

Reference Dependence and Monetary Incentive -Evidence from Major League Baseball-

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Many empirical studies have revealed the existence of reference-point dependent preference in field settings, including cases from professional sports players' decision making, whose performance is observed by the manager and evaluated. If there is some incentives that leads the individuals take "apparent" reference point dependent behavior, then we should rather think it is reference dependence of the managers that evaluates the players.. In this paper, I picked up the case of Major League Baseball players, which Pope and Simonsohn (2011) reported have reference-point dependent tastes about the number of their batting-average. I first confirmed the evidences that supports reference point dependent manipulation of batting-average and other performance indexes. Then I tested if there exist any monetary incentives that encourage the players to do so. My results are against this assumption: no overestimation of achieving the possible reference point was observed, which confirms that players adjust their aspiration level to reach their internal goal, even though there is no additional bonus that rewards their achievement.

1 Introduction

Reference-point dependence is one of the most important concepts to evaluate outcomes, and it affects agents' following economic behavior. Classical economic models assume that economic agents evaluate their choices/prospects according to the absolute value of (expected) return. On the other hand, Tversky and Kahneman (1992) introduced the be-

havioral assumption: reference-point dependent preference sets some target value of the outcomes and then subjects regard the possible outcome as gain or loss from the target. For example, workers feel happy if her/his wage goes up to \$15 per hour, but vice versa if it goes down to \$15, although the absolute value of \$15 is actually the same. In this case, s/he evaluates their new wage with the reference point of the previous one. In this paper, we picked up an empirical instance that this reference dependence occurs in nonmonetary outcomes, and tested if there exists any monetary incentives.

Prospect theory consists of two main characteristics: one is the probability weighting function, and the other is reference dependence, mentioned above. It enabled us to interpret phenomena that seem inconsistent with the traditional economic theory, and made it possible to understand them with some additional assumptions. Thus, a lot of following researches have been conducted in field and laboratory settings.

Reference dependence is also observed in the behavior of athletes. Pope and Schweizer (2011) found that professional golf players regard “per,” the standard number of shots determined according to the difficulty of each hole, as reference points. Also, Pope and Simonsohn (2011) tested the existence of the reference dependence in the Major League Baseball, a professional baseball league of America. When dealing with the case of the professional athletes, however, we must pay attention to how their salary contract is designed.

Suppose the case of the professional golf player. Golf is essentially competition of the total number of shots they needed to finish the whole tour, regardless of that of each single hole, or whether s/he saves per or not in the hole. Rank of order is determined according to the number of shots, and those with better scores are rewarded. Then, there appears a question that what if there is some monetary incentive to make effort to save per. Or it considers the following situation: when every time s/he saves per in each hole, then s/he can get some additional bonus separated from their total score. In this case, then, making effort to save per can be interpreted as sufficiently “rational” choice for the player,

although the observed behavior itself appears to be evidence of the reference dependence.

Dealing with reference dependence that occurs as some nonmonetary outcomes, therefore, we have to check if there exists any monetary incentives that modifies the benefit function of the individuals discontinuous, which leads to the behavior as if they have reference point dependent utility function. What we show in this paper is one of the empirical instances to be specified. We picked up the evidence from Major League Baseball (MLB) position players' behavior and considered the questions above.

MLB position players seems to have some reference points, about their batting performance indexes: .300 of batting average is one of the possible ones. Pope and Simonsohn (2011) have shown, that there exists bunching just above .300 of the distribution of this index. However, it might not be sufficient to prove that this is caused by reference point dependent utility function about their performance indexes: team managers may assign some monetary incentives for the players to adjust their aspiration level to meet those points, as we described in the example above.

Our most important contribution is this: to conduct analysis that reveals the observed behavior to be in fact reference dependence. First, we tested if there is any evidence for the manipulation around the round numbers, by using McCrary (2008)'s method to test manipulation. not only .300 of batting average, both for other points of batting-average and other indexes, such as homerun or stolen-bases. Confirming this, then we made examinations to answer the question, "Is this observed manipulation truly driven by the reference dependence of the players?" We applied regression analysis using the data of the players' salary.

Our paper found three important results. First, our examination for manipulation supported the previous study. There observed evidence to show there in fact exists seemingly reference point dependent behavior, where .300 of batting-average works as a reference point. Similar results were obtained about other round number of batting-average, and other batting indexes such as on-base percentage or homerun. Second, we found that

there does not exist any monetary incentive for them: for their fixed part of the salary contract, they do not receive any additional payment by achieving their internal goals. This result confirmed that observed behavior is actually the reference dependence of the players. Furthermore, we complement our results, discussing some alternative interpretation about the form of monetary incentives: the part of incentivised contract, and relation with contract length. And the third is that the indexes where players' reference dependence occurs vary through the history of the MLB: events that rotate how to design the contracts or evaluation about the importance of each performance index are likely to affect the behavior of the players.

This paper proceeds as follows. In Section 2, we review some literature and verify the standpoint of my paper. Section 3 describes the data we have. Section 4 presents the theoretical framework and empirical way to specification, and make some conjecture. Section 5 shows the results of the analysis. Discussion about some alternative interpretation and non-statistical data are included in Section 6. Finally, Section 7 concludes the paper.

2 Literature Review

Tversky and Kahneman (1992) mentioned reference point dependence as one of the two distinct respects of their prospect theory. The most primitive form of reference dependent utility function is:

$$u(x|r) = \begin{cases} x - r & \text{if } x \geq r \\ \lambda(x - r) & \text{if } x < r \end{cases}$$

where x denotes a certain outcome, and r is one of the reference points. This agent evaluates the outcome by the difference from the reference point. In addition, they assume "loss-aversion" of the individual, or $\lambda > 1$. Those who have this type of utility function,

than they regard same absolute amount of outcome in different way, depending on s/he faces gain or loss situation. “Diminishing sensitivity,” which is concave in facing gain and convex in facing loss is an advanced form of this specification.

Diecidue and Van de Ven (2008)’s “aspiration level” model added discontinuity assumption: that is, a utility function that “jumps” at the reference point. When there exists jump in their utility function, then individuals try to manipulate outcome level, paying additional cost which was not accepted in the standard “smooth” form of utility function. As mentioned below, I exploit this assumption to my model in this paper.

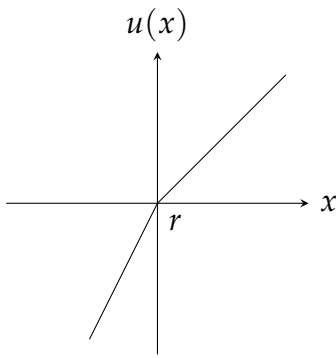


Figure 1: primitive gain-loss function

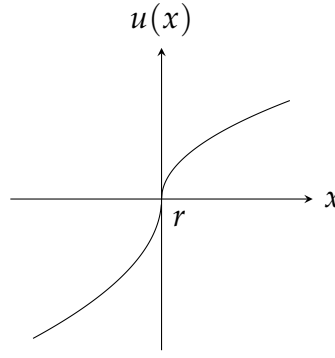


Figure 2: diminishing sensitivity

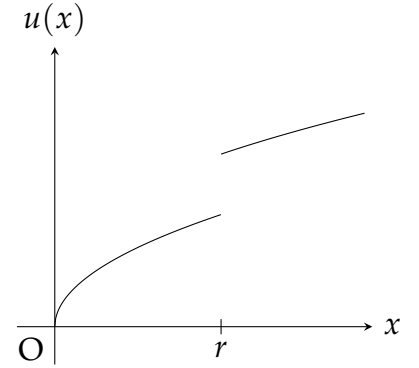


Figure 3: jump at the reference point

Individuals with such reference-point dependent utility try to adjust their effort level so as to achieve their internal target, or reference point. There is a number of empirical literature that specifies the existence of reference dependence in the field or lab studies. Farber(2008) applied this model to the labor supply of New York cab drivers to show that as soon as they reached daily target sales, they considers to quit, even when they reached it early in each day. Jones(2018) made analysis on the system of American tax payment. He showed that individuals try to manipulate their real payment by substituting it by donation or other charitable action, and that especially when facing losses, they make more effort. This observation is also caused loss-aversion, with the reference point of zero-payment threshold.

Reference dependence also occurs in the cases of sports. One of the most well-known papers among them is Pope and Schweizer (2011). They obtained the data of professional golf players, to point out that in each hole, players behave as they take “per” as the reference point: Specifically, they succeed their putts, significantly better when the putt was one to save per than when it was one to get “eagle” or “birdie.” Similarly, Allen *et al* (2016) specified the existence of reference point dependence of marathon runners, using data about the finish time of enormous number of race in the United States. In this case, the distribution of finishing time has excess mass around every hour and 30 minutes. Note that these cases are common in that even if they achieve their internal goals, they do not receive any monetary reward for their success. Professional golf players are awarded according to the total number of shots through the whole tour, not to the number of pars they saved.

Pope and Simonsohn (2011) mention a seemingly similar case. They picked up three empirical evidence of round number as reference points: SAT (a standardized test for college admission in the United States) scores, laboratory experiment, and baseball. In their section of baseball, they picked up the evidence of Major League Baseball (MLB) players to claim that they suggest their reference dependent preference by the distribution of their performance index: batting-average (AVG). According to their paper, the position players (batters) pay attention to their batting-average (AVG), especially to finish each season with their batting average of just above .300. They obtained MLB season individual AVG data from 1975 to 2008 and observed position players (= players except for pitchers) with at least 200 at-bats in each season. Then, they found that their distribution of the batting-average has excess mass just above .300, which reveals the existence of manipulation there. Furthermore, they found that players with batting-average of just below .300 are more likely to hit a base-hit and less likely to get a base-on-balls. Both base-hits and base-on-balls avoid the batter from being gotten out, so for the team he belongs to, base-on-balls also have important value to win the game. However, batting-average does not

count base-on-balls as the element to raise the number (For the definition of performance statistics, see Appendix), so they prefer getting hit to base-on-balls. Thus, observed behavior they claims is sufficient evidence that shows the existance of round-number reference point dependent preference of the MLB players.

It is true that there is observed behavior similar to the cases of Pope and Schweizer (2011) or Allen *et al* (2016). However, one important thing we have to take care of is there exists procedure of contract between the player and the team manager: those who evaluate the player. In other words, it may owe to their monetary value function, not by the preference of themselves. This is the main contribution of my research.

Pope and Simonsohn stated in their own paper that they conducted analysis only for batting-average, and following research is to be made. So I first follow this: "Is batting-average unique case?" Then, I test if there exists monetary incentive for the player. If the team managaers they belong to adopt a system of salary that discontinuously "jump" by a certain performance index reaching to the point where manipulation is observed, then for the players it can be interpreted as rational choice, given the contract design. In general, players with their performance index just above these cutoff point and those just below the point have almost same ability as a baseball player. At least, it is natural to think there is no reason to treat players discontinuously better, only because he achieve the cutoff. Then, it is interpreted that it is rather the team manager than the players themselves who have the reference-point dependent preference, which makes the players encouraged to meet their goals.

On the other hand, if there exists no evidence that team managers evaluate the players by the achievement of the cutoff, then we can say that the observed behavior is truly drawn by their own reference dependence. In addition, the consistency is so strong that even there exists no rational reason, they try to reach there. Analysing this and verify which hypothesis is my main contribution of this paper.

3 Data Description

In order to make empirical research, I need information about players' performance, contracts and other details. Then, I generated panel data that contains these specific information from some open data-source. Each sample is obtained by unit of a single season, but due to lack of open source of information about contracts, time range of the data used in each analysis is unbalanced.

Performance are obtained from baseball fan website: *fangraphs* and *Baseball Reference*. Since the regulation of "qualified plate-appearances," the cutoff number to get batting title such as batting-average, was introduced in 1957, I collected dataset from this year. Stats in each season contains that of only during the regular season, not that of Spring-Training or postseason games. The full-sample is $N = 54469$.

