

Tutorial 8 - Hypothesis Testing

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How to test a hypothesis





How to test a hypothesis

Steps you have to do as a researcher:

1. State your **null-hypothesis** H_0 and the **alternative hypothesis** H_1
 - a. H_0 is a statement about the population parameter and is the conservative statement
2. Decide **significance level** α
3. Think about your statistical model, develop your test
 - a. sample space, probability measure, set of possible outcomes, test statistic
4. Only then get the data
5. Calculate a **p-value**, making the assumption that H_0 is true
6. **Decide** to reject or to not reject hypothesis H_0



Important concepts



Test statistics

Quantity derived from the data, that reduces the data to one value that can be used to perform the hypothesis test. E.g. sample mean $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$, sample variance $s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$
number of heads for coin toss $\sum_{i=1}^n x_i$.

More abstract: quantify, within observed data, behaviour that would distinguish the null from the alternative hypothesis.

Has a probability distribution, which is used to compute p-values for the null hypothesis.



Test statistics - example: coin flip

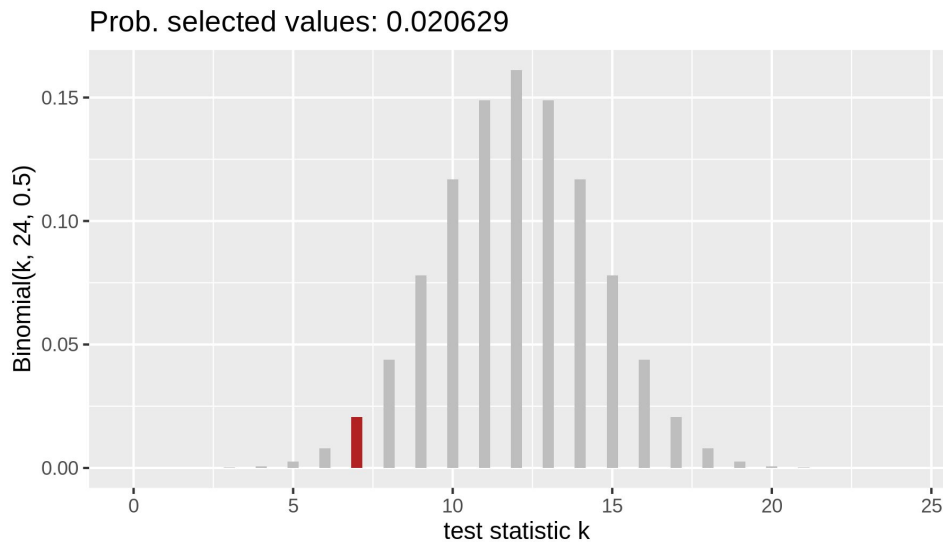
Test whether a coin is fair, flip 24 times. Raw data: c(0, 1, 1, 1, 0, ..., 0, 1, 1, 1, 1).

If there is interest in the probability of obtaining a head, only the number $k = \sum_{i=1}^{24} x_i$ out of the 24 flips that were heads needs to be recorded.

What generated k ?

$$k \sim \text{Binom}(\theta = 0.5, n = 24)$$

Note that this **test statistic** reduces a set of 24 numbers to a single numerical summary.





Sampling distribution

The probability distribution that generates the values that the test statistic can take on is called **sampling distribution**.

In our previous example is the binomial distribution the sampling distribution: $k \sim \text{Binom}(\theta = 0.5, n = 24)$
Using this sampling distribution, it is possible to compute a p-value for the null hypothesis that the coin is fair.

To sum up:

NULL HYPOTHESIS → DATA → TEST STATISTIC → SAMPLING DISTRIBUTION → P-VALUE



Model with Binomial Distribution





Example: Coin Toss Experiment

- We assume the distribution for the experiment: **Binomial** -> **discrete probability distribution, series of similar and independent events, each of which has exactly two possible outcomes ("success" or "failure") used to model binary data** → is used to model the likelihood function
- We assume the parameter we want to make inferences by: θ
- We specify **N = number of trials, $p = \theta$ = probability of success for each trial, k = observed successes ("heads")**

We use this information to obtain the test statistics & sampling distribution!



Set Hypotheses

- wish to make inferences about: θ = probability of success for each trial
- A possible research hypothesis for our example would be: **Is the coin fair?** (two-sided)

$$H_0 = \theta = 0.5$$

$$H_1 = \theta \neq 0.5 \text{ (departs from } H_0 \text{ left and right)}$$

- conceptually, we assume H_0 is true for the population
- null hypothesis is assumed to be true until evidence indicates otherwise
- researchers work to reject or disprove the null hypothesis



Example: Coin Toss Experiment

We observe:

coin is flipped **N= 24** times

k=7 times the coin successfully flipped “heads”

p(here θ) is the probability that the outcome will occur at any particular coin toss (**assuming H_0 , we assume $\theta = 0.5$, is a fair coin**)

Sampling distribution and the test statistic (k) are now specified $k \sim \text{Binomial}(\theta, N)$.

