

RH: Missing data and topology in total evidence approach

## **Effect of missing data on topological inference using a total evidence approach**

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## Abstract

Living species represent a marginal part of all species that have ever lived. Ignoring fossil taxa may lead to misinterpretation of macroevolutionary patterns and processes such as trends in species richness, biogeographical history or paleoecology. This fact has led to an increasing consensus among scientists that both living and fossil taxa must be included in macroevolutionary studies. One approach called Total Evidence, uses molecular data from living taxa and morphological data from both living and fossil taxa to infer phylogenies with both living and fossil taxa at the tips. Although the Total Evidence approach seems very promising, it requires a lot of data and is therefore likely to suffer from missing data issues which may affect its ability to infer correct phylogenies.

In this study we assess the effect of missing data on tree topologies inferred from Total Evidence matrices. Using simulations we investigate three major factors that directly affect the completeness of the morphological part of the matrix: (1) the proportion of living taxa with no morphological data, (2) the amount of missing data in the fossil record, and (3) the overall number of morphological characters in the matrix.

We find that, when using a clade conservative metric such as Robinson-Foulds distance, Bayesian consensus trees recovers the right topology better than Maximum Likelihood or than a subset of Bayesian posterior tree distributions and this regardless of the amount of missing data (minimal tree similarity of 0.69 in the worst scenario). Regarding the data, good topological recovery is not related to the amount of missing data *per se* but to the amount of data overlap. The two factors that influences the most this data overlap are the parameters 1 (proportion of coded living taxa) and 3 (number of morphological characters).

(Keywords: missing data, Total Evidence, Bayesian, Maximum Likelihood, topology)

## INTRODUCTION

Although most species that have ever lived are now extinct (Novacek and Wheeler 1992; Raup 1993), the majority of macroevolutionary studies focus solely on living species (e.g. Meredith et al. 2011; Jetz et al. 2012). Ignoring fossil taxa may lead to misinterpretation of macroevolutionary patterns and processes such as the timing of diversification events (e.g. Pyron 2011), relationships among lineages (e.g. Manos et al. 2007) or niche occupancy (e.g. Pearman et al. 2008). This has led to increasing consensus among scientists that fossil taxa must be included in macroevolutionary studies (Jackson and Erwin 2006; Quental and Marshall 2010; Dietl and Flessa 2011; Slater and Harmon 2013; Fritz et al. 2013). However, to do this we need to be able to place living and fossil taxa into the same phylogenies; a task that remains difficult despite recent methodological developments (e.g. Pyron 2011; Ronquist et al. 2012a; Matzke 2014).

Up to now, three main approaches have been used to place both living and fossil taxa into phylogenies. These approaches differ mainly in whether they treat fossil taxa as tips or as nodes in the phylogeny, and in which part of the available fossil data is used (i.e. the age of the fossil only or both its age and morphology). Classical cladistic methods use matrices containing morphological data from both living and fossil taxa and treat each taxon as a tip in the phylogeny. Relationships among the taxa are then inferred using optimality criteria such as maximum parsimony (Simpson 1945). This approach is commonly used by paleontologists but it ignores the additional molecular data available from living species and does not allow use of probabilistic methods for dealing with phylogenetic uncertainty. Neontologists, on the other hand, more commonly use probabilistic approaches (e.g. Maximum Likelihood or Bayesian methods) based on matrices containing only molecular data from living species.

Because fossil taxa do not usually have available DNA, fossils are used as nodes rather than tips in these phylogenies and their occurrence date are used to time calibrate phylogenies (Zuckerkandl and Pauling 1965). There have been great improvements in the theory and application of these two approaches (e.g. Bapst 2013; Stadler and Yang 2013; Heath et al. 2013) as well as much debate about the "best" approach to use (e.g. Spencer and Wilberg 2013; Wright and Hillis 2014). However neither approach uses all the available data.

A final approach, known as the Total Evidence method, uses matrices containing molecular data from living taxa and morphological data from both living and fossil taxa (Eernisse and Kluge 1993). This approach treats every taxon as a tip in the phylogeny, uses the occurrence age of the fossils to time calibrate the phylogeny (tip-dating method), and allows the use of probabilistic methods for estimating phylogenetic uncertainty (Ronquist et al. 2012a). Total Evidence methods have been successfully applied to empirical data (e.g. Pyron 2011; Ronquist et al. 2012a; Schrago et al. 2013; Slater 2013; Beck and Lee 2014), and are becoming an increasingly popular way of adding fossil taxa to phylogenies. However, although the Total Evidence approach seems very promising, there is one big drawback in using this approach: it requires a lot of data that can be difficult (or impossible) to collect. The morphological data for living taxa is rarely collected when molecular data is available (e.g. O'Leary et al. 2013 *vs.* Meredith et al. 2011), and, for fossil taxa, the scarcity of the fossil record only allow to collect the data available (for example, in vertebrates, the hardest parts of the skeleton; Sansom and Wills 2013). Therefore Total Evidence matrices are likely to contain a lot of missing data that may affect the method's ability to infer correct topologies, branch lengths and support values (Salamin et al. 2003).

The effect of missing data on phylogenetic inferences has been widely studied (Wiens 2003, 2006; Wiens and Moen 2008; Lemmon et al. 2009; Roure and Philippe 2011;

Sansom and Wills 2013; Pattinson et al. 2014; Sansom 2014). Missing molecular data has been seen by some authors as an issue because it can, in some part of the tree, decrease phylogenetic signal (i.e. the evolutionary information contained within the matrix allowing to infer topology and branch length), especially when using large matrices (Lemmon et al. 2009). However, this may not be a major issue because phylogenetic signal is easily increased by: (i) including a "modest" number ( 50) of highly-covered genes (i.e. a number of genes that are available for all taxa ; Roure and Philippe 2011); (ii) adding a greater number of taxa (especially slowly-evolving taxa or taxa close to the outgroup; Roure and Philippe 2011); and (iii) choosing more appropriate models of sequence evolution (Wiens 2006; Wiens and Moen 2008; Roure and Philippe 2011). Similarly, missing morphological data might be seen as either a major or minor issue for accurately inferring phylogenies depending on the study in question (Wiens 2003; Sansom and Wills 2013; Pattinson et al. 2014). Because soft-tissue characters are rarely preserved in the fossil record, missing data is mainly found in these characters, and is therefore not randomly distributed which can lead to biased placement of fossil taxa in phylogenies (e.g. Sansom and Wills 2013 but see Pattinson et al. 2014). However, the phylogenetic signal is not related to the amount of missing data *per se* but to the number of informative characters for each taxon, therefore missing data is less of an issue than the number of shared informative characters (Kearney 2002; Wiens 2003; Pattinson et al. 2014).

Although missing data does not appear be a major problem in molecular and morphological matrices separately (Wiens 2003, 2006; Wiens and Moen 2008; Roure and Philippe 2011; Pattinson et al. 2014), it may become more of an issue in Total Evidence matrices containing both molecular and morphological data for living and fossil species. This may be particularly problematic as fossil taxa (generally) do not have molecular data, resulting in a large section of missing data. Until now, no attempt has

been made to study the impact of this issue on phylogenetic inference from Total Evidence methods. In this study, we are interested in assessing the effect of missing data on the ability to recover a correct topology. Even though branch length is a second aspect of phylogenetic inferences equally important, topology is the first and most straightforward aspect reflecting phylogenetic signal (i.e. topological changes are discrete opposed to branch length changes are continuous). Also, interestingly, the effect of Total Evidence method has not been formally assessed in previous studies using fixed topology (Ronquist et al. 2012a; Schrago et al. 2013; Slater 2013; Beck and Lee 2014).

Here we use a simulation approach to assess the effect of missing data on tree topologies inferred from Total Evidence matrices. Since the molecular part of a Total Evidence matrix acts like a "classical" molecular matrix containing only the living taxa (Ronquist et al. 2012a), the effect of missing data on such matrices is well known (Wiens 2006; Wiens and Moen 2008; Roure and Philippe 2011). Therefore, we focus only on missing data in the morphological part of the matrix. We investigate three major parameters that directly affect the completeness of the morphological part of the matrix: (i) the proportion of living taxa with no morphological data; (ii) the proportion of missing data in the fossil taxa; and (iii) the proportion of missing morphological characters for both living and fossil taxa in the matrix. We remove data from a Total Evidence matrix by changing the values of these three parameters and then assess how this affects the topology of trees inferred using Maximum Likelihood and Bayesian methods. We chose these parameters because they reflect empirical biases in data availability. The advent of molecular phylogenetics means that morphological data for living species is rarely collected, and few people have the skills to identify characters needed for detailed phylogenetic analysis. Missing data in fossil taxa is very common due to preservation biases (Sansom and Wills 2013), and the overall number of

characters depends on the effort of the people identifying them (e.g. O'Leary et al. 2013). We find that the ability of recovering the correct topology can mainly be increased when using the Bayesian methods. Also minimizing the proportion of living taxa with no morphological data and the proportion of missing morphological characters improves the ability to recover the correct more than minimizing the proportion of missing data in the fossil record.

## METHODS

To explore how missing data in the morphological sections of Total Evidence matrices influences tree topology, we used the following protocol (note that we explain each step in detail below this general outline; Fig. 1).

### 1. Generating the matrix:

We randomly generated a birth-death tree (hereafter called the "true" tree) and used it to infer a matrix containing both molecular and morphological data for living and fossil taxa (hereafter called the "complete" matrix).

### 2. Removing data:

We removed data from the morphological part of the "complete" matrix to simulate the effects of missing data by modifying three parameters (i) the proportion of living taxa with no morphological data ( $M_L$ ), (ii) the proportion of missing data in the fossil taxa ( $M_F$ ) and (iii) the proportion of missing morphological characters ( $M_C$ ). The resulting matrices are called hereafter "missing-data" matrices.

### 3. Inferring phylogenies:

We inferred phylogenetic trees from the "complete" matrix and from the "missing-data" matrices resulting in one tree generated from a matrix containing no missing data (hereafter called the "best" tree) and multiple trees inferred from matrices with missing morphological data (hereafter called the "missing-data" trees). Phylogenies were inferred via both Maximum Likelihood and Bayesian approaches.

### 4. Comparing topologies:

We compared the "best" tree to the "missing-data" trees to assess the influence of each parameter ( $M_L$ ,  $M_F$ ,  $M_C$ ) and their interactions on the topologies of our phylogenies

We repeated the 4 steps 50 times to account for variation of our random parameters in the simulations.

