### 432 Class 09

https://thomaselove.github.io/432-2023/

2023-02-14

# Today's Agenda

- A New NHANES Example
- Logistic Regression Analyses in Project A
  - Establishing a Research Question
  - Identifying / Tidying Outcome and Candidate Predictors
  - Dealing with Missing Data
  - Building a "Main Effects" Model and Plotting Effects
  - Considering Non-Linear Terms
  - Fitting an "Augmented" Model and Plotting Effects
  - Summarizing/Presenting a Final Model
    - In-Sample and Validated Model Summaries
    - Selecting Model Y or Model Z
    - Describing a Meaningful Effect (see Notes, Chapter 22)
    - ROC curve for the Final Model
    - Nomogram for the Final Model

## Today's R Setup

```
knitr::opts chunk$set(comment = NA)
library(nhanesA)
library(broom)
library(caret)
library(janitor)
library(knitr)
library(mosaic)
library(naniar)
library(pROC)
library(rms)
library(simputation)
library(tidyverse)
theme set(theme bw())
```

Section 1

The Data

### **NHANES** Data

We'll use data from the 2011-2012 administration of NHANES here, just because I don't want to overlap with any studies people might be thinking about for Project B.

Data bases we'll use are:

- **DEMO\_G** for Demographic Variables
- CDQ\_G for Cardiovascular Health Questionnaire
- HSQ\_G for Current Health Status Questionnaire
- BPX\_G for Blood Pressure Examination Results
- BMX\_G for Body Measures Examination Results
- MCQ\_G for Medical Conditions Questionnaire

### Pulling the Data, I

```
demo_g <- nhanes("DEMO_G") |> tibble(); dim(demo_g)
[1] 9756
         48
cdq_g <- nhanes("CDQ_G") |> tibble(); dim(cdq_g)
[1] 3603 17
hsq g <- nhanes("HSQ G") |> tibble(); dim(hsq g)
[1] 8956 14
```

## Pulling the Data, II

```
bpx_g <- nhanes("BPX_G") |> tibble(); dim(bpx_g)
[1] 9338 27
bmx_g <- nhanes("BMX_G") |> tibble(); dim(bmx_g)
[1] 9338 26
mcq g <- nhanes("MCQ G") |> tibble(); dim(mcq g)
```

[1] 9364 92

### Merging the Tibbles

```
df_mlist <- list(demo_g, cdq_g, hsq_g, bpx_g, bmx_g, mcq_g)
nh_merge <- df_mlist |>
  reduce(left_join, by = "SEQN") # reduce is from purrr
dim(nh_merge)
```

[1] 9756 219

We had 262 variables in our original seven tibbles, but that counts the SEQN variable seven times, and we only have it once in our nh\_merge tibble.

Now, which of these 256 variables are we actually going to use?

# Variables We'll Use Today

NHANES	Description	Source
SEQN	Identifying code	All 6 files
CDQ010	Shortness of breath on stairs/inclines?	CDQ_G
RIDAGEYR	Age in years at screening	DEMO_G
HSD010	General health condition $(E/VG/G/F/P)$	HSQ_G
LBXSCLSI	Serum Chloride (mmol/L)	BIOPRO_G
BPXDI1	Diastolic BP (first reading, in mm Hg)	BPX_G
BMXBMI	Body Mass Index $(kg/m^2)$	BMX_G
MCQ010	Ever been told you have asthma	MCQ_G

## Selecting Today's Variables

[1] 3234

SEQN should be unique for each row in the data.

```
identical(nrow(nh_today), n_distinct(nh_today$SEQN))
```

[1] TRUE

2 CDQ010 was 1 for Yes, 2 for No. We want 1 for Yes, 0 for No.

```
nh_today <- nh_today |> mutate(sbreath = 2 - CDQ010)
nh_today |> count(CDQ010, sbreath) # sanity check
```

```
nh_today <- nh_today |> select(-CDQ010) |>
relocate(sbreath, .after = "SEQN")
```

- Age should be between 40 and 79 years
- ullet Body mass index should be between 12.4 and 82.1 kg/ $m^2$
- Oiastolic BP should be between 30 and 120 mm Hg (treat values below 30 as NA)

```
df_stats(~ age + bmi + dbp, data = nh_today) |>
rename(var = response) |> kable(digits = 1)
```

var	min	Q1	median	Q3	max	mean	sd	n	missing
age	40.0	48.0	57.0	65.0	79.0	57.3	10.7	3234	0
bmi	13.6	24.7	28.3	32.9	82.1	29.5	6.9	3062	172
dbp	0.0	66.0	74.0	80.0	120.0	73.1	12.4	2854	380

