1

Random Numbers

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Abstract—This solution manual provides solutions and link to codes used for generation of random numbers.

1 Uniform Random Variables

1.1 Generate 10^6 samples of U using a C program and save into a file called uni.dat .

Solution:

Download the following C code and run it to generate samples of U.

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/1.1/1.1.c

1.2 Load the uni.dat file into python and plot the empirical CDF of *U* using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr\left(U \le x\right) \tag{1.1}$$

Solution: Code used to plot empirical CDF of *U*

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/1.2/1.2.py

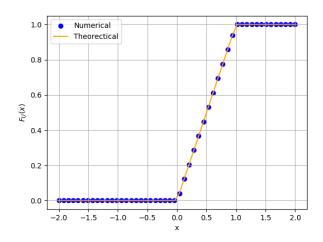


Fig. 1.2: CDF of *U*

1.3 Find a theoretical expression for $F_U(x)$. Solution:

$$F_U(x) = \begin{cases} 0 & x < 0 \\ x & 0 \le x \le 1 \\ 1 & x > 1 \end{cases}$$
 (1.2)

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/1.3/1.3.py

1.4 Find mean and variance of *U*.

Solution:

Mean of Random Variable U is given by,

$$E[U] = \frac{1}{N} \sum_{i=1}^{N} U_i$$
 (1.3)

and its Variance is given by,

$$E[[U - E[U]]^2] = E[U^2] - [E[U]]^2$$
 (1.4)

Using the above two formulas, we get,

Mean of
$$U = 0.500031$$
 (1.5)

Variance of
$$U = 0.083247$$
 (1.6)

Download and run the C code for mean and variance

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/1.4/1.4.c

1.5 Verify your result theoretically that

$$E[U^k] = \int_{-\infty}^{\infty} x^k dF_U(x) \tag{1.7}$$

Solution:

For k = 1,

$$E[U] = \int_{-\infty}^{\infty} x dF_U(x)$$
 (1.8)

$$E[U] = \int_{0}^{1} x d(x)$$
 (1.9)

$$E[U] = 0.5 (1.10)$$

Similarly, for k = 2,

$$E[U^2] = 0.3333 \tag{1.11}$$

Variance =
$$0.0833$$
 (1.12)

Thus the simulated and theoretical values of mean and variance of U are approximately equal.

2 Central Limit Theorem

2.1 **Solution:**

Download the following C code and run it to generate samples of X.

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/2.1/2.1.c

2.2 Solution:

The plot was generated by running the following Python code

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/2.2./2.2.py

The CDF is a non-decreasing function with its range between 0 and 1. It will be continuous if PDF is finite.

2.3 Load gau.dat in python and plot the empirical PDF of *X* using the samples in gau.dat. The PDF of *X* is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \tag{2.1}$$

What properties does the PDF have?

Solution:

The empirical PDF of X is plotted by the Python code

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/2.3/2.3.py The PDF takes non-negative values and area under its curve is 1.

2.4 Find the mean and variance of *X* by writing a C program.

Solution:

Run the following C file

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/2.4/2.4.c

$$E[X] = 0.000630 \tag{2.2}$$

$$var[X] = 1.000149$$
 (2.3)

2.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, (2.4)$$

repeat the above exercise theoretically.

Solution:

$$E[X] = \int_{-\infty}^{\infty} x p_X(x) dx$$
 (2.5)

$$E[X] = \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \qquad (2.6)$$

Taking $\frac{x^2}{2} = t$,

$$E[X] = -\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp(-t) dt \qquad (2.7)$$

$$E[X] = 0 (2.8)$$

To calculate variance,

$$var[X] = E[(X - E[X])^{2}]$$
 (2.9)

$$var\left[X\right] = E\left[X^2\right] \tag{2.10}$$

$$var[X] = \int_{-\infty}^{\infty} x^2 p_X(x) dx$$
 (2.11)

$$var[X] = \int_{-\infty}^{\infty} x^2 \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \quad (2.12)$$

