

1 General Notations

- N : Number of users
- n : Number of items, $n = \begin{cases} 2m - 1, & \text{if } n \text{ is odd} \\ 2m, & \text{otherwise} \end{cases}$
- \mathcal{P}_n : the space of permutation of n items
- $\mathbf{R}^1, \dots, \mathbf{R}^N$: full rankings given by the users
- $\mathbf{R}^j \in \mathcal{P}_n = \{R_1^j, \dots, R_n^j\} \sim \text{Mallows}(\boldsymbol{\rho}^0, \alpha^0)$
- $P(\boldsymbol{\rho} | \mathbf{R}^1, \dots, \mathbf{R}^N, \alpha^0)$: Mallows posterior
- $\{i_1, \dots, i_n\}$: a ranking of n items that determines the sequence following which the items are to be sampled. i.e. $i_j = k$ indicates that item j is the k -th item is to be sampled
- $\{o_1, \dots, o_n\}$: an ordering of n items that corresponds to $\{i_1, \dots, i_n\}$ s.t. $i_{o_k} = k$. $\{o_1, \dots, o_n\}$ and $\{i_1, \dots, i_n\}$ have a one-to-one relationship
- $\sum_{\{i_1, \dots, i_n\} \in \mathcal{P}_n} q(\tilde{\boldsymbol{\rho}} | i_1, \dots, i_n, \alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) \cdot g(i_1, \dots, i_n | \dots)$: pseudolikelihood that approximates the Mallows posterior
- $q(\tilde{\boldsymbol{\rho}} | i_1, \dots, i_n, \alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) = q(\tilde{\boldsymbol{\rho}} | o_1, \dots, o_n, \alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N)$
 $= q(\tilde{\rho}_{o_1} | \alpha^0, o_1, R_{o_1}^1, \dots, R_{o_1}^N) \cdot q(\tilde{\rho}_{o_2} | \alpha^0, o_2, \tilde{\rho}_{o_1}, R_{o_2}^1, \dots, R_{o_2}^N) \cdot \dots \cdot$
 $q(\tilde{\rho}_{o_{n-1}} | \alpha^0, o_{n-1}, \tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{n-2}}, R_{o_{n-1}}^1, \dots, R_{o_{n-1}}^N) \cdot q(\tilde{\rho}_{o_n} | \alpha^0, o_n, \tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{n-1}}, R_n^1, \dots, R_n^N)$
 $= q(\tilde{\rho}_{o_1} | \alpha^0, o_1, R_{o_1}^1, \dots, R_{o_1}^N) = \frac{\exp\{-\frac{\alpha_0}{n} \sum_{j=1}^N d(R_{o_1}^j, \tilde{\rho}_{o_1})\} \mathbb{1}_{\tilde{\rho}_{o_1} \in \{1, \dots, n\}}}{\sum_{\tilde{r}_{o_1} \in \{1, \dots, n\}} \exp\{-\frac{\alpha_0}{n} \sum_{j=1}^N d(R_{o_1}^j, \tilde{r}_{o_1})\}}$
 $= q(\tilde{\rho}_{o_k} | \alpha^0, o_k, \tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{k-1}}, R_{o_k}^1, \dots, R_{o_k}^N) = \frac{\exp\{-\frac{\alpha_0}{n} \sum_{j=1}^N d(R_{o_k}^j, \tilde{\rho}_{o_k})\} \mathbb{1}_{\tilde{\rho}_{o_k} \in \{1, \dots, n\} \setminus \{\tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{k-1}}\}}}{\sum_{\tilde{r}_{o_k} \in \{1, \dots, n\} \setminus \{\tilde{\rho}_{o_1}, \dots, \tilde{\rho}_{o_{k-1}}\}} \exp\{-\frac{\alpha_0}{n} \sum_{j=1}^N d(R_{o_k}^j, \tilde{r}_{o_k})\}}$
for $k = 2, \dots, n$