Salary data are obtained from American media site: *USA TODAY*. I collected salary data of the position players who are registered in MLB Roaster at the beginning of each season, and merged with the play stats in the previous year, because salary is as usually determined based on the performance of the previous season. Data was available only that since 1987, and as I need data of the next year about the salary, latest season is 2017. The aggregated number of the panel is $N = 13226$.

Then, for precise research, I made some sample-restriction. First of all, I have to distinguish the data with both play stats and contract, with that with only play stats. In section of specifying excess mass, I utilize the latter one (I call this Sample A), and in analyzing the monetary incentive, I use the former one (Sample B).

In considering rate indexes, we have to restrict the sample to the batters with a certain number of attendances, since these indexes of those with too little at-bats are enormously affected by the result of every single plate-appearances, and so team managers might not see them in evaluating performance. At Pope and Simonsohn (2011), the threshold was set at 200 at-bats, but in my research, I substitute at-bats with plate-appearances, since at-

Table 1: Summary Statistics for Sample A

Statistic	N	Mean	St. Dev.	Min	Pctl(25)	Pctl(75)	Max
PA	18,143	456.477	152.836	200	320	591	778
HR	18,143	11.811	9.747	0	4	17	73
RBI	18,143	51.882	26.912	4	31	69	165
SB	18,143	7.846	10.869	0	1	10	130
AVG	18,143	.264	.032	.135	.242	.285	.394
OBP	18,143	.331	.039	.174	.305	.356	.609
Age	18,143	28.506	4.042	18	25	31	46
H	18,143	108.941	42.933	29	72	143	262
OPS	18,143	.738	.106	.382	.665	.805	1.422

bats exclude the plate appearances with base-on-balls and hit-by-pitch, even though they attended the game and appear to the plate. Restricting the sample to the players with at least 200 plate-appearances yields $N = 18143$ for Sample A and $N = 8915$ for Sample B.

Play stats are tagged by players' ID and season, including batting indexes tested bunching: batting-average, on-base percentage, base-hits, stolen-base, homerun, runs-batted-in, and plate appearances. Moreover, for the analysis of monetary incentives, I collected other indexes that control the performance: slugging-average, OPS, fWAR, BATTING, FIELDING, BaseRun, and WPA. Salary data includes annual salary at the year: for the players with plural-year contract, it is calculated by deviding total payment of the contract by the years. It also contains the player-specific characteristics: age, position (catcher, firstbase-man, ... right fielder), team they signed, and possession of the right of free agency.

The summery statistics are descripted in Table 1 and Table 2.

4 Theoretical Frameworks and Way of Specification

4.1 Frameworks

MLB position players try to maximize their performance, for two main goals: one is to contribute to their team winning more games, and the other is to improve their living

Table 2: Summary Statistics for Sample B

Statistic	N	Mean	St. Dev.	Min	Pctl(25)	Pctl(75)	Max
+WPA	8,915	8.715	3.471	2.030	5.820	11.430	19.160
-WPA	8,915	-8.270	2.610	-15.050	-10.420	-6.060	-2.740
AB	8,915	419.820	133.528	162	305	537	716
AVG	8,915	.268	.031	.146	.248	.289	.394
Salary	8,915	3,487,838	4,487,344	62,500	512,750	4,658,334	29,200,000
Age	8,915	28.714	3.901	19	26	31	46
Bat	8,915	3.257	16.139	-44.200	-7.300	11.100	116.800
BsR	8,915	.092	2.712	-12.600	-1.200	1.200	14.300
FA	8,915	.168	.374	0	0	0	1
Fld	8,870	.304	7.482	-36.100	-4.000	4.400	37.000
H	8,915	114.232	42.481	30	78	148	262
HR	8,915	13.446	10.213	0	6	19	73
OBP	8,915	.337	.038	.174	.311	.360	.609
OPS	8,915	.761	.104	.382	.691	.825	1.422
PA	8,915	471.946	150.890	200	342	605.5	778
RBI	8,915	56.339	27.621	5	35	74	165
SB	8,915	8.534	10.851	0	1	11	109
SLG	8,915	.425	.075	.208	.372	.472	.863
fWAR	8,915	1.860	1.950	-3.100	.500	2.900	12.700

standard. That is, getting better contract and playing as long as possible is one of the most important object to play for the team. Performance index is clear and valid benchmarks to signal their ability to the team managers. Baseball is so rich in such indexes, so we can say these are good reliable proxy for their ability as a player. We assume a utility function of the player/manager as $u(x)$ for the value of a certain index x . Players' self-evaluation and his salary for the next season reflects this value.

Then, here I describe the nature of the indexes. Baseball batting indexes are roughly divided into two types. One is that simply indicates the number of a ceratain plays, and another is calculated using these numbers. Here I call the former "cumulative index," and the latter "rate statistics." Cumulative index are for example "base-hit" or "homerun," or "stolen-bases." They are irreversible and so once they reach a certain number then it cannot be decrease. On the other hand, rate indexes can be fluctuate, so in order to keep this kind of indexes above their goal, they have to continue to make effort or stop attending game and wait for the end of the season. Batting-average is sorted of this type, and On-base percentage or OPS are also kind of them. Note that even though I call this

for “rate” index for simplicity, it contains all the index that can both go up and down according to the player’s performance, such as OPS or slugging-average. See Appendix for specific definition of each index.

Anyway, the important assumption is that players evaluate their performance index using referene-dependent utility function of themselves or the team manager, and manipulate it in order to maximize their utility. This aspiration results in odd distribution of the index at the end of the season: excess mass around the possible referene point. In this paper, following Allen *et al* (2016), I exploit the assumption of the utility function of Diecidue and Van De Ven (2008)’s “aspiration level” model. The model presents the form of the utility function with “notch,”

$$\lim_{\epsilon \rightarrow 0} v_r(r + \epsilon) \neq \lim_{\epsilon \rightarrow 0} v_r(r - \epsilon)$$

This form of utility function is discontinuous at the reference point r .

In my paper, it is not determined who has such a utility function. At least one of the these two agents: players and managers has reference dependent feature in their utility. Players pay attention to if each index are above the reference points. In Allen *et al* (2016), it is assumed that discontinuous utility function occurs as excess mass around the reference points. In my paper, I suppose there may be another story. That is, discontinuous utility function of the managers emerges as discontinuous evaluation of the players: salary scheme jumps, depending on whether a ceertain performance index is above or below the reference point. Then, it modifies the income function of the players discontinuous, which results in bunching as the best response of them.

4.2 Empirical Method

4.2.1 Test for Bunching

First, I test if there is observed any behavior that seems to be related to reference dependent preference. As I explained in the previous chapters, the existence is verified by the observation of bunching, or excess mass around the possible reference point. To specify the excess mass, I apply the method of regression discontinuity design (RDD).

RDD is a way to measure the effect of a treatment, such that whether the treatment is assigned or not depends on the threshold of a certain variable (called “running variable”). Then, comparing the samples just above and just below the threshold is sufficient examination of the treatment, since they are in almost same states except for the existence of the treatment.

However, there is an important assumption for this specification to be valid (Lee and Lemieux, 2010): continuity of the running variable around the threshold. In other words, individuals must not be able to manipulate the running variable so as to be above the cutpoint. This is because if there exists manipulation, then there occurs selection bias problem, that those who try to be assigned the treatment can adjust their running variable. Therefore, there are some empirical way to test the manipulation of a variable, which is the very method I apply in my analysis.

One of the frequently applied methods for this specification is McCrary(2008)’s local linear density estimation. I avail this to my specification of bunching.

This estimation of the manipulation at the cutoff point c , proceeds in two steps: undersmoothing histograms and local linear smoothing. Undersmoothing means to determine the binsize b of the running variable: in this paper, the index argued.

$$g(R_i) = \left\lfloor \frac{R_i - c}{b} \right\rfloor b + \frac{b}{2} + c \in \left\{ \dots, c - 5\frac{b}{2}, c - 3\frac{b}{2}, c - \frac{b}{2}, c + \frac{b}{2}, c + 3\frac{b}{2}, c + 5\frac{b}{2}, \dots \right\}$$

where $\lfloor a \rfloor$ denotes the greatest integer less than a , and $g(R_i)$ stands for the height, or frequency of each bin. Then, we make local approximation above and below c , the possible value of reference dependence point. Bandwidth is calculated to minimize the estimation error. Then, we estimate the frequency at the cutoff point, $\hat{f}(r)$, by fitting the estimated density function below and above the cutoff, \hat{f}^+ and \hat{f}^- , respectively. Finally, we take the difference between $\ln \hat{f}^+$ and $\ln \hat{f}^-$ to calculate the statistics θ . With θ and its estimator of standard deviation, t -tests can be constructed to specifying manipulation.

4.2.2 Monetary Incentive

Then, We examine the existence of the monetary incentive. We employ the method of sharp regression discontinuity design (RDD): local linear regression around the possible reference point as the cutoff.

$$w_{it} = \beta_0 X_{it} + \beta_1 \text{ABOVE}_{it} + \beta_2 Z_{it}$$

For each player i in the season t , w_{it} is log annual salary in next season $t + 1$. X_{it} and Z_{it} are the value of performance index (batting-average, on-base percentage,...) and other player-specific characters (age, team he signed, position,...), respectively. ABOVE_{it} is an indicator of the achievement of their goals for the index. That is, if he finished his season with a certain index with above the possible reference point, it takes 1, and otherwise 0. This is the variable in interest. As well as OLS, team, position, or individual fixed effect model estimation was also conducted.

where the bandwidth was selected according to the polynomial approximation (Imbens, Guido and Kalyanaraman, 2009). As I mentioned above, RDD can be invalid when individuals can manipulate the running variable, so this may not appropriate in this analysis since I confirm there exists manipulation in the running variable: performance index. However, as I show below, this analysis found no evidence of discontinuous design of

contract. Therefore, I apply this result as a complimentary evidence of my research. We pay attention to whether for players around the possible reference point, it is worth manipulating their indexes or not, so in each analysis, we restrict our sample to the players there. The primitive model to check this is below:

Also, to check the robustness of my analysis, I tried alternative methodology: include interaction term, and methodology of regression discontinuity design. Interaction term analysis include that of X_{it} and $ABOVE_{it}$, the coefficient of which indicates additional return on their achievement.

To restrict sample around sample, it is ad hoc way to choose the range of players in the sample.

5 Result

5.1 Excess Mass Around The Reference Point

First, I show the results that verifies bunching. Table 3 includes the summary of McCrary (2008)'s manipulation tests about the performance indexes that say there is some manipulation occurs. Consistent with Pope and Simonsohn (2011), there actually occurs excess mass around .300, and .250 of batting-average. Also, manipulation were also observed in some of other numbers of other indexes: .350 of on-base percentage, 20 of homeruns, 100 of runs-batted-in, 30 and 40 of stolen-bases, and 200 of base-hits. Table shows parts of results that denies the existance of excess mass.