We know that,

$$\int_{-\infty}^{\infty} x^2 \exp\left(-\frac{x^2}{2}\right) = \sqrt{2\pi}$$
 (2.13)

$$var\left[X\right] = 1\tag{2.14}$$

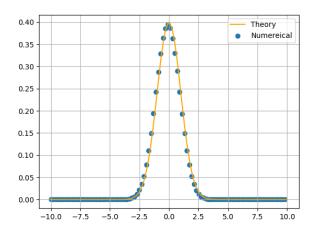


Fig. 2.5: PDF of *X*

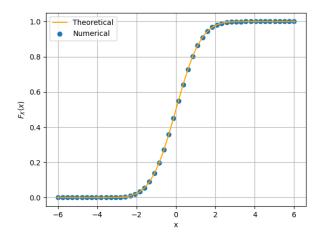


Fig. 2.5: CDF of *X*

Theoretical expression for X,

$$F_X(x) = 1 - Q(\frac{x - E[X]}{var[X]})$$
 (2.15)

$$F_X(x) = 1 - Q(x) (2.16)$$

The Python codes plot the CDF and PDF of X

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/2.5/2.5.py

wget https://github.com/TYCN129/AI1110-Assignments/blob/main/Manual %201/2.5/2.5 2.py

3 From Uniform to Other

3.1 Generate samples of

$$V = -2\ln(1 - U) \tag{3.1}$$

and plot its CDF.

Solution:

Download and run the following code to generate samples of V

wget https://github.com/TYCN129/AI1110– Assignments/blob/main/Manual %201/3.1/3.1.c

The following Python code plots CDF of V

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/3.1/3.1.py

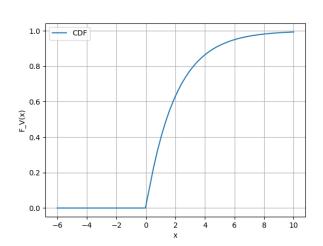


Fig. 3.1: The empirical CDF of V

3.2 Find a theoretical expression for $F_V(x)$. Solution:

$$F_V(x) = Pr(V \le x) \tag{3.2}$$

$$F_V(x) = Pr(-2\ln(1-U) \le x)$$
 (3.3)

$$F_V(x) = Pr\left(\ln\left(1 - U\right) \ge -\frac{x}{2}\right) \tag{3.4}$$

$$F_V(x) = Pr\left(1 - U \ge \exp\left(-\frac{x}{2}\right)\right)$$
 (3.5)

$$F_V(x) = Pr\left(U \le 1 - \exp\left(-\frac{x}{2}\right)\right)$$
 (3.6)

$$F_V(x) = F_U\left(1 - \exp\left(-\frac{x}{2}\right)\right) \tag{3.7}$$

$$F_{V}(x) = \begin{cases} 0 & 1 - \exp\left(-\frac{x}{2}\right) < 0\\ 1 - \exp\left(-\frac{x}{2}\right) & 0 \le 1 - \exp\left(-\frac{x}{2}\right) \le 1\\ 1 & 1 - \exp\left(-\frac{x}{2}\right) > 1 \end{cases}$$
(3.8)

This simplifies to

$$F_V(x) = \begin{cases} 0 & x < 0 \\ 1 - \exp\left(-\frac{x}{2}\right) & x \ge 0 \end{cases}$$
 (3.9)

The following python code plots the theoritical CDF

wget https://github.com/TYCN129/AI1110-Assignments/blob/main/Manual %201/3.2/3.2.py

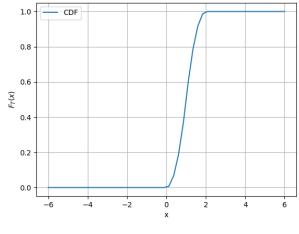


Fig. 4.2: CDF of *T*

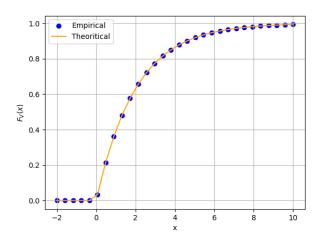


Fig. 3.2: CDF of *V*

4 Triangular Distribution

4.1 Generate

$$T = U_1 + U_2 (4.1)$$

Solution:

Download and run the following C code to generate tri.dat file.

