Exercise 4

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1 Group

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```
[]: import numpy as np
import matplotlib.pyplot as plt
import pandas as pd
from matplotlib.widgets import Slider

plt.rcParams["xtick.labelsize"] = 15
plt.rcParams["ytick.labelsize"] = 15
```

2 Task A

2.0.1
$$p(x|\theta) = \pi_1 \frac{1}{\sqrt{2\pi\sigma_{c_1}^2}} exp(-\frac{1}{2\sigma_{c_1}^2}(x-\mu_{c_1})^2) + \pi_2 \frac{1}{\sqrt{2\pi\sigma_{c_2}^2}} exp(-\frac{1}{2\sigma_{c_2}^2}(x-\mu_{c_2})^2)$$

```
[]: # Generate samples for two combined gaussian distributions.

# We want to have random samples. Therefore draw N numbers of uniform

distribution and if the

# value is below pi_1 (=pi[0]) we draw a value from the first Gaussian and

above from the second Gaussian

def gen_samples(N, pi: list, mu: list, sigma: list):

assert all([len(1) == 2 for 1 in [pi, mu, sigma]]), "Lists need to have the

clength 2."

assert round(sum(pi), 6) == 1.0, "Sum of pi_i must equal 1"

randoms = np.random.random(N)

samples = []
```

```
for r in randoms:
    if r <= pi[0]:
        samples.append(np.random.normal(mu[0], sigma[0], 1))
    else:
        samples.append(np.random.normal(mu[1], sigma[1], 1))

return np.array(samples).flatten()</pre>
```

```
[]: # Generative parameters
N = 200

mu1_gen = -2
pi1_gen = 0.4
sigma1_gen = 1

mu2_gen = 4
pi2_gen = 0.6
sigma2_gen = np.sqrt(10)
```

```
[]: samples = gen_samples(N, [pi1_gen, pi2_gen], [mu1_gen, mu2_gen], [sigma1_gen, u -sigma2_gen])
```

3 Task B

```
[]: mu1_init = -5
    pi1_init = 0.5
    sigma1_init = 1
    mu2_init = 5
    pi2_init = 0.5
    sigma2_init = np.sqrt(10)
```

```
[]: sum([pi1_init, pi2_init]) == 1.0
```

[]: True

```
assert round(sum(pi), 6) == 1.0, "Sum of pi_i must equal 1"

prob = 0
for p, m, s in zip(pi, mu, sigma):
    prob += p * gaussian_pdf(x, m, s)

return prob
```

```
[]: class MSTEPError(Exception):
         pass
     class EMAlgorithm:
         11 11 11
         Class that implements the EM algorithm for the 1 dimensional case.
         This can be applied for gaussian mixture models that consists of K_{\sqcup}
      \hookrightarrow different gaussians.
         parameters:
         sample func: Function to sample the data (generative model)
         N: Number of data points
         K_gaussians: Number of mixture components
         initial_mu: Initial mean value
         initial_sigma: Initial value for the standard deviation
         conv_thr: Convergence threshold for the log-likelihood
         max_steps: maximum number of iterations
         HHHH
         def __init__(self, samples, K_gaussians, initial_mu, initial_sigma,
                       initial_pi, conv_thr=0.000001, max_steps=100):
             self.K_gaussians = K_gaussians
             self.x = samples
             self.sigma = initial_sigma
             self.mu = initial_mu
             self.N= self.x.shape[0]
             self.r = None
             self.conv_thr = conv_thr
             self.max_steps = max_steps
             self.pi = np.array(initial_pi)
         def _expectation_step(self):
```

```
Performs one expectation step. Formulas see script p. 29, Headline "EM」
\hookrightarrow - Algorithm for the mixture of Gaussians model"
       denominator = self._calc_r_denomin()
       self.r = [gaussian pdf(self.x, self.mu[i], self.sigma[i]) * self.pi[i] /
→ denominator
                 for i in range(self.K_gaussians)]
   def calc r denomin(self):
       Calculate denominator for the update rule of r.
       denom = np.array([gaussian_pdf(self.x, self.mu[i], self.sigma[i]) *__
→self.pi[i]
                          for i in range(self.K_gaussians)]).sum(axis=0)
       return denom
   def _maximization_step(self):
       Performs one maximization step. Formulas see script top of p. 30
       HHHH
       self.pi = [1 / self.N * self.r[i].sum()
                    for i in range(self.K_gaussians)]
       self.mu = [(self.x * self.r[i]).sum() / self.r[i].sum()
                    for i in range(self.K_gaussians)]
       self.sigma = np.sqrt([(((self.x - self.mu[i]) ** 2) * self.r[i]).sum() /