### *Generating the matrix*

First we randomly generated a "true" tree of 50 taxa in R v3.0.2 (R Core Team 2014) using the package *diversitree* v0.9-6 (FitzJohn 2012). We generated the tree using a birth death process by sampling speciation ( $\lambda$ ) and extinction ( $\mu$ ) rates from a uniform distribution but maintaining  $\lambda > \mu$  (Paradis 2011). We implemented a rejection sampling algorithm to select only trees with 25 living and 25 fossil taxa to ensure that we had enough taxa of each type for our missing data simulations to work. We then added an outgroup to the tree, using the mean branch length of the tree to separate the outgroup from the rest of the taxa, and with the branch length leading to the outgroup set as the sum of the mean branch length and the longest root-to-tip length of the tree.

Next, we generated a molecular and a morphological matrix from the "true" tree. The molecular matrix was inferred from the "true" tree using the R package *phyclust* v0.1-14 (Chen 2011). The matrix contained 1000 character sites for 51 taxa and was generated using the seqgen algorithm (Rambaut and Grassly 1997) and using the HKY model (Hasegawa et al. 1985) with random base frequencies and transition/transversion rate of 2 (Douady et al. 2003). The substitution rates were distributed following a gamma distribution with an alpha ( $\alpha$ ) shape of 0.5 (Yang 1996). We chose a low value of  $\alpha$  to reduce the number of sites with high substitution rates, thus avoiding too much homoplasy and a decrease in phylogenetic signal. We selected

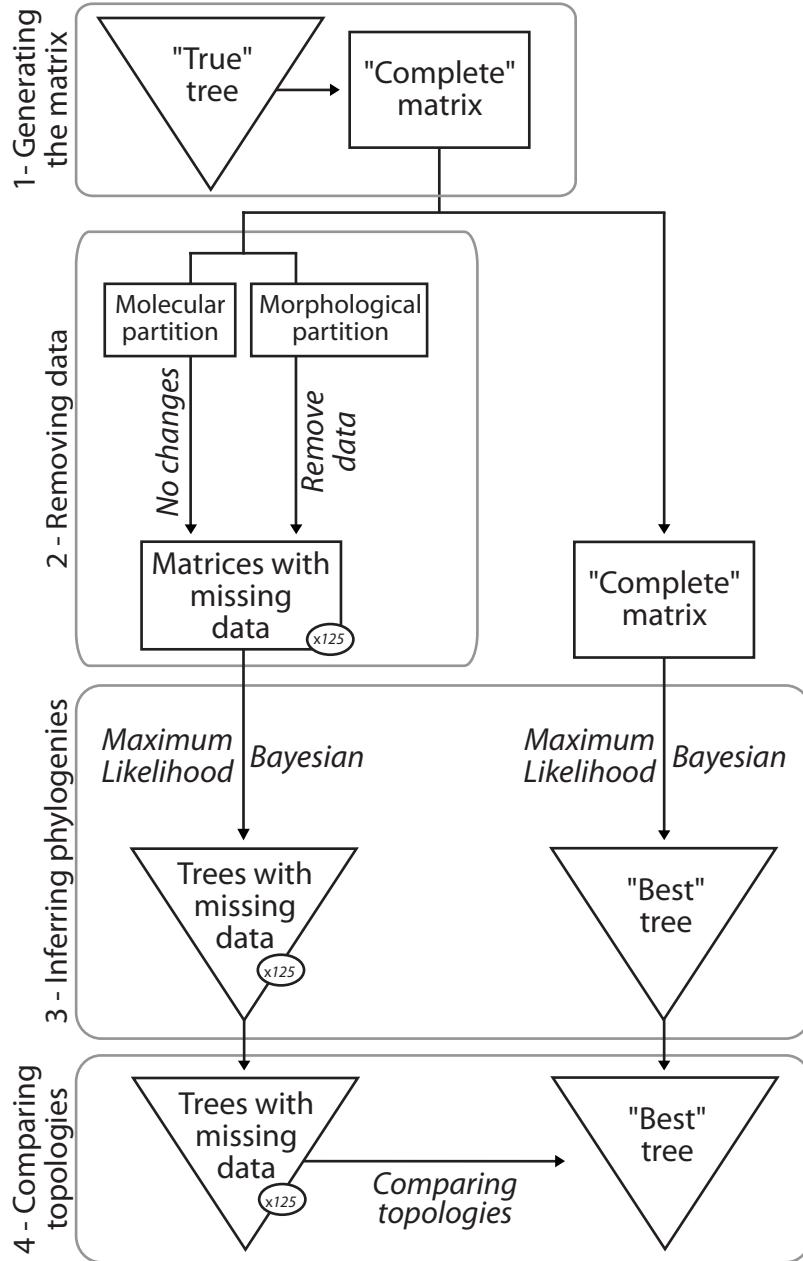


Figure 1: Protocol outline. (1) We randomly generated a birth-death tree (the "true" tree) and used it to infer a matrix with no missing data (the "complete" matrix). (2) We removed data from the morphological part of the "complete" matrix resulting in 125 "missing-data" matrices. (3) We built phylogenetic trees from each matrix using both Maximum Likelihood and Bayesian methods. (4) We compared the "missing-data" trees to the "best" tree. We repeated the 4 steps 50 times.

the parameters above to generate data with no special assumption about how the characters evolved, and to reduce the computational time required if these parameters were estimated rather than defined in the tree building part of the analysis (even with the parameters defined, total computational time for the whole analysis was around 150 CPU years). All the molecular information for fossil taxa was replaced by missing data ("?").

We inferred the morphological matrix using the R package ape v3.0-11 (Paradis et al. 2004) to generate a matrix of 100 character sites for 51 taxa. We assigned the number of character states (either two or three) for each morphological character by sampling with a probability of 0.85 for two states characters and 0.15 for three state characters. These probabilities were selected using the overall distribution of character states extracted from 100 published empirical morphological matrices (See Supplementary Material Section 1). We then ran an independent discrete character simulation for each character using the "true" tree with the character's randomly selected number of states (two or three) and assuming an equal rate of change (i.e. evolutionary rate) from one character state to an other (Pagel 1994). This method allows us to have only two parameters for each character: the number of states and the evolutionary rate. For each character, the evolutionary rate was sampled from a gamma distribution with  $\alpha = 0.5$ . We used low evolutionary rate parameters to avoid homoplasy in the morphological part of the matrix and create a clear phylogenetic signal (Wagner 2000; Davalos et al. 2014; Wright and Hillis 2014).

Finally, we combined the morphological and molecular matrices obtained from the "true" tree. Hereafter we call this the "complete" matrix: the matrix with no missing data except for the molecular data of the fossil taxa.

### *Removing data*

We modified the "complete" matrix to get matrices with missing data by randomly replacing data with "?" in the morphological part of the matrices according to the following parameters:

1. The proportion of living taxa with no morphological data ( $M_L$ ): 0%, 10%, 25%, 50% or 75%. This parameter illustrates the number of living taxa that are present in the molecular part of the matrix but not in the morphological part. This reflects the fact that because of the increasing availability of DNA sequences for living taxa, detailed morphological data is scarce.
2. The proportion of missing data in the fossil taxa ( $M_F$ ): 0%, 10%, 25%, 50% or 75%. This parameter illustrates the quality of the fossil record.
3. the proportion of missing morphological characters for both living and fossil taxa ( $M_C$ ): 0%, 10%, 25%, 50% or 75%. This parameter illustrates the number of available morphological characters for both living and fossil taxa.

In practice, each parameter represents a different way of removing data from the matrix:  $M_L$  removes rows from the living taxa's data;  $M_F$  removes cells from the fossil taxa's data; and  $M_C$  removes columns across both living and fossil taxa's data. Note that  $M_L$  is different to  $M_F$  not only because of the region of the matrix affected: for  $M_L$ , all the morphological data of a percentage of living taxa is removed, but for  $M_F$ , a percentage of the data is removed at random from across the whole of the morphological matrix for fossil taxa.

We tested all parameters combinations resulting in 125 ( $5^3$ ) matrices. Note that one of these combinations has no missing data so is equivalent to the "complete" matrix, thus we have one effectively complete matrix in our 125 "missing-data" matrices. Because some parameter combinations introduce a lot of missing data (e.g.

$M_L=75\%$ ,  $M_F=75\%$  and  $M_C=75\%$ ), some matrices contained fossil taxa without any data at all. When this occurred we repeated the random deletion of characters until every taxa had at least 5% data across the whole morphological part of the matrix.

### *Building phylogenies*

From the resulting matrices we generated two types of trees, the "best" tree inferred from the "complete" matrix and the "missing-data" trees inferred from the 125 matrices with various amounts of missing data. The "true" tree was used to generate the "complete" matrix and reflects the "true" evolutionary history in our simulations. The "best" tree, on the other hand, is the best tree we can build using state-of-the-art phylogenetic methods. In real world situations, the "true" tree is never available to us because we cannot know the true evolutionary history of a clade (except in very rare circumstances, e.g. Rozen et al. 2005). Therefore, here we focus on comparing the trees inferred from the matrices with missing data to the "best" tree, rather than the "true" tree, as the "best" tree is generally what biologists have to work with.

*Maximum Likelihood.*— The "best" tree and the "missing-data" trees were inferred using RAxML v8.0.20 (Stamatakis 2014). For the molecular data, we used the GTR +  $\Gamma_4$  model (Tavaré 1986; default GTRGAMMA in RAxML v8.0.20; Stamatakis 2014) for the molecular data. For the morphological data, we used the implemented Markov  $k$  state model (Lewis 2001) assuming an equal state frequency and a unique overall substitution rate ( $\mu$ ) following a gamma distribution of the rate variation with four distinct categories (Mk +  $\Gamma_4$ ; -K MK option in RAxML v8.0.20; Stamatakis 2014).

In order to measure the phylogenetic signal of our simulations, we first ran a fast bootstrap analysis with 500 replicates on the "complete" matrix. We removed all the simulations that had a median bootstrap support lower than 50 as a proxy for weak

phylogenetic signal (Zander 2004). We repeated this selection until we obtained 50 sets of simulations (i.e. 50 "complete" and  $50 \times 125$  "missing-data" matrices) with a relative good phylogenetic signal (median bootstrap  $> 50$ ).

On these selected simulations, we used the fast bootstrap algorithm and performed 1000 bootstraps per tree inference to assess the topological support (Pattengale et al. 2010). When using these parameters, it took 8 CPU years to build 50 sets of 125 bootstrapped Maximum Likelihood trees (2.30GHz clock speed nodes).