## Replace dbp values below 30 with NA

3 NA 380

nh today |> count(dbp < 30)

```
nh_today <- replace_with_na_at(nh_today, "dbp", ~ .x < 30)</pre>
```

```
favstats(~ dbp, data = nh_today) |> kable(digits = 1)
```

min	Q1	median	Q3	max	mean	sd	n	missing
30	66	74	80	120	73.4	11.6	2842	392

ullet asthma should be a two-level factor (currently 1= Yes, 2= No, 9= Don't Know, which we'll treat as missing)

```
nh_today |> count(asthma)
```

# Recoding Asthma as a factor with 3 missing values

```
nh_today <- nh_today |>
  mutate(asthma = fct_recode(
    factor(asthma), "Yes" = "1", "No" = "2", NULL = "9"),
    asthma = fct_relevel(asthma, "No"))

nh_today |> tabyl(asthma) |> adorn_pct_formatting()
```

```
asthma n percent valid_percent
No 2779 85.9% 86.0%
Yes 452 14.0% 14.0%
<NA> 3 0.1% -
```

Self-reported overall health should be a five-level factor

sroh	n	percent	valid_percent
E	238	7.4%	8.5%
VG	684	21.2%	24.5%
G	1119	34.6%	40.1%
F	612	18.9%	21.9%
Р	136	4.2%	4.9%
NA	445	13.8%	-

# Which Variables are Missing?

#### miss\_var\_summary(nh\_today)

```
# A tibble: 7 x 3
 variable n_miss pct_miss
 <chr>
                    <dbl>
           <int>
1 sroh
            445 13.8
2 dbp
             392 12.1
3 bmi
             172 5.32
4 asthma
               3 0.0928
5 SEQN
                    0
6 sbreath
                   0
7 age
                    0
```

# How Many Missing Values?

#### miss\_case\_table(nh\_today)

```
# A tibble: 4 x 3
 n_miss_in_case n_cases pct_cases
           <int>
                   <int>
                             <dbl>
                   2555
                            79.0
              0
                     475
                            14.7
3
                     75
                         2.32
              3
4
                     129
                             3.99
```

### Updated Codebook

We have 3234 rows and 7 in the nh\_today data now. 2555, or 79% of the rows have complete data on these 7 variables.

Name	Description	NHANES Source
SEQN	Identifying code	All 6 files
sbreath	Shortness of breath on stairs/inclines?	CDQ_G (CDQ010)
age	Age in years at screening	DEMO_G (RIDAGEYR)
sroh	Self-reported health $(E/VG/G/F/P)$	HSQ_G (HSD010)
dbp	Diastolic BP (1st reading, in mm Hg)	BPX_G (BPXDI1)
bmi	Body Mass Index (kg $/m^2$ )	BMX_G (BMXBMI)
asthma	Ever been told you have asthma?	MCQ_G (MCQ010)

• Inclusions/Exclusions: Valid (1 or 0) response to sbreath, age between 40 and 79 years, inclusive.

### Section 2

# Project A Tasks

## Establishing a Research Question

How effectively can we predict whether or not an adult subject has experienced "shortness of breath when hurrying on the level or walking up a slight hill" on the basis of their age, self-reported overall health, diastolic blood pressure, body mass index and whether or not they have been told they have asthma?

- Our data come from NHANES 2011-12, and describe a total of 3234 (unweighted) adult (ages 40-79) subjects.
- We will not use survey weights in this work.

## Identifying / Tidying Outcome

Our outcome is the subject's response to the following question:

### Have you had shortness of breath either when hurrying on the level or walking up a slight hill?

This was asked of adults ages 40 years and up, as question CDQ010 on the  $CDQ_G$  questionnaire in NHANEZS 2011-12, and we've included subjects who gave either a Yes or No response.

```
nh_today |> tabyl(sbreath) |> adorn_totals() |>
adorn_pct_formatting()
```

```
sbreath n percent
0 2219 68.6%
1 1015 31.4%
Total 3234 100.0%
```

## Identifying Candidate Predictors

The five predictors we will examine for this outcome are age, sroh, dbp, bmi and asthma.