→ self.r[i].sum()
                    for i in range(self.K_gaussians)])
       if np.any(self.r == 0.0) or np.any(self.mu == 0.0) or np.any(self.sigma_{\perp}
\Rightarrow== 0.0) or np.any(self.pi == 0.0):
           raise MSTEPError
   def _calculate_gaussian_mixture(self):
       11 11 11
```

```
Calculate the probability of x for each gaussian in the gaussian \Box
\hookrightarrow mixture.
       Recall script p. 26 formula (95). For one x_i in the vector x we
⇒calculate the likelihood given
       all Gaussians c of the total K_{gaussians}. Therefore first each column _{\sqcup}
→of the matrix (np.array[]) corresponds
       to one x_1 and each row in this column is the probability density u
\hookrightarrow p(x,c/theta) or each c.
       After that we take the sum along axis=0 / axis=rows. Thus, we add all_{\sqcup}
⇒rows on top of each other which means
       sum p(x,c|theta) other all c. This is now p.
       For the log-likelihood take the log of p and the sum other the vector, \Box
\hookrightarrow thus sum other all x_i in x. Which is done
       in function _calculate_log_likelihood.
       p = np.array([self.pi[i] * gaussian_pdf(self.x, self.mu[i], self.
→sigma[i])
                      for i in range(self.K gaussians)]).sum(axis=0)
       return p
   def _calculate_log_likelihood(self):
       Calculate the log-likelihood. p is a vector with p(x_i|theta) for all u
\hookrightarrow x_i in vector x.
       Take log of p(x_i|theta) and sum other all x_i.
       p = self._calculate_gaussian_mixture()
       log_likelihood = np.log(p).sum()
       return log_likelihood
   def run(self, print_=True, return_=True, conv=False):
       11 11 11
       Run the EM algorithm.
       parameters:
       print_: Whether to print the log-likelihood in each step
       return_: Whether to return the list of log-likelihoods
```

```
conv: Whether to consider the convergence of the log-likelihood to stop_{\sqcup}
\hookrightarrow the iterations. If False,
             only the maximum number of steps is considered.
       converged = False
       step = 0
       log_likelihood = self._calculate_log_likelihood()
       log_likelihood_list = [log_likelihood]
       mu_parameters = [self.mu]
       sigma_parameters = [self.sigma]
       pi_parameters = [self.pi]
       if print_:
               print(f"step {step}: {log_likelihood}")
       # The actual algorithm
       while not converged and step < self.max_steps:</pre>
           self._expectation_step()
           self._maximization_step()
           prev_log_likelihood = log_likelihood
           log_likelihood = self._calculate_log_likelihood()
           """ Save parameters in array or each iteration """
           log_likelihood_list.append(log_likelihood)
           mu_parameters.append(self.mu)
           sigma_parameters.append(self.sigma)
           pi_parameters.append(self.pi)
           step += 1
           if print_:
               print(f"step {step}: {log_likelihood}")
           """ Abort early if additional convergence criterium is met even_{\sqcup}
→ though max_steps were not yet performed """
           if np.abs(log_likelihood - prev_log_likelihood) < self.conv_thr and__
converged = True
       if return :
           return dict(mu= mu_parameters, sigma=sigma_parameters, __
→pi=pi_parameters, log_likelihood=log_likelihood_list)
   def return_final_parameters(self):
```

```
Return the final parameters for the gaussian mixture.
             return {"pi": self.pi, "mu": self.mu, "sigma": self.sigma}
[]: EM = EMAlgorithm(samples, 2, [mu1_init, mu2_init], [sigma1_init, sigma2_init],
      →[pi1_init, pi2_init])
[ ]: parameters = EM.run(print_=False)
[]: %matplotlib widget
     # run this cell before you run the cell containing the Slider (below). If you
      →run all cells, then the slider is already closed again since the next cell
      →runs "matplotlib inline".
[]: # Create interactive widget. You can use the slider to show the pdf after each
     \rightarrowEM iteration
     # Furthermore the correct pdf is in orange and the underlying histogram of \Box
     \rightarrow datapoints for the
     # EM algorithm is shown.
     fig, ax = plt.subplots(figsize=(6, 8))
     x = np.linspace(-10, 15, 1000)
     mu = parameters["mu"]
     sigma = parameters["sigma"]
     pi = parameters["pi"]
     ax.hist(samples, bins=51, density=True, alpha=0.7, label="Histogram of $x n$")
     ax.plot(x, pdf_combined(x, pi=[0.4, 0.6], mu=[-2, 4], sigma=[1, np.sqrt(10)]),_{\sqcup}
     →label="Correct PDF", c="C0")
     ax.set_xlabel("Datapoint $x_n$", fontsize=15)
     ax.set_ylabel("Probability", fontsize=15)
     line, = ax.plot(x, pdf_combined(x, pi[0], mu[0], sigma[0]), label="Approximated_u
     →PDF", C="C1")
     ax.legend()
     # Creation of adjustable Slider
     ax sld = plt.axes([0.25, 0.001, 0.65, 0.03])
     SLD = Slider(
         ax=ax_sld,
         label="Iteration number",
         valmin=0,
         valmax=len(mu)-1,
```

