^2}}$$

高斯噪声模型根据最大似然估计法求参数w,b: (优化最小二乘)





$$D = \left\{ (\chi_{1}, \chi_{1}), \dots, (\chi_{N}, \chi_{N}) \right\}$$

$$u_{3}^{+} b^{+} = \underset{v=1}{\operatorname{argmax}} \underbrace{\prod_{i=1}^{N} p(y_{0}|\chi_{v}) = \underset{v=1}{\operatorname{argmax}} \log \left[\underbrace{\prod_{i=1}^{N} p(y_{0}|\chi_{v})}_{v_{0}} \right]$$

$$= \underset{v=1}{\operatorname{argmax}} \underbrace{\sum_{i=1}^{N} \log p(y_{0}|\chi_{v}) = \underset{v=1}{\operatorname{argmax}} \underbrace{\sum_{i=1}^{N} \log \left[\underbrace{\bigcup_{i=1}^{N} p(y_{0}|\chi_{v})}_{v_{0}} \right]}_{2\sigma^{2}}$$

$$= \underset{v=1}{\operatorname{argmax}} \underbrace{\sum_{i=1}^{N} - \log \left[\underbrace{\bigcup_{i=1}^{N} p(y_{0}|\chi_{v})}_{v_{0}} \right]^{2}}_{v_{0}^{2}}$$

$$= \underset{v=1}{\operatorname{argmax}} \underbrace{\sum_{i=1}^{N} - \left(\underbrace{\bigcup_{i=1}^{N} q(y_{0}|\chi_{v})}_{v_{0}^{2}} \right)^{2}}_{v_{0}^{2}}$$

$$= \underset{v=1}{\operatorname{argmax}} \underbrace{\sum_{i=1}^{N} - \left(\underbrace{\bigcup_{i=1}^{N} q(y_{0}|\chi_{v})}_{v_{0}^{2}} \right)^{2}}_{v_{0}^{2}}$$

$$= \underset{v=1}{\operatorname{argmax}} \underbrace{\sum_{i=1}^{N} - \left(\underbrace{\bigcup_{i=1}^{N} q(y_{0}|\chi_{v})}_{v_{0}^{2}} \right)^{2}}_{v_{0}^{2}}$$

$$\max \log p(D) <=> \min \sum_{i=1}^{N} (w^{\tau} \times_{i} + b - y_{i})^{2}$$

高斯误差模型下的最大似然

线性回归的最小二乘

● 与w,b无关的参数忽略--在求导过程中为0

假设有N维向量 $x_1, x_2, ..., x_N, x_i \in R^n$ 观测值Y: $y_1, y_2, ..., y_N, y_i \in R^n$

定义线性方程: $y_i = w^T x_i, w \in R^n$

拟合误差: $e_i = y_i - w^T x_i$

假设误差符合标准正太分布: $e_i \sim N(0,1)$ 即概率密度函数: $e_i \sim \frac{1}{\sqrt{2\pi}} e^{-\frac{e_i^2}{2}}$

似然函数

似然函数
$$L = \ln\left[\frac{1}{\sqrt{2\pi}}e^{-\frac{e_1^2}{2}} \frac{1}{\sqrt{2\pi}}e^{-\frac{e_2^2}{2}} \cdots, \frac{1}{\sqrt{2\pi}}e^{-\frac{e_N^2}{2}}\right]$$

= $-N\ln\sqrt{2\pi} - \frac{1}{2}(e_1^2 + e_2^2 + \cdots + e_N^2)$

最小化误差

最大化
$$L$$
等价于
最小化 $(e_1^2 + e_2^2 + \dots + e_N^2)$
 $J = \min(y_1 - w^T x_1)^2 + (y_2 - w^T x_2)^2 + \dots + (y_N - w^T x_N)^2$

求解权重最优解

$$J = \min(y_1 - w^T x_1)^2 + (y_2 - w^T x_2)^2 + \dots + (y_N - w^T x_N)^2$$

$$\frac{\partial J}{\partial w} = (y_1 - w x_1^T) x_1 + \dots + (y_N - w x_N^T) x_N = 0$$

$$w(\sum_{i=1}^N x_i^T x_i) = \sum_{i=1}^N x_i y_i$$

$$w = (\sum_{i=1}^N x_i y_i) (\sum_{i=1}^N x_i^T x_i)^{-1}$$
对比之前 $a = (x^T x)^{-1} x^T Y$