- \mathbf{o}^0 : a set of ordering that corresponds to $\boldsymbol{\rho}^0$ s.t. $\rho^{0^{-1}}(m) = o_m^0$
- Define the “v-function” $f_v(\cdot)$ such that $f_v(\boldsymbol{\rho}^0) = \mathcal{V}_{\boldsymbol{\rho}^0}$, where

$$- \mathcal{V}_{\boldsymbol{\rho}^0} = \begin{cases} \{\mathbf{r} \in \mathcal{P}_n : r_{o_m^0} = 1, r_{o_{m \pm k}^0} \in \{2k, 2k+1\}, k = 1, \dots, m-1\}, & \text{if } n \text{ is odd} \\ \{\mathbf{r} \in \mathcal{P}_n : \{r_{o_{m-k}^0}, r_{o_{m+k+1}^0}\} \in \{2k+1, 2k+2\}, k = 0, \dots, m\}, & \text{if } n \text{ is even} \end{cases}$$

2 Theorems and Lemmas

2.1

Lemma 2.1.1 *Given there are odd number of items, i.e. $n = 2m - 1$. $\forall \alpha^0 \in (0, \infty)$,*

1. $\mathbb{E}(R_{o_m^0} | \boldsymbol{\rho}_0, \alpha^0) = \rho_{o_m^0}^0 = m$
2. $\forall j \in [1, m-2], j < \mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_{j+1}^0} | \boldsymbol{\rho}^0, \alpha^0] < m$
3. $\forall j \in [m+2, 2m-1], m < \mathbb{E}[R_{o_{j-1}^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < j$

Similarly, if n is even, i.e. $n = 2m$, $\forall \alpha^0 \in (0, \infty)$,

1. $\forall j \in [1, m-1], j < \mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_{j+1}^0} | \boldsymbol{\rho}^0, \alpha^0]$
2. $\forall j \in [m+2, 2m], \mathbb{E}[R_{o_{j-1}^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < j$

Note that for both cases, it satisfies that $\forall 1 \leq j < k \leq n$ and $\forall \alpha > 0$, $\mathbb{E}[R_{o_j^0} | \boldsymbol{\rho}^0, \alpha^0] < \mathbb{E}[R_{o_k^0} | \boldsymbol{\rho}^0, \alpha^0]$

Lemma 2.1.2 *As $N \rightarrow \infty$, $\frac{1}{N} \sum_{j=1}^N R_i^j \rightarrow \mathbb{E}[R_i | \boldsymbol{\rho}^0, \alpha^0]$, $\forall i = 1, \dots, n$*

Definition 1 *Given a vector of length n , i.e. $\{x_1, \dots, x_n\}$, the function $\text{rank}(x_1, \dots, x_n)$ is defined as $\text{rank}(x_1, \dots, x_n) = \{r_1, \dots, r_n\}$ such that $x_{(r_k)} = x_k$, $\forall k = 1, \dots, n$*

Theorem 2.1.3 *As $N \rightarrow \infty$, and $\forall \alpha > 0$,*

$$\text{rank}(\frac{1}{N} \sum_{j=0}^N R_1^j, \dots, \frac{1}{N} \sum_{j=0}^N R_n^j) \rightarrow \text{rank}(\mathbb{E}[R_1 | \boldsymbol{\rho}^0, \alpha_0], \dots, \mathbb{E}[R_n | \boldsymbol{\rho}^0, \alpha_0]) = \boldsymbol{\rho}^0$$

To rephrase, as N approaches infinity, the Mallows consensus parameter $\boldsymbol{\rho}^0$ can be inferred from the data by taking the marginal mean for each item and then apply the rank function to these marginal means.

2.2

Theorem 2.2.1 *For a function g defined on \mathcal{P}_n which can depend on $\boldsymbol{\rho}^0$,*

$$\begin{aligned} & \arg \min_{g \in \mathcal{D}_{\rho^0}} \lim_{N \rightarrow \infty} KL(P(\rho|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) || \sum_{\{i_1, \dots, i_n\} \in \mathcal{P}_n} q(\tilde{\rho}|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N, i_1, \dots, i_n) g(i_1, \dots, i_n|\rho^0)) \\ & = g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0}), \\ & \text{where} \end{aligned}$$

- \mathcal{D}_{ρ^0} is the set of all distributions on \mathcal{P}_n , which can depend on ρ^0
- $g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0})$ is a distribution whose density is concentrated on ρ^0 , defined as
$$\begin{cases} g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0}) = |\mathcal{V}_{\rho^0}|^{-1} > 0, & \text{if } \{i_1, \dots, i_n\} \in \mathcal{V}_{\rho^0} \\ g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0}) = 0, & \text{if } \{i_1, \dots, i_n\} \notin \mathcal{V}_{\rho^0} \end{cases}, \text{ where } |\mathcal{V}_{\rho^0}| = \begin{cases} 2^{m-1}, & \text{if } n \text{ is odd} \\ 2^m, & \text{otherwise} \end{cases}$$