*Bayesian.*— The "best" tree and the "missing-data" trees were inferred using MrBayes v3.2.1 (Ronquist et al. 2012b). We partitioned the data to treat the molecular part as a non-codon DNA partition and the morphological part as a multi-state morphological partition. The molecular evolutionary history was inferred using the HKY model with a transition/transversion ratio of 2 (Douady et al. 2003) and a gamma distribution for the rate variation with four distinct categories (HKY +  $\Gamma_4$ ). For the morphological data, we used the Markov  $k$  state model (Lewis 2001), with equal state frequency and a unique overall substitution rate ( $\mu$ ) with four distinct rates categories (Mk +  $\Gamma_4$ ). We chose these models to be consistent with the parameters used to generate the "complete" matrix.

Each Bayesian tree was estimated using two runs of four chains each for a maximum of  $50 \times 10^6$  generations. We used the average standard deviation of split frequencies (ASDS) as a proxy to estimate the convergence of the chains and used a stop rule when the ASDS went below 0.01 (Ronquist et al. 2012b). The effective sample size (ESS) was also checked on a random sub-sample of runs in each simulation to ensure that  $ESS >> 200$  (Drummond et al. 2006). For each run, we removed 25% of the iterations as burn-in. We used the following priors for each tree (see Supplementary Material S1): (i) the "true" trees topology as a starting tree (with a starting value for each branch length of 1), (ii) an exponential prior on the shape of the gamma

distribution of  $\alpha = 0.5$  for both partitions, and (iii) a transition/transversion ratio prior of two sampled from a strong beta distribution ( $\beta(80,40)$ ).

We used these prior to speed up the Bayesian estimation process. These priors biased the way the Bayesian process calculated branch lengths by giving non-random starting points and boundaries for parameter estimation, however, here we are focusing on the effect of missing data on tree topology and not branch lengths. Even using these priors, it took 140 CPU years to build 50 sets of 125 Bayesian trees (2.30GHz clock speed nodes).

### *Comparing topologies*

We compared the topology of the "missing-data" trees to the "best" tree to measure the effect of the three parameters  $M_L$ ,  $M_F$  and  $M_C$  on tree topology. We used the Robinson-Foulds distance (Robinson and Foulds 1981) to assess the amount conserved clade positions and the Triplets distance (Dobson 1975) to assess amount of wildcard taxa (Kearney 2002). We then used Normalized Tree Similarity index (Bogdanowicz et al. 2012) to generalize our results for any  $n$  number of taxa. These metrics are described in detail below.

*Robinson-Foulds distance.*— Robinson-Foulds distance (Robinson and Foulds 1981), or "path difference", measures the number of shared clades across two trees. The metric reflects the distance between the distributions of tips among clades in the two trees (Robinson and Foulds 1981 ; see Supplementary Material S2). This metric is bounded between 1 when the two trees are identical and  $n - 2$  (for two trees with  $n$  taxa) when there is not one single shared clade between both trees. This metric is sensitive to the exact clade conservation: if the trees are composed of two clades of three taxa  $((((a,b),c),((d,e),f)))$ , the swap of any two taxa will lead to a maximal score of the Robinson-Foulds distance indicating a bad tree similarity.

*Triplets distance.*— The Triplets distance (Dobson 1975) measures the number of sub-trees made up of three taxa that differ between two given trees (Critchlow et al. 1996 ; see Supplementary Material S2). This metric measures the position of each taxon and clade towards it's closest neighbours. It is bounded between 0 when the two trees are identical and  $\binom{n}{4}$  (for two trees with  $n$  taxa) when there is not one single position of taxon/clade identical between both trees. Therefore this metric sensitive to the conservation of wildcard taxas.

*Normalized Tree Similarity.*— We used the Normalized Tree Similarity index,  $NTS_m$  (Bogdanowicz et al. 2012) to be able to compare the two metrics for any  $n$  taxa. This index allows to scale the value of any metric  $m$  (either Robinson-Foulds or Triplets distance in our study) to the expected value of the metric  $m$  when comparing two random trees (see Supplementary Material S2). When  $NTS_m=1$ , the two trees are strictly identical, when  $NTS_m=0$  the trees are no more different than expected when comparing two random trees and when  $NTS_m<0$ , the difference between the two trees is greater than when comparing two random trees. In our study we used the  $NTS_m$  index as a proxy for topology: a high score of this index (i.e. towards 1) means that the topology is highly conserved between the two trees; on the opposite, a low score of this index (i.e. towards 0) means that the topological difference between the two trees is as much as expected when comparing two random trees. Note that both tree comparisons are representing different aspects of topological similarity and do not perform equaly (see discussion below and Kuhner and Yamato 2014).

*Tree comparisons.*— For the Maximum Likelihood and Bayesian consensus trees we performed pairwise comparisons between the "best" tree and each "missing-data" tree using both the Robinson-Foulds and Triplets metrics with the TreeCmp java script (Bogdanowicz et al. 2012). For each metric, we then normalized the value using the

Normalized Tree Similarity scaled by the mean value of 1000 pairwise random tree comparisons for the metric in question and  $n = 51$  taxa (see Supplementary Material Section 2).

Also, to take into account the uncertainty of tree inference in both Maximum Likelihood and Bayesian (i.e node support), we ran 1000 random pairwise comparison between respectively the bootstrapped trees from the Maximum Likelihood analysis and the posterior tree distribution of the Bayesian analysis. In the same way that we compared a single "missing-data" tree to the "best" tree (whether the trees are Maximum Likelihood or Bayesian consensus): we randomly selected 1000 trees from the "missing-data" tree sets (either the Bootstrapped trees or the posterior tree distribution) and did a pairwise comparison with 1000 randomly selected trees from the "best" tree set.

For each of the 125 "missing-data" tree, we obtained Robinson-Foulds and Triplets distance distributions from either 50 or  $50 \times 1000$  pairwise comparisons. We then calculated the mode and the 50% and 95% confidence intervals from the resulting distribution using the `hdrcde` R package v3.1 (Hyndman et al. 2013).

In order to investigate the effect of the parameters and/or the methods used in this simulations, we measured the similarity among the different distributions using the Bhattacharyya Coefficient (Bhattacharyya 1943). The Bhattacharyya Coefficient is the probability of overlap between two distributions, ranging from 0 to 1 (Bhattacharyya 1943 - see Supplementary Material Section 2). For each method (Maximum likelihood trees, Bayesian consensus trees, Bootstraps and Bayesian posterior trees) and metric (Robinson-Foulds and Triplets distance), we used this coefficient in two ways in order to:

1. Assess the effect of the method on all parameters:

We performed a pairwise comparison of each parameter between each pair of methods. This resulted in 125 comparisons per method and per metric (each parameter combination within one method compared to the same parameter combination in the other method). We then compared the distribution of these pairwise comparisons to see the global difference between two methods: if the methods have really similar results, we would expect the distribution to be cluster around 1 (high probability of overlap between the distributions) and if the methods are really dissimilar around 0 (low probability of overlap - Fig. 4).

## 2. Assess the effect of the parameters within a method:

We performed a pairwise comparison per method and per metric for each combination of parameters. This resulted in 7875 pairwise comparisons per method and per metric (triangle of a  $125 \times 125$  matrix). We represented these results as a triangular matrix with the values of each pairwise comparison coloured according to the value of the Batthacharyya Coefficient (coloured in green when the distributions overlap completely and in red when they don't overlap at all- Fig. 5).

# RESULTS

## *Effect of missing data on topology*

As expected from the literature, the amount of missing data in the morphological part of the total evidence matrix does decrease the ability to recover the right topology, regardless of the parameter, the method or the metric used (e.g. Roure and Philippe 2011; Sansom and Wills 2013; Pattinson et al. 2014). However each variable does not

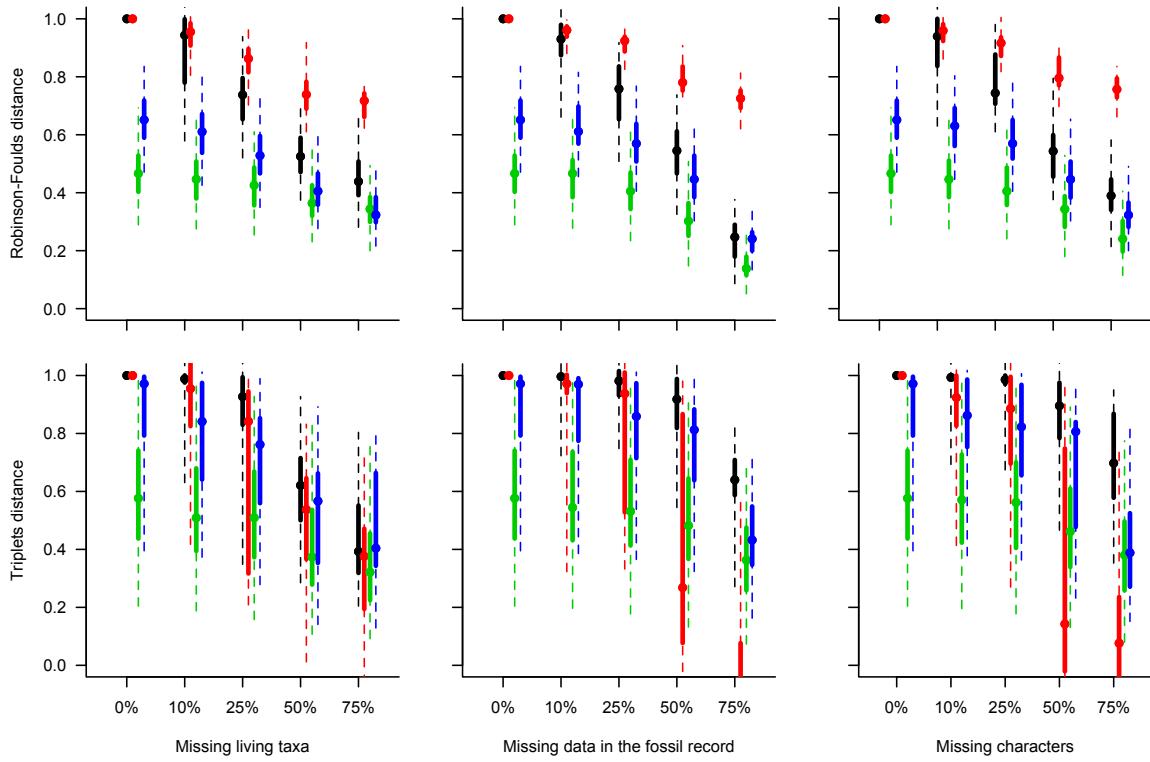


Figure 2: Comparison between the effect of missing data and the tree inference method on topology. The amount of missing data for each parameter is represented on the x axis. The topology is represented on the y axis, both using Robinson-Foulds distance (upper row) and Triplets distance (lower row). Points represent the modal value of each distribution ; thick solid and thin dashed lines represents respectively the 50% and 95% confidence intervals or the distributions. The Maximum Likelihood trees are represented in black, the Bayesian consensus trees in red, the bootstrap trees in green and the posterior tree distribution in blue.

affect the topology in the same way (Fig. 2). Regarding the conservation of entire clades (i.e. the Ronbinson-Foulds distance), the Bayesian consensus trees outperform the other methods and the amount of missing data in the living taxa ( $M_L$ ) decreases the most rapidly the clade conservation when using this method. However, when looking at the position of wildcard taxa (i.e. using the Triplets distance), the Maximum Likelihood outperforms the other methods but with wider distributions overlap (Fig. 2).