```
valinit=0,
valfmt="%i"
)

def update(val):
    line.set_ydata(pdf_combined(x, pi[int(val)], mu[int(val)], sigma[int(val)]))
    fig.canvas.draw_idle()

SLD.on_changed(update)

Canvas(toolbar=Toolbar(toolitems=[('Home', 'Reset original view', 'home',
    'home'), ('Back', 'Back to previous ...

<ipython-input-62-a348d90784b9>:17: MatplotlibDeprecationWarning: Case-insensitive properties were deprecated in 3.3 and support will be removed two minor releases later
```

[]: 0

3.1 The Slider figure cannot be shown in the pdf. Please use the attached Jupyter Notebook

line, = ax.plot(x, pdf_combined(x, pi[0], mu[0], sigma[0]),

label="Approximated PDF", C="C1")

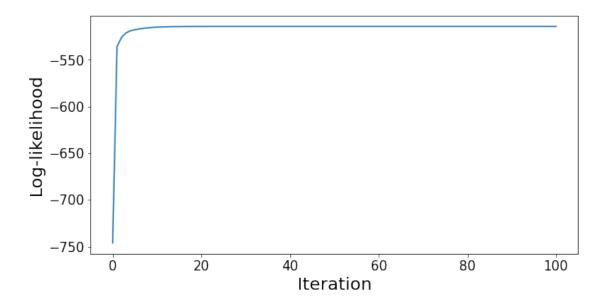
3.1.1 You can see that the pdf of the EM algorithm approaches the correct pdf with each iteration step. However, due to the limited N=200 datapoints the pdf won't match the correct pdf exactly. The underlying histogram does not represent the correct pdf exactly and therefore we get small deviations from the EM-pdf to the correct pdf

```
[]: | %matplotlib inline
```

```
[]: # Plot the evolution of the log_likelihood
log_likelihood = parameters['log_likelihood']

fig, ax = plt.subplots(figsize=(10, 5))

ax.plot(log_likelihood)
ax.set_xlabel("Iteration", fontsize=20)
ax.set_ylabel("Log-likelihood", fontsize=20)
plt.show()
```



