

# MISSING DATA: TREE-BASED MULTIPLE IMPUTATION METHODS

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## Motivation: Valid Inference under Missingness

Proper multiple Imputation (MI) aims at valid inference with incomplete data by accounting for total variance. This mitigates information loss due to missing Data. Flexible methods such as PMM, CART, and miceRanger are widely used in complex, non-linear settings, instead of standard MICE[1]. In our simulation study, we examine whether these flexible methods are proper.[2]

**Good predictions  $\neq$  Valid inference**

## What is "proper" Multiple Imputation (MI)

MI aims at **valid inference  $\rightarrow$  correct total uncertainty  $T_j$**  for each Parameter  $j$ :

$$T_j = W_j + (1 + 1/M)B_j \quad (1)$$

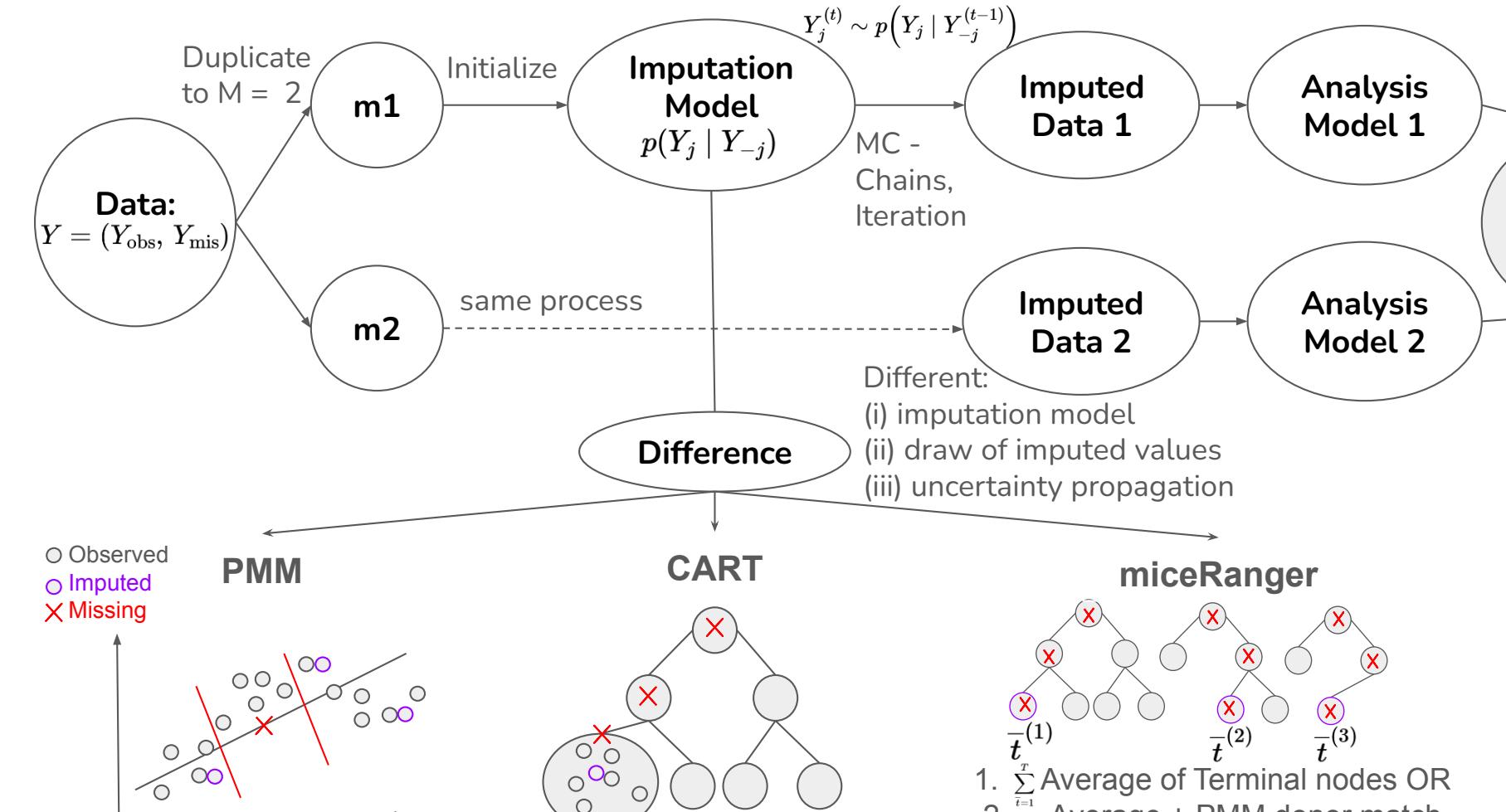
$W_j$  = within-imputation variance (analysis model)

$B_j$  = between-imputation variance (missing-data uncertainty).