That is to say, for a set of distributions g , which are defined on the space of permutation of n items \mathcal{P}_n , the distribution g^* that minimizes the KL-divergence between the Mallows posterior and the pseudolikelihood defined above, is a uniform distribution with its density concentrated on \mathcal{V}_{ρ^0}

2.3

For a given $N < \infty$, define $\hat{\rho}^0$ as $\text{rank}(\frac{1}{N} \sum_{j=0}^N R_1^j, \dots, \frac{1}{N} \sum_{j=0}^N R_n^j)$ and $\mathcal{V}_{\hat{\rho}^0} = f_v(\hat{\rho}^0)$.

Theorem 2.3.1 $\exists \sigma \geq 0$ and $g'(i_1, \dots, i_n|\mathcal{V}_{\hat{\rho}^0}, \sigma)$ such that

$$\begin{aligned} & KL(P(\rho|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) || \sum_{\{i_1, \dots, i_n\} \in \mathcal{P}_n} q(\tilde{\rho}|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N, i_1, \dots, i_n) g^*(i_1, \dots, i_n|\mathcal{V}_{\rho^0})) \geq \\ & KL(P(\rho|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N) || \sum_{\{i_1, \dots, i_n\} \in \mathcal{P}_n} q(\tilde{\rho}|\alpha^0, \mathbf{R}^1, \dots, \mathbf{R}^N, i_1, \dots, i_n) g'(i_1, \dots, i_n|\mathcal{V}_{\hat{\rho}^0}, \sigma)) \end{aligned}$$

where $g'(i_1, \dots, i_n|\mathcal{V}_{\hat{\rho}^0}, \sigma) = \sum_{\hat{v} \in \mathcal{V}_{\hat{\rho}^0}} \{g^*(\hat{v}|\mathcal{V}_{\hat{\rho}^0}) \int_{\mathbf{x}} \mathcal{F}_r(i_1, \dots, i_n|x_1, \dots, x_n) \prod_{i=1}^n \mathcal{N}(x_i|\hat{v}_i, \sigma) d\mathbf{x}\}$, and

- $\hat{v} \sim g^*(\hat{v}|\mathcal{V}_{\hat{\rho}^0})$
- $x_i \sim \mathcal{N}(x_i|\hat{v}_i, \sigma)$ for $i = 1, \dots, n$
- $i_1, \dots, i_n \sim \mathcal{F}_r(i_1, \dots, i_n|x_1, \dots, x_n)$, where $\mathcal{F}_r = \begin{cases} 1, & \text{if } \{i_1, \dots, i_n\} = \text{rank}(x_1, \dots, x_n) \\ 0, & \text{otherwise} \end{cases}$

As N is limited, ρ^0 and therefore, \mathcal{V}_{ρ^0} usually cannot be accurately inferred from the data. We can however, sample for i_1, \dots, i_n by sampling for each item i from a univariate Gaussian distribution centered on \hat{v}_i with a fixed variance σ for all items, and then obtain a ranking using the rank function. By introducing the variance, a smaller KL divergence from the Mallows posterior can be achieved.

2.4

Theorem 2.4.1 *With the usage of $g'(i_1, \dots, i_n | \mathcal{V}_{\hat{\rho}^0}, \sigma)$, the value of σ that minimizes the KL-divergence between the Mallows posterior and the resulted pseudolikelihood is*

$$\sigma = \begin{cases} 0, & \text{if } \delta(\alpha^0, n, N) \leq \delta^* \\ f(\alpha^0, n, N), & \text{otherwise} \end{cases}$$

In other words, σ should be 0 when $\delta(\alpha^0, n, N) \geq \delta^*$. Beyond this point, the optimal choice of σ should be greater than 0, and it follows a function $f(\alpha^0, n, N)$.

2.5

Theorem 2.5.1 *As $N \rightarrow \infty$, $\sigma = 0 \ \forall \alpha > 0$ and ≥ 1*