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/4.1/4.1.c

4.2 Find the CDF of T.

Solution:

The following code plots the CDF of T

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/4.2/4.2.py

4.3 Find the CDF of *T*. **Solution:**

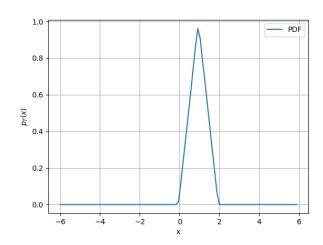


Fig. 4.3: PDF of *T*

The following code plots the PDF of T

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/4.3/4.3.py

4.4 Find the theoritical expressions for PDF and CDF of *T*.

Solution:

$$p_T(x) = p_{U_1 + U_2}(x) = p_{U_1}(x) * p_{U_2}(x)$$
 (4.2)

$$p_T(x) = \int_{-\infty}^{\infty} p_{U_1}(\tau) p_{U_2}(x - \tau) d\tau$$
 (4.3)

$$p_{T}(x) = \int_{-\infty}^{\infty} p_{U_{1}}(\tau) p_{U_{2}}(x - \tau) d\tau$$
 (4.3)
$$p_{T}(x) = \int_{0}^{1} p_{U_{2}}(x - \tau) d\tau$$
 (4.4)

$$p_T(x) = \begin{cases} 0 & x \le 0\\ \int_0^x 1d\tau & 0 < x < 1\\ \int_{x-1}^1 1d\tau & 1 \le x < 2\\ 0 & x > 2 \end{cases}$$
(4.5)

$$p_T(x) = \begin{cases} 0 & x \le 0 \\ x & 0 < x < 1 \\ 2 - x & 1 \le x < 2 \\ 0 & x > 2 \end{cases}$$
 (4.6)

Expression for CDF can be obtained by integrating $p_T(x)$ w.r.t. X

$$F_T(x) = \begin{cases} 0 & x \le 0\\ \frac{x^2}{2} & 0 < x < 1\\ -\frac{x^2}{2} + 2x - 1 & 1 \le x < 2\\ 1 & x > 2 \end{cases}$$
(4.7)

4.5 Verify the results through a plot. **Solution:**

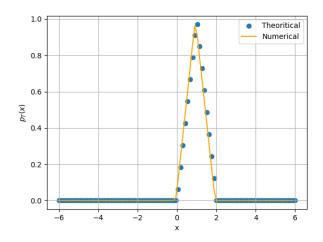


Fig. 4.5: Theoretical PDF of T

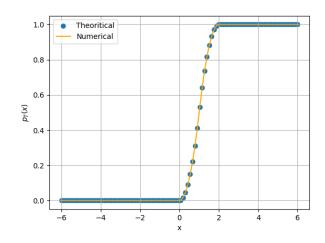


Fig. 4.5: The CDF of T

PDF and CDF plotted by the Python codes

wget https://github.com/TYCN129/AI1110-Assignments/blob/main/Manual %201/4.5/4.5 1.py

wget https://github.com/TYCN129/AI1110-Assignments/blob/main/Manual %201/4.5/4.5 1.py

5 Maximul Likelihood

5.1 Generate equiprobable $X \in \{1, -1\}$.

Solution:

Download and run the following to generate X.

wget https://github.com/TYCN129/AI1110-Assignments/blob/main/Manual %201/5.1/5.1 X.c

5.2 Generate

$$Y = AX + N, (5.1)$$

where A = 5 dB, and $N \sim 01$.

Solution:

Download and run the following code to gen-

wget https://github.com/TYCN129/AI1110-Assignments/tree/main/Manual %201/5.1/5.1 Y.c

5.3 Plot *Y* **Solution:**

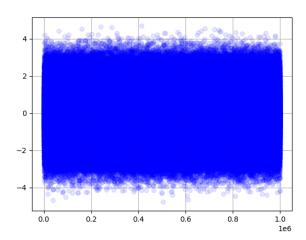


Fig. 5.3: Plot of *Y*

The Python code plots Y.