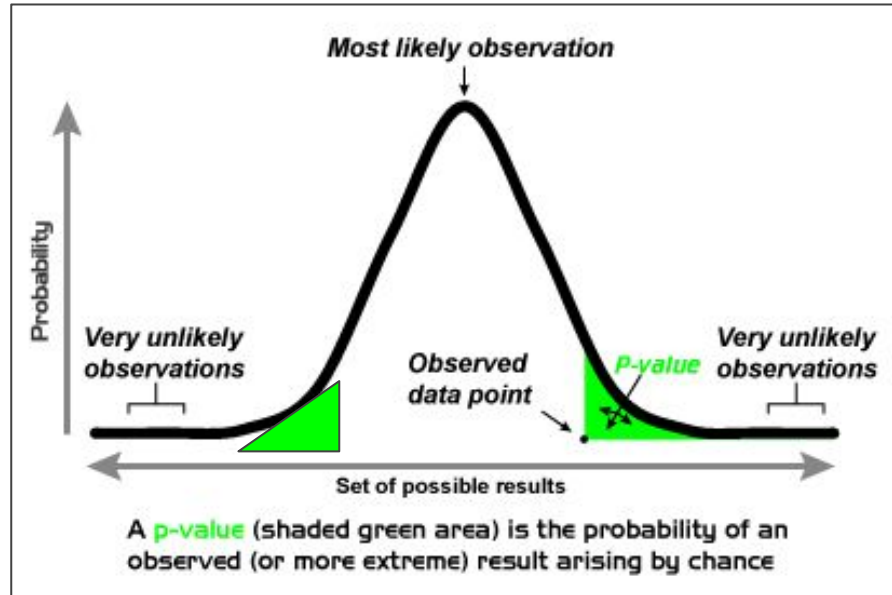
Lets test our hypothesis!



Hypotheses testing

- we need to set a null hypothesis, i.e., a **value θ of coin bias θ** that we would like to **collect evidence against** $\rightarrow H_0 = \theta = 0.5$
- if empirical observations are sufficiently **unlikely** from the point of view of the **null-hypothesis H_0** , this should be treated as **evidence against the null-hypothesis**
- a measure of **how unlikely** data is in the light of H_0 is **the p-value**
- to obtain a p-value: **what is the probability of observing more extreme values (in this case: to both ends) compared to what we sampled ($k=7$) and therefore count as more extreme evidence against the chosen null hypothesis?**
- which values for k are less or equally as probable compared to our sample of the test statistic $k=7$?

Area under curve =
represents 100% of
all possible events



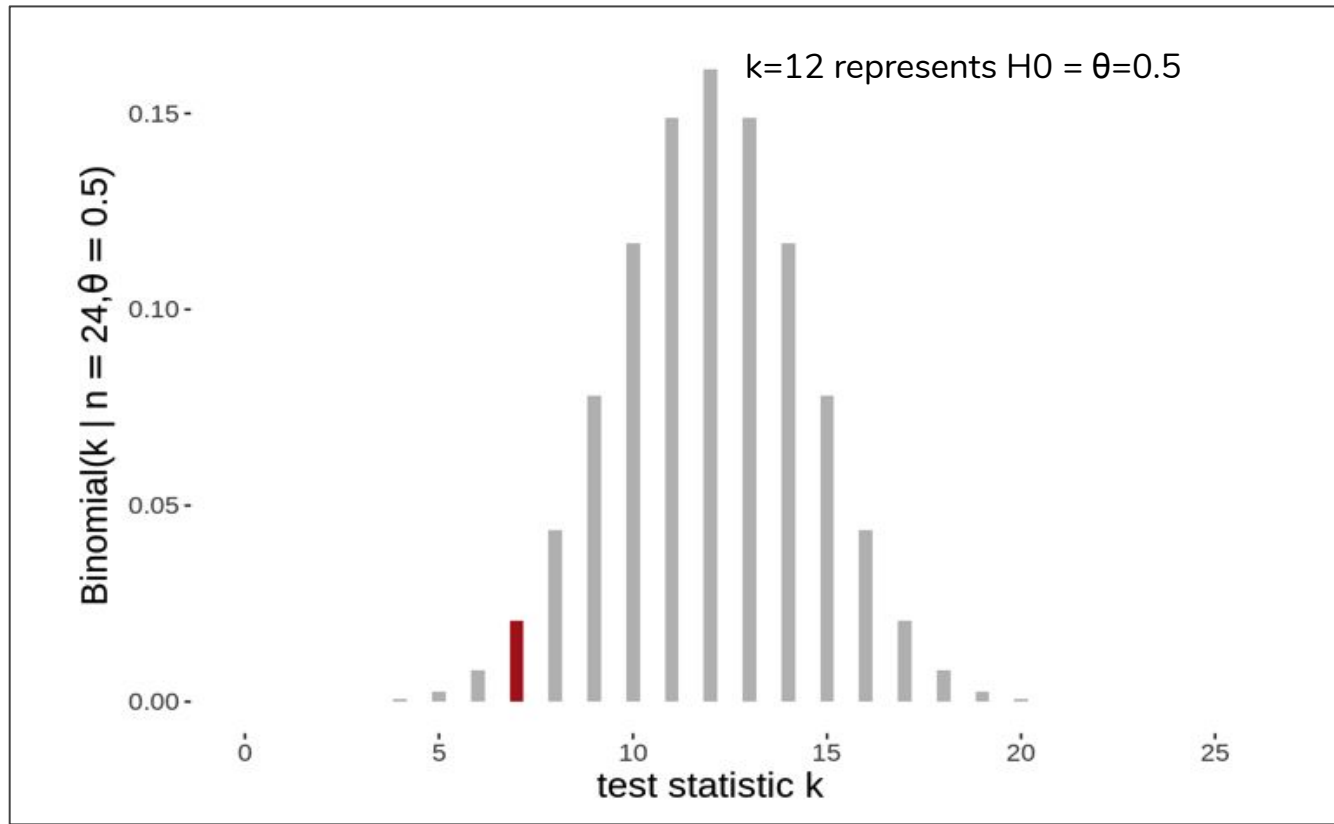
If we assume H_0 to be true, values to the left and right side of the curve become increasingly unlikely