When looking at the effect of all parameters combinations on the Bayesian consensus trees, this method outperforms Maximum Likelihood for the Ronbinson-Foulds distance and plateaus to a minimal modal tree similarity of 0.69 regardless the amount of missing data (Fig. 3 - see also Supplementary Material Section 3). However, regarding the Triplets distance, Bayesian consensus trees seems to perform poorly but with great confidence intervals overlaps (Fig. 2 and Fig. 3). These larger confidence intervals are due to the lower ability of the Triplet method to highlight precise differences in topology (Kuhner and Yamato 2014). Results for the global trend comparison between the uncertainty methods (Bootstrap and Bayesian posterior tree distribution) are available in the supplementary materials (see Supplementary Material Section 3).

### *Effect of the method on topology*

When looking at the comparisons between the methods using the Robinson-Foulds distance, the Bayesian consensus trees have the least overlap with respectively the Bayesian posterior trees and the Bootstrap (Fig. 4-A). When regarding at the Triplets distance, the two method that have the most distribution are the Bootstraps and the Bayesian posterior trees (Fig. 4-B). These results are due to the difference in comparing topologies between Maximum Likelihood and Bayesian consensus trees and Bootstraps and Bayesian posterior trees: the two last ones are

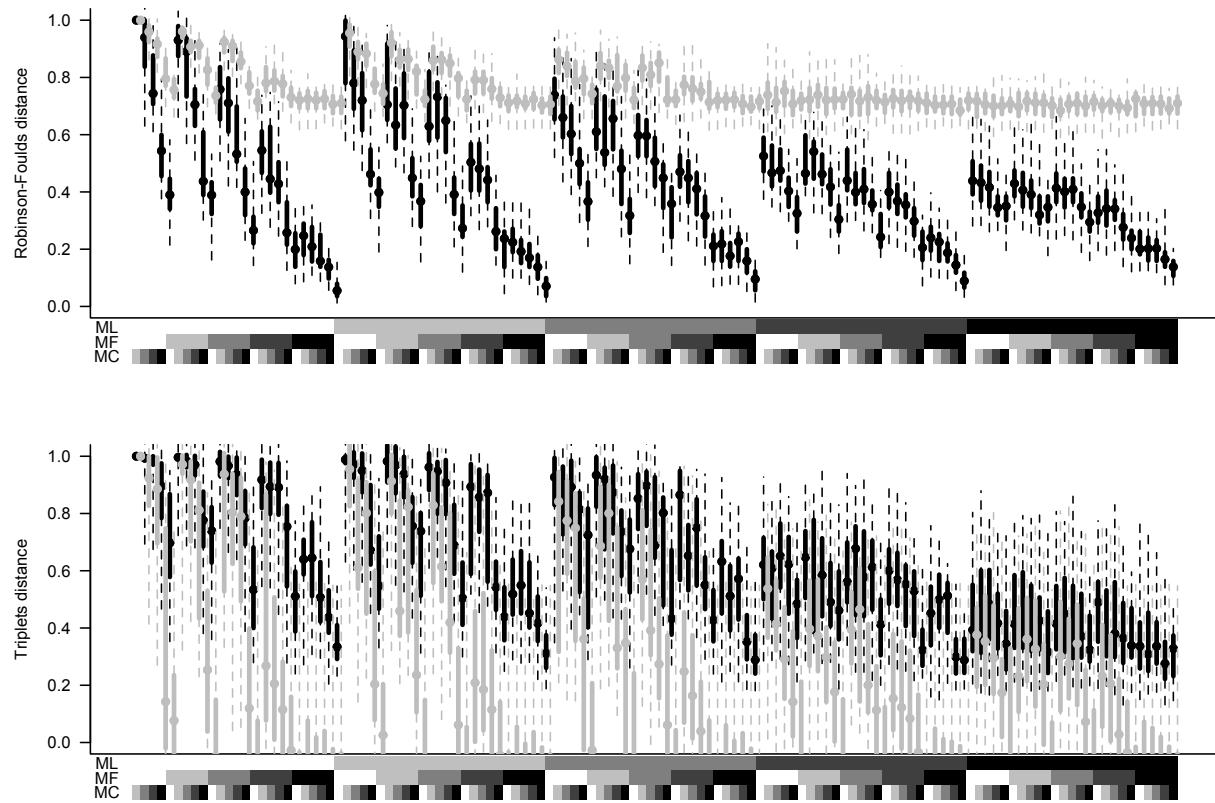


Figure 3: Trend of the effect of missing data on topology on the Maximum Likelihood and the Bayesian consensus trees. The amount of missing data per parameter ( $M_L$ ,  $M_F$  and  $M_C$ ) is represented along the x axis. The colour gradient from white to black represents respectively, 0%, 10%, 25%, 50% and 75% of missing data for each parameter. The topology is represented on the y axis, both using Robinson-Foulds distance (upper row) and Triplets distance (lower row). Points represent the modal value of each distribution ; thick solid and thin dashed lines represents respectively the 50% and 95% confidence intervals or the distributions. The Maximum Likelihood trees are represented in black and the Bayesian consensus trees in grey.

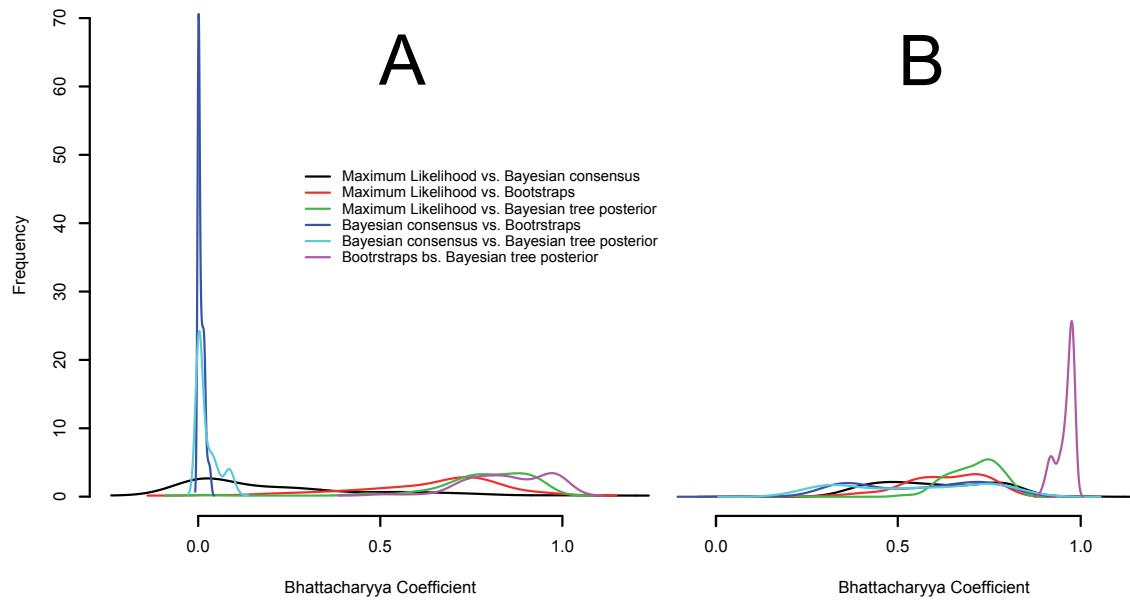


Figure 4: Distribution of the Bhattacharyya Coefficients between methods. Curves represents the kernel density estimations of the Bhattacharyya Coefficients between the four different methods. A. Results for the Normalised Robinson-Foulds distance. B. Results with the for Normalised Triplets distance. Pikes around the two different extremes values of the Bhattacharyya Coefficients show the two pairs of most dissimilar methods (Bayesian consensus vs. Bootstraps and Bayesian consensus vs. Bayesian posterior trees) for the Normalised Robinson-Foulds distance ; and the most similar methods (Bootstraps and Bayesian posterior trees) for the Normalised Triples methods.

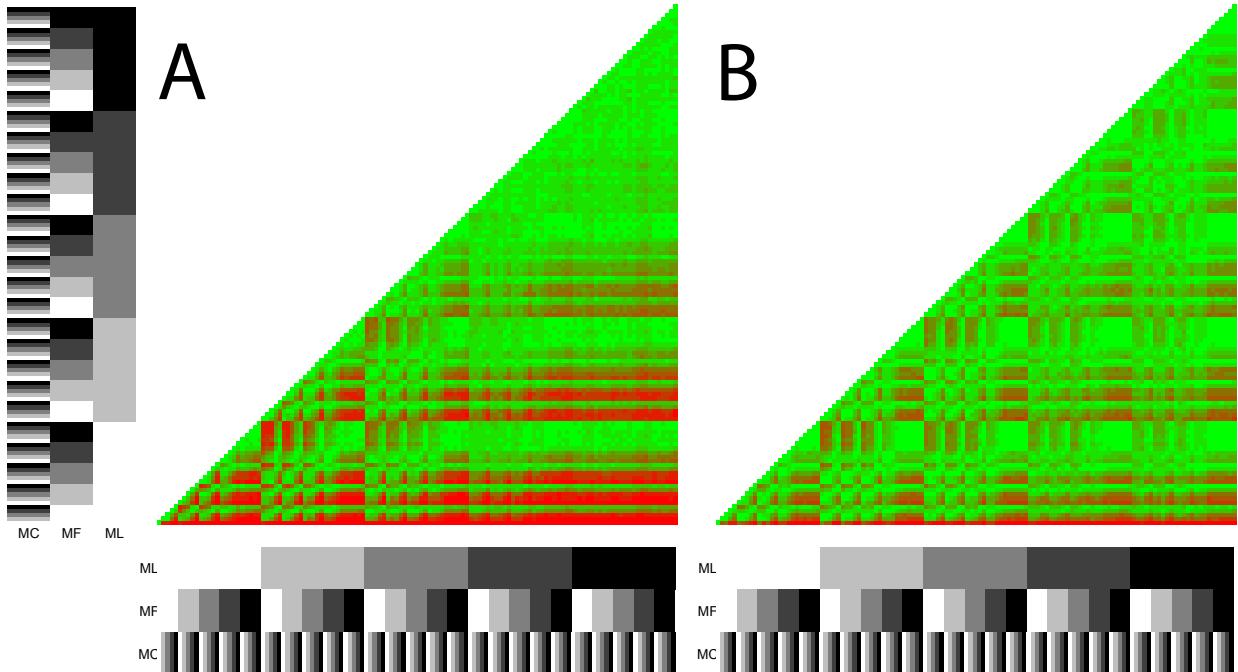


Figure 5: Pairwise Bhattacharyya Coefficients within the Bayesian consensus trees. The pairwise trees comparisons are represented on both axis. The colour gradient from white to black represents respectively, 0%, 10%, 25%, 50% and 75% of missing data for each parameter. The matrix represents the values of pairwise Bhattacharyya Coefficients going from green (1) to red (0). A. Results for the Normalised Robinson-Foulds distance. B. Results with the for Normalised Triplets distance.

based on a random pairwise comparison of a sub-sample of 1000 trees. This is likely to add noise to the data.