Name	Description	Missing?
SEQN	Identifying code	None
sbreath	Shortness of breath on stairs/inclines?	None
age	Age in years at screening	None
sroh	Self-reported health $(E/VG/G/F/P)$	445
dbp	Diastolic BP (1st reading, in mm Hg)	392
bmi	Body Mass Index (kg $/m^2$ )	172
asthma	Ever been told you have asthma?	3

## Dealing with Missing Data

We have excluded all cases with missing sbreath so our outcome is complete.

We will assume MAR for the remaining missing values and then use single imputation both to:

- ullet build a Spearman  $ho^2$  plot
- fit our models Y and Z

If you wanted to use multiple imputation in the project, I would do that at the end, by refitting the "winning" model and summarizing those results (after imputation) only as part of your **Final Model** materials. I might also use aregImpute() to get my single imputation, although I won't here.

# Single Imputation via simputation

```
set.seed(43212345)
nh_today_i <- nh_today |> data.frame() |>
  impute_rhd(asthma ~ age) |>
  impute_rlm(dbp ~ age + asthma) |>
  impute_rlm(bmi ~ dbp + age + asthma) |>
  impute_cart(sroh ~ age + bmi) |>
  as_tibble()

n_miss(nh_today_i) # should now have no missing data
```

[1] 0

# Resulting nh\_today\_i tibble

#### nh\_today\_i

```
# A tibble: 3,234 x 7
   SEQN sbreath age sroh dbp bmi asthma
  <dbl> <dbl> <dbl> <fct> <dbl> <fct> <dbl> <fct>
 1 62164
              0
                   44 G
                             56 23.2 No
 2 62172
              0
                   43 G
                             70 33.3 No
 3 62177
              0
                   51 G
                             68
                                   20.1 No
4 62179
              0
                   55 VG
                             78 27.6 No
 5 62182
                   75 G
                           69.6 28.5 No
 6 62191
              0
                   70 G
                             68
                                   28.5 No
              1
 7 62199
                   57 VG
                             70
                                   28
                                       No
                   42 VG
8 62200
              0
                             88
                                   27.6 No
 9 62201
              0
                   58 G
                             73.3 28.6 No
10 62209
              1
                   62 F
                             60
                                   26
                                        No
\# ... with 3,224 more rows
```

### Section 3

Model Y: The "Main Effects"

# Building a "Main Effects" Model and Plotting Effects

First, we'll assume MAR and do our analysis on the (singly) imputed data nh\_today\_i)

# modY\_si results (from lrm fit)

```
modY_si
Logistic Regression Model
 lrm(formula = sbreath \sim age + sroh + dbp + bmi + asthma, data = nh_today_i,
     x = TRUE, y = TRUE
                        Model Likelihood
                                                Discrimination
                                                                   Rank Discrim.
                              Ratio Test
                                                       Indexes
                                                                         Indexes
              3234
                      LR chi2
                                   426.39
                                                R2
                                                         0.174
                                                                           0.718
 Obs
  0
              2219
                      d.f.
                                        8
                                               R2(8,3234)0.121
                                                                           0.436
                                                                   Dxv
  1
              1015
                      Pr(> chi2) < 0.0001
                                             R2(8,2089.3)0.181
                                                                           0.436
                                                                   gamma
 max |deriv| 2e-10
                                                Brier
                                                         0.187
                                                                           0.188
                                                                   tau-a
            Coef
                    S.E.
                           Wald Z Pr(>|Z|)
 Intercept
            -4.1132 0.4878 -8.43
                                   < 0.0001
 age
             0.0149 0.0039
                            3.79
                                   0.0002
 sroh=VG
             0.4849 0.2337
                            2.07
                                   0.0380
 sroh=G
             1.1641 0.2179
                            5.34
                                   < 0.0001
 sroh=F
             1.8890 0.2269
                            8.33
                                   < 0.0001
 sroh=P
             1.9771 0.2780
                            7.11
                                   < 0.0001
            -0.0062 0.0038 -1.63
                                  0.1033
 dbp
 bmi
             0.0544 0.0062
                            8.79
                                   < 0.0001
 asthma=Yes 0.7981 0.1111
                            7.18
                                   < 0.0001
```

# Key Fit Summary Statistics for Model Y (modY\_si)

```
temp <- modY_si$stats
temp["C"]</pre>
```

0.7181358

```
temp["R2"]
```

R2

0.1735251

The Nagelkerke  $\mathbb{R}^2$  for this model is 0.718 and the C statistic is 0.174.

