

# Quiz: Directed Graphical Models & Inference

STA414/2104 - Winter 2026

## 1. HMMs with Identity Transition

In the Forward Algorithm for Hidden Markov Models (HMMs), suppose the transition matrix  $A$  is the identity matrix ( $A = I$ ), such that  $A_{ii} = 1$  and  $A_{ij} = 0$  for  $i \neq j$ .

Which of the following statements best describes the behavior of the filtered marginal  $\alpha_t(j) = p(z_t = j | x_{1:t})$ ?

- (a) It depends only on the current observation  $x_t$ , ignoring all past observations.
- (b) It is proportional to the product of the initial probability  $\pi(j)$  and the emission probabilities for that state across all time steps up to  $t$ .
- (c) It remains constant over time, regardless of the observations  $x_{1:t}$ .
- (d) It becomes a uniform distribution because the state cannot change.

*Correct Answer: (b)*

*Rationale: The Forward Algorithm updates the filtered marginal recursively:*

$$\alpha_t(j) \propto p(x_t | z_t = j) \sum_i p(z_t = j | z_{t-1} = i) \alpha_{t-1}(i)$$

*With  $A = I$ , the transition term  $p(z_t = j | z_{t-1} = i)$  is 1 if  $i = j$  and 0 otherwise. This simplifies the recursion to:*

$$\alpha_t(j) \propto p(x_t | z_t = j) \alpha_{t-1}(j)$$

*Unrolling this recursion back to  $t = 1$ , we get:*

$$\alpha_t(j) \propto p(z_1 = j) \prod_{k=1}^t p(x_k | z_k = j)$$

*Thus, the belief state for  $j$  is proportional to the initial prior times the likelihood of generating the entire observed sequence from that single, constant state  $j$ .*

## 2. Rejection Sampling Acceptance Probability

Recall from the lecture that the probability of accepting a specific candidate sample  $x$  (where  $x \sim q(x)$  and  $u|x \sim \text{Unif}[0, cq^*(x)]$ ) is given by:

$$\mathbb{P}(u \leq p^*(x)|x) = \frac{p^*(x)}{cq^*(x)}$$

where  $p(x) = \frac{p^*(x)}{Z_p}$  and  $q(x) = \frac{q^*(x)}{Z_q}$ .

Given this formula, what is the probability of acceptance,  $\mathbb{P}(\text{accept proposal } x)$ ?

- (a)  $\frac{p^*(x)}{cq^*(x)}$
- (b)  $\frac{1}{c}$
- (c)  $\frac{Z_p}{cZ_q}$
- (d)  $c\frac{Z_p}{Z_q}$

*Correct Answer: (c)*

*Rationale: The marginal probability of acceptance is found by integrating the conditional acceptance probability with respect to the proposal distribution  $q(x)$ :*

$$\mathbb{P}(\text{accept}) = \int \mathbb{P}(\text{accept}|x)q(x)dx = \int \frac{p^*(x)}{cq^*(x)}q(x)dx$$

*Substituting the definitions  $p^*(x) = Z_p p(x)$  and  $q^*(x) = Z_q q(x)$ :*

$$\mathbb{P}(\text{accept}) = \int \frac{Z_p p(x)}{cZ_q q(x)} q(x)dx = \frac{Z_p}{cZ_q} \int p(x)dx$$

*Since  $\int p(x)dx = 1$ , the final acceptance probability is  $\frac{Z_p}{cZ_q}$ .*

### 3. Quantitative Curse of Dimensionality

Consider a Rejection Sampling setup where the target  $p(\mathbf{x})$  and proposal  $q(\mathbf{x})$  factorize over  $D$  independent dimensions (i.e.,  $p(\mathbf{x}) = \prod_{i=1}^D p(x_i)$ ) (with the same marginal distribution for all  $x_i$ ). Suppose that for a single dimension ( $D = 1$ ), the optimal constant  $c$  results in an acceptance rate of 50% (0.5).

If you use this same sampling strategy in  $D = 10$  dimensions, what is the expected acceptance rate?

- (a) It remains 50% because the dimensions are independent.
- (b) 5%, decreasing linearly with dimension ( $0.5 \times 10^{-1}$ ).
- (c)  $\approx 0.1\%$ , because the acceptance rate scales as  $0.5^{10}$ .
- (d) 0%, because rejection sampling is theoretically impossible in dimensions  $D > 3$ .

*Correct Answer: (c)*

*Rationale:* The acceptance rate is given by  $\frac{Z_p}{cZ_q}$ . In this independent setup, the total acceptance probability is the product of the acceptance probabilities for each dimension. Alternatively, the constant  $c$  required to bound the joint distribution is the product of the constants for each dimension:  $c_{\text{total}} = (c_1)^D$ . Since the acceptance rate in 1D is  $1/c_1 = 0.5$ , the acceptance rate in  $D$  dimensions is  $(1/c_1)^D = 0.5^{10} \approx \frac{1}{1024} \approx 0.1\%$ . This illustrates why acceptance rates are “exponentially small in dimension”.

#### 4. Efficiency of Rejection Sampling

You are designing a Rejection Sampling algorithm to sample from a target distribution  $p(x)$  using a proposal distribution  $q(x)$ . Recall that you must find a constant  $c$  such that  $cq(x) \geq p(x)$  for all  $x$ .

Consider two different proposal distributions:

- **Proposal A:**  $q_A(x)$  is a “broad” distribution (e.g., a Gaussian with very large variance) that covers  $p(x)$  but is much flatter.
- **Proposal B:**  $q_B(x)$  is a “spiky” distribution (e.g., a Gaussian with very small variance) that matches the peak of  $p(x)$  well but decays much faster than  $p(x)$  in the tails.

Which of the following outcomes is most likely?

- (a) **Proposal A** is valid but inefficient because the required constant  $c$  will be very large, resulting in a high rejection rate.
- (b) **Proposal B** is the best choice because it fits the high-probability regions of  $p(x)$  tightly, minimizing the rejection rate.
- (c) **Proposal B** is valid, but it will result in biased samples because it ignores the tails of  $p(x)$ .
- (d) **Proposal A** is invalid because a flat distribution cannot upper-bound a peaked distribution.

*Correct Answer: (a)*

*Rationale: Rejection sampling requires  $cq(x) \geq p(x)$  everywhere.*

- For **Proposal B** (tight/light-tailed), if  $q(x)$  decays faster than  $p(x)$  in the tails, the ratio  $p(x)/q(x)$  goes to infinity as  $|x| \rightarrow \infty$ . This makes it impossible to find a finite constant  $c$  that satisfies the inequality everywhere, rendering **Proposal B invalid**.
- For **Proposal A** (broad/heavy-tailed), the inequality can be satisfied. However, because  $q_A(x)$  is much flatter than  $p(x)$ , the gap between  $cq_A(x)$  and  $p(x)$  will be massive in most regions. Since the acceptance rate is  $1/c$  (assuming normalized densities) or  $Z_p/cZ_q$ , a large  $c$  leads to extremely rare acceptance.

*Thus, Proposal A is valid but computationally wasteful.*