```
[]: # Compare parameters of EM algorithm to generative parameters
    final_parameters_task_b = EM.return_final_parameters()
    # Round final parameters
    pi_rounded = [f'%.3f' % elem for elem in final_parameters_task_b["pi"]]
    mu rounded = [f'%.3f' % elem for elem in final_parameters_task_b["mu"]]
    sigma_rounded = [f'%.3f' % elem for elem in final_parameters_task_b["sigma"]]
    comparison = []
    comparison.append({"Generative parameters": str([pi1_gen, pi2_gen]), __
     →"EM-algorithm parameters": str(pi_rounded), "label": "[Pi_1, Pi_2]"})
    comparison.append({"Generative parameters": str([mu1 gen, mu2 gen]),
     → "EM-algorithm parameters": str(mu_rounded), "label": "[Mu_1, Mu_2]"})
    comparison.append({"Generative parameters": str([sigma1_gen,__
     →round(sigma2_gen,3)]), "EM-algorithm parameters": str(sigma_rounded),
     df = pd.DataFrame(comparison).set_index("label")
    display(df)
```

Generative parameters EM-algorithm parameters

3.1.2 We can see that the parameters by the EM-algorithm are quite close to the generative parameters

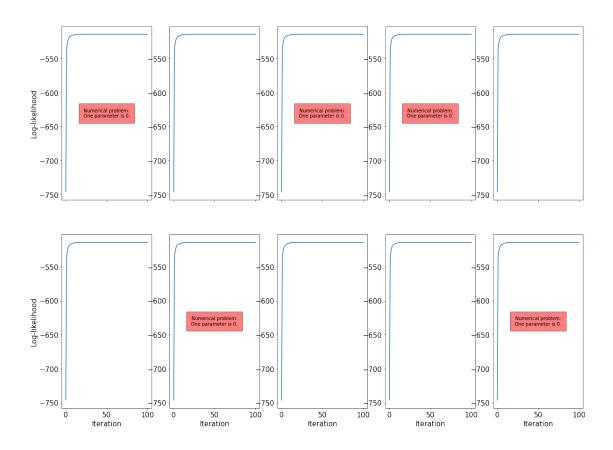
4 Task C

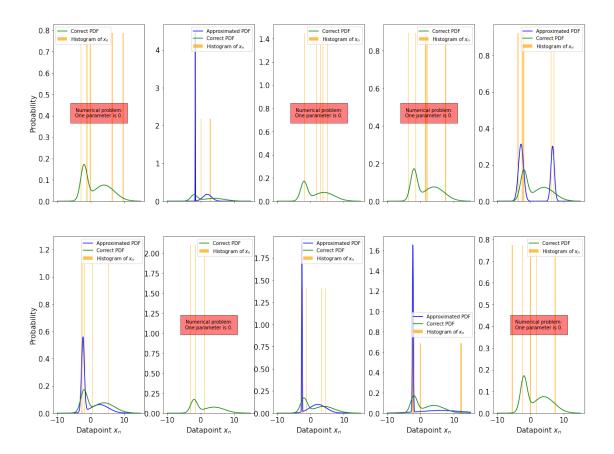
- 4.0.1 With so few datapoints we run into problems to compute the parameters of the gaussian distribution. When we run into problems, we don't (and cannot) plot the curves.
- 4.0.2 This can be explained as follows. If we take a look at the formulas on p. 30 top and especially $\mu_c^{new} = \frac{\sum_n r_c^{(n)} x^{(n)}}{\sum_n r_c^{(n)}}$ and $(\sigma_c^{new})^2 = \frac{\sum_n r_c^{(n)} (x^{(n)} \mu_c^{new})^2}{\sum_n r_c^{(n)}}$
- 4.0.3 After some Debugging we found, that some σ_c are 0 and this results in a problem because σ is in the gaussian distribution in the denominator of the exponential term. Thus, we would devide by 0.
- 4.0.4 The question is: Why do we have $sigma_c = 0$? According to the formulas this happens if $r_c^{(n)}$ is 0 or $x^{(n)} \mu_c^{new}$ is 0 for all c.
- 4.0.5 1. Take a look at the formula of $r_c^{(n)}$. If the x_i value is too far away from the intial gaussian guess, the nominator gets 0 and therefore $r_c^{(n)}$ is 0. This happens because our 5 datapoints are randomly generated. It can happen, that these datapoints are far far away from the intial guess and the exponential function yields very small numbers (~1e-30).
- 4.0.6 2. This gets problematic if we compute $\mu_c^{(n)}$. Assume only 4 out of 5 datapoints yields to very small $r_c^{(n)}$. Then the summation "picks out" only one $x^{(n)}$ namely x_i . So $\mu_c^{(new)} \approx x_i$
- 4.0.7 3. In the formula for $(\sigma_c^{(new)})^2$ we take the difference of $\mu_c^{(new)}$ and the $x^{(n)}$ vector and get approximately $\sigma_c^{(new)} \approx 0$. The other terms are approximately also 0 because $r_c^{(n)}$ is nearly 0. And now we divide by 0 (in one of the next iterations if sigma gets even smaller) in the exponential function.
- 4.0.8 In summary: If we have few datapoints and many of them lie far away from the initial gaussian guess. The $\mu_c^{(new)}$ and $r_c^{(n)}$ just get the right value that $\sigma_c^{(new)}$ is 0.

