

## 1 General Notations

- $N$ : Number of users
- $n$ : Number of items,  $n = \begin{cases} 2m - 1, & \text{if } n \text{ is odd} \\ 2m, & \text{otherwise} \end{cases}$
- $\mathbf{R}^1, \dots, \mathbf{R}^N$ : full rankings given by the users
- $\mathbf{R}^j = \{R_1^j, \dots, R_n^j\} \sim \text{Mallows}(\boldsymbol{\rho}^0, \alpha^0)$
- $P(\boldsymbol{\rho}|\mathbf{R}^1, \dots, \mathbf{R}^N, \alpha^0)$ : Mallows posterior
- $\{i_1, \dots, i_n\}$ : a ranking of  $n$  items that determines the sequence following which the items are to be sampled. i.e.  $i_j = k$  indicates that item  $j$  is the  $k$ -th item is to be sampled
- $\{o_1, \dots, o_n\}$ : an ordering of  $n$  items that corresponds to  $\{i_1, \dots, i_n\}$  s.t.  $i_{o_k} = k$ .  $\{o_1, \dots, o_n\}$  and  $\{i_1, \dots, i_n\}$  have a one-to-one relationship
- $\sum_{\{i_1, \dots, i_n\} \in \mathcal{P}_n} q(\tilde{\boldsymbol{\rho}}|i_1, \dots, i_n, \alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) \cdot g(i_1, \dots, i_n | \dots)$  : pseudolikelihood that approximates the Mallows posterior
- $q(\tilde{\boldsymbol{\rho}}|i_1, \dots, i_n, \alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N)$   
 $= q(\tilde{\rho}_{o_1}|\alpha^0, o_1, R_{o_1}^1, \dots, R_{o_1}^N) \cdot q(\tilde{\rho}_{o_2}|\alpha^0, o_2, \tilde{\rho}_{o_1}, R_{o_2}^1, \dots, R_{o_2}^N) \cdot \dots \cdot$   
 $q(\tilde{\rho}_{o_n}|\alpha^0, o_n, \tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{n-1}}, R_n^1, \dots, R_n^N)$   
 $- q(\tilde{\rho}_{o_1}|\alpha^0, o_1, R_{o_1}^1, \dots, R_{o_1}^N) = \frac{\exp\{-\frac{\alpha_0}{n} \sum_{j=1}^N d(R_{o_1}^j, \tilde{\rho}_{o_1})\} \mathbb{1}_{\tilde{\rho}_{o_1} \in \{1, \dots, n\}}}{\sum_{\tilde{r}_{o_1} \in \{1, \dots, n\}} \exp\{-\frac{\alpha_0}{n} \sum_{j=1}^N d(R_{o_1}^j, \tilde{r}_{o_1})\}}$   
 $- q(\tilde{\rho}_{o_k}|\alpha^0, o_k, \tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{k-1}}, R_{o_k}^1, \dots, R_{o_k}^N) = \frac{\exp\{-\frac{\alpha_0}{n} \sum_{j=1}^N d(R_{o_k}^j, \tilde{\rho}_{o_k})\} \mathbb{1}_{\tilde{\rho}_{o_k} \in \{1, \dots, n\} \setminus \{\tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{k-1}}\}}}{\sum_{\tilde{r}_{o_k} \in \{1, \dots, n\} \setminus \{\tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{k-1}}\}} \exp\{-\frac{\alpha_0}{n} \sum_{j=1}^N d(R_{o_k}^j, \tilde{r}_{o_k})\}}$   
for  $k = 2, \dots, n$
- $\boldsymbol{\rho}^0 \leftrightarrow \mathbf{o}^0$  s.t.  $\rho^{0^{-1}}(m) = o_m^0$

- Define the “v-function”  $f_v(\cdot)$  such that  $f_v(\boldsymbol{\rho}^0) = \mathcal{V}_{\boldsymbol{\rho}^0}$ , where

$$- \mathcal{V}_{\boldsymbol{\rho}^0} = \begin{cases} \{\mathbf{r} \in \mathcal{P}_n : r_{o_m^0} = 1, r_{o_{m \pm k}^0} \in \{2k, 2k+1\}, k = 1, \dots, m-1\}, & \text{if } n \text{ is odd} \\ \{\mathbf{r} \in \mathcal{P}_n : \{r_{o_{m-k}^0}, r_{o_{m+k+1}^0}\} \in \{2k+1, 2k+2\}, k = 0, \dots, m\}, & \text{if } n \text{ is even} \end{cases}$$