If the p -value falls within a sufficiently unlikely area, this is taken as evidence that H_0 cannot be true for a population



- value that occurs at the **peak (k=12)** represent $H_0 = \theta = 0.5$ → typically, H_0 states that there is **no effect**
- as **values for k move further away** from the peak, it represents **larger effect sizes (in refuting H_0)**
- when **H_0 is true** for the population, obtaining samples that exhibit **large effects** becomes **less likely**, which is why the probabilities for k values taper off to the sides of the curve, further from $\theta = 0.5$

y-axis
shows
probability
to observe
certain
measurements



Sampling distribution shows the **probability associated with observed data $k=7$ highlighted in red**. Displays the probabilities of obtaining test statistic values when the null hypothesis is correct ($\theta=0.5$).



Calculate p-value Two-Tailed

- To obtain a p-value:
- **sum up** all **probabilities** for observing values equal to or smaller than empirically sampled test statistic k ($=7$)
- In other words: **sum over** all possible orders of coin-toss-outcome-values with probabilities equal to or less than probability of observing $k=7$

Calculate in R Two-Tailed- handwritten function

```
```{r, echo=T}
exact p-value for k=7 with N=24 and null-hypothesis theta = 0.5

k_obs <- 7
N <- 24
theta_0 <- 0.5
#args: vecofquantiles, num oftrials, prob of success each trial
tibble(lh = dbinom(0:N, N, theta_0)) %>%

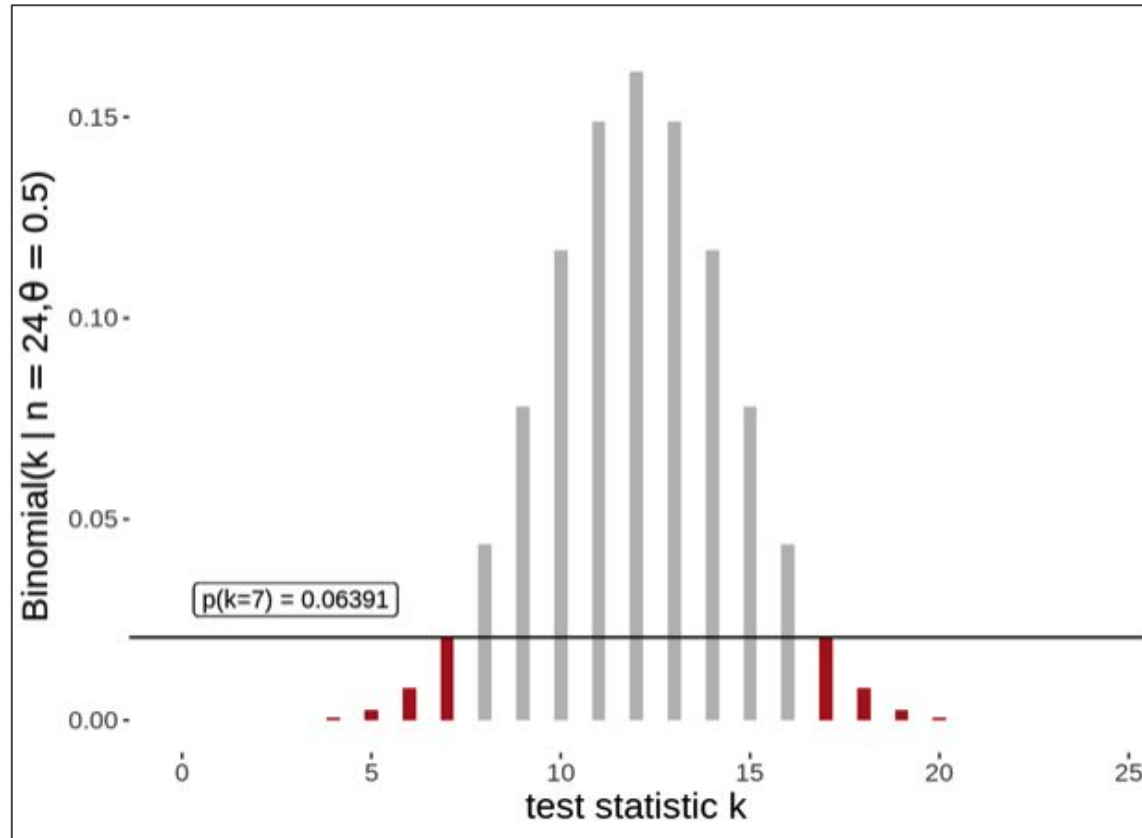
Use filter() to choose rows/cases where conditions are true
logical condition: lh smaller or equal to dbinom(k_obs, N, theta_0)
filter(lh <= dbinom(k_obs, N, theta_0)) %>%
#sum of all values of lh
pull(lh) %>% sum %>% round(5)

```
```

```
[1] 0.06391
```



y-axis shows
probability to
observe certain
measurements



Sampling distribution shows the values that **need to be summed over in red**. p-value for the observation of $k=7$ successes in $N=24$ coin flips. Displays the probabilities of obtaining test statistic values when the null hypothesis is correct ($\theta=0.5$).