```
EM = EMAlgorithm(samples, 2, [mu1_init, mu2_init], [sigma1_init,__

⇒sigma2_init], [pi1_init, pi2_init])
   try:
       parameters = EM.run(print_=False)
   except MSTEPError:
       ax_log[i // 5, i % 5].text(0.5, 0.5, "Numerical problem:\nOne parameter_u
\hookrightarrowis 0.",
                                   bbox={'facecolor': 'red', 'alpha': 0.5, __
\rightarrow'pad': 10},
                                   transform=ax_log[i // 5, i % 5].transAxes,__
⇔ha="center", va="center")
       ax_pdf[i // 5, i % 5].text(0.5, 0.5, "Numerical problem:\nOne parameter_
\hookrightarrowis 0.".
                                   bbox={'facecolor': 'red', 'alpha': 0.5,_
\rightarrow 'pad': 10},
                                   transform=ax_pdf[i // 5, i % 5].transAxes, u
approximated = False
   final_parameters = EM.return_final_parameters()
   pi = final_parameters["pi"]
   mu = final parameters["mu"]
   sigma = final_parameters["sigma"]
   if i // 5 == 1:
       ax_log[i // 5, i % 5].set_xlabel("Iteration", fontsize=15)
       ax_pdf[i // 5, i % 5].set_xlabel("Datapoint $x_n$", fontsize=15)
   if i % 5 == 0:
       ax_log[i // 5, i % 5].set_ylabel("Log-likelihood", fontsize=15)
       ax_pdf[i // 5, i % 5].set_ylabel("Probability", fontsize=15)
   if approximated:
       ax_pdf[i // 5, i % 5].plot(x, pdf_combined(x, pi, mu, sigma),_
→label="Approximated PDF", color="blue")
   ax_pdf[i // 5, i \% 5].plot(x, pdf_combined(x, pi=[0.4, 0.6], mu=[-2, 4],_u

→sigma=[1, np.sqrt(10)]), label="Correct PDF", color="green")
   ax_pdf[i // 5, i % 5].hist(samples, bins=51, density=True, alpha=0.7, u
→label="Histogram of $x_n$", color="orange")
   ax_log[i // 5, i % 5].plot(log_likelihood)
   ax_pdf[i // 5, i % 5].legend()
```





5 Task D

→10)))

```
sigma_start = list(zip(np.random.uniform(0.2, 3, 10), np.random.uniform(0.2, 6, u=10)))
pi_start = np.random.uniform(size=10)
pi_start = list(zip(pi_start, 1 - pi_start))

[]: N = 200

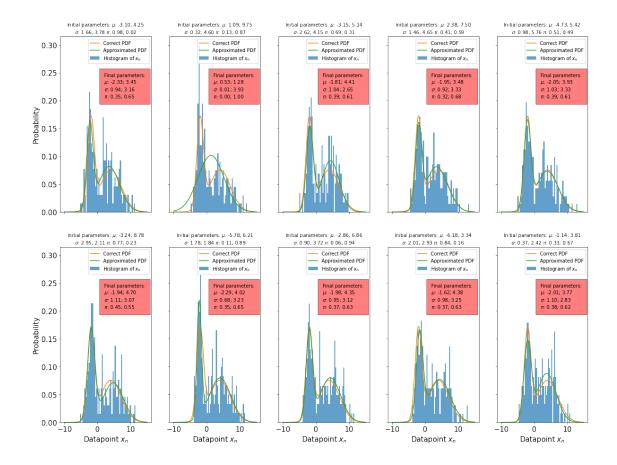
#fig_log, ax_log = plt.subplots(ncols=5, nrows=2, figsize=(20, 15), u=sharex=True, sharey=True)
fig_pdf, ax_pdf = plt.subplots(ncols=5, nrows=2, figsize=(20, 15), sharex=True, u=sharey=True)

for i in range(10):
```

[]: mu_start = list(zip(np.random.uniform(-8, 3, 10), np.random.uniform(3.1, 10,__

```
samples = gen_samples(N, [pi1_gen, pi2_gen], [mu1_gen, mu2_gen],__
→[sigma1_gen, sigma2_gen])
  EM = EMAlgorithm(samples, 2, mu_start[i], sigma_start[i], pi_start[i])
  parameters = EM.run(print_=False)
  final parameters = EM.return final parameters()
  pi = final parameters["pi"]
  mu = final_parameters["mu"]
  sigma = final_parameters["sigma"]
  title = "Initial parameters: $\mu$: {:.2f}; {:.2f} \n$\sigma$: {:.2f}; {:.
→*pi_start[i])
  ax_pdf[i // 5, i % 5].set_title(title, fontsize=10)
  results = "Final parameters:\n$\mu$: {:.2f}; {:.2f}\n$\sigma$: {:.2f}; {:.
\rightarrow 2f\n\pi\: {:.2f}; {:.2f}".format(*mu, *sigma, *pi)
   ax pdf[i // 5, i % 5].text(0.48, 0.65, results, bbox={'facecolor': 'red', __
→ 'alpha': 0.5, 'pad': 10}, transform=ax pdf[i // 5, i % 5].transAxes)
  #ax log[i // 5, i % 5].plot(log likelihood)
  if i // 5 == 1:
       #ax log[i // 5, i \% 5].set xlabel("Iteration", fontsize=15)
       ax_pdf[i // 5, i % 5].set_xlabel("Datapoint $x_n$", fontsize=15)
   if i % 5 == 0:
       \#ax_log[i // 5, i \% 5].set_ylabel("Loglikelyhood", fontsize=15)
      ax_pdf[i // 5, i % 5].set_ylabel("Probability", fontsize=15)
  ax_pdf[i // 5, i % 5].hist(samples, bins=51, density=True, alpha=0.7,__
→label="Histogram of $x_n$")
   ax_pdf[i // 5, i \% 5].plot(x, pdf_combined(x, pi=[0.4, 0.6], mu=[-2, 4],__