## 2 Theorems and Lemmas

### 2.1

**Lemma 2.1.1** *Given there are odd number of items, i.e.  $n = 2m - 1$ .  $\forall \alpha^0 \in (0, \infty)$ ,*

1.  $\mathbb{E}(R_{o_m^0} | \boldsymbol{\rho}_0, \alpha^0) = \rho_m^0 = m$
2.  $\forall j \in [1, m-2], j < \mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_{j+1}^0} | \boldsymbol{\rho}^0, \alpha^0] < m$
3.  $\forall j \in [m+2, 2m-1], m < \mathbb{E}[R_{o_{j-1}^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < j$

*Similarly, if  $n$  is even, i.e.  $n = 2m$ ,  $\forall \alpha^0 \in (0, \infty)$ ,*

1.  $\forall j \in [1, m-1], j < \mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_{j+1}^0} | \boldsymbol{\rho}^0, \alpha^0]$
2.  $\forall j \in [m+2, 2m], \mathbb{E}[R_{o_{j-1}^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < j$

*Note that for both cases, it satisfies that  $\forall 1 \leq j < k \leq n$  and  $\forall \alpha > 0$ ,  $\mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_k^0} | \boldsymbol{\rho}^0, \alpha^0]$*

**Lemma 2.1.2** *As  $N \rightarrow \infty$ ,  $\frac{1}{N} \sum_{j=1}^N R_i^j \rightarrow \mathbb{E}[R_i | \boldsymbol{\rho}^0, \alpha^0]$ ,  $\forall i = 1, \dots, n$*

**Definition 1** *Given a vector of length  $n$ , i.e.  $\{x_1, \dots, x_n\}$ , the rank function  $\text{rank}(x_1, \dots, x_n)$  is defined as  $\text{rank}(x_1, \dots, x_n) = \{r_1, \dots, r_n\}$  such that  $x_{(r_k)} = x_k$ ,  $\forall k = 1, \dots, n$*

**Theorem 2.1.3** *As  $N \rightarrow \infty$ , and  $\forall \alpha > 0$ ,*

$$\text{rank}(\frac{1}{N} \sum_{j=0}^N R_1^j, \dots, \frac{1}{N} \sum_{j=0}^N R_n^j) \rightarrow \text{rank}(\mathbb{E}[R_1 | \boldsymbol{\rho}^0, \alpha_0], \dots, \mathbb{E}[R_n | \boldsymbol{\rho}^0, \alpha_0]) = \boldsymbol{\rho}^0$$

That rephrase it, as  $N$  approaches infinity, the Mallows consensus parameter  $\boldsymbol{\rho}^0$  can be inferred from the data by taking the marginal mean for each item and then apply the rank function to these marginal means.

### 2.2

**Theorem 2.2.1** *For a function  $g$  defined on the space of permutation  $\mathcal{P}_n$  which can depend on  $\boldsymbol{\rho}^0$ ,*

$$\begin{aligned} & \arg \min_{g \in \mathcal{D}_{\rho^0}} \lim_{N \rightarrow \infty} KL(P(\rho|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) || \sum_{\{i_1, \dots, i_n\} \in \mathcal{P}_n} q(\tilde{\rho}|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N, i_1, \dots, i_n) g(i_1, \dots, i_n|\rho^0)) \\ & = g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0}), \\ & \text{where} \end{aligned}$$

- $\mathcal{D}_{\rho^0}$  is the set of all distributions on the space of permutation  $\mathcal{P}_n$ , which can depend on  $\rho^0$
- $g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0})$  is a distribution whose density is concentrated on  $\rho^0$ , defined as
$$\begin{cases} g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0}) = |\mathcal{V}_{\rho^0}|^{-1} > 0, & \text{if } \{i_1, \dots, i_n\} \in \mathcal{V}_{\rho^0} \\ g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0}) = 0, & \text{if } \{i_1, \dots, i_n\} \notin \mathcal{V}_{\rho^0} \end{cases}, \text{ where } |\mathcal{V}_{\rho^0}| = \begin{cases} 2^{m-1}, & \text{if } n \text{ is odd} \\ 2^m, & \text{otherwise} \end{cases}$$