### glm version of this same fit

```
modY_si_g
Call: glm(formula = sbreath ~ age + sroh + dbp + bmi + asthma, family = binomial(link = logit),
   data = nh_today_i)
Coefficients:
(Intercept)
                             srohVG
                                          srohG
                                                     srohF
                                                                   srohP
                                                                                 dbp
                   age
  -4.113185
             0.014921
                           0.484919
                                       1.164107
                                                   1.888952 1.977082 -0.006171
       bmi
              asthmaYes
  0.054416
             0.798134
Degrees of Freedom: 3233 Total (i.e. Null); 3225 Residual
Null Deviance:
Residual Deviance: 3598
                              AIC: 3616
```

```
glance(modY si g) |> select(AIC, BIC) |> kable(dig = 1)
```

AIC BIC 3615.7 3670.4

# Tidied Table of Model Y (Exponentiated) Coefficients

Here's the code: result is on next slide.

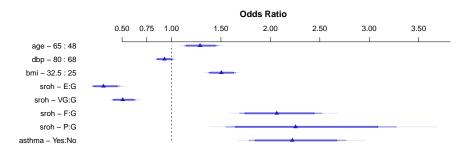
```
tidy(modY_si_g, exponentiate = TRUE,
    conf.int = TRUE, conf.level = 0.90) |>
select(term, estimate, std.error,
    low90 = conf.low, high90 = conf.high,
    p = p.value) |> kable(digits = 3)
```

# Tidied Table of Model Y (Exponentiated) Coefficients

term	estimate	std.error	low90	high90	р
(Intercept)	0.016	0.488	0.007	0.036	0.000
age	1.015	0.004	1.008	1.022	0.000
srohVG	1.624	0.234	1.117	2.415	0.038
srohG	3.203	0.218	2.267	4.652	0.000
srohF	6.612	0.227	4.607	9.737	0.000
srohP	7.222	0.278	4.609	11.519	0.000
dbp	0.994	0.004	0.988	1.000	0.103
bmi	1.056	0.006	1.045	1.067	0.000
asthmaYes	2.221	0.111	1.850	2.667	0.000

### Model Y Effects Plot on Odds Ratio Scale

#### plot(summary(modY\_si))

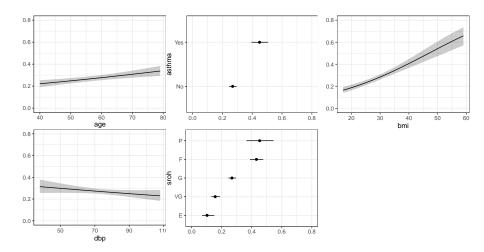


# Details of Effects Plot (Model Y)

```
> summary(modY_si)
            Effects
                                  Response : sbreath
Factor
                 Low High Diff. Effect
                                        S.E. Lower 0.95 Upper 0.95
                 48
                     65.0 17.0
                                 0.253660 0.066962
                                                    0.12242
                                                               0.384900
age
                                                    1.13020
 Odds Ratio
                 48
                     65.0 17.0 1.288700
                                                               1.469500
                                                NA
dbp
                 68
                     80.0 12.0
                                -0.074048 0.045455 -0.16314
                                                               0.015043
                 68
                     80.0 12.0
 Odds Ratio
                                0.928630
                                                NA
                                                    0.84947
                                                               1.015200
                 25
                     32.5 7.5
                                 0.408120 0.046411
                                                    0.31716
                                                               0.499080
bmi
 Odds Ratio
                 25
                     32.5
                           7.5
                                1.504000
                                                    1.37320
                                                               1.647200
                                                NA
                  3
                                -1.164100 0.217900 -1.59120
                                                               -0.737040
sroh - E:G
                     1.0
                            NA
                                 0.312200
                                                    0.20369
                                                                0.478530
 Odds Ratio
                     1.0
                            NA
                                                NA
                      2.0
                                -0.679190 0.116900 -0.90831
                                                               -0.450060
sroh - VG:G
                            NA
 Odds Ratio
                                 0.507030
                                                    0.40320
                                                               0.637590
                     2.0
                            NA
                                                NA
sroh - F:G
                      4.0
                            NA
                                 0.724850 0.101520
                                                    0.52588
                                                               0.923810
                                                    1.69190
 Odds Ratio
                      4.0
                            NA
                                 2.064400
                                                NA
                                                                2.518900
                     5.0
                                 0.812980 0.190200
                                                    0.44019
                                                               1.185800
sroh - P:G
                            NA
 Odds Ratio
                      5.0
                            NA
                                 2.254600
                                                NA
                                                    1.55300
                                                                3.273200
                     2.0
                                 0.798130 0.111120
                                                    0.58035
                                                               1.015900
asthma - Yes:No
                            NA
 Odds Ratio
                      2.0
                                 2.221400
                                                    1.78670
                                                                2.761900
                            NA
                                                NA
```