sigma=[1, np.sqrt(10)]), label="Correct PDF")
   ax_pdf[i // 5, i % 5].plot(x, pdf_combined(x, pi, mu, sigma),__
→label="Approximated PDF")
  ax_pdf[i // 5, i % 5].legend()
```



5.1 Overall, the different approximations look quite similiar (except for subplot 2). Even in the case of initial parameters that differ significantly from the ones that were used for the data generation, the final parameters approximate the generative parameters quite well. Moreover, the second subplot emphasizes that you can reach local minima which lead to incorrect approximations. In this case pi_2 is 0 and therefore the result consists only of one gaussian.

6 Task E

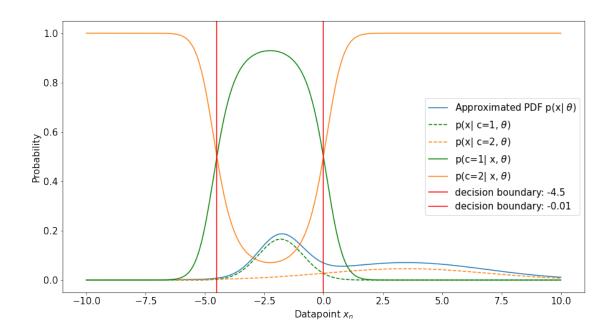
```
[]: def intersect(p1, p2):
    idxs = np.argwhere(np.diff(np.sign(p1 - p2))).flatten()
    return idxs

[]: final_parameters_task_b

[]: {'pi': [0.394425798858007, 0.6055742011419932],
    'mu': [-1.8085265673912105, 3.4694114471905064],
    'sigma': array([0.94977382, 3.41271131])}
```

```
[]: x = np.linspace(-10, 10, 10000)
     p_1 = gaussian_pdf(x, final_parameters_task_b["mu"][0],
     →final_parameters_task_b["sigma"][0]) * final_parameters_task_b["pi"][0]
     p_2 = gaussian_pdf(x, final_parameters_task_b["mu"][1],__
     →final_parameters_task_b["sigma"][1]) * final_parameters_task_b["pi"][0]
     posterior_1 = p_1 / (p_1 + p_2)
     posterior_2 = p_2 / (p_1 + p_2)
     dec_bound = intersect(posterior_1, posterior_2)
[]: dec bound
[]: array([2751, 4995], dtype=int64)
[]: fig, ax = plt.subplots(figsize=(15, 8))
     ax.plot(x, pdf_combined(x, **final_parameters_task_b), label=r"Approximated PDF_u
     \rightarrow p(x| $\theta)")
     ax.plot(x, p_1, linestyle="--", label=r"p(x| c=1, $\theta$)", color="green")
     ax.plot(x, p_2, linestyle="--", label=r"p(x| c=2, $\theta)", color="C1")
     ax.plot(x, posterior_1, linestyle="-", label=r"p(c=1| x, $\theta$)",__
     ax.plot(x, posterior_2, linestyle="-", label=r"p(c=2| x, $\theta$)", color="C1")
     ax.axvline(x[dec_bound[0]], color="red", label=f"decision boundary:__
     \rightarrow {round(x[dec_bound[0]], 2)}")
     ax.axvline(x[dec_bound[1]], color="red", label=f"decision boundary:u
     \rightarrow {round(x[dec_bound[1]], 2)}")
     ax.legend(fontsize=15)
     ax.set_xlabel("Datapoint $x_n$", fontsize=15)
     ax.set_ylabel("Probability", fontsize=15)
```

[]: Text(0, 0.5, 'Probability')



6.1 Based on the lecture we know that the decision boundaries are equal to the intersection points of the posterior probabilities. (eq. 67)