For batting-average, extending sample size from Pope and Simonsohn (2011) yields similar results: Players manipulate their batting-average. The difference between the estimated frequency according to the approximation below .300 and that of above .300 was significant at .1% ($z = 7.442$) level. In addition, bunching occurs also in .250 ($z = 5.061$, $p < 0.1\%$). This is inconsistent with Pope and Simonsohn (2011), in that there was no bunch-

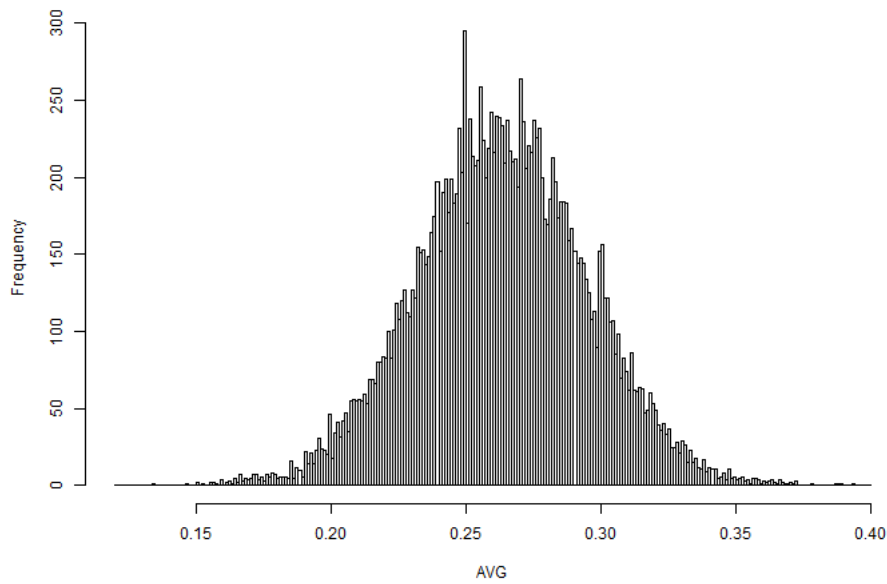


Figure 4: Histogram of Batting-Average

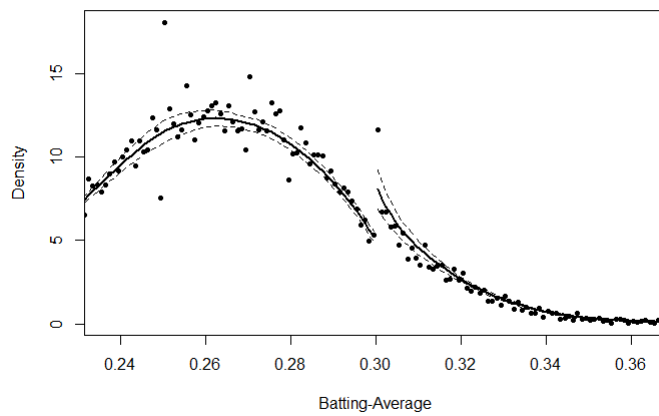


Figure 5: Discontinuity at .300 of AVG

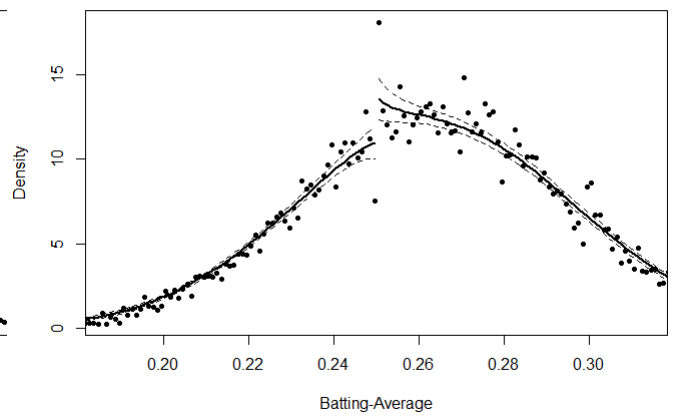


Figure 6: Discontinuity at .250 of AVG

index	type	cutpoint	binsize	bandwidth	θ	z
AVG	rate	.300	.001	.019	.499 (.067)	7.442***
		.250	.001	.024	.212 (.042)	5.061***
OBP	rate	.350	.001	.024	.139 (.049)	2.854**
HR	cumulative	20	1	5.309	.259 (.075)	3.465***
RBI	cumulative	100	4	15.423	.311 (.094)	3.295***
SB	cumulative	30	1	10.000	.529 (.124)	4.274***
		40	1	11.505	.481 (.174)	2.764**
PA	cumulative	500	1	.003	.160 (.063)	2.515*
H	cumulative	200	1	18.922	.453 (.178)	2.547 *

Note ***: $p < 0.1\%$, **: $p < 1\%$, *: $p < 5\%$.
Bandwidth is optimized following the method of McCrary(2008).

Table 3: Test for Manipulation, leastPA = 200

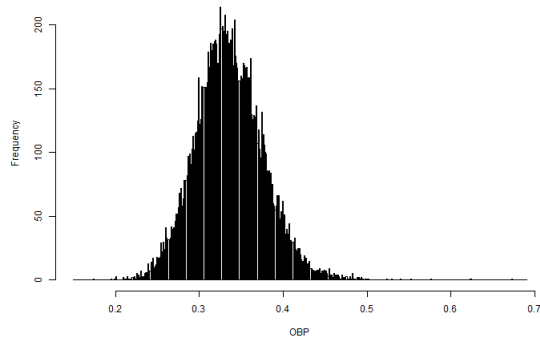


Figure 7: Histogram of On-Base Percentage

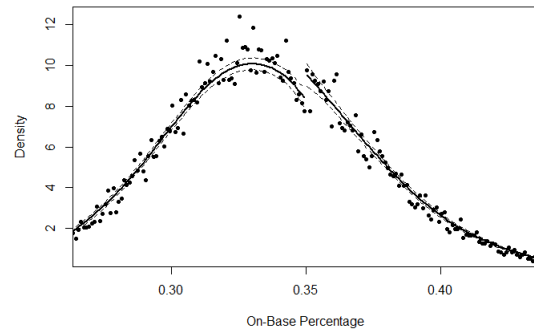


Figure 8: Discontinuity at .350 of OBP

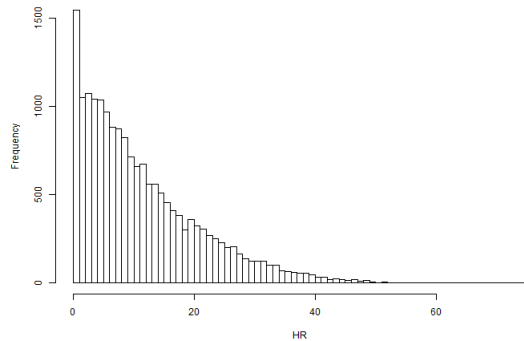


Figure 9: Histogram of Homerun

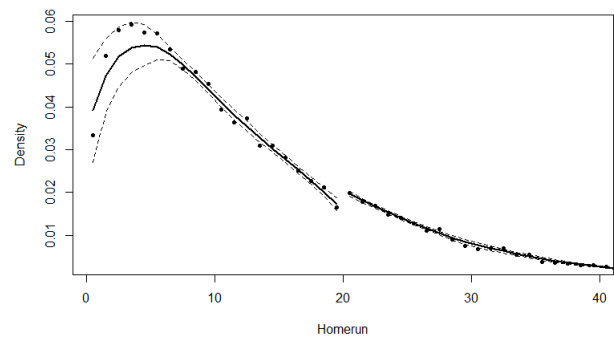


Figure 10: Discontinuity at 20 of HR

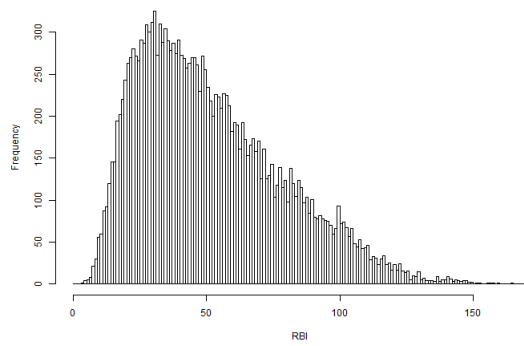


Figure 11: Histogram of Runs-Batted-In

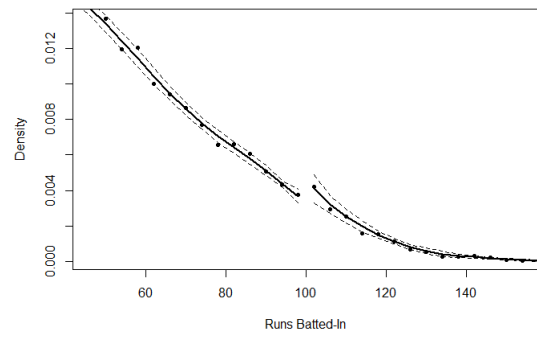


Figure 12: Discontinuity at 100 of RBI

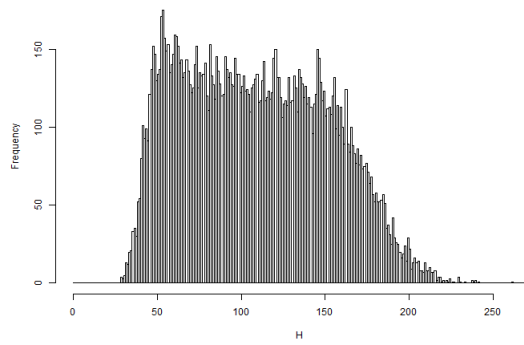


Figure 13: Histogram of Base-Hit

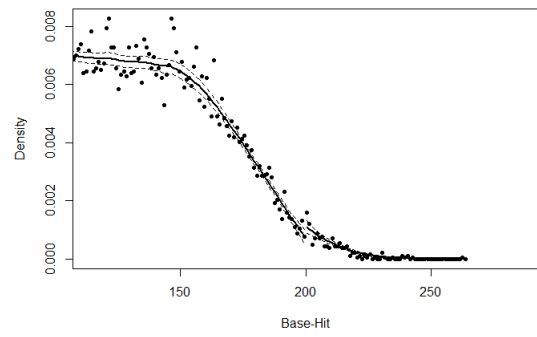


Figure 14: Discontinuity at 200 of Base-Hit

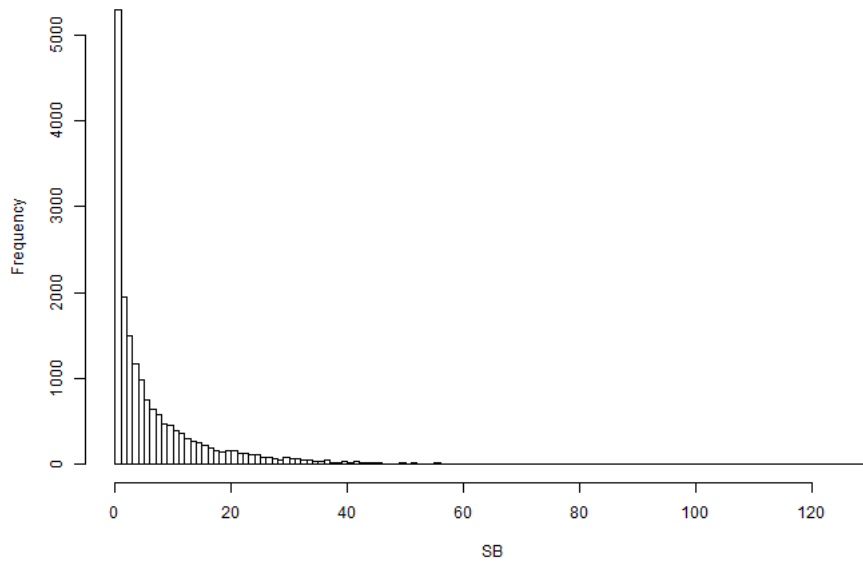


Figure 15: Histogram of Stolen-Base

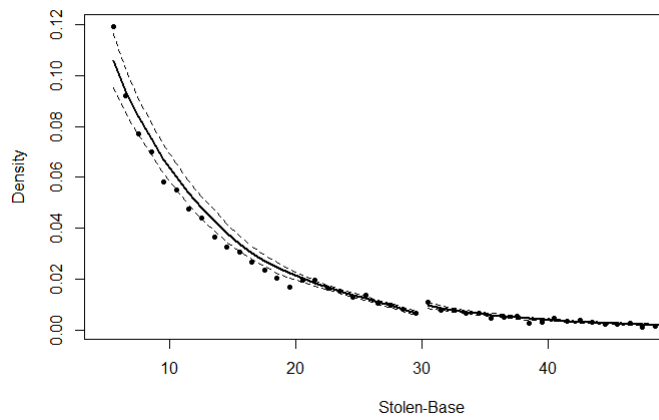


Figure 16: Discontinuity at .300 of AVG

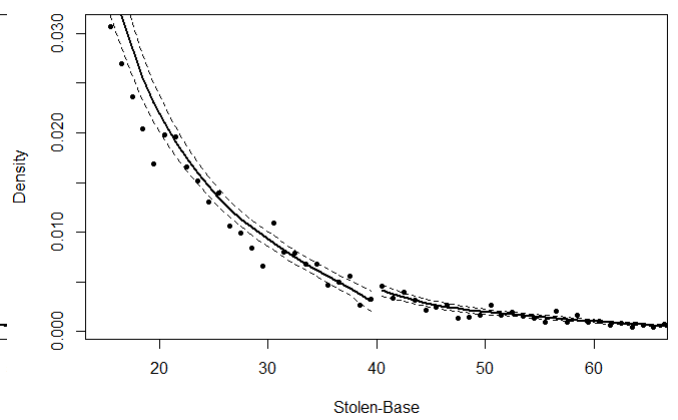


Figure 17: Discontinuity at .250 of AVG

ing observed in any other round numbers of batting-average, so it can be related to that the sample size is extended to further old ages. I specifically analyzed this in Section 6.4. Furthermore, changing binsize yields similar results.