### *Effect of the missing data parameters on topology*

When looking at the pairwise Bhattacharyya Coefficients between all the parameters combinations for Bayesian consensus trees, the number of missing living taxa ( $M_L$ ) results the lowest probabilities of overlap (right lower corner in Fig. 5-A). However, when looking at the Triplets distance, because of the high overlap in

confidence intervals, there is no major region with low probability of overlap. One should note however that the bottom line of the matrix where the Bhattacharyya Coefficients are very low is due to comparing the distribution of the pairwise comparison of the "best" tree against the "missing-data" tree with no missing data ( $M_L=0\%$ ,  $M_C=0\%$ ,  $M_F=0\%$ ) , leading to a normalised Triplets score of 1 every time (with no variance, see Fig. 2 and Fig. 3). The pairwise comparisons for the other methods (Maximum Likelihood, Bootstraps and Bayesian posterior trees) are available in the supplementary materials (see Supplementary Material Section 3).

## DISCUSSION

The results of our simulations are in adequacy with recent analysis on the effect of missing data in morphological matrices (Pattinson et al. 2014; Wright and Hillis 2014; Sansom 2014). Our results are showing that the ability to recover the correct topology decreases with the increase of missing data, regardless the metric, the method or the parameters considered. However, within a method and a metric, the parameters used to remove data have different effects on the ability to recover the correct topology. Additionally, the effect of missing data on topological recovery varies with different methods to infer topology and different metrics to compare topologies. Thus, we found that using the topologies of Bayesian consensus trees performed consistently better than other methods for recovering the correct topology (Fig. 3 and Fig. 5) and that both the amount of available characters ( $M_C$ ) and the amount of data for the living taxa ( $M_L$ ) are crucial missing data parameters (Fig. 5).

As proposed in previous studies, the amount of missing data is not a problem *per se* as long as enough data overlaps in the matrix (e.g. Kearney 2002; Wiens 2003; Roure and Philippe 2011; Pattinson et al. 2014). To improve the ability to recover the

"correct" topology in a Total Evidence framework, one can efficiently increase the amount of overlapping data by reducing the amount of missing data in each of our three selected parameters. In our simulations, the overlap in data is greatly reduced when different parameters are combined with a high amount of missing data (e.g.  $M_L=75\%$ ,  $M_F=75\%$  and  $M_C=75\%$  - Fig. 2 and Fig. 3). In this scenario the overall number of missing characters ( $M_C$ ) affects the proportion of overlapping data by reducing the overall size of the matrix (Fig. 3). However it is important to note that in the pairwise Bhattacharyya Coefficient analysis, the area with the lowest probability of distributions overlap is led by comparing the trees with no missing data for the living taxa to the trees with 75% missing data for the living data (right lower corner in Fig. 5).

This variable effect of the missing data parameters on the ability to recover the "correct" topology (Fig. 5) can be explained by the way our parameters are illustrating the sources of missing data in empirical matrices: (1) non coded-living taxa ( $M_L$ ), (2) the quality of the fossil record ( $M_F$ ) and (3) the amount of coded characters ( $M_C$  - see Methods). We have suggestions for practical ways to reduce this missing data problem in empirical matrices:

1.  $M_L$  (non coded living taxa): The vast majority of known living species have specimens preserved in natural history museums. This is a significant resource and more research effort is needed to code morphological characters in these specimens.
2.  $M_F$  (quality of the fossil record): The amount of missing data in the fossil record solely depends on the preservation quality of the fossils and can not be actively improved. However, gaps in the matrix can be filled with the discoveries exceptional preserved fossil (e.g. Ni et al. 2013). Fortunately, this parameter is the most difficult to improve but it also has the least influence on whether the

simulations recover the "correct" topology (Fig. 5).

3.  $M_C$  (amount of coded characters): The amount of time it takes to build a morphological matrix increases directly with the number of characters involved. Improvement in both decreasing the time taken to build the matrix and increasing the number of characters can be done through vast collaborative data collection projects through web portals such as *morphobank* (OLeary and Kaufman 2011).

Fixing these three major issues can lead to a sensible increase in the ability to recover the "correct" topology.

Nonetheless, the variation in the ability to recover the "correct" topology does not only depend on the missing data in morphological part of the matrix but also heavily on the method used to infer the topology (Fig. 2 and Fig. 3). Previous studies have shown some superiority on methods using simple probabilistic models such as the Mk model (Lewis 2001) over cladistic methods (Wright and Hillis 2014 but see Spencer and Wilberg 2013). However, this is the first study, to our knowledge, that compares the performance of the Mk model (Lewis 2001) between Maximum Likelihood and Bayesian inference in a Total Evidence framework. When using the clade conservative Robinson-Foulds distance (see Kuhner and Yamato 2014 and discussion below), the Bayesian consensus trees greatly outperforms the Maximum Likelihood trees in the ability to recover the correct tree topology (Fig. 3). In fact, regardless of the amount of missing data, this method always displays the highest normalised Robinson-Foulds score (always  $> 0.69$  - see Supplementary Material Section 3). Also, when looking at the pairwise Bhattacharyya Coefficients within each method, the Bayesian consensus tree one display the highest probability of overlap between the results in any case (see Supplementary Material Section 3). When comparing this to the probability of overlap between the methods, it is counter intuitive to note that the Bayesian consensus tree

shows the highest dissimilarity with the Bayesian posterior trees from which the Bayesian consensus tree is build (Fig. 5-A). Note that it is also highly dissimilar to the Bootstraps. However that can be linked to the fact that both Bootstraps and Bayesian posterior trees are measured as  $125 \times 1000$  random pairwise comparisons instead of just 125 tree comparisons for the Bayesian consensus trees. The fact that the 1000 pairwise trees to be compared for each of the 125 comparisons are chosen randomly, it can add noise to the results for these two methods (see methods above). Also, in the case of the lower performances of the Bayesian consensus trees when using the Triplets distance, one can note the high overlap of the confidence intervals (Fig. 2 and Fig. 5-B).

This difference between the metrics can be explained by their respective ability to measure the differences between two trees (Kuhner and Yamato 2014). As shown in Fig. 2 and Fig. 3, the contrast between the results of the two metrics can be explained by the actual nature of both metrics (see Supplementary Material S2):

1. The Robinson-Foulds metric is a conservative metric, which is more sensible to single taxon displacement because it will count clades as similar only if they are composed of the same taxa with the same topologies (Robinson and Foulds 1981). When getting closer to the root of the tree, displacement of single taxon makes the clades not being exactly identical any more even if the clade still contains all the other taxa in both trees. This metric illustrates therefore the conservation of clades among the compared trees (i.e. when the normalised Robinson-Foulds distance is close to 1, most of the clades in the two trees are identical).
2. On the other hand, the Triplets method is measuring the position of each taxon towards two other reference taxa (Critchlow et al. 1996). It will penalise only trees when taxa get removed furthest from their original position. This metric therefore illustrates the amount of wildcard taxa (Kearney 2002) between the two compared

trees (i.e. when the score is closed to 1, few wildcard taxa are present in the tree).

Therefore, in our study, both metrics are showing two different aspects of tree topology. For example, regarding the Bayesian consensus trees, complete clade are conserved most of the time (Fig. 3-top) but with an increasing amount of missing data, wildcard taxa get more and more displaced (Fig. 3-bottom). Because we are comparing the Bayesian consensus tree to a single tree topology, the probability of recovering the exact position of a wildcard taxa than in the "best" tree is really low and even reducing with the amount of missing data. Regarding this fundamental differences, the Robinson-Foulds metric performs better in showing a distance from the true topology than the Triplets distance (Kuhner and Yamato 2014). Therefore, the results for the Bayesian consensus trees can be preferred on the Maximum Likelihood trees because - when using the Robinson-Fould distance - they display a higher ability to recover the "correct" topology (Fig. 2).

The variations in our results could also be explained by another methodological caveat: the performance of our simulation protocol. In order to reduce the computational time of our analysis (150 CPU years), we had to run our sensitivity analysis on modest size matrices (1000 molecular data and 100 morphological data). This matrix size is two orders of magnitude lower than some matrices used in published analysis (e.g. Springer et al. 2012; Ni et al. 2013) but still ranges in the sizes of smaller matrices used in other ones (e.g. Sallam et al. 2011; Kelly et al. 2014). Also, simulating morphological characters is more complex than molecular ones. The underlying pattern of their evolution are often more complex and ruled by more parameters than molecular characters (i.e the number of character states, the states frequencies, the substitution matrix and the statistical model used) (Pagel 1994; Wagner 2000; Lewis 2001; Wright and Hillis 2014). Morphological characters studies involve many potential statistical pitfalls (e.g. independent characters violation, rate variation -

Davalos et al. (2014)) and especially (i) incongruence with molecular signal and (ii) homoplasy (Wagner 2000).

1. First, morphological data can display a different signal than molecular data, especially in small matrices (Wagner 2000). This might lead to a controversial phylogenetic signal in the overall matrix and lower down the support values. However, regarding empirical data studies, most of the groups shows fairly congruent morphological and molecular phylogenetic signal (e.g. Lee et al. 2013).
2. Secondly, in this study, we made the assumption that theoretically, morphological characters are randomly distributed on an organism however it seems clear that empirical morphological data does not act randomly (Sansom and Wills 2013; Pattinson et al. 2014). However, following our simulation assumption of random character distributions, if they accumulate through time in the same way as the majority of the molecular characters then, homoplasic characters are expected to appear randomly through time (Davalos et al. 2014). Therefore, homoplasy is expected to be more important (by chance) in bigger morphological matrices (Davalos et al. 2014). After reaching a critical amount of morphological characters, adding new ones increases homoplasy (Wagner 2000).