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual %201/5.2/5.2.py

5.4 Guess how to estimate X from Y

Solution:

Estimate of $X = \hat{X}$ is found out from Y by,

$$\hat{X} = \begin{cases} -1 & Y < 0\\ 1 & Y \ge 0 \end{cases} \tag{5.2}$$

5.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1)$$
 (5.3)

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1)$$
 (5.4)

Solution:

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1)$$
 (5.5)

$$= 0.4998032$$
 (5.6)

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1)$$
 (5.7)

$$= 0.4995446$$
 (5.8)

5.6 Find P_e assuming X has equiprobable symbols. **Solution:**

$$P_{e|0} = Pr(\hat{X} = -1|X = 1) \tag{5.9}$$

$$= Pr(AX + N < 0|X = 1)$$
 (5.10)

$$= Pr(N < -A) \tag{5.11}$$

$$= \int_{-\infty}^{-A} \frac{1}{\sqrt{2\pi}} \exp^{\frac{-x^2}{2}}$$
 (5.12)

$$= \int_{A}^{\infty} \frac{1}{\sqrt{2\pi}} \exp^{\frac{-x^2}{2}} = Q_N(A)$$
 (5.13)

where, $Q_N(A)$ is the Q-function of Normal distribution. Similarly,

$$P_{e|1} = Q_N(A) \tag{5.14}$$

Thus,

$$P_e = P_{e|0} \times Pr(X = 1) + P_{e|1} \times Pr(X = -1)$$
(5.14)

$$P_e = \frac{1}{2} \times P_{e|0} + \frac{1}{2} \times P_e e|1$$
 (5.16)

$$P_e = \frac{P_{e|0} + P_{e|1}}{2} \tag{5.17}$$

$$P_e = Q_N(A) \tag{5.18}$$

5.7 Verify by plotting theoretical P_e by varying A from 0 to 10 dB

Solution:

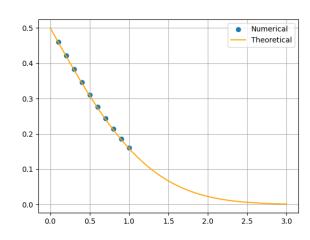


Fig. 5.7: Plot of P_e with varying A

wget https://github.com/TYCN129/AI1110-Assignments/tree/main/Manual%201/5.7

5.8 Now, consider a threshold δ while estimating X from Y. Find the value of δ that maximizes the theoretical P_e .

Solution:

Defining estimate of X as,

$$\hat{X} = \begin{cases} 1 & Y > \delta \\ -1 & Y < \delta \end{cases}$$
 (5.19)

$$P_{e|0} = Pr(\hat{X} = -1|X = 1)$$
 (5.20)
= $Pr(Y < \delta|X = 1)$ (5.21)
= $Pr(AX + N < \delta|X = 1)$ (5.22)

$$= Pr(N < \delta - A) \tag{5.23}$$

$$= \int_{-\infty}^{\delta - A} \frac{1}{\sqrt{2\pi}} e^{\frac{-x^2}{2}}$$
 (5.24)

$$= \int_{A-\delta}^{\infty} \frac{1}{\sqrt{2\pi}} e^{\frac{-x^2}{2}}$$
 (5.25)

$$= Q_N(A - \delta) \tag{5.26}$$

Similarly,

$$P_{e|1} = Q_N(A + \delta) \tag{5.28}$$

Therefore,

$$P_e = P(X = 1) \times P_{e|0} + P(X = -1) \times P_{e|1}$$
(5.29)

$$=\frac{Q(A-\delta)+Q_N(A+\delta)}{2}$$
 (5.30)

To maximize theoretical P_e , we differentiate equation (5.30) w.r.t δ .

$$\frac{dP_e}{d\delta} = \frac{1}{2} \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta - A)^2}{2}} - \frac{1}{\sqrt{2\pi}} e^{-\frac{(A + \delta)^2}{2}} \right) \quad (5.31)$$

$$0 = \frac{1}{2} \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta - A)^2}{2}} - \frac{1}{\sqrt{2\pi}} e^{-\frac{(A + \delta)^2}{2}} \right) \quad (5.32)$$