### Prediction Plot for Model Y

### ggplot(Predict(modY\_si, fun = plogis))



### Confusion Matrix for Model Y

How well does our Model Y classify subjects using a decision rule at 0.5?

### sbreath

pred	0	1	Total
Predict No SB	2040	726	2766
Predict SB	179	289	468
Total	2219	1015	3234

• What fraction of our predictions are correct with this decision rule?

# Summaries of Classification Accuracy, 1

# sbreath pred 0 1 Total Predict No SB 2040 726 2766 Predict SB 179 289 468 Total 2219 1015 3234

- Accuracy is (2040 + 289) / 3234 = 0.720
  - 72.0% of this model's predictions were accurate.
- **Sensitivity** is 289 / 1015 = 0.285
  - 28.5% of those who actually were short of breath are predicted to be short of breath.
- **Specificity** is 2040 / 2219 = 0.919
  - 91.9% of those who actually weren't short of breath were predicted not to be short of breath.

# Summaries of Classification Accuracy, 2

# sbreath pred 0 1 Total Predict No SB 2040 726 2766 Predict SB 179 289 468 Total 2219 1015 3234

- Positive Predictive Value (PPV) is 289 / 468 = 0.618
  - 61.8% of those predicted to be short of breath actually were short of breath.
- Negative Predictive Value (NPV) is 2040 / 2766 = 0.738
  - 73.8% of those predicted to not be short of breath actually were not short of breath.

# Using the caret package to get a confusion matrix

```
cmY <- confusionMatrix(
  data = factor(modY_aug$.fitted >= 0.5),
  reference = factor(modY_aug$sbreath == 1),
  positive = "TRUE")
```

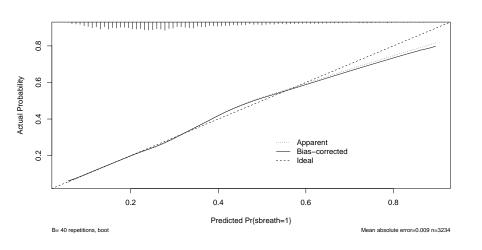
Result on the next slide.

## Confusion Matrix Output (from caret) for Model Y

Confusion Matrix and Statistics Reference Prediction FALSE TRUE 2040 726 FALSE TRUE 179 289 Accuracy: 0.7202 95% CI: (0.7043, 0.7356) No Information Rate: 0.6861 P-Value [Acc > NIR] : 1.393e-05 Kappa: 0.239 Mcnemar's Test P-Value : < 2.2e-16 Sensitivity: 0.28473 Specificity: 0.91933 Pos Pred Value: 0.61752 Neg Pred Value: 0.73753 Prevalence : 0.31385 Detection Rate: 0.08936 Detection Prevalence: 0.14471 Balanced Accuracy: 0.60203 'Positive' Class : TRUE

### Calibration Plot for Model Y

plot(calibrate(modY\_si))

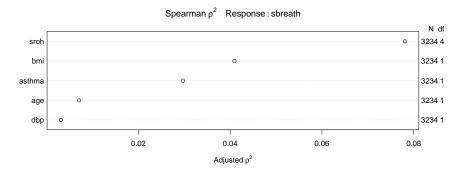


### Section 4

Model Z: The "Augmented" model

## Considering Non-Linear Terms

- Use 3-6 additional degrees of freedom to account for non-linearity, and add 1-3 non-linear terms.
- $\bullet$  We'll start with the Spearman  $\rho^2$  plot...