On-base percentage, on the other hand, showed similar tendency in .350, although the significance level was $z = 2.854$ (significant at 5% level). Pope and Simonsohn (2011) reported that they are less likely to get base-on-balls when facing marginal range of batting-average and so they feel of more importance in batting-average than on-base percentage. However, when they face marginal point of on-base percentage, they may try to get more base-on-balls in order to achieve .350.

Bunching also occurs in cumulative indexes. Cumulative ones are irreversible, and so players may feel these indexes are easier to manipulate. As same as batting-average and on-base percentages, however, such manipulation was observed in only limited numbers: not all the round numbers. For Homerun ($z = 3.465, p < 0.1\%$), bunching occurred only in 20: there may be diminishing sensitivity: 20 is located on above the 75 percentiles of the whole Sample A. Also, stolen-base and base-hit have shown bunching, at 30 ($z = 4.274, p < 0.1\%$) and 40 ($z = 2.764, p < 1\%$) of stolen-base and 200 in base-hit ($z = 2.547, p < 5\%$). Stealing-bases are skill that are talented to limited number of the players, but succeeds with the probability of 60% to almost 100%, so those who are evaluated by their number of stolen-bases, it may certain be relatively accessible number to manipulate. Base-hit is manipulated by similar to batting-average, but for this index, as same as other cumulative ones, it cannot be valid to be replaced in the last plate-appearance, which seems to lower the confidence level of the z -statistics. And surprisingly such manipulation occurs also in runs-batted-in. Compared to other indexes, runs-batted-in is harder to manipulate, since the number of that depends on the performance of his teammates, and the number they can earn at a single plate-appearance varies from 1 to 4. As in Table 3, the test said that there occurs the evidence in plate-appearances. However, it may insufficient because the optimized bandwidth was smaller than one, even though

the number of PAs takes only integers. Setting bandwidth larger than 1, then the result went insignificant.

Similar to the rate indexes, there occur manipulation in only some parts of round numbers. That is, it should be considered that the reference points are not determined only because they are round numbers as Pope and Simonsohn (2011) argued. The possible numbers of reference points found in this section are kinds of well-known indicators of the skill of the players. That is, the nature of the reference points are likely to be close to “per” in Pope and Schweizer (2011) rather than round numbers; or well, these numbers are monetarily incentivised goals by the team managers. So next, I examine whether there is any monetary bonus in their contract.

5.2 Existence of Monetary Incentive

Table 4: RDD Test for Monetary Incentives

index,cutpoint	Other Control	bw type	bandwidth	Observations	Estimate	Std. Error	z
AVG, .300	No	LATE	.084	8514	.047	.061	.773
		Half-BW	.042	5599	.088	.075	1.174
		Double-BW	.170	8915	.067	.056	1.184
	Yes	LATE	.045	5930	.034	.056	.615
		Half-BW	.023	3005	.061	.077	.788
		Double-BW	.090	8605	.016	.045	.354
AVG, .250	No	LATE	.036	6110	.019	.068	.286
		Half-BW	.018	3496	.015	.092	.161
		Double-BW	.072	8539	.034	.054	.636
	Yes	LATE	.048	7271	.070	.052	1.340
		Half-BW	.024	4402	.066	.069	.953
		Double-BW	.096	8810	.075	.044	1.713
HR, 20	No	LATE	3.32	1315	.071	.175	.406
		Half-BW	1.66	562	.073	.127	.576
		Double-BW	6.64	2582	-.004	.109	-.034
	Yes	LATE	3.30	1307	-.002	.141	-.015
		Half-BW	1.65	560	.030	.102	.299
		Double-BW	6.61	2558	-.032	.088	-.364
OBP, .350	No	LATE	.044	6440	-.038	.065	-.592
		Half-BW	.021	3542	-.076	.089	-.849
		Double-BW	.087	8656	-.029	.051	-.570
	Yes	LATE	.045	6525	-.013	.049	-.272
		Half-BW	.022	3673	-.055	.069	-.807
		Double-BW	.089	8637	.004	.039	.107
RBI, 100	No	LATE	4.08	393	.072	.289	.250
		Half-BW	2.04	228	.282	.400	.707
		Double-BW	8.16	714	.008	.185	.043
	Yes	LATE	4.04	390	.018	.209	.086
		Half-BW	2.02	227	-.042	.324	.130
		Double-BW	8.07	708	.056	.127	.435
H, 200	No	LATE	4.08	393	.072	.289	.250
		Half-BW	2.04	228	.282	.400	.707
		Double-BW	8.16	714	.008	.185	.043
	Yes	LATE	4.04	390	.018	.209	.086
		Half-BW	2.02	227	-.042	.324	.130
		Double-BW	8.07	708	.056	.127	.435
SB, 30	No	LATE	3.39	282	.962	.372	2.585**
		Half-BW	1.70	134	.920	.263	3.492***
		Double-BW	8.16	714	.008	.185	2.941**
	Yes	LATE	3.40	282	.379	.297	1.271
		Half-BW	1.70	134	.290	.249	1.163
		Double-BW	6.79	533	.408	.180	2.260*
SB, 40	No	LATE	3.16	134	-1.276	.453	-2.818**
		Half-BW	1.58	56	-.736	.383	-1.924
		Double-BW	6.32	245	-.712	.313	-2.274*
	Yes	LATE	3.16	134	-.346	.396	-.875
		Half-BW	1.58	56	-.313	.429	-.730
		Double-BW	6.33	245	-.115	.244	-.472

Note:

***: $p < 0.1\%$, **: $p < 1\%$, *: $p < 5\%$.

Bandwidth is optimized following the method of Imbens-Kalyanaram.

Table 5: Regression on Log-Salary, Including Interaction Term: around .300

Dependent variable: <i>Sal</i>								
	(1)	(2)	OLS (3)	(4)	(5)	(6)	<i>fdm</i> (7)	(8)
Constant	11.166*** (.423)	9.728*** (.399)	-6.616*** (.665)	-5.049*** (.673)	-5.161*** (.669)			
AVG	11.513*** (1.537)	11.604*** (1.405)	11.620*** (1.209)	4.232*** (1.212)	4.083*** (1.204)	4.221*** (1.201)	3.808** (1.189)	3.752** (1.412)
AVG_300	-.169 (1.050)	-.455 (.954)	-.413 (.821)	-.261 (.789)	-.213 (.784)	-.142 (.780)	-.069 (.706)	-.279 (.918)
FLD		.004 (.002)	.006*** (.002)	.007*** (.002)	.007*** (.002)	.007*** (.002)	.008*** (.002)	.005** (.002)
BsR		-.014** (.005)	.009* (.005)	.003 (.004)	.003 (.004)	.003 (.005)	.020*** (.005)	-.018** (.005)
AGE			1.005*** (.039)	.983*** (.037)	.984*** (.037)	.997*** (.037)		
AGE_sq			-.015*** (.001)	-.014*** (.001)	-.014*** (.001)	-.014*** (.001)		
AVG:AVG_300	.663 (3.429)	1.689 (3.118)	1.428 (2.681)	.930 (2.577)	.791 (2.560)	.540 (2.549)	.160 (2.312)	1.077 (3.000)
WPA				X	X	X	X	X
EA dummy					X	X	X	X
Season dummies		X	X	X	X	X	X	X
Fixed effects						Position	Individual	Position
Observations	5,960	5,930	5,930	5,930	5,930	5,930	5,930	5,930
R ²	.035	.215	.420	.464	.471	.478	.744	.276
Adjusted R ²	.035	.210	.416	.461	.468	.473	.660	.271
Residual Std. Error	1.286 (df = 5956)	1.164 (df = 5894)	1.001 (df = 5892)	.962 (df = 5890)	.956 (df = 5889)	.950 (df = 5880)	.764 (df = 4459)	1.119 (df = 5882)
F Statistic	71.983*** (df = 3; 5956)	46.134*** (df = 35; 5894)	115.152*** (df = 37; 5892)	130.803*** (df = 39; 5890)	131.204*** (df = 40; 5889)			
***p<0.05; **p<0.01; *p<0.001								
Note:								

* p<0.05; ** p<0.01; *** p<0.001

Table 6: Regression on Log-Salary, Including Interaction Term: around .250

	Dependent variable:					
	Sal					
	(1)	(2)	OLS (3)	(4)	(5)	(6)
Constant	12.039*** (.501)	10.292*** (.457)	-6.061*** (.632)	-5.288*** (.632)	-5.513*** (.624)	
AVG	8.003*** (2.145)	9.795*** (1.926)	8.623*** (1.636)	2.156 (1.621)	1.906 (1.599)	1.554 (1.592)
AVG_250	-.684 (.618)	-.250 (.555)	-.597 (.471)	-.771 (.459)	-.783 (.452)	-.923* (.450)
FLD		.003 (.002)	.004** (.002)	.006*** (.002)	.005*** (.001)	.006*** (.001)
BsR		-.027*** (.005)	-.001 (.004)	-.006 (.004)	-.006 (.004)	-.006 (.005)
AGE			1.021*** (.034)	1.013*** (.033)	1.019*** (.033)	1.030*** (.033)
AGE_sq			-.015*** (.001)	-.015*** (.001)	-.015*** (.001)	-.015*** (.001)
AVG:AVG_250	2.836 (2.520)	1.098 (2.262)	2.591 (1.922)	3.383 (1.871)	3.409 (1.845)	3.957* (1.836)
WPA				X	X	X
FA dummy					X	X
Season dummies		X	X	X	X	X
Fixed effects						
Observations	7,307	7,271	7,271	7,271	7,271	7,271
R ²	.032	.231	.445	.476	.490	.735
Adjusted R ²	.031	.227	.442	.473	.487	.655
Residual Std. Error	1.271 (df = 7303)	1.135 (df = 7235)	.964 (df = 7233)	.937 (df = 7231)	.924 (df = 7230)	.758 (df = 5590)
F Statistic	79.391*** (df = 3; 7303)	62.031*** (df = 35; 7235)	156.664*** (df = 37; 7233)	168.232*** (df = 39; 7231)	173.678*** (df = 40; 7230)	.918 (df = 7221)

Note:

* p<0.05; ** p<0.01; *** p<0.001

Table 7: Regression on Log-Salary, Including Interaction Term: around .350

	Dependent variable:							
	OLS				felm			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	10.793*** (.516)	9.117*** (.472)	-7.002*** (.666)	-6.993*** (.689)	-7.081*** (.683)			
OBP	10.357*** (1.574)	11.155*** (1.421)	10.183*** (1.228)	5.580*** (1.258)	5.445*** (1.248)	5.346*** (1.242)	6.739*** (1.212)	7.368*** (1.467)
OBP_350	-.132 (.888)	-.625 (.801)	.103 (.691)	.080 (.676)	.103 (.670)	.130 (.667)	-.832 (.620)	-.625 (.788)
FLD		.006** (.002)	.007*** (.002)	.008*** (.002)	.007*** (.002)	.008*** (.002)	.007*** (.002)	.007*** (.002)
BsR		-.022*** (.005)	.002 (.004)	.001 (.004)	.001 (.004)	-.0003 (.005)	.022*** (.005)	-.020*** (.005)
AGE			1.016*** (.036)	1.021*** (.035)	1.021*** (.035)	1.037*** (.035)		
AGE_sq			-.015*** (.001)	-.015*** (.001)	-.015*** (.001)	-.015*** (.001)		
OBP:OBP_350	.356 (2.516)	1.783 (2.269)	-.237 (1.960)	-.188 (1.915)	-.274 (1.898)	-.321 (1.889)	2.166 (1.760)	1.790 (2.232)
WPA				X	X	X	X	X
FA dummy					X	X	X	X
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	6,656	6,620	6,620	6,620	6,620	6,620	6,620	6,620
R ²	.036	.231	.427	.453	.462	.470	.733	.260
Adjusted R ²	.035	.227	.424	.450	.459	.466	.650	.255
Residual Std. Error	1.267 (df = 6652)	1.135 (df = 6584)	.980 (df = 6582)	.957 (df = 6580)	.949 (df = 6579)	.943 (df = 6570)	.764 (df = 5042)	1.114 (df = 6572)
F Statistic	81.971*** (df = 3; 6652)	56.520*** (df = 35; 6584)	132.495*** (df = 37; 6582)	139.690*** (df = 39; 6580)	141.515*** (df = 40; 6579)			

Note:

* p<0.05; ** p<0.01; *** p<0.001

Table 8: Regression on Log-Salary, Including Interaction Term: around 20 HR

	Dependent variable:						
	OLS			Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Constant	14.429*** (1.092)	12.592*** (1.045)	-6.385*** (1.381)	-5.374*** (1.425)	-5.750*** (1.409)		
HR	.014 (.061)	.036 (.057)	.076 (.046)	.076 (.045)	.074 (.045)	.083 (.045)	.070 (.054)
HR_20	-1.154 (1.393)	-.474 (1.313)	.008 (1.055)	-.142 (1.035)	-.051 (1.022)	-.003 (1.025)	.993 (1.248)
FLD		.005 (.004)	.010*** (.003)	.010** (.003)	.010*** (.003)	.010*** (.003)	.006 (.004)
BsR		-.032** (.011)	.001 (.009)	-.011 (.009)	-.012 (.009)	-.022* (.010)	-.007 (.015)
AGE			1.109*** (.073)	1.107*** (.072)	1.123*** (.071)	1.127*** (.071)	
AGE_sq			-.016*** (.001)	-.016*** (.001)	-.016*** (.001)	-.016*** (.001)	
HR:HR_20	.061 (.073)	.025 (.069)	-.005 (.055)	.0004 (.054)	-.003 (.054)	-.007 (.054)	-.052 (.065)
WPA				X	X	X	X
FA dummy							
Season dummies		X	X	X	X	X	X
Fixed effects							
Observations	1,315	1,307	1,307	1,307	1,307	Position 1,307	Individual 1,307
R ²	.009	.172	.467	.488	.501	.508	.806
Adjusted R ²	.007	.149	.452	.473	.485	.489	.610
Residual Std. Error	1.203 (df = 1311)	1.114 (df = 1271)	.894 (df = 1269)	.877 (df = 1267)	.867 (df = 1266)	.864 (df = 1257)	.755 (df = 650)
F Statistic	3.882** (df = 3; 1311)	7.543*** (df = 35; 1271)	30.087*** (df = 37; 1269)	31.023*** (df = 39; 1267)	31.805*** (df = 40; 1266)		

* p<0.05; ** p<0.01; *** p<0.001

Table 9: Regression on Log-Salary, Including Interaction Term: around 30 SB

	Dependent variable:							
	OLS				Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	13.466*** (3.672)	8.006* (3.191)	-4.126 (3.586)	-3.729 (3.651)	-2.949 (3.675)			
SB	.029 (.132)	.167 (.114)	-.036 (.099)	-.041 (.099)	-.052 (.099)	-.067 (.096)	.075 (.185)	.145 (.113)
SB_30	12.468** (4.569)	13.567*** (3.972)	2.544 (3.531)	2.234 (3.526)	1.579 (3.541)	.816 (3.464)	7.200 (6.586)	12.144** (3.947)
FLD		.003 (.007)	.007 (.006)	.007 (.006)	.007 (.005)	.008 (.005)	.013 (.011)	.005 (.006)
BAT		.028*** (.003)	.023*** (.003)	.015** (.006)	.014* (.006)	.017** (.006)	.019 (.010)	.021** (.007)
AGE			1.142*** (.174)	1.150*** (.174)	1.105*** (.176)	1.132*** (.172)		
AGE_sq			-.018*** (.003)	-.018*** (.003)	-.017*** (.003)	-.017*** (.003)		
SB_SB_30	-.391* (.158)	-.444** (.137)	-.071 (.122)	-.061 (.122)	-.039 (.122)	-.013 (.119)	-.230 (.229)	-.397** (.136)
WPA				X	X	X	X	X
FA dummy		X		X	X	X	X	X
Season dummies			X					
Fixed effects								
Observations	282	282	282	282	282	Position 282	Individual 282	Position 282
R ²	.075	.472	.628	.633	.636	.675	.915	.523
Adjusted R ²	.065	.397	.571	.573	.576	.608	.651	.430
Residual Std. Error	1.229 (df = 278)	.988 (df = 246)	.833 (df = 244)	.830 (df = 242)	.828 (df = 241)	.796 (df = 233)	.751 (df = 68)	.960 (df = 235)
F Statistic	7.559*** (df = 3; 278)	6.276*** (df = 35; 246)	11.110*** (df = 37; 244)	10.684*** (df = 39; 242)	10.537*** (df = 40; 241)			

* p<0.05; ** p<0.01; *** p<0.001

Table 10: Regression on Log-Salary, Including Interaction Term: around 40 SB

Dependent variable: <i>Sal</i>								
	(1)	(2)	OLS (3)	(4)	(5)	(6)	<i>fdm</i> (7)	(8)
Constant	.097 (6.865)	2.795 (6.575)	-6.186 (6.345)	-5.258 (6.781)	-3.914 (6.974)			
SB	.392* (.182)	.287 (.174)	.191 (.155)	.183 (.158)	.148 (.163)	.148 (.163)	.679* (.265)	.187 (.182)
SB.40	20.483* (8.497)	7.339 (8.171)	5.634 (7.281)	5.543 (7.430)	5.082 (7.461)	7.669 (7.511)	37.152** (12.783)	7.632 (8.367)
FLD		.003 (.009)	.005 (.008)	.005 (.008)	.005 (.008)	.008 (.009)	.021 (.020)	.005 (.009)
BAT		.025*** (.005)	.023*** (.004)	.020* (.009)	.020* (.009)	.021* (.009)	-.001 (.020)	.017 (.010)
AGE			.763** (.239)	.751** (.242)	.751** (.243)	.745** (.245)		
AGE.sq			-.011** (.004)	-.011** (.004)	-.011** (.004)	-.011* (.004)		
SB.SB.40	-.537* (.218)	-.205 (.210)	-.152 (.187)	-.149 (.191)	-.134 (.192)	-.198 (.194)	-.956** (.328)	-.203 (.215)
WPA				X	X	X	X	X
EA dummy								
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	134	134	134	134	134	134	134	134
R ²	.062	.519	.634	.636	.638	.665	.949	.559
Adjusted R ²	.041	.353	.499	.490	.488	.494	.645	.348
Residual Std. Error	1.158 (df = 130)	.951 (df = 99)	.837 (df = 97)	.845 (df = 95)	.846 (df = 94)	.841 (df = 88)	.705 (df = 19)	.955 (df = 90)
F Statistic	2.875* (df = 3; 130)	3.137*** (df = 34; 99)	4.674*** (df = 36; 97)	4.363*** (df = 38; 95)	4.257*** (df = 39; 94)			
Note: *p<0.05; **p<0.01; ***p<0.001								

*p<0.05; **p<0.01; ***p<0.001

Table 11: Regression on Log-Salary, Including Interaction Term: around 100 RBI

	Dependent variable:						
	OLS			Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Constant	17.552* (7.013)	13.364* (6.189)	-3.577 (5.256)	-3.767 (5.174)	-3.797 (5.183)		
RBI	-.023 (.072)	.001 (.063)	-.017 (.052)	-.007 (.051)	-.007 (.051)	-.011 (.050)	.049 (.078)
RBI_100	-11.458 (8.459)	-3.512 (7.492)	-2.428 (6.115)	-1.810 (5.971)	-1.692 (6.001)	-2.447 (5.966)	.582 (8.865)
FLD		-.0002 (.005)	.004 (.004)	.004 (.004)	.004 (.004)	.007 (.004)	-.004 (.007)
BsR		-.035* (.014)	.001 (.012)	-.008 (.012)	-.008 (.012)	-.016 (.013)	.034 (.023)
AGE			1.185*** (.127)	1.195*** (.125)	1.195*** (.125)	1.187*** (.126)	
AGE_sq			-.018*** (.002)	-.018*** (.002)	-.018*** (.002)	-.018*** (.002)	
RBI_RBI_100	.115 (.086)	.036 (.076)	.026 (.062)	.019 (.060)	.018 (.061)	.026 (.060)	-.007 (.090)
WPA				X	X	X	X
FA dummy		X				X	X
Season dummies			X	X	X	X	X
Fixed effects							
Observations	393	390	390	390	390	390	390
R ²	.015	.345	.569	.592	.592	.614	.895
Adjusted R ²	.007	.281	.523	.546	.545	.559	.670
Residual Std. Error	1.033 (df = 389)	.880 (df = 354)	.717 (df = 352)	.699 (df = 350)	.700 (df = 349)	.690 (df = 340)	.841 (df = 342)
F Statistic	1.984 (df = 3; 389)	5.333*** (df = 35; 354)	12.547*** (df = 37; 352)	13.012*** (df = 39; 350)	12.654*** (df = 40; 349)		

*p<0.05; **p<0.01; ***p<0.001

Note:

Table 12: RDD Test for Discontinuity

index,cutpoint	Other Control	bw type	bandwidth	Observations	Estimate	Std. Error	z
AVG, .300	No	LATE	.025	503	-.175	.197	-.888
		Half-BW	.013	252	-.307	.302	-1.016
		Double-BW	.052	1043	-.180	.141	-1.271
	Yes	LATE	.026	509	-.253	.138	-1.832
		Half-BW	.013	266	-.209	.212	-.986
		Double-BW	.052	1038	.199	.102	-1.938
AVG, .250	No	LATE	.056	1366	.074	.102	.721
		Half-BW	.028	910	.147	.133	1.099
		Double-BW	.114	1501	.067	.090	.735
	Yes	LATE	.058	1367	.084	.082	1.020
		Half-BW	.029	923	.149	.107	.398
		Double-BW	.117	1480	.070	.072	.964
HR, 20	No	LATE	3.48	211	-.302	.300	-1.007
		Half-BW	1.74	96	-.123	.226	-.543
		Double-BW	6.96	387	.045	.203	.224
	Yes	LATE	3.50	206	-.273	.296	-.924
		Half-BW	1.75	95	-.156	.278	-.560
		Double-BW	7.00	439	-.098	.174	-.565
OBP, .350	No	LATE	.045	1103	.034	.129	.262
		Half-BW	.023	597	-.106	.172	-.620
		Double-BW	.092	1469	.024	.105	.225
	Yes	LATE	.043	1044	.021	.107	.196
		Half-BW	.021	566	-.085	.153	-.558
		Double-BW	.086	1435	.016	.084	.194
RBI, 100	No	LATE	4.90	50	-.100	.559	-.179
		Half-BW	2.45	30	-.095	.949	-.101
		Double-BW	9.80	102	.256	.333	.770
	Yes	LATE	4.93	49	.195	.433	.449
		Half-BW	2.46	30	-1.360	1.295	-1.050
		Double-BW	9.86	100	.398	.160	2.481*

Note:

***: $p < 0.1\%$, **: $p < 1\%$, *: $p < 5\%$.
Bandwidth is optimized following the method of Imbens-Kalyanaraman.

6 Alternative Interpretation and Some Evidence

In section 5, my analysis said that there in fact exists manipulation in some of the batting indexes, but no evidence observed in their contracts, which supports the assumption these observations are driven by the reference dependence of the players themselves. Here I consider some possible alternative and additional discussion about my result.