**One sentence:** MI is *proper* if it captures the correct total uncertainty  $T_j$ , so pooled SEs and confidence intervals are valid.

## Models and Methodological Differences

MICE follows a Markov chain (FCS) procedure with the idea of “perturb everything” to reflect missing-data uncertainty. The graphic below displays the process of imputing with standard mice and where the methods we investigate differ.[2][3][1]



**PMM, CART, RF :** more flexible model assumptions. Easier + better nonlinear.

**Distinction:** Imputation Model  $\neq$  Analysis Model

- **PMM:**[4][6][7] Fit  $Y_{\text{mis}} \sim Y_{\text{obs}}$   $\rightarrow$  compute  $\hat{y}_i$   $\rightarrow$  select the  $k$  donors minimizing  $d_{il} = |\hat{y}_i - y_l|$   $\rightarrow$  randomly draw observed  $y_l$ .
- **CART (only the differing step):** Fit tree  $Y_{\text{mis}} \sim Y_{\text{obs}}$   $\rightarrow$  route case  $i \rightarrow$  sample observed  $y$  from the terminal node where it lands (not from other nodes).
- **miceRanger:**[4][5]
  - Works alongside mice; pooling via Rubin's rules.
  - Fit RF  $Y_{\text{mis}} \sim Y_{\text{obs}}$   $\rightarrow$  average tree predictions  $\rightarrow$  directly impute ensemble prediction (no donor sampling by default).

Figure 1: Heat map of the average missing entries depending on  $X_9$  and  $X_{10}$ . An  $i$ -th grid square contains  $a_i$  values, of which  $b_i$  are NA. The colour visualises the averaged  $\frac{b_i}{a_i}$  over several NA simulations. The existence of a color gradient implies a MAR-process.

## Data Generating Process

The focus lays on a multivariate Normaldistribution with  $\forall k : X_k \sim \mathcal{N}(0, 1); \varepsilon \sim \mathcal{N}(0, 1)$ . The model is based on that of [6] and [7].

$$Y = 0.5 \cdot \sum_{i=1}^{10} X_i + 0.5 \cdot X_3^2 + X_1 X_2 + X_4 X_5 + \varepsilon \quad (2)$$

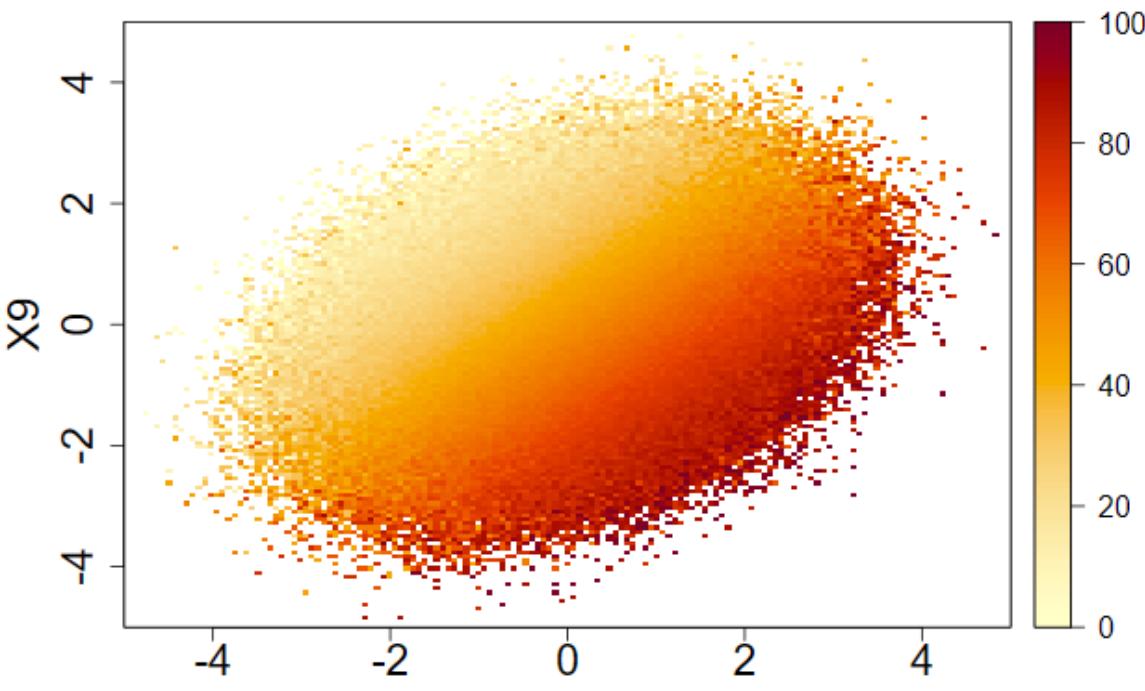
with correlation  $\rho$  given as:

$$\rho(X_i, X_j) = \begin{cases} 0.5, & \text{falls } i \neq j; \quad i, j \in \{1; \dots; 5\} \\ 0.3, & \text{falls } i \neq j; \quad i, j \in \{6; \dots; 10\} \end{cases} \quad (3)$$