## Fitting an "Augmented" Model

We'll include the interaction of sroh and bmi, which will add 4 df, and a restricted cubic spline with three knots in bmi (which will add one more), and stop there.

# modZ\_si results (from 1rm fit)

```
> modZ si
Logistic Regression Model
 1rm(formula = sbreath ~ age + sroh + rcs(bmi, 3) + sroh %ia%
     bmi + dbp + asthma, data = nh_today_i, x = TRUE, y = TRUE)
                       Model Likelihood
                                               Discrimination
                                                                 Rank Discrim.
                             Ratio Test
                                                      Indexes
                                                                       Indexes
0bs
             3234
                     IR chi2
                                 428.25
                                               R_2
                                                        0.174
                                                                         0.719
 0
             2219
                     d.f.
                                     13
                                             R2(13.3234)0.120
                                                                 Dxy
                                                                         0.438
             1015
                     Pr(> chi2) < 0.0001
                                           R2(13,2089,3)0,180
                                                                 aamma
                                                                         0.438
 max Iderivl 4e-09
                                               Brier
                                                        0 187
                                                                         0.189
                                                                 tau-a
              Coef
                      S.E.
                             Wald Z Pr(>|Z|)
              -4.2765 1.1558 -3.70 0.0002
 Intercept
               0.0148 0.0039 3.75
                                    0.0002
 age
 sroh=VG
               0.4052 1.1722 0.35
                                    0.7295
 sroh=G
               1.5974 1.0686 1.49
                                    0.1350
               2.1630 1.0985 1.97
                                    0.0489
sroh=F
               2.9370 1.2318 2.38
                                    0.0171
sroh=P
 bmi
               0.0596 0.0387 1.54
                                    0.1235
 bmi'
               0.0095 0.0207 0.46
                                    0.6461
sroh=VG * bmi 0.0023 0.0402 0.06 0.9539
sroh=G * bmi -0.0151 0.0368 -0.41 0.6811
              -0.0100 0.0376 -0.27 0.7907
sroh=F * bmi
              -0.0323 0.0414 -0.78 0.4352
sroh=P * bmi
              -0.0061 0.0038 -1.62 0.1047
 dbp
                                    < 0.0001
 asthma=Yes
               0.8017 0.1111 7.21
```

### glm version of Model Z

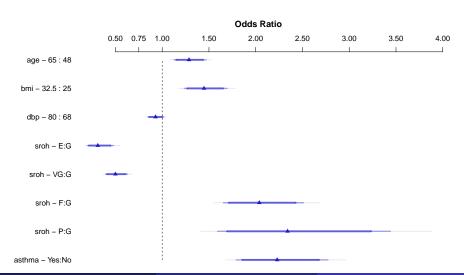
```
modZ_si_q
Call: glm(formula = sbreath ~ age + sroh + rcs(bmi, 3) + sroh %ia%
   bmi + dbp + asthma, family = binomial(link = logit), data = nh today i)
Coefficients:
               (Intercept)
                                                                              srohVG
                                                                                                            srohG
                 -4 276453
                                               0.014819
                                                                            0 405248
                                                                                                         1 597374
                                                                      rcs(bmi, 3)bmi
                                                                                                  rcs(bmi, 3)bmi'
                     srohE
                                                  srohP
                  2.163030
                                               2.937031
                                                                            0.059582
                                                                                                         0.009525
sroh %ia% bmisroh=VG * bmi
                             sroh %ia% bmisroh=G * bmi
                                                          sroh %ia% bmisroh=F * bmi
                                                                                       sroh %ia% bmisroh=P * bmi
                  0.002326
                                              -0.015125
                                                                           -0.009987
                                                                                                        -0.032268
                                              asthmaYes
                       dbp
                 -0.006146
                                               0.801692
Degrees of Freedom: 3233 Total (i.e. Null); 3220 Residual
Null Deviance:
                    4024
Residual Deviance: 3596
                                AIC: 3624
```