6.1 Incentivised Contract

One most possible explanation that may interpret my result is the incentivised design of the contract. So far, I checked monetary incentive for the player, analyzing only the fixed parts of the contract. However, players often sign contracts with additional bonus according to their performances. Even though the results in Section 5.2 did not support the existence of incentives in fixed salary, it may occurs as this addition rewards. Here I present that this story is hard to be applied, showing the specific contracts of the some players.

Table shows the specific contents of the MLB position players' contracts with incentivised bonus. It is clear that there is no conditions which pay additional salary depending on whether he reaches certain round numbers of their performance indexes, except for plate-appearances.

6.2 Misunderstanding by the Players

Misunderstandings by the player is also a possible way to interpret the results.

6.3 Contract Length

Skilled players usually get contracts with plural-year package.

6.4 By-Time Analysis

My paper used data from wide range of time: 62 years for bunching, 31 years for monetary incentives. Through such a long time, techniques of the players or the quality of instruments must have evolved, which leads to change in mean or the standard value of the indexes. Also, it is natural to think there may have been a lot of change in the design of the contract they agreed. Here I consider time-variable elements in my analysis. Specifically, there are two main possible effect that changes the contract design: one is the relative market power of the players, and another is change in relative importance of each performance index.

Relative market power has direct relation to the contract. Before the system of free agency was introduced, players are forbidden to move to other teams without permission by the team they belong to. '94 strike by the Players Association of Major League Baseball, against the team owners to request improvement of their treatment, also may have great influence on their contracts (See Appendix about the specific information about free agency and Strike).

Relative importance captures the change in evaluation of each index. Through the history of baseball, there have been invented a lot of indexes that measures the performance/ability of the player, and it has been argued which index is the best to evaluate them. One of the most important revoution was the publication of '*Moneyball*'(2003), written by Michael Lewis. In this book, he described that batting-average is not as appropriate measure: there is more close correlation with total runs the team earns in the season in on-base percentage. In practice, Oakland Athletics applied strategy to form the member of the team, and won the playoff. This story was widely spread and changed the sense of view about the baseball index. Then, one possible question occurs: "Does the tendency of manipulation/discontinuous contract design also change through the history of baseball?"

In this analysis, I replicate the methodologies conducted in the previous sections, but sorting the sample into periods below:

1. Before Free Agency (1957 - 1975)
2. After Free Agency and Before Strike (1986 - 1994)
3. After Strike and Before *Moneyball* (1995 - 2003)
4. After *Moneyball* (2004 - 2017)

Sample B does not include data from 1957 to 1986, so in the section of monetary incentive, We conducted tests for only three parts except for "Before Free Agency." From here, we mention the three important batting indexes: batting-average, on-base percentage, and homerun.

6.4.1 Bunching

Table 13 shows the results of the McCrary (2008)'s manipulation tests, for each grouped samples. Compared to the full-sample analysis conducted in Section 5.1, we observed partly different results for each index.

Table 13: Manipulation Test for the Grouped Sample by Time

index, cutpoint		'57-'75	'75-'93	'94-2003	2004-	full sample
AVG, .300	bw	.023	.020	.022	.019	.019
	θ	.573	.566	.310	.403	.499
		(.146)	(.120)	(.130)	(.120)	(.067)
	z	3.934***	4.732***	2.393*	3.376***	7.442***
AVG, .250	bw	.028	.028	.032	.027	.024
	θ	.250	.151	.306	.121	.212
		(.080)	(.069)	(.094)	(.0760)	(.042)
	z	3.149**	2.188*	3.242**	1.595	5.061***
OBP, .350	bw	.031	.030	.036	.030	.024
	θ	.137	.149	-.035	.137	.139
		(.089)	(.081)	(.093)	(.082)	(.049)
	z	1.538	1.846	-.380	1.672	2.854**
HR, 20	bw	6.313	6.677	10.165	7.273	5.309
	θ	.222	.214	.145	.315	.259
		(.150)	(.123)	(.129)	(.112)	(.075)
	z	1.479	1.751	1.117	2.819**	3.465***

Note

***: $p < 0.1\%$, **: $p < 1\%$, *: $p < 5\%$.

Bandwidth is optimized following the method of McCrary(2008).

Table 14: RDD for the Grouped Sample by Time

index, cutpoint	bw, type		'87-'94	'95-2003	2004-	full sample
AVG, .300	LATE	bw	.024	.042	.030	.045
		Obs.	697	1806	1872	5930
		estimate	-.034	.064	.066	.034
			(.137)	(.092)	(.103)	(.056)
		z	-.250	.697	.637	.615
AVG, .250	LATE	bw	.036	.043	.075	.048
		Obs.	1482	1806	3991	7271
		estimate	.154	.064	.076	.070
			(.084)	(.092)	(.060)	(.052)
		z	1.825	.697	1.277	1.340
HR, 20	LATE	bw	4.183	3.685	2.46	3.30
		Obs.	341	371	475	1307
		estimate	-.255	-.348	.343	-.002
			(.228)	(.218)	(.264)	(.141)
		z	-1.122	-1.600	1.300	-.015
OBP, .350	LATE	bw	.031	.025	.027	.045
		Obs.	1098	1281	2042	6525
		estimate	.109	-.151	-.030	-.013
			(.106)	(.120)	(.093)	(.049)
		z	1.031	-1.262	-.323	-.272

Note:

***: $p < 0.1\%$, **: $p < 1\%$, *: $p < 5\%$.

Bandwidth is optimized following the method of Imbens-Kalyanaraman.

Table 15: RDD for the Grouped Sample by Time, Only for FA Players

index, cutpoint	bw, type		'87-'94	'95-2003	2004-	full sample
AVG, .300	LATE	bw	.060	.032	.039	.026
		Obs.	218	229	354	509
		estimate	-.026 (.247)	-.309 (.182)	-.186 (.182)	-.253 (.138)
		z	-1.08	-1.700	-1.020	-1.832
AVG, .250	LATE	bw	.018	.023	.078	.058
		Obs.	123	227	716	1367
		estimate	.425 (.281)	.293 (.230)	.047 (.103)	.084 (.082)
		z	1.512	1.272	-.448	1.020
HR, 20	LATE	bw	5.35	3.504	3.566	3.50
		Obs.	47	70	102	206
		estimate	.004 (.284)	-.701 (.492)	-.337 (.513)	-.273 (.296)
		z	-1.600	-1.423	-.657	-.924
OBP, .350	LATE	bw	.034	.042	.031	.043
		Obs.	154	344	395	1044
		estimate	.080 (.291)	-.174 (.179)	.115 (.188)	.021 (.107)
		z	.276	-.971	.616	.196

Note:

***: $p < 0.1\%$, **: $p < 1\%$, *: $p < 5\%$.

Bandwidth is optimized following the method of Imbens-Kalyanaraman.

Table 16: Regression on Log-Salary, Including Interaction Term: around .300

	Dependent variable:							
	OLS				Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	12.194*** (2.026)	11.645*** (1.919)	-1.113 (2.322)	.922 (2.306)	.917 (2.305)			
AVG	5.130 (7.089)	4.726 (6.702)	9.079 (5.977)	.891 (5.839)	.530 (5.841)	-404 (5.837)	5.004 (5.919)	-4.048 (6.589)
AVG_300	-3.292 (3.601)	-2.067 (3.398)	-1.271 (3.023)	-1.856 (2.910)	-2.054 (2.911)	-1.456 (2.906)	-2.129 (2.879)	-1.695 (3.290)
FLD		.0005 (.005)	.001 (.004)	.002 (.004)	.002 (.004)	.002 (.004)	.001 (.005)	.002 (.005)
BsR		.077*** (.019)	.083*** (.017)	.062*** (.017)	.063*** (.017)	.062*** (.017)	.041 (.021)	.054** (.020)
AGE			.699*** (.099)	.690*** (.096)	.695*** (.096)	.692*** (.096)		
AGE_sq			-.010*** (.002)	-.010*** (.002)	-.010*** (.002)	-.010*** (.002)		
AVG:AVG_300	11.428 (11.958)	7.355 (11.288)	4.333 (10.043)	6.289 (9.665)	6.955 (9.670)	5.047 (9.653)	6.977 (9.588)	6.200 (10.927)
WPA				X	X	X	X	X
FA dummy								
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	703	697	697	697	697	697	697	697
R ²	.031	.174	.348	.400	.402	.413	.849	.246
Adjusted R ²	.027	.159	.335	.386	.387	.393	.697	.223
Residual Std. Error	1.026 (df = 699)	.955 (df = 684)	.850 (df = 682)	.816 (df = 680)	.816 (df = 679)	.812 (df = 673)	.573 (df = 348)	.919 (df = 675)
F Statistic	7.539*** (df = 3, 699)	11.988*** (df = 12, 684)	26.031*** (df = 14, 682)	28.385*** (df = 16, 680)	26.867*** (df = 17, 679)			

*p<0.05; **p<0.01; ***p<0.001

Table 17: Regression on Log-Salary, Including Interaction Term: around .300

	Dependent variable:							
	OLS				Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	10.414*** (.717)	9.677*** (.718)	-8.332*** (1.250)	-5.706*** (1.242)	-5.673*** (1.227)			
AVG	13.708*** (2.594)	14.833*** (2.583)	13.943*** (2.280)	4.183 (2.269)	3.672 (2.242)	3.604 (2.238)	.837 (2.182)	3.979 (2.569)
AVG_300	-1.285 (1.696)	-1.084 (1.679)	-.554 (1.482)	-.969 (1.407)	-1.009 (1.390)	-1.241 (1.387)	-.816 (1.226)	-1.631 (1.593)
FLD		.002 (.003)	.005 (.003)	.006* (.003)	.006* (.003)	.006* (.003)	.007* (.003)	.004 (.003)
BsR		.020 (.013)	.040*** (.011)	.031** (.011)	.031** (.011)	.033** (.011)	.029* (.012)	.018 (.013)
AGE			1.142*** (.073)	1.092*** (.069)	1.088*** (.068)	1.107*** (.069)		
AGE_sq			-.017*** (.001)	-.017*** (.001)	-.016*** (.001)	-.017*** (.001)		
AVG:AVG_300	4.259 (5.555)	3.544 (5.497)	1.861 (4.852)	3.258 (4.607)	3.470 (4.551)	4.216 (4.544)	2.874 (4.023)	5.344 (5.218)
WPA				X	X	X	X	X
FA dummy					X	X	X	X
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	1,878	1,867	1,867	1,867	1,867	Position	Individual	Position
R ²	.062	.099	.299	.370	.385	.392	.766	.197
Adjusted R ²	.060	.093	.294	.364	.379	.384	.657	.187
Residual Std. Error	1.166 (df = 1874)	1.148 (df = 1853)	1.014 (df = 1851)	.962 (df = 1849)	.950 (df = 1848)	.947 (df = 1842)	.707 (df = 1271)	1.087 (df = 1844)
F Statistic	41.081*** (df = 3; 1874)	15.715*** (df = 13; 1853)	52.680*** (df = 15; 1851)	63.750*** (df = 17; 1849)	64.330*** (df = 18; 1848)			