Calculate in R Two-Tailed- built-in function

```
binom.test(  
  x = 7,      # observed successes  
  n = 24,     # total nr. observations  
  p = 0.5     # null hypothesis  
)  
##  
## Exact binomial test  
##  
## data: 7 and 24  
## number of successes = 7, number of trials = 24, p-value = 0.06391  
## alternative hypothesis: true probability of success is not equal to 0.5  
## 95 percent confidence interval:  
##  0.1261521 0.5109478  
## sample estimates:  
## probability of success  
##           0.2916667
```



Calculate p-value One-Tail

Another possible research hypothesis for our experiment: **Is the coin biased towards “heads”?**

$$H_0 = \theta > 0.5$$

$$H_1 = \theta \leq 0.5$$

Now we wish to calculate the p-value for data for this model!

- only seeks to measure effect into one direction from H_0 (and from the curve)
- what would count as the most extreme evidence against H_0 ?

- We need to adjust what we consider



Calculate p-value One-Tail

- **to obtain a p-value:** What are the probabilities of observing values less than or equal to test statistic $k=7$? Sum them!
- values on the **right hand side from $\theta = 0.5$ will not serve as evidence against H_0 because all values there are $\theta > 0.5$**
- the associated p-value must be calculated using a **one-sided test, only considering values on the left side of the curve**



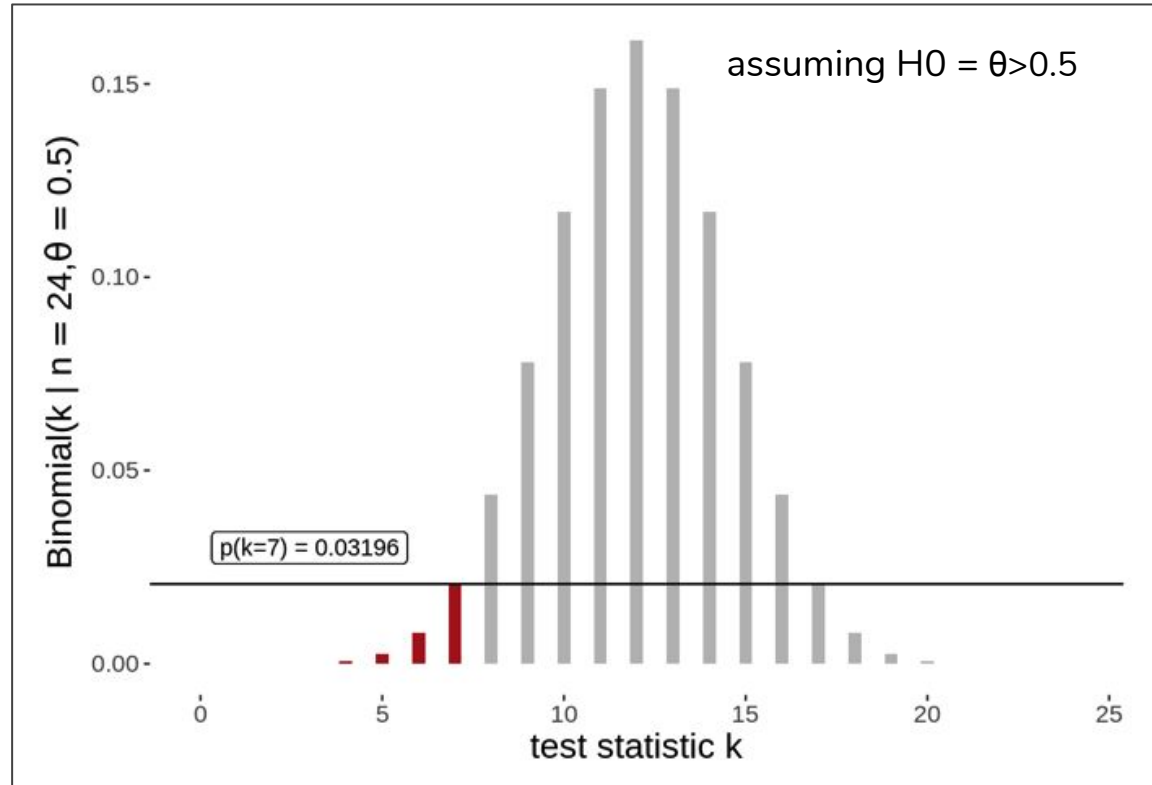
Calculate in R One-Tailed - handwritten function

```
k_obs <- 7
N <- 24
theta_0 <- 0.5
# exact p-value for k=7 with N=24 and null-hypothesis theta > 0.5
dbinom(0:k_obs, N, theta_0) %>% sum %>% round(5)
## [1] 0.03196
```

- doubling the p value for one-tailed test results in the p-value for two-tailed test for **symmetrical** sampling distributions!



y-axis shows
probability to
observe certain
measurements



Sampling distribution shows p-value for the observation of $k=7$ successes. Displays the probabilities of obtaining test statistic values when null hypothesis is correct ($\theta > 0.5$).
Sum over all probabilities for observing values smaller than or equal to k only on the left-hand side.



