## Fast method in evaluating density of multivariate Gaussian

▶ Given data  $x \in \mathbb{R}^d$ , the likelihood that it comes from a multivariate Gaussian density with mean vector  $\mu \in \mathbb{R}^d$  and covariance matrix  $\Sigma \in \mathbb{R}^{d \times d}$  is

$$\mathcal{N}(x;\mu,\Sigma) = \frac{1}{\sqrt{(2\pi)^d \det(\Sigma)}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$

- ▶ The most expensive part to compute this is to evaluate  $\Sigma^{-1}$ , which has a complexity  $\mathcal{O}(d^3)$ .
- Moreover, when  $\Sigma$  is rank-deficient, i.e., there are close-to-zero eigenvalues, computing  $\Sigma^{-1}$  will return NAN (you cannot invert the matrix)

- Now let's make it faster and avoid the numerical issues by compute using "low-rank approximation"
- $\Sigma = U\Lambda U^T \begin{tabular}{ll} ** the U is the eigen vector you need to take that instead of sigma. that was my mistake!!! \\ ** the U is the eigen vector you need to take that instead of sigma. The property of the proper$

where  $\Lambda = \mathsf{diag}\{\lambda_1, \dots, \lambda_d\}$  and the eigenvalues are ordered

$$\lambda_1 \geq \lambda_2 \geq \ldots \geq \lambda_d$$

▶ The rank-r approximation (r < d) of  $\Sigma$  is

$$\tilde{\Sigma} = \tilde{U} \tilde{\Lambda} \tilde{U}^T$$

where  $\tilde{U}$  is a d-by-r matrix formed by the first r columns of U,  $\tilde{\Lambda} = \text{diag}\{\lambda_1, \ldots, \lambda_r\}$ .

▶ Typically we will choose r such that at least  $\lambda_r > 0$ 

▶ Now compute transform of data and parameters

$$\tilde{x} = \tilde{U}^T x$$
$$\tilde{\mu} = \tilde{U}^T \mu$$

- ▶ Compute  $\tilde{\Lambda}^{-1} = \text{diag}\{\lambda_1^{-1}, \dots, \lambda_r^{-1}\}$
- Note that

$$\det(\Sigma) = \prod_{i=1}^{d} \lambda_i, \quad \det(\tilde{\Sigma}) = \prod_{i=1}^{r} \lambda_i$$

Finally, the density calculated by replacing  $\Sigma$  with  $\tilde{\Sigma}$  is:

$$\mathcal{N}(x; \mu, \Sigma) \approx \frac{1}{\sqrt{(2\pi)^d \prod_{i=1}^r \lambda_i}} \exp\left\{-\frac{1}{2} \sum_{i=1}^r \frac{(\tilde{x}_i - \tilde{\mu}_i)^2}{\lambda_i}\right\}$$

where  $\tilde{x}_i$  and  $\tilde{\mu}_i$  denote the ith entry of  $\tilde{x}$  and  $\tilde{\mu}$ , respectively.

Note: you can play with different r to have a good tradeoff between accuracy and speed Note that above we have used the following basic identity from linear algebra

$$\tilde{\Sigma}^{-1} = \tilde{U}\tilde{\Lambda}^{-1}\tilde{U}^T$$

and

$$(x - \mu)^T \Sigma^{-1} (x - \mu)$$
  
=  $(x - \mu)^T \tilde{U} \tilde{\Lambda}^{-1} \tilde{U}^T (x - \mu)$ 

$$= [\tilde{U}^T(x-\mu)]^T \tilde{\Lambda}^{-1} [\tilde{U}^T(x-\mu)]$$
$$= [\tilde{x} - \tilde{\mu}]^T \tilde{\Lambda}^{-1} [\tilde{x} - \tilde{\mu}]$$

$$= [x - \mu]^T \Lambda^{-1} [x - \mu_i]^2$$
 $= \sum_{i=1}^r \frac{(\tilde{x}_i - \tilde{\mu}_i)^2}{\lambda_i}$ 

## Avoiding numerical issues in GMM-EM

Note that in evaluating E-step

$$\tau_k^i = \frac{\pi_k \mathcal{N}(x_i | \mu_k, \Sigma_k)}{\sum_{k'=1}^K \pi_{k'} \mathcal{N}(x_i | \mu_{k'}, \Sigma_{k'})}$$

where the normal distributional density  $\mathcal{N}(\cdot|\cdot,\cdot)$  appeared both in numerical and denominator

Multivariate normal density

$$\mathcal{N}(X|\mu_k, \Sigma_k) := \frac{1}{|\Sigma|^{1/2} (2\pi)^{d/2}} \exp\left(-\frac{1}{2} (X - \mu)^T \Sigma^{-1} (X - \mu)\right)$$

The term  $(2\pi)^{d/2}$  can be very large when d is large

- ▶ So we can simplify the calculation without calculating  $(2\pi)^{d/2}$
- ▶ Steps to evaluate E-step

For  $k = 1, \dots, K$ 

Use low-rank approximation to compute

$$m_k = (X - \mu_k)^T \Sigma_k^{-1} (X - \mu_k)$$
  
 $d_k = |\Sigma_k|^{-1/2}$ 

Compute

$$\hat{\tau}_k^i = \pi_k d_k \exp\left(-\frac{1}{2}m_k\right)$$

## **Normalize**

$$C = \sum_{k=1}^{K} \hat{\tau}_k^i$$
$$\tau_k^i = \hat{\tau}_k^i / C$$