One way to address this issue would be to run a similar analysis using real data. The amount of studies using Total Evidence data type matrices is steadily increasing (e.g. Ronquist et al. 2012a; Slater 2013; Beck and Lee 2014) and one could analysis these total evidence matrices in the same framework as this study. However, the true topology (obtained from a matrix with no missing data) is unlikely to be known. A second way to address this issue, and especially to incorporate the variation of signal in morphological character would be to simulate morphological characters at different rates (Wright and Hillis 2014) or to apply patterns of missing data directly obtained

from the fossil record (Pattinson et al. 2014). Nevertheless, our simulation already show clear results of the effect of missing data on topology in a Total Evidence framework. Also, our results show how recovering the correct topology can be improved by using Bayesian consensus trees in combination with a maximum amount of characters and a of living taxa with morphological data.

Despite the various caveats discussed above, we answered our main question on how does Total Evidence matrices perform in recovering topology with different methods and different amounts of missing data. Our results shows that this method performance decreases mainly with a reduction in the overlap of morphological data. In order to address this eventual problem, we propose to increase the data overlap by coding living species and to have a large morphological matrix. Also, our results shows disparity among methods and a clear superiority of the Bayesian consensus tree topology over the Maximum Likelihood one.

## CONCLUSION

Using Total Evidence matrices with missing data is not a problem for recovering the correct topology as long as enough morphological data is overlapping between the fossil and living taxa. In order to optimize good topological recovery, we propose to use the Bayesian consensus tree over the Maximum Likelihood tree or over a random subset of the Bayesian posterior trees distribution. Also, in order to improve the overlap in data in the morphological part of the matrix, we advice to use all the morphological characters available and to code them for a maximum of living taxa present in the matrix.

## SUPPLEMENTARY MATERIAL

Supplementary material (code, analysis and full results) can be found in the Dryad data repository at <http://dx.doi.org/10.5061/dryad.XXXX>.

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## SUPPLEMENTARY MATERIAL

### *Tree Building*

#### SUPPLEMENTARY MATERIAL SECTION 1

##### *Morphological characters states*

In order to obtain a realistic probabilistic value for of  $k$  characters states for each simulated morphological character, we downloaded 100 random morphological characters (with more than 100 characters each) from TreeBASE database (<http://treebase.org/>) published between 1985 and 2013 and covering 19 taxonomic classes (Chordata, Arthropoda, Annelida, Angiosperm, Gymnosperm and Pteridophyta). We selected a total of 22563 characters ranging from 2 to 10 states. We calculated the proportion of characters with 2, 3, 4, 5, 6, 7, 8, 9 or 10 states. We then sampled 22563  $k$  values between 2 and 10 with the same proportion of characters from the empirical data. We then used a simple t-test to check if our simulation was equal to the empirical data. In this study, we only simulated characters with 2 or 3 states because of the high proportion of ordered characters encountered on characters with more than 3 states and the difficulties of simulate biologically sensible ordered characters.

##### *Tree Building Software settings*

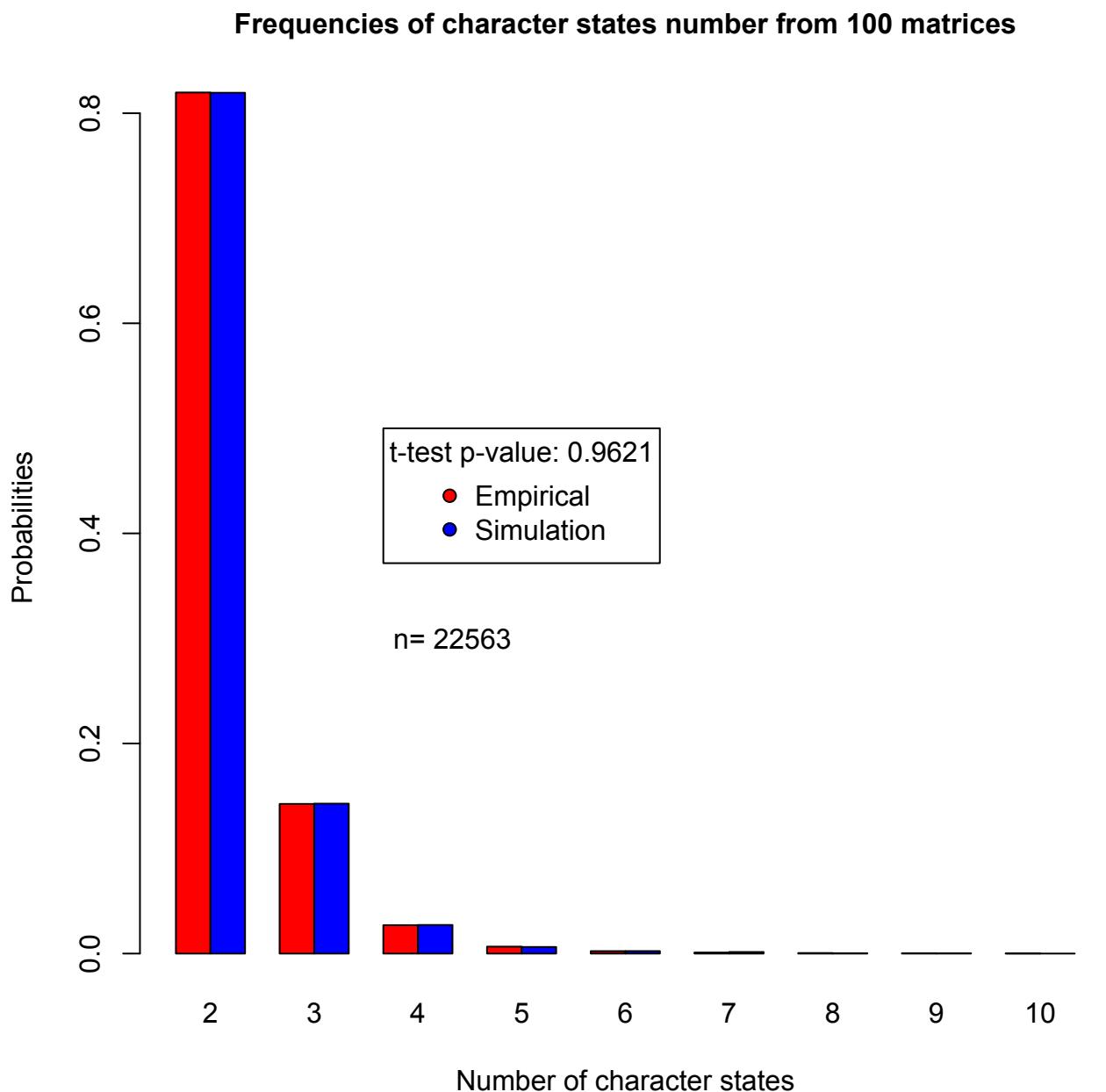


Figure 6: Character states distribution in empirical matrices. Characters states number distribution extracted from 100 random morphological matrices downloaded from TreeBase.

*Maximum Likelihood - RAxML v8.0.20 (Stamatakis 2014).—*

Model:

Molecular data:

GTR +  $\Gamma_4$  (-m GTRGAMMA)

Morphological data:

Mk +  $\Gamma_4$  (-K MK)

Support:

Rapid Bootstrap algorithm (LSR), 1000 replicates

*Bayesian - MrBayes v3.2.1 (Ronquist et al. 2012b).—*

Priors:

Molecular data:

rates distribution shape ( $\alpha$ ) = 0.5

Transition/Transversion ratio = 2 ( $\beta(80,40)$ )

Starting tree: "True" tree topology with each branch length = 1

Morphological data:

rates distribution shape ( $\alpha$ ) = 0.5

Models:

Molecular data: HKY +  $\Gamma_4$

Morphological data: Mk +  $\Gamma_4$

MCMC:

2 runs

4 chains per run

generations  $\geq 50 \times 1^6$

sample frequency  $= 1050 \times 1^3$

ASDS diagnosis frequency  $= 50 \times 1^3$

ASDS < 0.01

ESS >> 200

Burnin = 25%

### *Tree Comparisons*

## SUPPLEMENTARY MATERIAL SECTION 2

### *Robinson-Foulds distance details*

Robinson-Foulds distance (*RF*; Robinson and Foulds 1981), or "path difference", measures the number of shared clades across two trees. The metric reflects the distance between the distributions of tips among clades in the two trees (Robinson and Foulds 1981) and can be expressed as following:

$$RF_{x,y} = N_x + N_y - 2C_{x,y} \quad (1)$$

Where  $C_{x,y}$  is the number of clades in common in the two trees.  $C$  is equal to one if the two trees have the same  $n$  taxa; and  $C = n - 2$  when none of the  $n$  taxa are shared between the trees. This metric is more sensitive to taxon displacement than Triples distance (i.e. if one taxon moves out of a clade, then the clades are no longer considered similar; Critchlow et al. (1996); Johnson and Soltis (1998); Wiens (2003)). The minimal value of  $C$  is equal to 1 if the two trees have the same  $n$  taxa; the maximal value in  $C = n - 2$ . For a fully unresolved tree (star tree)  $N=1$  and for a fully resolved tree (binary tree)  $N = n - 2$ . The minimal and maximal topological distance for taxa is:

$$RF_{min} = 1 + 1 - 2C_{x,y} \quad (2)$$

And:

$$RF_{max} = 2(n - 2) - 2 \quad (3)$$

One can then rescale  $RF.scaled$  by using the maximal and minimal value for any  $n$  taxa:

$$RF.scaled_{x,y} = \frac{RF_{x,y} - RF_{min}}{RF_{max}} \quad (4)$$

This metric is more sensitive to taxa displacement than the Triplet distance (Critchlow et al. 1996; Johnson and Soltis 1998; Wiens 2003) and therefore a low value will show a good clade conservation between two trees and a high value will show a bad recovery of common clades.