The simulation runs with following parameters:  $n = 1000$ ;  $M = 10$ ;  $R = 300$  and 25 iterations

## Missingness Model

The generation of the Missing Data follows a Missing at Random (MAR)-mechanism. The missingness indicator  $R_l$  is determined by a Bernoulli function, which probability depends on a logit function with arguments  $X_9$  and  $X_{10}$ .



$X_{i,l}$  depicts a value of the  $i^{th}$  column and  $l^{th}$  row. As recommended in [8], the parameters were chosen such that it results in a missingness of around 34% in the data.

$$z_l = -0.85 + 0.5 \cdot (X_{4,l} - X_{5,l}) \quad (4)$$

$$\sigma(z_l) = \frac{1}{1 - \exp(z_l)} \quad (5)$$

$$R_l \sim \text{Ber}(\sigma(z_l)) \quad (6)$$

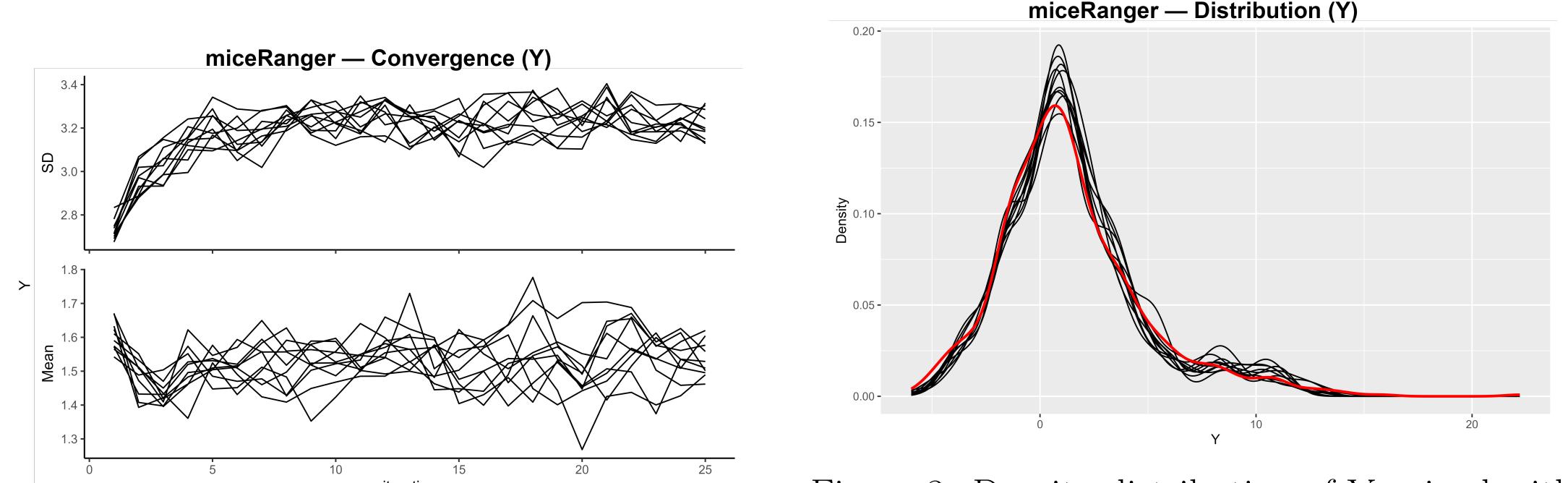


Figure 2: Convergence diagnostic of  $Y$  gained with miceRanger for one simulation. After a small Burn-in of around 5 iterations, the chains dont show systematic divergence indicating visual evidence for meaningful results. PMM and CART depict similar behaviour.