*p<0.05; **p<0.01; ***p<0.001

Table 18: Regression on Log-Salary, Including Interaction Term: around .300

	Dependent variable:							
	OLS				Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	14.252*** (1.163)	13.060*** (1.155)	-4.948*** (1.392)	-3.925** (1.415)	-3.852** (1.409)			
AVG	2.257 (4.111)	5.079 (4.067)	7.693* (3.399)	1.969 (3.333)	2.193 (3.320)	1.946 (3.312)	.332 (3.374)	-1.304 (4.010)
AVG_300	-.732 (2.357)	-.034 (2.327)	-1.023 (1.945)	-.166 (1.880)	.129 (1.874)	.275 (1.869)	-1.570 (1.757)	.395 (2.264)
FLD		.0001 (.004)	.006 (.003)	.007* (.003)	.007* (.003)	.008* (.003)	.002 (.004)	.003 (.004)
BsR		-.031*** (.008)	-.0001 (.006)	-.005 (.006)	-.004 (.006)	-.006 (.007)	.013 (.009)	-.037*** (.008)
AGE			1.034*** (.067)	1.026*** (.065)	1.013*** (.065)	1.032*** (.065)		
AGE_sq			-.015*** (.001)	-.015*** (.001)	-.014*** (.001)	-.015*** (.001)		
AVG:AVG_300	3.198 (7.782)	.745 (7.683)	3.577 (6.421)	.711 (6.209)	-.269 (6.187)	-.759 (6.169)	5.210 (5.813)	-.736 (7.473)
WPA				X	X	X	X	X
FA dummy					X	X	X	X
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	1,880	1,872	1,872	1,872	1,872	1,872	1,872	1,872
R ²	.013	.056	.342	.387	.392	.403	.776	.124
Adjusted R ²	.011	.047	.335	.379	.384	.393	.640	.109
Residual Std. Error	1.273 (df = 1876)	1.250 (df = 1853)	1.044 (df = 1851)	1.009 (df = 1849)	1.005 (df = 1848)	.998 (df = 1839)	.769 (df = 1165)	1.209 (df = 1841)
F Statistic	8.204*** (df = 3; 1876)	6.103*** (df = 18; 1853)	48.083*** (df = 20; 1851)	52.968*** (df = 22; 1849)	51.809*** (df = 23; 1848)			

*p<0.05; **p<0.01; ***p<0.001

Table 19: Regression on Log-Salary, Including Interaction Term: around .250

Dependent variable: Sal								
	(1)	(2)	OLS (3)	(4)	(5)	(6)	feim (7)	(8)
Constant	11.622*** (1.047)	11.080*** (1.022)	-1.923 (1.306)	-1.853 (1.315)	-2.264 (1.314)			
AVG	6.412 (4.422)	6.379 (4.298)	4.973 (3.690)	-329 (3.640)	-057 (3.625)	-302 (3.615)	3.633 (3.479)	.208 (4.244)
AVG_250	-743 (1.333)	-662 (1.293)	-601 (1.110)	-790 (1.081)	-677 (1.076)	-681 (1.074)	-281 (1.016)	-872 (1.261)
FLD		.004 (.003)	.003 (.003)	.005* (.003)	.005 (.003)	.005 (.003)	.006* (.003)	.007* (.003)
BsR		.027 (.017)	.031* (.015)	.015 (.014)	.015 (.014)	.012 (.015)	.002 (.017)	.010 (.017)
AGE			.811*** (.067)	.815*** (.065)	.834*** (.065)	.836*** (.065)		
AGE_sq			-.012*** (.001)	-.012*** (.001)	-.012*** (.001)	-.012*** (.001)		
AVG:AVG_250	3.425 (5.390)	3.103 (5.233)	2.981 (4.491)	3.770 (4.371)	3.306 (4.355)	3.323 (4.343)	1.649 (4.120)	4.007 (5.100)
WPA				X	X	X	X	X
FA dummy					X	X	X	X
Season dummies		X	X	X	X	X	X	Position
Fixed effects								
Observations	1,521	1,510	1,510	1,510	1,510	1,510	1,510	
R ²	.038	.112	.347	.383	.389	.396	.770	.166
Adjusted R ²	.036	.105	.340	.377	.382	.387	.642	.154
Residual Std. Error	1.004 (df = 1517)	.969 (df = 1497)	.832 (df = 1495)	.809 (df = 1493)	.805 (df = 1492)	.802 (df = 1486)	.613 (df = 972)	.942 (df = 1488)
F Statistic	19.883*** (df = 3; 1517)	15.715*** (df = 12; 1497)	56.641*** (df = 14; 1495)	57.955*** (df = 16; 1493)	55.828*** (df = 17; 1492)			
Note:	*p<0.05; **p<0.01; ***p<0.001							

*p<0.05; **p<0.01; ***p<0.001

Note:

Table 20: Regression on Log-Salary, Including Interaction Term: around .250

Dependent variable:									
Sal									
	(1)	(2)	OLS (3)	(4)	(5)	(6)	<i>feim</i> (7)	(8)	
Constant	10.767*** (1.070)	10.285*** (1.057)	-7.056*** (1.403)	-5.992*** (1.384)	-6.050*** (1.347)				
AVG	12.553** (4.553)	12.610** (4.498)	10.263** (3.905)	2.621 (3.818)	2.395 (3.717)	1.963 (3.705)	5.558 (3.607)	4.376 (4.415)	
AVG.250	.703 (1.270)	.448 (1.253)	- .148 (1.088)	-.289 (1.049)	-.183 (1.021)	-.322 (1.018)	1.007 (.965)	.359 (1.213)	
FLD		.006 (.003)	.006* (.003)	.007** (.003)	.006* (.003)	.006* (.003)	.007* (.003)	.007* (.003)	
BsR		-.016 (.014)	.011 (.012)	.0003 (.012)	.004 (.012)	.004 (.012)	.019 (.012)	-.020 (.014)	
AGE			1.117*** (.073)	1.119*** (.071)	1.107*** (.069)	1.135*** (.069)			
AGE.sq			-.017*** (.001)	-.017*** (.001)	-.016*** (.001)	-.017*** (.001)			
AVG:AVG.250	-2.668 (5.206)	-1.685 (5.136)	.700 (4.459)	1.339 (4.299)	.980 (4.185)	1.537 (4.171)	-3.777 (3.963)	-1.228 (4.970)	
WPA				X	X	X	X	X	X
FA dummy			X	X	X	X	X	X	X
Season dummies									
Fixed effects									
Observations	1,934	1,927	1,927	1,927	1,927	1,927	1,927	1,927	Position
R ²	.038	.079	.307	.357	.391	.398	.764	.144	
Adjusted R ²	.036	.073	.302	.351	.385	.391	.646	.134	
Residual Sd. Error	1.174 (df = 1930)	1.151 (df = 1913)	.999 (df = 1911)	.963 (df = 1909)	.938 (df = 1908)	.933 (df = 1902)	.712 (df = 1282)	1.112 (df = 1904)	
F Statistic	25.188*** (df = 3; 1930)	12.648*** (df = 13; 1913)	56.445*** (df = 15; 1911)	62.241*** (df = 17; 1909)	67.973*** (df = 18; 1908)				
*p<0.05; **p<0.01; ***p<0.001									

*p<0.05; **p<0.01; ***p<0.001

Table 21: Regression on Log-Salary, Including Interaction Term: around .250

	Dependent variable:						
	Sal						
	(1)	(2)	OLS (3)	(4)	(5)	(6)	fdm (7)
Constant	13.563*** (.550)	12.692*** (.553)	-5.106*** (.846)	-4.164*** (.850)	-4.177*** (.841)		
AVG	3.865 (2.379)	5.769* (2.389)	5.847** (1.971)	-.646 (1.961)	-1.249 (1.941)	-2.190 (1.941)	-3.109 (1.853)
AVG.250	-1.645* (.652)	-1.423* (.645)	-1.371* (.539)	-1.575** (.526)	-1.717*** (.520)	-1.979*** (.518)	-1.574** (.484)
FLD		.001 (.003)	.004 (.002)	.006** (.002)	.006* (.002)	.006** (.002)	.005* (.002)
BsR		-.036*** (.006)	-.006 (.005)	-.009 (.005)	-.008 (.005)	-.010 (.005)	.017** (.006)
AGE			1.068*** (.048)	1.053*** (.046)	1.056*** (.046)	1.072*** (.046)	
AGE_sq			-.016*** (.001)	-.015*** (.001)	-.015*** (.001)	-.015*** (.001)	
AVG:AVG.250	6.660* (2.688)	5.787* (2.659)	5.801** (2.222)	6.738** (2.166)	7.256*** (2.144)	8.315*** (2.135)	6.612*** (1.993)
WPA				X	X	X	X
FA dummy					X	X	X
Season dummies		X	X	X	X	X	X
Fixed effects							
Observations	4,008	3,991	3,991	3,991	3,991	Position 3,991	Individual 3,991
R ²	.039	.079	.358	.394	.407	.418	.733
Adjusted R ²	.038	.075	.355	.391	.404	.413	.644
Residual Std. Error	1.233 (df = 4004)	1.209 (df = 3972)	1.010 (df = 3970)	.981 (df = 3968)	.971 (df = 3967)	.963 (df = 3958)	.750 (df = 2995)
F Statistic	53.573*** (df = 3; 4004)	19,000*** (df = 18; 3972)	110.623*** (df = 20; 3970)	117.455*** (df = 22; 3968)	118.457*** (df = 23; 3967)		

Note:

*p<0.05; **p<0.01; ***p<0.001

Table 22: Regression on Log-Salary, Including Interaction Term: around .350

Dependent variable: Sal								
	(1)	(2)	OLS (3)	(4)	(5)	(6)	fdm (7)	(8)
Constant	9.531*** (1.440)	9.212*** (1.377)	-1.887 (1.643)	-1.753 (1.676)	-1.986 (1.672)			
OBP	11.761** (4.320)	10.736** (4.125)	10.050** (3.743)	4.791 (3.708)	4.747 (3.695)	4.464 (3.695)	8.332* (3.756)	5.986 (4.149)
OBP_350	2.549 (2.347)	.982 (2.246)	1.389 (2.039)	1.382 (1.971)	1.492 (1.964)	1.407 (1.966)	.372 (1.851)	.809 (2.207)
FLD		.006 (.004)	.004 (.003)	.006 (.003)	.006 (.003)	.005 (.003)	.006 (.003)	.007 (.004)
BsR		.046** (.017)	.054*** (.016)	.035* (.015)	.036* (.015)	.033* (.016)	.041* (.018)	.031 (.018)
AGE			.693*** (.077)	.712*** (.074)	.724*** (.074)	.725*** (.074)		
AGE_sq			-.010*** (.001)	-.010*** (.001)	-.010*** (.001)	-.010*** (.001)		
OBP:OBP_350	-7.008 (6.683)	-2.664 (6.393)	-3.751 (5.803)	-3.731 (5.608)	-4.036 (5.589)	-3.781 (5.594)	-1.182 (5.285)	-2.166 (6.279)
WPA				X	X	X	X	X
FA dummy								
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	1,108	1,098	1,098	1,098	1,098	Position 1,098	Individual 1,098	Position 1,098
R ²	.037	.146	.300	.348	.354	.359	.791	.188
Adjusted R ²	.035	.137	.291	.339	.343	.345	.646	.172
Residual Std. Error	1.000 (df = 1104)	.949 (df = 1085)	.860 (df = 1083)	.831 (df = 1081)	.828 (df = 1080)	.827 (df = 1074)	.608 (df = 646)	.929 (df = 1076)
F Statistic	14.308*** (df = 3; 1104)	15.464*** (df = 12; 1085)	33.159*** (df = 14; 1083)	36.120*** (df = 16; 1081)	34.756*** (df = 17; 1080)			
*p<0.05; **p<0.01; ***p<0.001								