Calculate in R One-Tailed- built-in function

```
binom.test(  
  x = 7,      # observed successes  
  n = 24,     # total nr. observations  
  p = 0.5,    # null hypothesis  
  alternative = "less" # the alternative to compare against is  $\theta < 0.5$   
)  
##  
## Exact binomial test  
##  
## data: 7 and 24  
## number of successes = 7, number of trials = 24, p-value = 0.03196
```



Significance of p-values

- **Fisher:** p-values as quantitative measures of strength of **evidence against the null hypothesis**:

if you get a result that is barely significant, there is chance that you falsely rejected H_0 !
Same with results that are barely non-significant, maybe we are falsely keeping H_0 !
- we say a test result is **significant** if the p-value of the observed data is **lower than a specified α**
- we fix the **α** -level of significance with common values **$\alpha \in \{0.05, 0.01, 0.001\}$**
- commonly, a **significant test** results is interpreted as the signal to **reject the null hypothesis**, to render it false
- if your **test statistic** falls in **either critical region**, your **sample data** are sufficiently **incompatible with the null hypothesis** and observing this value is sufficiently unlikely in light of the null hypothesis so that you can reject it for the population!

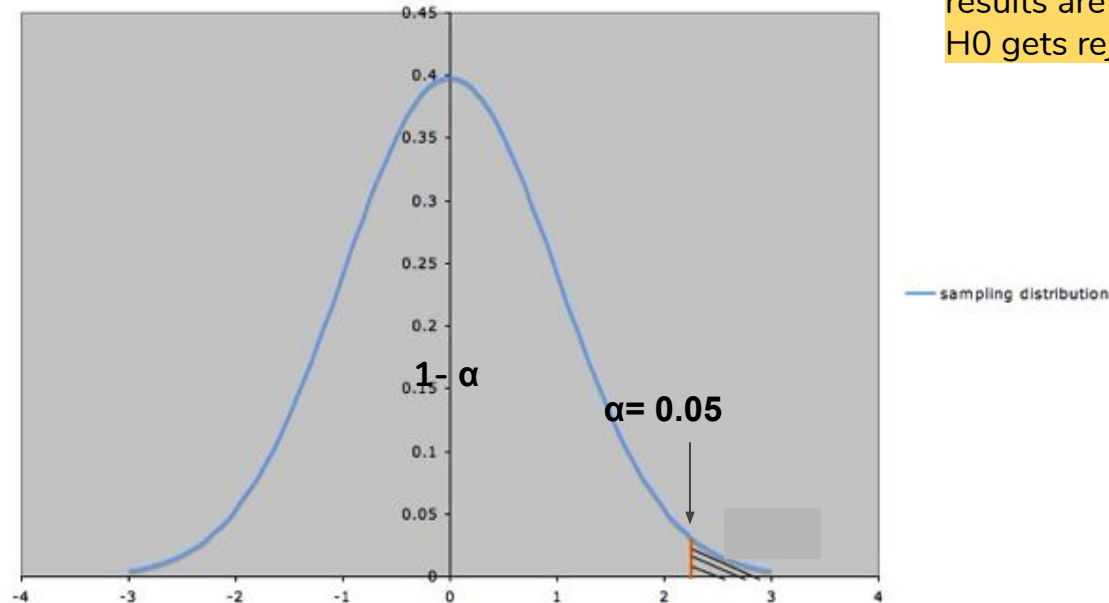
Significance of p-values

Sampling dist assuming H0

sampling distribution

If p-value falls below α , results are significant → H0 gets rejected!

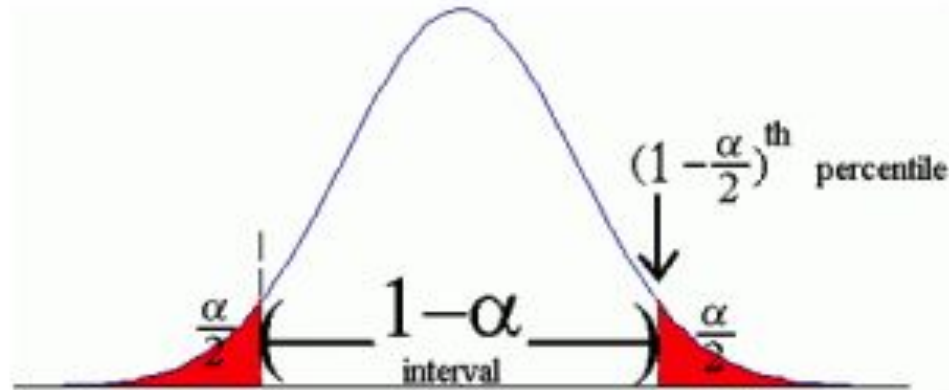
$\alpha=0.05$ translates to 0.5% chance of observing values for test statistic k in this area when randomly sampling and a 95% chance it will be in the $\alpha-1$ area under the curve



Source:

<https://www.statisticshowto.datasciencecentral.com/what-is-an-alpha-level/>

$1 - \alpha$
represents a
confidence
interval =
certainty/prob
ability any
random
sampled k will
be amongst
these values



The two red tails are the alpha level, divided by two (i.e. $\alpha/2$).



