

## Quiz 1 - Solutions

### Maximum Likelihood and Sufficient Statistics

1. Let  $\hat{\theta}_{\text{MLE}}$  be the maximum likelihood estimator for a parameter  $\theta$ . If  $g$  is a one-to-one function, what is the maximum likelihood estimator for  $\eta = g(\theta)$ ?

**Solution:**  $g(\hat{\theta}_{\text{MLE}})$ .

**Derivation:** Recall the definition of the MLE for the original parameter  $\theta$ :

$$\hat{\theta}_{\text{MLE}} = \underset{\theta}{\operatorname{argmax}} \mathcal{L}(\theta; \mathcal{D})$$

where  $\mathcal{L}(\theta; \mathcal{D}) = p(\mathcal{D}|\theta)$  is the likelihood function, given the observations  $\mathcal{D}$ .

Now consider the re-parameterization  $\eta = g(\theta)$ . Since  $g$  is one-to-one, we can invert it to write  $\theta = g^{-1}(\eta)$ . The likelihood function for the new parameter  $\eta$ , denoted  $\mathcal{L}^*(\eta; \mathcal{D})$ , is simply the original likelihood evaluated at the corresponding  $\theta$ :

$$\mathcal{L}^*(\eta; \mathcal{D}) = \mathcal{L}(g^{-1}(\eta); \mathcal{D})$$

The MLE for  $\eta$  is defined as the value that maximizes this new likelihood:

$$\hat{\eta}_{\text{MLE}} = \underset{\eta}{\operatorname{argmax}} \mathcal{L}^*(\eta; \mathcal{D}) = \underset{\eta}{\operatorname{argmax}} \mathcal{L}(g^{-1}(\eta); \mathcal{D})$$

Since  $\mathcal{L}(\theta; \mathcal{D})$  achieves its maximum at  $\theta = \hat{\theta}_{\text{MLE}}$ , the composite function  $\mathcal{L}(g^{-1}(\eta); \mathcal{D})$  achieves its maximum when the input to  $\mathcal{L}$ , which is  $g^{-1}(\eta)$ , equals  $\hat{\theta}_{\text{MLE}}$ :

$$g^{-1}(\hat{\eta}_{\text{MLE}}) = \hat{\theta}_{\text{MLE}} \implies \hat{\eta}_{\text{MLE}} = g(\hat{\theta}_{\text{MLE}})$$

2. Let  $X_1, \dots, X_n$  be i.i.d. uniform random variables on the interval  $[0, \theta]$ . Which of the following is the sufficient statistic for  $\theta$ ?

**Solution:** The maximum of the sample  $X_{(n)} = \max(X_1, \dots, X_n)$ .

The likelihood function for a Uniform distribution on  $[0, \theta]$  is:

$$L(\theta) = \prod_{i=1}^n \frac{1}{\theta} \mathbb{I}(0 \leq x_i \leq \theta) = \frac{1}{\theta^n} \mathbb{I}(0 \leq \max(x_i) \leq \theta)$$

By the Fisher-Neyman Factorization Theorem, the likelihood depends on the data  $x$  only through the maximum value  $X_{(n)}$ . Therefore,  $X_{(n)}$  is the sufficient statistic.

3. If  $T(X)$  is a sufficient statistic for  $\theta$ , which of the following statements regarding the Maximum Likelihood Estimator (MLE)  $\hat{\theta}$  is always true?

**Solution:** The MLE  $\hat{\theta}$  is always a function of sufficient statistic  $T(X)$ .

By the Factorization Theorem, the likelihood can be written as  $\mathcal{L}(\theta; X) = h(X)g_\theta(T(X))$ . When maximizing  $\mathcal{L}(\theta; X)$  with respect to  $\theta$ , the term  $h(X)$  is constant and does not influence the location of the maximum. The maximization depends solely on  $g_\theta(T(X))$ . Therefore, the resulting estimator  $\hat{\theta}$  will depend on the data only through  $T(X)$ .

4. Which of the following probability distributions does not belong to the one-parameter Exponential Family?

**Solution:** Uniform distribution on the interval  $[0, \theta]$  for  $\theta > 0$ .

Either we recognize the other families are exponential families, or we use the following fact. For a distribution to belong to the exponential family, its support (the set of  $x$  with density satisfying  $p(x) > 0$ ) must not depend on the parameter  $\theta$ . Indeed, the parameter only intervenes in an exponential factor that cannot be equal to 0.  $p(x)$  being equal to 0 is then equivalent to the other factor being equal to 0, and it does not depend on  $\theta$ .

- Poisson, Bernoulli, and Normal (with known variance) have supports that are independent of their parameters.
- The Uniform distribution on  $[0, \theta]$  has support  $0 \leq x \leq \theta$ , which depends directly on  $\theta$ . Thus, it is not in the exponential family.

5. Consider a model for three binary random variables  $(x_1, x_2, x_3)$  where  $x_i \in \{0, 1\}$ . The joint probability mass function is given by:

$$p(x_1, x_2, x_3) = \frac{1}{Z(\theta)} \exp(\theta(x_1 x_2 + x_2 x_3))$$

Which of the following represents the sufficient statistic for  $\theta$ ?

**Solution:**  $T(x_1, x_2, x_3) = x_1 x_2 + x_2 x_3$ .

This distribution is already written in the canonical exponential family form:

$$p(x|\theta) = h(x) \exp(\eta(\theta) \cdot T(x) - A(\theta))$$

Here, the natural parameter is  $\theta$ , and the term multiplying it in the exponent is the sufficient statistic  $T(x) = x_1 x_2 + x_2 x_3$ . We also have  $h(x) = 1$  and  $A(\theta) = \log Z(\theta)$ .