*p<0.05; **p<0.01; ***p<0.001

Note:

Table 23: Regression on Log-Salary, Including Interaction Term: around .350

Dependent variable: Sal								
	(1)	(2)	OLS (3)	(4)	(5)	(6)	<i>felm</i> (7)	(8)
Constant	5.281* (2.095)	4.395* (2.066)	-10.795*** (2.189)	-10.617*** (2.185)	-10.635*** (2.151)			
OBP	26.125*** (6.221)	27.294*** (6.122)	21.250*** (5.482)	15.192** (5.390)	15.684** (5.306)	15.472** (5.317)	8.583 (5.469)	22.958*** (6.097)
OBP_350	2.589 (3.080)	3.439 (3.036)	4.426 (2.714)	4.651 (2.627)	4.763 (2.585)	4.693 (2.598)	-1.519 (2.561)	3.863 (2.986)
FLD		.005 (.004)	.009** (.003)	.007* (.003)	.008* (.003)	.008* (.003)	.007 (.004)	.004 (.004)
BsR		-.0002 (.015)	.021 (.014)	.019 (.013)	.020 (.013)	.021 (.014)	.042** (.015)	.003 (.016)
AGE			1.078*** (.084)	1.079*** (.082)	1.055*** (.080)	1.077*** (.081)		
AGE_sq			-.016*** (.001)	-.016*** (.001)	-.016*** (.001)	-.016*** (.001)		
OBP:OBP_350	-8.245 (8.815)	-10.629 (8.688)	-12.975 (7.766)	-13.572 (7.516)	-13.929 (7.398)	-13.714 (7.434)	3.634 (7.347)	-11.835 (8.545)
WPA				X	X	X	X	X
FA dummy								
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	1,285	1,281	1,281	1,281	1,281	Position	Individual	Position
R ²	.031	.084	.269	.317	.339	.342	.781	.129
Adjusted R ²	.029	.074	.261	.308	.330	.329	.622	.113
Residual Std. Error	1.165 (df = 1281)	1.139 (df = 1267)	1.018 (df = 1265)	.985 (df = 1263)	.969 (df = 1262)	.970 (df = 1256)	.728 (df = 743)	1.115 (df = 1258)
F Statistic	13.839*** (df = 3; 1281)	8.922*** (df = 13; 1267)	31.083*** (df = 15; 1265)	34.517*** (df = 17; 1263)	35.987*** (df = 18; 1262)			
*p<0.05; **p<0.01; ***p<0.001								

*p<0.05; **p<0.01; ***p<0.001

Note:

Table 24: Regression on Log-Salary, Including Interaction Term: around .350

	Dependent variable:							
	OLS				Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	13.575*** (1.653)	12.636*** (1.622)	-6.244*** (1.687)	-6.180*** (1.741)	-6.325*** (1.728)			
OBP	3.505 (4.914)	5.128 (4.806)	9.240* (4.004)	5.879 (4.020)	6.451 (3.990)	7.316 (3.961)	6.127 (3.983)	2.449 (4.826)
OBP_350	-2.396 (2.526)	-1.247 (2.468)	.506 (2.056)	.985 (2.027)	1.377 (2.013)	1.736 (1.999)	-2.266 (1.961)	-593 (2.436)
FLD		.010* (.004)	.012*** (.003)	.012*** (.003)	.012*** (.003)	.011*** (.003)	.007* (.003)	.010** (.004)
BsR		-.030*** (.007)	.001 (.006)	.001 (.006)	.002 (.006)	.004 (.007)	.023** (.009)	-.027** (.008)
AGE			1.052*** (.066)	1.047*** (.065)	1.039*** (.064)	1.067*** (.064)		
AGE_sq			-.015*** (.001)	-.015*** (.001)	-.015*** (.001)	-.015*** (.001)		
OBP*OBP_350	7.097 (7.210)	3.920 (7.045)	-1.321 (5.868)	-2.730 (5.787)	-3.892 (5.746)	-4.923 (5.709)	6.108 (5.607)	2.030 (6.954)
WPA				X	X	X	X	X
FA dummy								
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	2,075	2,063	2,063	2,063	2,063	2,063	2,063	2,063
R ²	.013	.077	.361	.381	.391	.404	.772	.113
Adjusted R ²	.011	.069	.355	.375	.384	.395	.639	.100
Residual Std. Error	1.258 (df = 2071)	1.221 (df = 2044)	1.016 (df = 2042)	1.001 (df = 2040)	.993 (df = 2039)	.985 (df = 2030)	.761 (df = 1300)	1.201 (df = 2032)
F Statistic	8.919*** (df = 3, 2071)	9.436*** (df = 18, 2044)	57.728*** (df = 20, 2042)	57.141*** (df = 22, 2040)	56.931*** (df = 23, 2039)			

*p<0.05; **p<0.01; ***p<0.001

Table 25: Regression on Log-Salary, Including Interaction Term: around 20 HR

	Dependent variable:						
	OLS			Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Constant	11.805*** (1.091)	11.894*** (.993)	-.144 (1.697)	.298 (1.819)	.303 (1.824)		
HR	.125* (.063)	.076 (.057)	.040 (.048)	.044 (.048)	.044 (.048)	.035 (.048)	.050 (.061)
HR_20	2.620 (1.547)	.239 (1.409)	-.242 (1.193)	.007 (1.182)	.005 (1.184)	-.376 (1.196)	-.019 (1.384)
FLD		.0004 (.005)	.003 (.004)	.003 (.004)	.003 (.004)	.003 (.004)	.004 (.006)
BsR		.108*** (.031)	.110*** (.026)	.092*** (.026)	.092*** (.027)	.087** (.027)	-.003 (.046)
AGE			.774*** (.100)	.759*** (.100)	.758*** (.100)	.753*** (.101)	
AGE_sq			-.011*** (.002)	-.011*** (.002)	-.011*** (.002)	-.011*** (.002)	
HR:HR_20	-.141 (.080)	-.019 (.073)	.009 (.062)	-.005 (.061)	-.005 (.061)	.014 (.062)	-.010 (.073)
WPA				X	X	X	X
FA dummy		X		X	X	X	X
Season dummies			X				
Fixed effects							
Observations	348	341	341	341	341	341	341
R ²	.015	.261	.475	.490	.490	.500	.830
Adjusted R ²	.006	.234	.452	.465	.463	.464	.617
Residual Std. Error	.948 (df = 344)	.837 (df = 328)	.707 (df = 326)	.699 (df = 324)	.700 (df = 323)	.700 (df = 317)	.591 (df = 151)
F Statistic	1.727 (df = 3; 344)	9.641*** (df = 12; 328)	21.051*** (df = 14; 326)	19.443*** (df = 16; 324)	18.243*** (df = 17; 323)		

*p<0.05; **p<0.01; ***p<0.001

Table 26: Regression on Log-Salary, Including Interaction Term: around 20 HR

	Dependent variable:						
	OLS			Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Constant	13.328*** (1.798)	12.700*** (1.836)	-10.122*** (2.908)	-9.386*** (2.990)	-9.889*** (2.921)		
HR	.071 (.100)	.085 (.102)	.180* (.086)	.182* (.084)	.182* (.082)	.191* (.083)	.132 (.099)
HR.20	.324 (2.340)	.046 (2.400)	1.417 (2.014)	1.419* (1.984)	1.535 (1.937)	1.684 (1.967)	6.165** (2.361)
FLD		.004 (.006)	.010* (.005)	.011* (.005)	.013* (.005)	.013** (.005)	-.002 (.006)
BsR		.035 (.030)	.053* (.025)	.029 (.026)	.022 (.025)	.005 (.026)	-.034 (.033)
AGE			1.323*** (.159)	1.318*** (.156)	1.338*** (.152)	1.328*** (.155)	
AGE_sq			-.020*** (.003)	-.020*** (.003)	-.020*** (.003)	-.020*** (.003)	
HR*HR.20	-.021 (.122)	-.009 (.125)	-.089 (.105)	-.091 (.104)	-.097 (.101)	-.106 (.103)	-.307* (.123)
WPA				X	X	X	X
FA dummy							
Season dummies		X	X	X	X	X	X
Fixed effects							
Observations	374	371	371	371	371	Position 371	Position 371
R ²	.006	.046	.336	.365	.397	.409	.873
Adjusted R ²	-.002	.011	.308	.334	.366	.368	.663
Residual Std. Error	1.071 (df = 370)	1.066 (df = 357)	.892 (df = 355)	.875 (df = 353)	.854 (df = 352)	.852 (df = 346)	1.047 (df = 348)
F Statistic	.744 (df = 3, 370)	1.327 (df = 13, 357)	11.971*** (df = 15, 355)	11.929*** (df = 17, 353)	12.849*** (df = 18, 352)		

*p<0.05; **p<0.01; ***p<0.001

Table 27: Regression on Log-Salary, Including Interaction Term: around 20 HR

	Dependent variable:							
	OLS				Sal			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	19.239*** (3.381)	18.618*** (3.382)	-1.330 (3.119)	-1.373 (3.159)	-2.020 (3.139)			
HR	-229 (.184)	-198 (.184)	-182 (.135)	-141 (.133)	-123 (.132)	-125 (.132)	-103 (.190)	-151 (.181)
HR_20	-5.545 (3.871)	-4.512 (3.877)	-4.699 (2.840)	-4.220 (2.798)	-3.460 (2.785)	-3.979 (2.788)	-3.839 (3.955)	-4.799 (3.840)
FLD		.005 (.008)	.014* (.006)	.013* (.006)	.013* (.006)	.012* (.006)	-.0005 (.010)	.002 (.008)
BsR		-.042** (.016)	.002 (.012)	-.007 (.012)	-.008 (.012)	-.016 (.014)	-.006 (.025)	-.062*** (.018)
AGE			1.144*** (.123)	1.182*** (.121)	1.195*** (.120)	1.200*** (.121)		
AGE_sq			-.016*** (.002)	-.017*** (.002)	-.017*** (.002)	-.017*** (.002)		
HR:HR_20	.301 (.204)	.248 (.205)	.254 (.150)	.224 (.148)	.187 (.147)	.211 (.147)	.204 (.209)	.253 (.203)
WPA				X	X	X	X	X
FA dummy								
Season dummies		X	X	X	X	X	X	X
Fixed effects								
Observations	475	475	475	475	475	475	475	475
R ²	.009	.052	.494	.513	.522	.535	.894	.112
Adjusted R ²	.003	.015	.472	.489	.498	.501	.665	.052
Residual Std. Error	1.247 (df = 471)	1.240 (df = 456)	.908 (df = 454)	.893 (df = 452)	.885 (df = 451)	.882 (df = 442)	.723 (df = 150)	1.216 (df = 444)
F Statistic	1.491 (df = 3; 471)	1.388 (df = 18; 456)	22.182*** (df = 20; 454)	21.625*** (df = 22; 452)	21.436*** (df = 23; 451)			

Note: *p<0.05; **p<0.01; ***p<0.001

6.4.2 Monetary Incentive

6.5 Across Indexes: Priority of Manipulation

7 Conclusion

This paper have considered monetary incentives that exists behind the behavior that appears to be related to reference dependent preference. For the case of MLB, we conclude

Also, we may pay attention to the oppsite-side approach: whether there exists discontinuous design of the contracts that does not affect the players manipulation behavior.

8 Appendix

8.1 Proxies for Performance

To measure batter's performance, I use some *SABR Metrics's* statistics. *SABR* is short for *Society of American Baseball Research*, an American organization of analysing baseball. They try to evaluate players in more "efficient " way, or by statistics with more close correlation with the expected wins the team obtain.

8.2 Important Events Related to Section 6.4

9 References

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