### *Triplets distance details ( $T_{x,y}$ )*

Triples distance ( $T_{x,y}$ ; Dobson 1975) measures the number of sub-trees made up of three taxa (triplets) that differ between two given trees. Each triplet can be written as

$I_{ijk} = (ijk)$ . Where  $I_{ijk}$  is equal to zero if the the two triplets  $(ijk)$  are the same in the two trees otherwise  $I_{ijk}$  is equal to one. For any rooted binary tree there are only three possible combinations for each triplet:  $((j,k),i)$ ;  $((i,k),j)$ ; and  $((i,j),k)$ ; (Johnson and Soltis 1998). If the trees used are not fully binary, a fourth triplet combination is possible:  $(i,j,k)$ . We can calculate the triplet distance between two trees,  $S_n$ , as:

$$S_n = \sum_{ijk} I_{ijk} \quad (5)$$

Where:

$$\sum_{ijk} = \binom{n}{4} = \frac{n!}{4!(n-4)!} \quad (6)$$

And where  $n$  is the total number of taxa in both trees (modified from Critchlow et al. (1996)). If  $S_n = 0$ , the trees are identical; when  $S_n = \binom{n}{4}$ , the trees are as different as possible (i.e. every taxon has a different placement in the two trees). Because the

possible number of triplets per clade is a finite number, the probability of two random trees with the same  $n$  taxa to have the same triplet is:

$$P(I_{ijk} = 0) = \frac{1}{4} \quad (7)$$

Therefore one can calculate the probability of two random trees having the same triplets:

$$P(S_n = 0) = \sum_{ijk} P_{I_{ijk}=0} \quad (8)$$

$$P(S_n = 0) = \frac{n!}{4(3!(n-3)!)} \quad (9)$$

And in the same way:

$$P(S_n = 1) = \frac{3n!}{4(3!(n-3)!)} \quad (10)$$

### *Normalised Tree Similarity*

For any tree with  $n$  taxa compared using a tree distance metric  $m$ , Normalized Tree Similarity,  $NTS_m$  (Bogdanowicz et al. 2012), represents the similarity score for the two trees given the expected distance between two random Yule trees with  $n$  taxa. If  $\bar{d}_{m,n}(rand)$  is the average distance between two random Yule trees with  $n$  taxa and  $d_{m,n}(x,y)$  the distance between the two trees  $x$  and  $y$  containing each  $n$  taxa, then:

$$NTS_{m,n}(x,y) = \frac{\bar{d}_{m,n}(rand) - d_{m,n}(x,y)}{\bar{d}_{m,n}(rand)} \quad (11)$$

$NTS$  ranges from one to  $-\infty$ . For any  $m, n$ , when  $NTS = 1$ , the trees are identical, when  $NTS = 0$  the trees are no more different than expected by chance, and when  $NTS < 0$ , the trees are more different than expected when comparing two random trees.

### *Tree comparisons*

*Random tree comparison scaling.*— We used the comparison of 1000 random trees to obtain the mean comparison value  $\bar{d}_{m,n}(rand)$  for the NTS metric. We randomly generated two sets of 1000 trees of  $n$  taxa using the rmtree function of ape package (v3.0-11 Paradis et al. (2004)) that generates a given number of random Yule trees. We calculated the  $\bar{d}_{m,n}(rand)$  value using an approach similar to the RPCBTC (described below) by performing 1000 random pairwise comparisons using the TreeCmp java script (Bogdanowicz et al. 2012).

### Bhattacharyya Coefficient

The Bhattacharyya Coefficient calculates the probability of overlap of two distributions (Bhattacharyya 1943). When it is equal to 0, the probability of overlap of the distributions is null and when it is equal to 1, the two distributions are entirely overlapping. It is an elegant and easy to compute continuous measurement of probability of similarity between two distributions. The coefficient is calculated as the sum of the square root of the relative counts shared in a  $n$  number of bins among two distributions.

$$\text{Bhattacharyya Coefficient} = \sum_{i=1}^n \sqrt{\sum a_i \times \sum b_i} \quad (12)$$

Where, for  $a$  and  $b$ :

$$\sum a_i = \frac{\text{Number of counts in bin } i \text{ for the distribution } a}{\text{Total number of counts for the distribution } a} \quad (13)$$

The precision of the Bhattacharyya Coefficient is directly related to the number of  $n$  bins. If  $n$  is low, the overlap will be overestimated and if  $n$  is too high, the overlap will be underestimated. In this analysis, we determined the number of bins Silverman's rule of thumb which determines the bins to be equal to 0.9 times the minimum of the standard deviation and the interquartile range divided by 1.34 times the sample size to the negative one-fifth power (Silverman (1986, page 48, eqn (3.31))).

## *Codes*

All codes are available at: [https://github.com/TGuillerme/Total\\_Evidence\\_Method-Missing\\_data/tree/master/Functions](https://github.com/TGuillerme/Total_Evidence_Method-Missing_data/tree/master/Functions)

The tree comparison results analysis can be repeated for more details at:  
[https://github.com/TGuillerme/Total\\_Evidence\\_Method-Missing\\_data/tree/master/Analysis](https://github.com/TGuillerme/Total_Evidence_Method-Missing_data/tree/master/Analysis)

## *Additional Results*

### SUPPLEMENTARY MATERIAL SECTION 3

	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
Maximum likelihood-RF	0.06	0.26	0.40	0.41	0.50	0.95
Maximum likelihood-Tr	0.29	0.45	0.59	0.63	0.84	1.00
Bayesian consensus-RF	0.69	0.71	0.72	0.76	0.79	0.96
Bayesian consensus-Tr	-0.28	-0.11	0.17	0.19	0.37	0.98
Bootstraps-RF	0.06	0.18	0.27	0.26	0.34	0.46
Bootstraps-Tr	0.23	0.31	0.35	0.38	0.45	0.58
Bayesian posterior trees-RF	0.16	0.22	0.32	0.34	0.42	0.65
Bayesian posterior trees-Tr	0.24	0.35	0.40	0.50	0.67	0.98

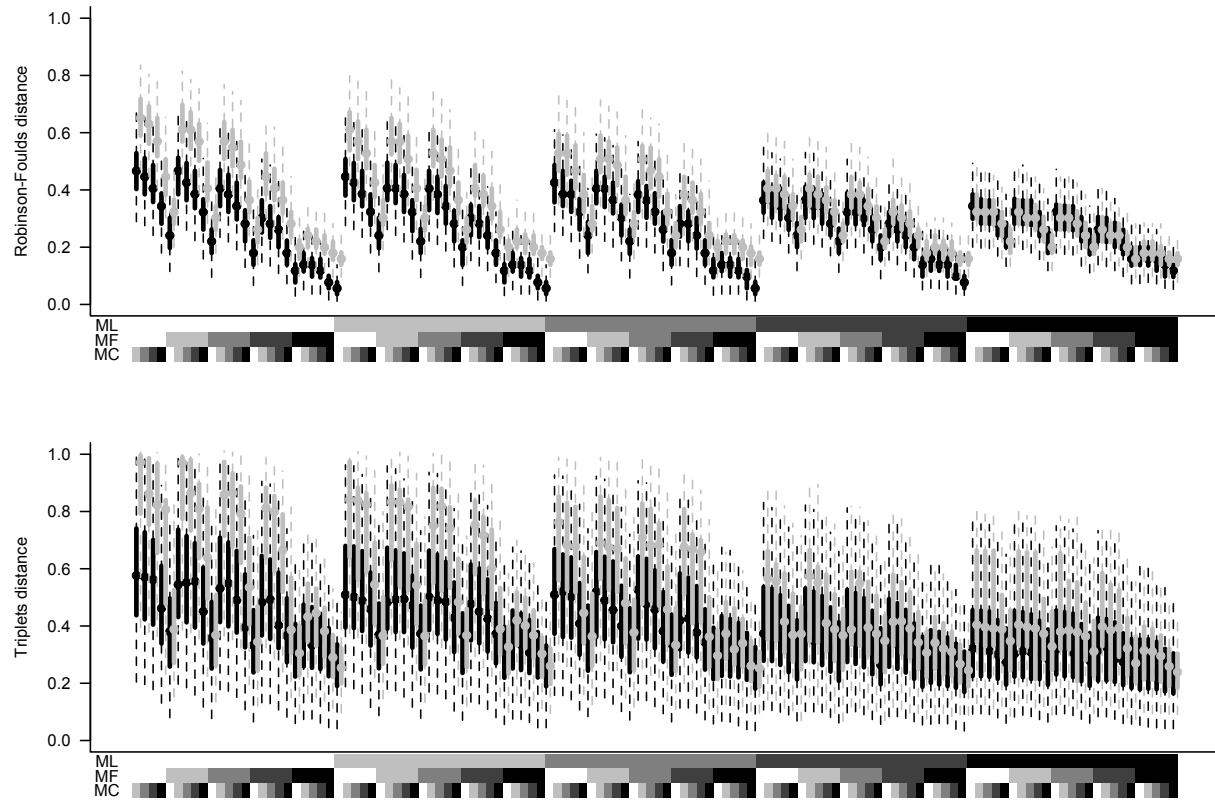


Figure 7: Trend of the effect of missing data on topological recovery on the Bootstraps and the Bayesian posterior trees distributions. The amount of missing data per parameter ( $M_L$ ,  $M_F$  and  $M_C$ ) is represented along the x axis. The colour gradient from white to black represents respectively, 0%, 10%, 25%, 50% and 75% of missing data. The topological recovery is represented on the y axis, both using Robinson-Foulds distance (upper row) and Triplets distance (lower row). Points represent the modal value of each distribution ; thick solid and thin dashed lines represents respectively the 50% and 95% confidence intervals or the distributions. The Bootstraps are represented in black and the Bayesian posterior trees distributions in grey.

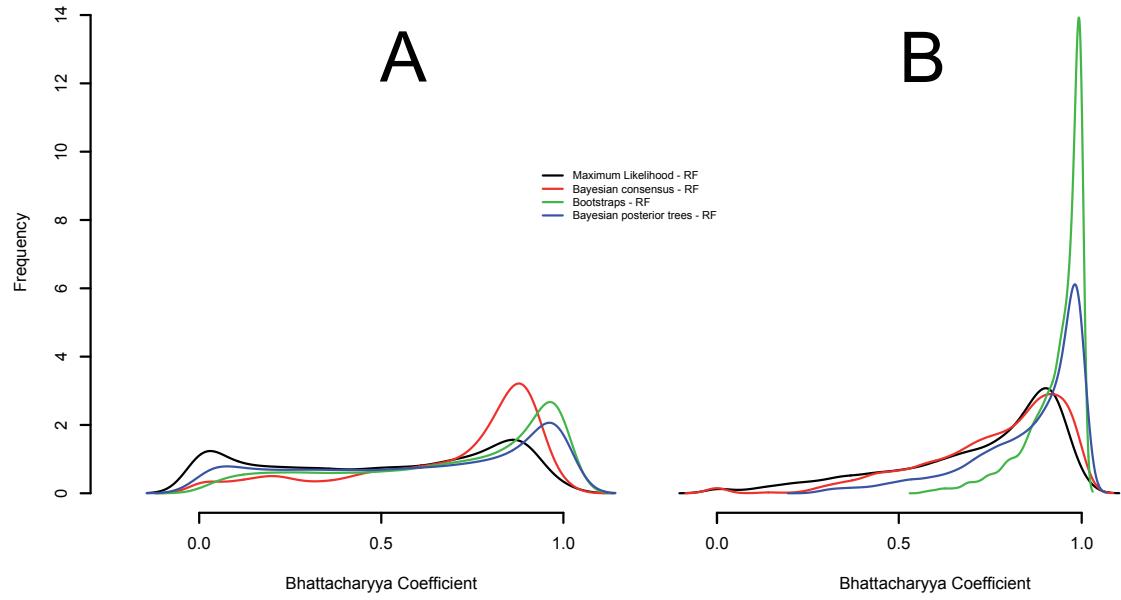


Figure 8: Distribution of the Pairwise Bhattacharyya Coefficients within each method.  
 A- Distribution of the coefficients when comparing Ronbinson-Foulds distances. B- Distribution of the coefficients when comparing Triplets distances.