Thus,

$$\delta = 0 \tag{5.33}$$

5.9 Repeat the above exercise when

$$p_X(0) = p \tag{5.34}$$

Solution:

$$P_e = P(X = 1) \times P_{e|0} + P(X = -1) \times P_{e|1}$$
(5.35)

$$P_e = pP_{e|0} + (1-p)P_{e|1} (5.36)$$

On differentiating both sides,

$$0 = \frac{1}{2} \left(\frac{p}{\sqrt{2\pi}} e^{-\frac{(\delta - A)^2}{2}} - \frac{(1 - p)}{\sqrt{2\pi}} e^{-\frac{(A + \delta)^2}{2}} \right) \quad (5.37)$$

Thus.

$$\delta = \frac{1}{2A} ln \left(\frac{(1-p)}{p} \right) \tag{5.38}$$

5.10 Repeat the above exercise using the MAP criterion.

Solution:

$$p_{Y}(y) = p_{Y|X=1}(y|1) \times P(X=1) +$$

$$p_{Y|X=-1}(y|-1) \times P(X=-1)$$

$$= p_{N+A}(y) \times p + p_{N-A}(y) \times (1-p)$$
(5.39)

Here, $p_Y(y)$ is the PDF of Y, p_{N+A} and p_{N-A} are the PDF of shifted normal distribution.

$$p_Y(y) = \frac{p}{\sqrt{2\pi}}e^{-\frac{(y+A)^2}{2}} + \frac{(1-p)}{\sqrt{2\pi}}e^{-\frac{(y-A)^2}{2}}$$
 (5.40)

Now, we need to find $p_{X|Y}(x|y)$ using the formula,

$$p_{X|Y}(x|y) = \frac{p_{Y|X}(y|x) \times p_X(x)}{p_Y(y)}$$
 (5.41)

For X = 1,

$$p_{X|Y}(1|y) = \frac{p_{Y|X}(y|1) \times p}{p_Y(y)}$$
 (5.42)

$$= \frac{\frac{p}{\sqrt{2\pi}}e^{-\frac{(y+A)^2}{2}}}{\frac{p}{\sqrt{2\pi}}e^{-\frac{(y+A)^2}{2}} + \frac{(1-p)}{\sqrt{2\pi}}e^{-\frac{(y-A)^2}{2}}}$$
(5.43)

$$=\frac{pe^{2Ay}}{pe^{2Ay}+(1-p)}$$
(5.44)

Similarly, for X = -1,

$$p_{X|Y}(-1|y) = \frac{p_{Y|X}(y|-1) \times (1-p)}{p_Y(y)}$$
 (5.45)

$$p_{X|Y}(-1|y) = \frac{p_{Y|X}(y|-1) \times (1-p)}{p_{Y}(y)}$$

$$= \frac{\frac{(1-p)}{\sqrt{2\pi}} e^{-\frac{(y-A)^{2}}{2}}}{\frac{p}{\sqrt{2\pi}} e^{-\frac{(y+A)^{2}}{2}} + \frac{(1-p)}{\sqrt{2\pi}} e^{-\frac{(y-A)^{2}}{2}}}$$
(5.46)

$$=\frac{(1-p)}{pe^{2Ay}+(1-p)}\tag{5.47}$$

• Case 1) $p_{X|Y}(1|y) > p_{X|Y}(-1|y)$

$$\frac{pe^{2Ay}e^{2Ay}}{pe^{2Ay} + (1-p)} > \frac{(1-p)}{pe^{2Ay} + (1-p)}$$
 (5.48)

$$e^{-2Ay} < \frac{p}{1-p} \tag{5.49}$$

$$y > \frac{1}{2A} ln \left(\frac{p}{1 - p} \right) \tag{5.50}$$

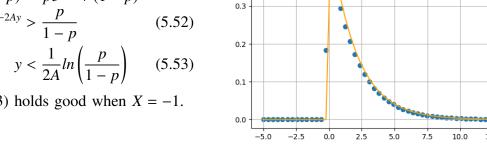
Equation (5.50) holds good when X = 1.