Errors of decision making

In hypothesis tests, two errors are possible:

- **Type I error: Supporting the alternative hypothesis when the null hypothesis is true (alpha error)**

Example: type I error corresponds to FDA approving a novel drug while it actually has no measured effect. H_0 = drug has no effect; H_1 = drug has desired effect

- **Type II error: Not supporting the alternative hypothesis when the alternative hypothesis is true (beta error)**

Example: type II error corresponds to FDA rejecting a novel drug although it has the desired effect. H_0 = drug has no effect; H_1 = drug has desired effect



Errors of decision making

try to minimize both errors
by determining optimal
level of significance

Statistical power = probability that
an effect will be discovered when
an effect actually exists

defined as $1 - \beta$ where β is the
probability of making a second type
of error

If the statistical power is high, the
probability of committing a Type II
error decreases

| | Reject H_0 | Fail to Reject H_0 |
|----------------------------|---|--|
| Reality:
H_0 is True | Type I error
(probability = α) | higher for lower alpha
Probability = $1 - \alpha$ |
| Reality:
H_0 is False | Power ($1 - \beta$) | Type II error
(probability = β) |



Law of large numbers



The law of large numbers

Imagine you gather a lot of samples from random variables X_1, \dots, X_n that all have the same expected value, i.e. $\mathbb{E}[X_i] = \mu$

→ e.g. $X_i \sim \mathcal{N}(10.5, 1)$, where $\mathbb{E}[X_i] = 10.5$ or e.g. $X_i \sim \mathcal{U}(0, 100)$, where $\mathbb{E}[X_i] = 50$

You then want to estimate this expected value based on your samples (because you don't know it yet)

You think the arithmetic mean could be a good estimator of μ : $\hat{T} = \frac{1}{n} \sum_{i=1}^n X_i$

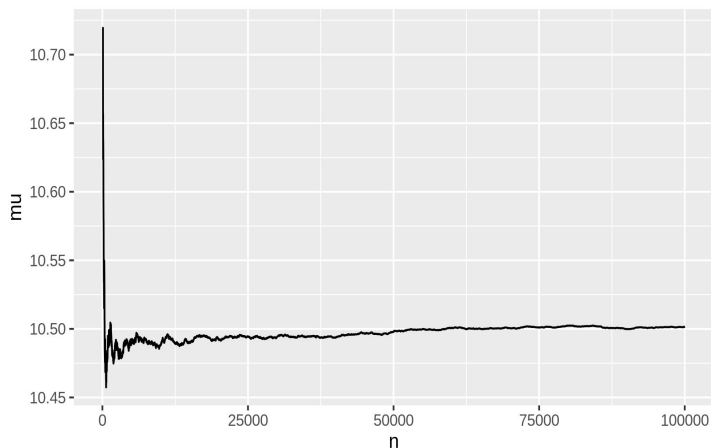
The law of large numbers assures that if you gather more and more samples and you compute their arithmetic mean, this arithmetic mean will be (almost surely) the expected value $\mathbb{E}[X_i]$



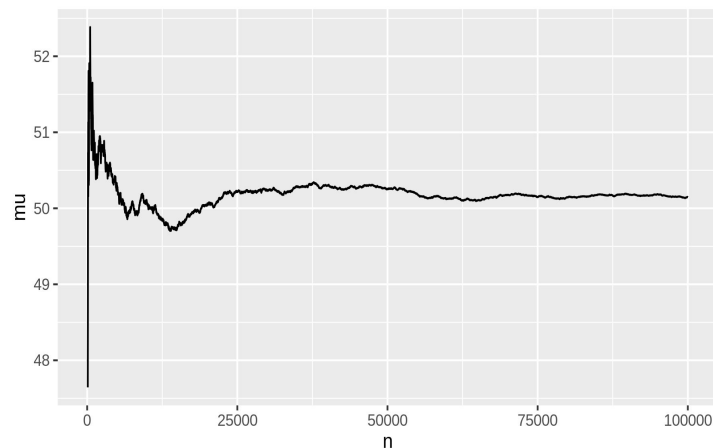
Law of large numbers

Arithmetic mean calculated for samples from

- $X_i \sim \mathcal{N}(10.5, 1)$



- $X_i \sim \mathcal{U}(0, 100)$



Arithmetic means and expected values of the distributions are the same for large n!

The central limit theorem



The Central limit theorem

Imagine you gather samples from random variables X_1, \dots, X_n that all have the same expected value and the same (finite) variance, i.e. $\mathbb{E}[X_i] = \mu, \text{Var}[X_i] = \sigma^2$.

Then for each bunch of samples, e.g. for each 10 samples, you compute their arithmetic mean.