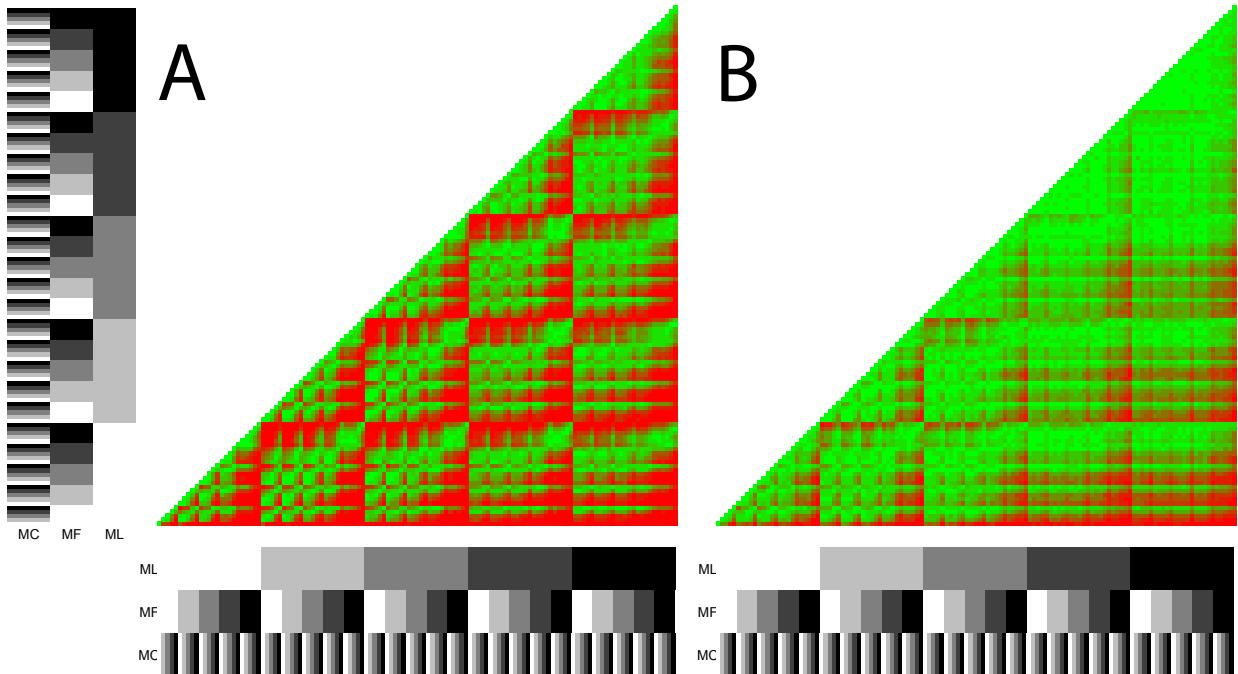


Figure 9: Pairwise Bhattacharyya Coefficients within the Maximum Likelihood trees. The pairwise trees comparisons are represented on both axes. The colour gradient from white to black represents respectively, 0%, 10%, 25%, 50% and 75% of missing data for each parameter. The matrix represents the values of pairwise Bhattacharyya Coefficients going from green (1) to red (0). A. Results for the Normalised Robinson-Foulds distance. B. Results with the Normalised Triplets distance.

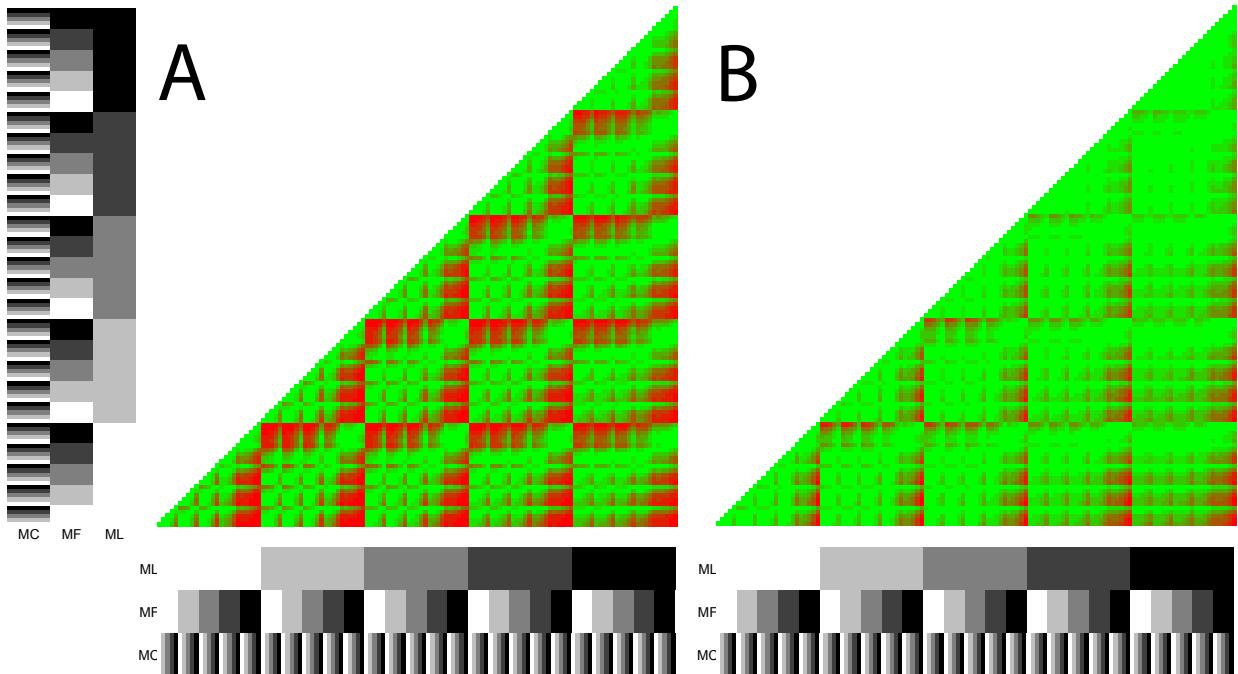


Figure 10: Pairwise Bhattacharyya Coefficients within the Bootstrap trees. The pairwise trees comparisons are represent on both axis. The colour gradient from white to black represents respectively, 0%, 10%, 25%, 50% and 75% of missing data for each parameter. The matrix represents the values of pairwise Bhattacharyya Coefficients going from green (1) to red (0). A. Results for the Normalised Robinson-Foulds distance. B. Results with the for Normalised Triplets distance.

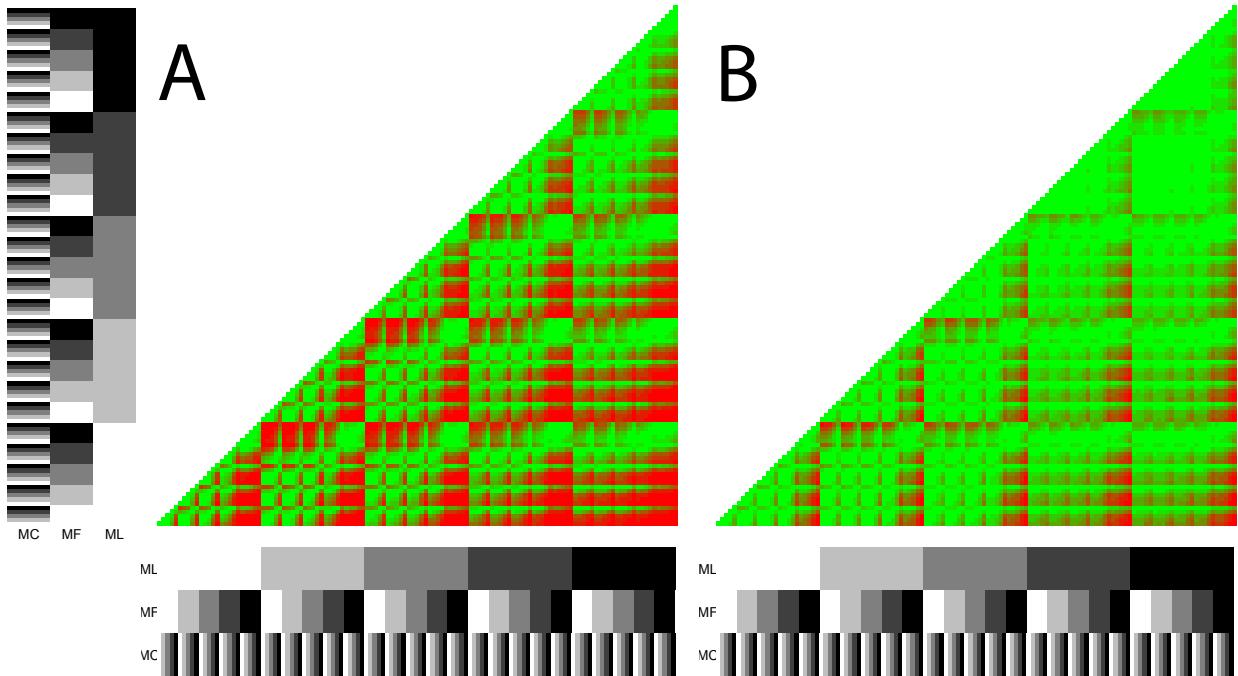


Figure 11: Pairwise Bhattacharyya Coefficients within the Bayesian posterior distribution trees. The pairwise trees comparisons are represent on both axis. The colour gradient from white to black represents respectively, 0%, 10%, 25%, 50% and 75% of missing data for each parameter. The matrix represents the values of pairwise Bhattacharyya Coefficients going from green (1) to red (0). A. Results for the Normalised Robinson-Foulds distance. B. Results with the for Normalised Triplets distance.