最终得到的w和最小二乘估计得到的结果是一致的。

逻辑回归-分类问题:

0-1问题:是否违约,情感分析(开心与否),广告点击率,疾病分析逻辑回归本身解决的是二分类问题:p(y=1|x)和p(y=0|x)建立逻辑回归模型要考虑定义域与值域的匹配问题,y与x的对应关系中,p(y|x)的取值只可能落在(0,1)里

逻辑函数:

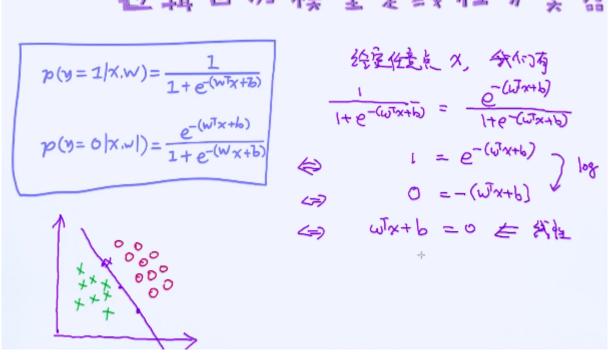
采用线性回归建模的思想来处理逻辑回归的映射关系:

对于特征向量×和二分类标签y,我们可以定义如下的条件概率:

$$p(y = 1 | x, w, b) = \frac{1}{1 + e^{-(w^{T}x + b)}}$$
$$p(y = 0 | x, w, b) = \frac{e^{-(w^{T}x + b)}}{1 + e^{-(w^{T}x + b)}}$$

证明逻辑回归是线性分类器:





采用条件独立(conditional independence)的假设,在这个假设的前提下,可 以把条件概率 $p(x_1,...,x_n|w)$ 分解成 $p(x_1|w)...p(x_n|w)$,构造极大似然函数。

目标函数(Objective Function)

假设我们拥有数据集: $D = \{(x_i, y_i)\}_{i=1}^n x_i \in \mathbb{R}^d, y_i \in \{0, 1\}$

而且我们已经定义了:

$$p(y|x,w,b) = p(y = 1|x,w,b)^{y}[1-p(y = 1|x,w,b)]^{1-y}$$

我们需要最大化的目标函数为(也叫做最大似然函数);

$$\widehat{w}_{MLE}$$
, $\widehat{b}_{MLE} = argmax_{w,b} \prod_{i=1}^{n} p(y_i | x_i, w, b)$

将两个条件概率公式合并: 当y=1时,后面那一项不起任何作用也就变成了第一个条件概率; 当y=0时,前面那一项不起作用,也就等同于第二个条件概率。

目标逐数

给定条件概率: $p(y|x,w,b) = p(y = 1|x,w,b)^y [1-p(y = 1|x,w,b)]^{1-y}$

$$argmin_{w,b} = \sum_{i=1}^{n} \log p(y_i|x_i,w,b)$$

$$|\omega_{i}(\alpha^{x},b^{y})| = |\omega_{j}(\alpha^{y}) + |\omega_{j}(b^{y})|$$

$$= x - |\omega_{j}(\alpha^{y})| + |\omega_{j}(a^{y})| + |\omega_{j}(a^{y$$

$$= \underset{\omega = b}{\operatorname{argmin}} - \sum_{i=1}^{n} \log \left[p(y_{i}=|X_{i}, \omega_{b})^{y_{i}} \cdot \left[1 - p(y_{i}=|X_{i}, \omega_{b}) \right]^{1-y_{i}} \right]$$

$$= \underset{\omega = b}{\operatorname{argmin}} - \sum_{i=1}^{n} y_{i} \log p(y_{i}=|X_{i}, \omega_{b}) + (1-y_{i}) \log \left[1 - p(y_{i}=|X_{i}, \omega_{b}) \right]$$

求解目标函数的最值:

- 导数为0:对逻辑回归不适用,导数表达式难以写出
- 迭代方法如梯度下降:不断更新参数,与lr参数相关。

梯度下降法:

方法2:梯度下降法

求使得f(w)值最小的参数w

初始化:
$$W^{1}$$

$$fort = 1.2...$$

$$W^{t+1} = W^{t} - \eta_{\nabla} f(W^{t})^{t}$$

例子: 求解函数 f(w) = 4w2 + 5w + 1 的最优解

$$w'=0$$
, $b=a1$, $f'(w)=bw+5$
 $w^2=w'-a1\cdot(8\cdot0+5)=0-a1\cdot5=-a5$
 $w^3=w^2-a-1(8\cdot(a5)+5)=-a5-a-1\cdot1=-a-6$
 $w^4=w^3-a-1(8\cdot(a6)+5)=-a6-a-1\cdota2=-a62$
 $w^5=w^4-a-1(8\cdot(a62)+5)=-a62-a-1\cdota=-a625$

对于逻辑函数的事导

$$G(X) = \frac{1}{1 + e^{-X}}$$

$$G'(X) = \frac{(y \cdot (1 + e^{X}) - 1 \cdot (1 + e^{X})'}{(1 + e^{-X})^{2}} = \frac{e^{-X}}{(1 + e^{X})^{2}}$$

$$= \frac{1}{1 + e^{X}} \cdot \frac{e^{X}}{1 + e^{X}} = \frac{1}{1 + e^{X}} \cdot \left[1 - \frac{1}{1 + e^{X}}\right]$$

$$= \sigma(X) \cdot \left[1 - \sigma(X)\right]$$

求解w参数:

逻辑回归的梯度下降法一求解》

$$p(y = 1 | X.W.b) = \frac{1}{1 + e^{-(w^{T}x+b)}} = \sigma(\overline{w}x+b)$$

$$argmin_{w,b} - \sum_{i=1}^{n} y_{i} log p(y_{i} = 1 | X_{i} w.b) + (1 - y_{i}) log(1 - p(y_{i} = 1 | X_{i} w.b)]$$

$$\lambda = arg_{w,b} - \sum_{i=1}^{n} y_{i} log \sigma(\overline{w}x+b) + (1 - y_{i}) log(1 - p(y_{i} = 1 | X_{i} w.b))$$

$$\lambda = arg_{w,b} - \sum_{i=1}^{n} y_{i} log \sigma(\overline{w}x+b) + (1 - y_{i}) log(1 - p(y_{i} = 1 | X_{i} w.b))$$

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求解b参数:

逻辑回归的梯度下降法一承解的

$$p(y = 1|X, W, b) = \frac{1}{1 + e^{-(W^{T}X + b)}}$$

$$argmin_{W,b} - \sum_{i=1}^{n} y_{i}log p(y_{i} = 1|X_{i}W, b) + (1 - y_{i})log(1 - p(y_{i} = 1|X_{i}W, b)]$$

$$= argmin_{W,b} - \sum_{i=1}^{n} y_{i}log \sigma(UX_{i} + b) + (1 - y_{i})log(1 - p(y_{i} = 1|X_{i}W, b)]$$

$$= argmin_{W,b} - \sum_{i=1}^{n} y_{i}log \sigma(UX_{i} + b) + (1 - y_{i})log(1 - p(y_{i} = 1|X_{i}W, b))$$

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$$= argmin_{W,b} - \sum_{i=1}^{n} y_{i}log \sigma(UX_{i} + b) + (1 - y_{i})log(1 - p(y_{i} = 1|X_{i}W, b)$$

$$= argmin_{W$$

迭代更新:

$$w' = w - Ir * f'(w)$$

$$b' = b - lr * f'(b)$$

f'(w,b)为损失函数

逻辑回归的梯度下降法

初始化: W1, b1

$$fort = 1, 2...$$

$$w^{t+1} = w^{t} - \eta_{t} \sum_{i=1}^{n} (\sigma(w^{T} \chi_{i} + b) - y_{i}) \chi_{i}$$

$$b^{t+1} = b^{t} - \eta_{t} \sum_{i=1}^{n} (\sigma(w^{T} \chi_{i} + b) - y_{i})$$

判断收敛:

- 相邻时间段计算当前的损失函数,损失函数变化很小或者不变即收敛
- 相邻时间段计算当前参数的值w,b,参数值变化很小或者不变即收敛