# Tidied Table of Model Z (Exponentiated) Coefficients

term	estimate	std.error	low90	high90
(Intercept)	0.014	1.156	0.002	0.092
age	1.015	0.004	1.008	1.022
srohVG	1.500	1.172	0.215	10.758
srohG	4.940	1.068	0.836	30.246
srohF	8.697	1.098	1.402	55.705
srohP	18.860	1.232	2.436	147.704
rcs(bmi, 3)bmi	1.061	0.039	0.995	1.132
rcs(bmi, 3)bmi'	1.010	0.021	0.976	1.045
sroh %ia% bmisroh=VG * bmi	1.002	0.040	0.937	1.072
sroh %ia% bmisroh=G * bmi	0.985	0.037	0.926	1.049
sroh %ia% bmisroh=F * bmi	0.990	0.038	0.930	1.055
sroh %ia% bmisroh=P * bmi	0.968	0.041	0.904	1.038
dbp	0.994	0.004	0.988	1.000
asthmaYes	2.229	0.111	1.857	2.677

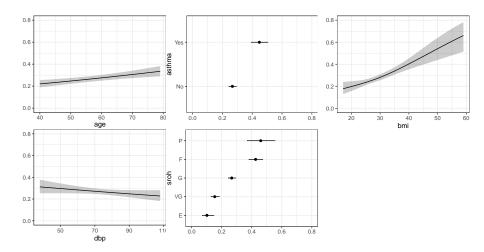
### Model Z Effects Plot on Odds Ratio Scale

plot(summary(modZ\_si))



### Prediction Plot for Model Z

### ggplot(Predict(modZ\_si, fun = plogis))



### Confusion Matrix for Model Z

How well does our Model Z classify subjects using a decision rule at 0.5?

# pred 0 1 Total Predict No SB 2040 725 2765 Predict SB 179 290 469 Total 2219 1015 3234

sbreath

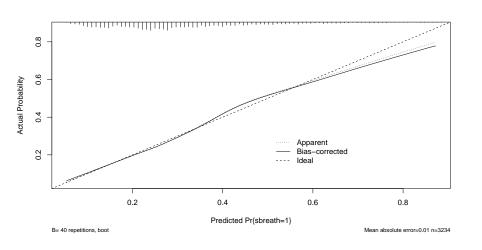
# Key Summaries of Classification Accuracy

# sbreath pred 0 1 Total Predict No SB 2040 725 2765 Predict SB 179 290 469 Total 2219 1015 3234

- **Sensitivity** is only slightly changed, to 290 / 1015 = 0.286
  - 28.6% of those who actually were short of breath are predicted to be short of breath.
- **Specificity** is still 2040 / 2219 = 0.919
  - 91.9% of those who actually weren't short of breath were predicted not to be short of breath.
- Positive Predictive Value (PPV) is 290 / 469 = 0.618
  - Again, 61.8% of those predicted to be short of breath actually were short of breath.

### Calibration Plot for Model Z

plot(calibrate(modZ\_si))



### Section 5

Summarizing/Presenting a Final Model

# Compare Models Y and Z on Key Summaries

```
temp1 <- bind_rows(glance(modY_si_g), glance(modZ_si_g)) |>
   mutate(model = c("Y", "Z")) |>
   select(model, AIC, BIC)

temp2 <- tibble(model = c("Y", "Z"),
   auc = c(modY_si$stats["C"], modZ_si$stats["C"]),
   r2_nag = c(modY_si$stats["R2"], modZ_si$stats["R2"]))

left_join(temp1, temp2, by = "model") |> kable()
```

model	AIC	BIC	auc	r2_nag
Y	3615.664	3670.398	0.7181358	0.1735251
Z	3623.808	3708.949	0.7188002	0.1742317

# ANOVA comparing Model Y to Z

> anova(modZ_si) Wald Statistics	Response: st	oreath	1
Factor	Chi-Square	d.f.	Р
age	14.08	1	0.0002
sroh (Factor+Higher Order Factors)	165.26	8	<.0001
All Interactions	1.79	4	0.7745
bmi (Factor+Higher Order Factors)	79.65	6	<.0001
All Interactions	1.79	4	0.7745
Nonlinear	0.21	1	0.6461
sroh * bmi (Factor+Higher Order Factors	) 1.79	4	0.7745
dbp	2.63	1	0.1047
asthma	52.03	1	<.0001
TOTAL NONLINEAR + INTERACTION	1.89	5	0.8642
TOTAL	352.90	13	<.0001
>			

# Validating Model Summaries (code)

```
set.seed(432123)
valY \leftarrow validate(modY si, B = 40)
valZ <- validate(modZ_si, B = 40)</pre>
val_1 <- bind_rows(valY[1,], valZ[1,]) |>
  mutate(model = c("Y", "Z"),
         AUC nominal = 0.5 + (index.orig/2),
         AUC_validated = 0.5 + (index.corrected/2)) |>
  select (model, AUC nominal, AUC validated)
val_2 <- bind_rows(valY[2,], valZ[2,]) |>
  mutate(model = c("Y", "Z"),
         R2_nominal = index.orig,
         R2_validated = index.corrected) |>
  select(model, R2_nominal, R2_validated)
```

See next slide for the result.