You collect all the arithmetic means (forming a sampling distribution). If you got a sufficient amount of means, then:

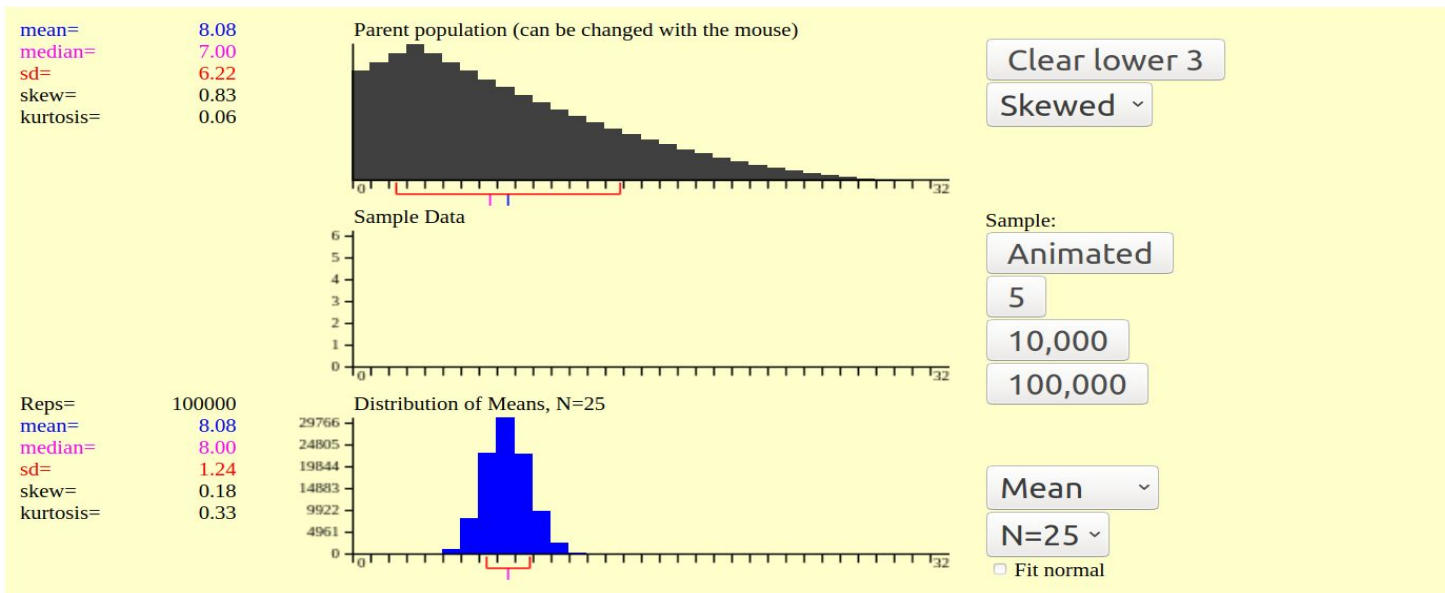
whatever the distribution of the random variables, the **sampling distribution** will be the Normal distribution $\mathcal{N}(\mu, \frac{\sigma^2}{n})$ if the overall sample size n is large enough.

If you normalize the sampling distribution, meaning you subtract the mean μ from each random variable and multiply by \sqrt{n} , it will be $\mathcal{N}(0, \sigma^2)$ -distributed.



Central limit theorem

[See for yourself!](#)



The χ^2 - test for goodness of fit



The χ^2 - test for goodness of fit testing

Pearson's χ^2 -test for **goodness of fit** tests whether an observed vector of counts is well explained by a given vector of predicted proportion.

“Goodness of fit” is a term used in model checking (a.k.a. model criticism, model validation, ...). In such a context, tests for goodness-of-fit investigate whether a model's predictions are compatible with the observed data.



The χ^2 - test for goodness of fit testing

We need:

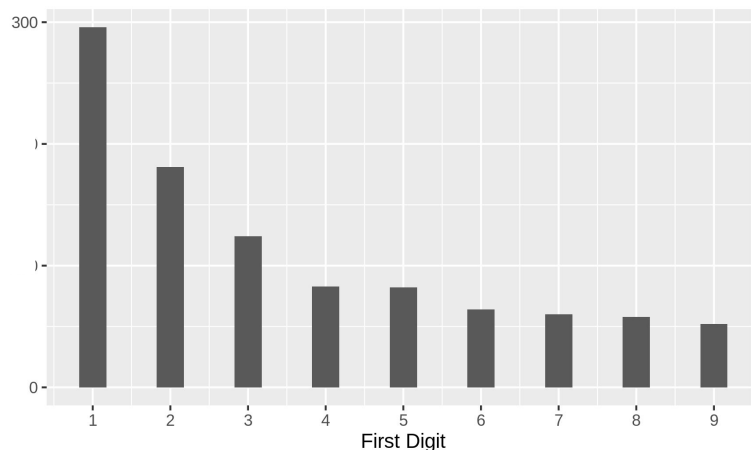
- Categorical data (each data observation falls into one of several unordered categories)
 - With k categories
- A null hypothesis H_0
- Vector of probabilities $\vec{p} = \langle p_1, \dots, p_k \rangle$ that correspond to our H_0 and gives the probability with which a single data observation falls into the i-th category.