That is to say, for a set of distributions  $g$ , which are defined on the space of permutation of  $n$  items  $\mathcal{P}_n$ , the distribution  $g^*$  that minimizes the KL-divergence between the Mallows posterior and the pseudolikelihood defined above, is a uniform distribution with its density concentrated on  $\mathcal{V}_{\rho^0}$

### 2.3

For a given  $N < \infty$ , define  $\hat{\rho}^0$  as  $\text{rank}(\frac{1}{N} \sum_{j=0}^N R_1^j, \dots, \frac{1}{N} \sum_{j=0}^N R_n^j)$  and  $\mathcal{V}_{\hat{\rho}^0} = f_v(\hat{\rho}^0)$ .

**Theorem 2.3.1**  $\exists \sigma \geq 0$  and  $g'(i_1, \dots, i_n|\mathcal{V}_{\hat{\rho}^0}, \sigma)$  such that

$$\begin{aligned} & KL(P(\rho|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) || \sum_{\{i_1, \dots, i_n\} \in \mathcal{P}_n} q(\tilde{\rho}|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N, i_1, \dots, i_n) g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0})) \geq \\ & KL(P(\rho|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) || \sum_{\{i_1, \dots, i_n\} \in \mathcal{P}_n} q(\tilde{\rho}|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N, i_1, \dots, i_n) g'(i_1, \dots, i_n|\mathcal{V}_{\hat{\rho}^0}, \sigma)) \end{aligned}$$

where

- $\hat{\mathbf{v}} \sim g^*(\hat{\mathbf{v}}|\mathcal{V}_{\hat{\rho}^0})$
- $x_i \sim \mathcal{N}(x_i|\hat{v}_i, \sigma)$  for  $i = 1, \dots, n$
- $i_1, \dots, i_n \sim \mathcal{F}_r(i_1, \dots, i_n|x_1, \dots, x_n)$ , where  $\mathcal{F}_r = \begin{cases} 1, & \text{if } i_1, \dots, i_n = \text{rank}(x_1, \dots, x_n) \\ 0, & \text{otherwise} \end{cases}$

$$\text{and } g'(i_1, \dots, i_n|\mathcal{V}_{\hat{\rho}^0}, \sigma) = \sum_{\hat{\mathbf{v}} \in \mathcal{V}_{\hat{\rho}^0}} \{g^*(\hat{\mathbf{v}}|\mathcal{V}_{\hat{\rho}^0}) \int_{\mathbf{x}} \mathcal{F}_r(i_1, \dots, i_n|x_1, \dots, x_n) \prod_{i=1}^n \mathcal{N}(x_i|\hat{v}_i, \sigma) d\mathbf{x}\}$$

As  $N$  is limited,  $\rho^0$ , and therefore,  $\mathcal{V}_{\rho^0}$  usually cannot be accurately inferred from the data. We can however, sample for  $i_1, \dots, i_n$  by sampling for each item  $i$  from a univariate Gaussian distribution centered on  $\hat{v}_i$  with a fixed variance  $\sigma$  for all items, and then obtain a ranking by using the rank function. By doing so, a smaller KL divergence from the Mallows posterior can be obtained, compared to not introducing the variations.

## 2.4

**Theorem 2.4.1** *With the usage of  $g'(i_1, \dots, i_n | \mathcal{V}_{\hat{\rho}^0}, \sigma)$ , the value of  $\sigma$  that minimizes the KL-divergence between the Mallows posterior and the resulted pseudolikelihood is*

$$\sigma = \begin{cases} 0, & \text{if } \delta(\alpha, n, N) \leq \delta^* \\ f(\alpha, n, N), & \text{otherwise} \end{cases}$$

In other words,  $\sigma$  should be 0 when  $\delta(\alpha, n, N) \geq \delta^*$ . Beyond this point, the optimal choice of  $\sigma$  should be greater than 0, and it follows a function  $f(\alpha, n, N)$ .

## 2.5

**Theorem 2.5.1** *As  $N \rightarrow \infty, \sigma = 0 \forall \alpha > 0, n \geq 1$*