# Validating Model Summaries

```
val <- left_join(val_1, val_2, by = "model")
val |> kable()
```

model	AUC_nominal	AUC_validated	R2_nominal	R2_validated
Y	0.7181438	0.7163131	0.1735251	0.1692934
Z	0.7187991	0.7132489	0.1742317	0.1632280

Which model should we choose?

# Describing a Meaningful Effect

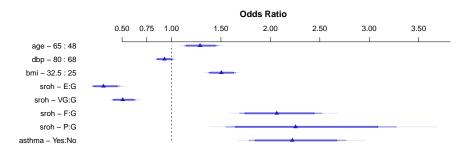
This is for you to do.

[W]rite a detailed and correct description of the effect of at least one predictor on your outcome for your chosen logistic regression model, providing all necessary elements of such a description, and link this directly to what the (effects) plot is telling you.

See Chapter 22 of the Notes for more details, and this is also the major task in several questions within Lab 4. The effects plot for Model Y is repeated in the next slide, and you'll want the actual summary as well as the plot so you can specify the numbers. We prefer you discuss a meaningful effect, should one exist. Pick an effect to describe that is interesting to you.

### Model Y Effects Plot on Odds Ratio Scale

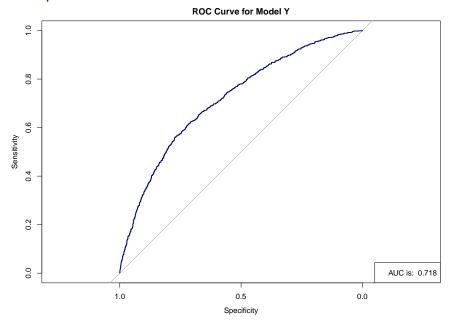
### plot(summary(modY\_si))



### ROC Calculations for Model Y

```
roc_modY <- roc(nh_today_i$sbreath ~
    predict(modY_si_g, type="response"), ci = TRUE)
roc modY
Call:
roc.formula(formula = nh today i$sbreath ~ predict(modY si g,
Data: predict(modY_si_g, type = "response") in 2219 controls
Area under the curve: 0.7181
95% CI: 0.6992-0.7371 (DeLong)
plot(roc modY, main = "ROC Curve for Model Y",
     lwd = 2, col = "salmon")
legend('bottomright',
   legend = paste("AUC is: ",round_half_up(auc(roc_modY),3)))
```

# ROC plot for Model Y

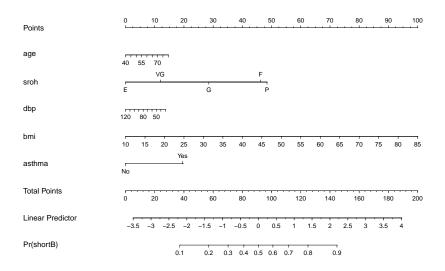


# Nomogram for the Final Model (Model Y)

#### Result on next slide.

- The final part of your summary of the final model should be a nomogram with a demonstration of a predicted probability associated with two new subjects of interest that differ in terms of some of the parameters in your model.
- Your predictions should describe two different subjects. You don't
  have to call them Harry and Sally, but it is helpful to give them actual
  names.

### Nomogram for Model Y



### Next Time

### More on Logistic Regression

- See section 20 for more on confusion matrices and ROC curves and some material on assessing assumptions through residual plots, all of which are in the context of logistic models fit with glm().
- See section 21 for more on using Spearman's  $\rho^2$  plot, Nagelkerke  $R^2$ , the C statistic, its relationship to Somers' d, validation and plotting the results, along with some thoughts on identifying influential points, mostly in the context of logistic models fit with lrm().
- See section 22 for some thoughts on estimating and interpreting effect sizes in logistic and in linear models. Some really useful tips to be found here.