An example

Raw data: c(143, 1, 34, 100, 923, 23, 42, 844, ..., 59, 66, 71, 2), counting **first digits** leads to:

Categorical Data with 9 categories: $\vec{n} = \langle 296, 181, 124, 83, 82, 64, 60, 58, 52 \rangle$, $N = 1000$

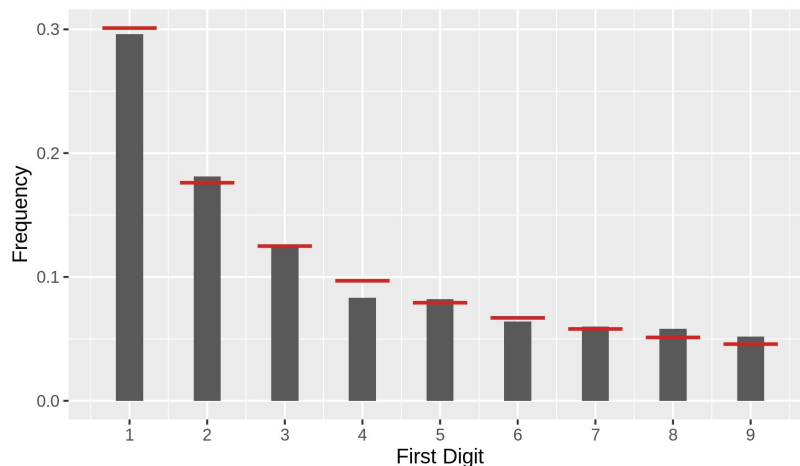


Vector of probabilities: $\vec{p} = \langle 0.301, 0.176, 0.125, 0.0969, 0.0792, 0.0669, 0.0580, 0.0512, 0.0458 \rangle$
(nine probabilities, one for each category, summing up to 1.)

An example

H_0 : the difference from the red bars to the data is not significant.

→ χ^2 - test for goodness of fit allows us to test whether this data could plausibly have been generated by (a model whose predictions are given by) the prediction vector.



$$\vec{p} = \langle 0.301, 0.176, 0.125, 0.0969, 0.0792, 0.0669, 0.0580, 0.0512, 0.0458 \rangle$$



Test statistic

$$\chi^2 = \sum_{i=1}^k \frac{(n_i - np_i)^2}{np_i}$$

To get the value of the test statistic for the χ^2 test, insert all the values:

- n_i is the i th entry of vector $\vec{n} = \langle 296, 181, 124, 83, 82, 64, 60, 58, 52 \rangle$
- Nine categories, so $k = 9$
- p_i is the i th entry of vector $\vec{p} = \langle 0.301, 0.176, 0.125, 0.0969, 0.0792, 0.0669, 0.0580, 0.0512, 0.0458 \rangle$
- $n = 1000$, all our data together

$$\chi^2 = \frac{(296-301)^2}{301} + \dots + \frac{(52-45.8)^2}{45.81} = 4.263105$$



In R

Manually:

```
n <- d$count
# proportion predicted
p <- c(0.301, 0.176, 0.125, 0.0969, 0.0792, 0.0669, 0.0580, 0.0512, 0.0458)
# expected number in each cell
e <- sum(n)*p
# chi-squared for observed data
chi2_observed <- sum((n-e)^2 * 1/e)
chi2_observed

p_value <- 1 - pchisq(chi2_observed, df = 8)
p_value
[[1]]
```

```
[1] 0.8326392
```

With in-build test:

```
```{r}
chisq.test(x = d$count, p = p)
```
```

Chi-squared test for given probabilities

```
data: d$count
X-squared = 4.2631, df = 8, p-value = 0.8326
```



Result & Interpretation

The common interpretation of our calculations would be to say that the test yielded no significant result, at least at the significance level of $\alpha=0.05$.

In a research paper we might report this results roughly as follows:

“The observed counts deviated not significantly from what is expected if each category (here: first digits) followed the specified probabilities (χ^2 -test, with $\alpha = 0.05$ $\chi^2 \approx 4.2631$, $df=8$ and $p \approx 0.8326$). We therefore conclude that there is no evidence to reject the hypothesis that our data conforms to the specifies probabilities.”



Homework hints





Homework Hints

Exercise 1: Addressing hypotheses about coin flips with hypothesis testing

- Similar procedure for all three cases: Think about that a point null hypothesis (e.g. $\theta = 0.5$) results in a two-sided test, and an interval hypothesis (e.g. $\theta \leq 0.5$) in a one-sided test



Exercise 2: Pearson's χ^2 -test of goodness of fit

- Think what the test does, what are your data, your expected probabilities?
- Check how to use the R function:

```
chisq.test {stats}
```

R Documentatio

Pearson's Chi-squared Test for Count Data

Description

`chisq.test` performs chi-squared contingency table tests and goodness-of-fit tests.

Usage

```
chisq.test(x, y = NULL, correct = TRUE,  
           p = rep(1/length(x), length(x)), rescale.p = FALSE,  
           simulate.p.value = FALSE, B = 2000)
```

Arguments

| | |
|----------------------|---|
| <code>x</code> | a numeric vector or matrix. <code>x</code> and <code>y</code> can also both be factors. |
| <code>y</code> | a numeric vector; ignored if <code>x</code> is a matrix. If <code>x</code> is a factor, <code>y</code> should be a factor of the same length. |
| <code>correct</code> | a logical indicating whether to apply continuity correction when computing the test statistic for 2 by 2 tables: one half is subtracted from all $ O - E $ differences; however, the correction will not be bigger than the differences themselves. No correction is done if <code>simulate.p.value = TRUE</code> . |
| <code>p</code> | a vector of probabilities of the same length of <code>x</code> . An error is given if any entry of <code>p</code> is negative. |



Exercise 3: Some claims about frequentist testing

- Frequentist statistics is all about repetitions. It never puts probabilities into parameter values.
- Check the lecture slides/ tutorial slides



Thank you for listening.

Questions?