• Case 2) $p_{X|Y}(1|y) < p_{X|Y}(-1|y)$

$$\frac{pe^{2Ay}}{pe^{2Ay} + (1-p)} < \frac{(1-p)}{pe^{2Ay} + (1-p)}$$
 (5.51)

$$e^{-2Ay} > \frac{p}{1-p} \tag{5.52}$$

Equation (5.53) holds good when X = -1.



0.4

Numerical

Theoretical

0.8

0.6

0.4

0.2

Fig. 6.1: PDF of *V*

Fig. 6.1: CDF of *V*

Numerical

Theoretical

6 Gaussian to Other

6.1 Let $X_1 \sim 01$ and $X_2 \sim 01$. Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \tag{6.1}$$

Solution:

Download and run the C code to generate V.

wget https://github.com/TYCN129/AI1110-Assignments/blob/main/Manual %201/6.1/6.1.c

wget https://github.com/TYCN129/AI1110-Assignments/blob/main/Manual%201/6.1/ CDF.py

wget https://github.com/TYCN129/AI1110-Assignments/blob/main/Manual%201/6.1/ PDF.py

6.2 If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \ge 0\\ 0 & x < 0, \end{cases}$$
 (6.2)

find α .

Solution:

Let $X_1 = R \sin \theta$ and $X_2 = R \cos \theta$

Jacobian matrix is given as follows

$$J = \begin{pmatrix} \frac{\delta x_1}{\delta r} & \frac{\delta x_1}{\delta \theta} \\ \frac{\delta x_2}{\delta r} & \frac{\delta x_2}{\delta \theta} \end{pmatrix}$$
 (6.3)

$$J = \begin{pmatrix} \cos \theta & -R \sin \theta \\ \sin \theta & R \cos \theta \end{pmatrix} \tag{6.4}$$

$$|J| = R \tag{6.5}$$

Also,

$$f_{r,\theta}(r,\theta) = |J| f_{X_1,X_2}(x_1, x_2)$$
 (6.6)

Since, X_1 and X_2 are independent,

$$f_{X_1,X_2}(x_1,x_2) = f_{X_1}(x_1)f_{X_2}(x_2)$$
 (6.7)

$$=\frac{1}{2\pi}e^{-\frac{(x_1^2+x_2^2)}{2}}\tag{6.9}$$

(6.8)

$$=\frac{1}{2\pi}e^{-\frac{r^2}{2}}\tag{6.10}$$

Put in Equation 6.6, we get,

$$f_{R,\theta}(r,\theta) = \frac{r}{2\pi}e^{-\frac{r^2}{2}}$$
 (6.11)

Now,

$$f_R(r) = \int_0^{2\pi} f_{R,\theta}(r,\theta)$$
 (6.12)

$$= \int_0^{2\pi} \frac{r}{2\pi} e^{-\frac{r^2}{2}} d\theta \tag{6.13}$$

$$= re^{-\frac{r^2}{2}} \tag{6.14}$$

CDF is given by,

$$F_R(r) = \int_0^r re^{-\frac{r^2}{2}}$$
 (6.15)

$$=1-e^{\frac{-r^2}{2}} \tag{6.16}$$

And,

$$V = X_1^2 + X_2^2 (6.17)$$

$$=R^2 \tag{6.18}$$

Now,

$$F_V(x) = \Pr(V \le x) \tag{6.19}$$

$$= \Pr\left(R^2 \le x\right) \tag{6.20}$$

$$= \Pr\left(R \le \sqrt{x}\right) \tag{6.21}$$

$$F_V(x) = F_R(\sqrt{x}) \tag{6.22}$$

$$= \begin{cases} 0, & x < 0 \\ 1 - e^{-\frac{x}{2}} & x \ge 0 \end{cases}$$
 (6.23)

$$\alpha = \frac{1}{2} \tag{6.24}$$

6.3 Plot the CDF and PDF of A.

$$A = \sqrt{V} \tag{6.25}$$

Solution:

Download the C code to generate $A = \sqrt{V}$

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual%201/6.3/ A.c

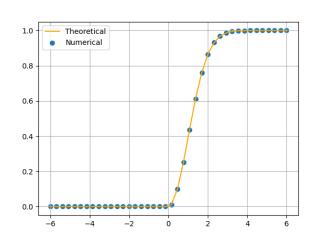


Fig. 6.3: CDF of *A*

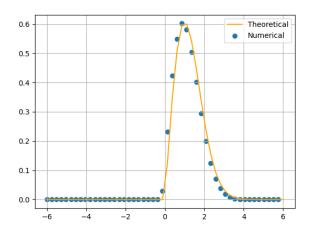


Fig. 6.3: PDF of *A*

7 CONDITIONAL PROBABILITY

7.1 Plot

$$P_e = \Pr(\hat{X} = -1|X = 1) \tag{7.1}$$

for

$$Y = AX + N, (7.2)$$

where A is Raleigh with $E\left[A^2\right] = \gamma, N \sim 01, X \in (-1,1)$ for $0 \le \gamma \le 10$ dB.

Solution:

Download and run the following C code to

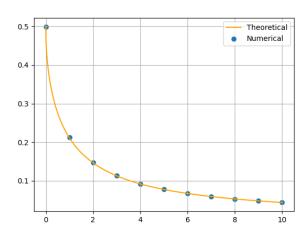


Fig. 7.1: P_e as a function for γ

generate Rayleigh distribution.

wget https://github.com/TYCN129/AI1110— Assignments/blob/main/Manual%201/7.1/ ral_gen.c

The Python code plots the plot of P_e vs γ

wget https://github.com/TYCN129/AI1110– Assignments/blob/main/Manual %201/7.1/7.1.py

7.2 Assuming that N is a constant, find an expression for P_e . Call this $P_e(N)$ Solution:

$$P_e = Pr(\hat{X} = -1|X = 1) \tag{7.3}$$

$$= Pr(A + N < 0) \tag{7.4}$$

$$= Pr(A < -N) \tag{7.5}$$

$$= \int_{-\infty}^{-N} \frac{1}{\sqrt{2\pi}} e^{\frac{-x^2}{2}} dx \tag{7.6}$$

$$= \int_{N}^{\infty} \frac{1}{\sqrt{2\pi}} e^{\frac{-x^2}{2}} dx \tag{7.7}$$

$$= Q_A(N) \tag{7.8}$$

$$= F_A(-N) \tag{7.9}$$

$$F_A(-N) = \begin{cases} 1 - e^{-\frac{N^2}{2}} & N \le 0\\ 0 & N > 0 \end{cases}$$
 (7.10)

7.3 For a function g,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)p_X(x) dx \qquad (7.11)$$

Find $P_e = E[P_e(N)]$.

Solution:

$$P_e = Pr(\hat{X} = -1|X = 1) \tag{7.12}$$

$$= Pr(A + N < 0) (7.13)$$

$$= Pr(A < -N) \tag{7.14}$$

$$= \int_{-\infty}^{0} Pr(A < -\tau) P_N(\tau) d\tau \tag{7.15}$$

$$= \int_{-\infty}^{0} (1 - e^{\frac{-\tau^2}{\gamma}}) \frac{1}{\sqrt{2\pi}} e^{-\frac{\tau^2}{2}} d\tau$$
 (7.16)

$$= \int_{-\infty}^{0} \frac{1}{\sqrt{2\pi}} e^{\frac{-r^2}{2}} - \int_{-\infty}^{0} \frac{1}{\sqrt{2\pi}} e^{\frac{r^2(\frac{2}{\tilde{\gamma}}+1)}{2}}$$
(7.17)

$$= \frac{1}{2} - \frac{1}{2} \frac{\sqrt{\gamma}}{\sqrt{2 + \gamma}} \tag{7.18}$$

7.4 Plot P_e in problems 7.1 and 7.3 on the same graph w.r.t γ . Comment.