## Results

- As you can see in figure 5 CART and RF less biased for interactions (still biased) and quadratic terms than PMM
- CART overestimates Between-Variance

$$B = \frac{1}{M-1} \sum_{m=1}^M (\hat{\theta}^{(m)} - \bar{\theta}_{\text{MI}})^2$$

- RF underestimates B

### Confidence Intervals

$$CI_{\text{MI}} = \bar{\theta}_{\text{MI}} \pm t_{\nu, 1-\alpha/2} \sqrt{W + \left(1 + \frac{1}{M}\right) B} \quad (7)$$

of RF too narrow

As you can see in figure 6 CART and RF achieve higher coverage rates

$$\widehat{\text{Coverage}} = \frac{1}{R} \sum_{r=1}^R \mathbb{1}_{\theta \in CI_{\text{MI}}^{(r)}} \quad (8)$$

or non-linear parameter, but still under-coverage

MI Performance and Properness Summary			
Predictive Accuracy and Rubin Diagnostics			
Metric	rf_ranger	cart	pmm
J RMSE	2.106	2.247	2.412
1 R <sup>2</sup>	0.663	0.608	0.537
Mean SE	0.064	0.108	0.116
$\lambda$	0.301	0.578	0.547
FMI	0.244	0.406	0.391
DF	118.700	31.100	34.700
Var (B)	0.001	0.007	0.007

Figure 4

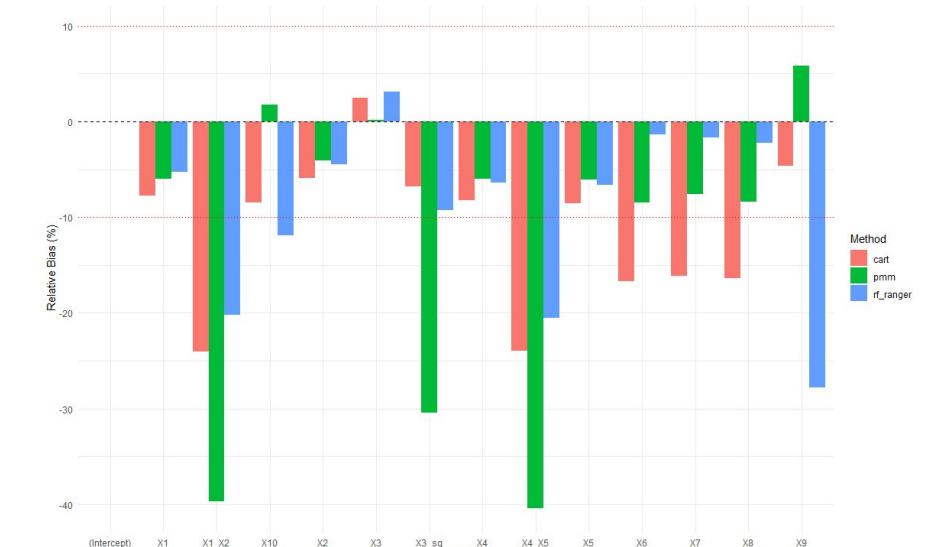


Figure 5: Comparison of rel. bias by parameter

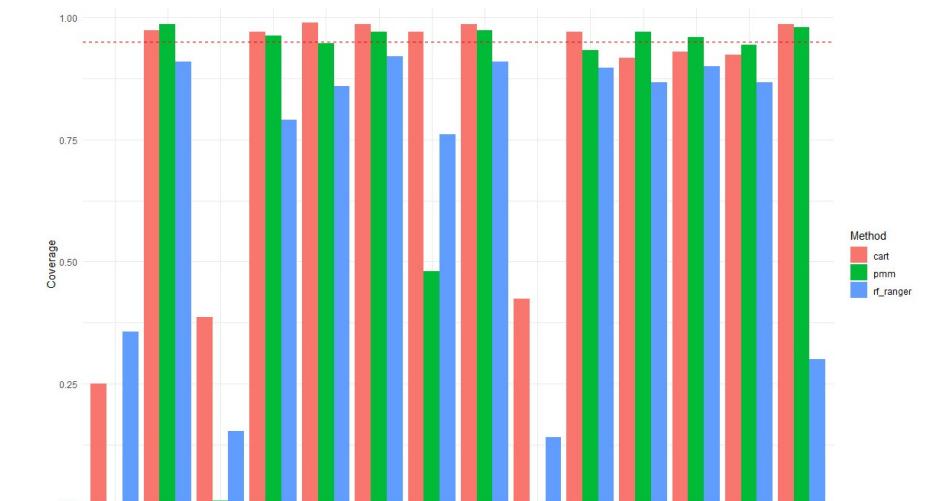


Figure 6: Comparison of coverage rates

## Conclusion

CART and RF can lead to lower bias and higher coverage rates for imputations of non-linear parameter. However, in this case, PMM, CART and RF are not 'proper' MI methods.

## References

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