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Unsupervised Learning: t-SNE

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Day 2

Dennis Wylie, UT Bioinformatics Consulting Group

May 22, 2018

Outline

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8 Feature Selection

PCA

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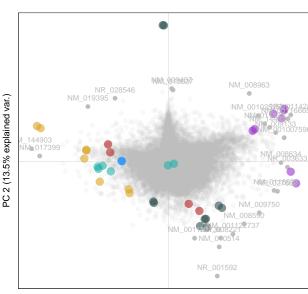
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- Variable
- lymphatic
- nervous
- circulatory
- digestive/excretory
- other
- respiratory

PC1 (18.5% explained var.)

Single principal component

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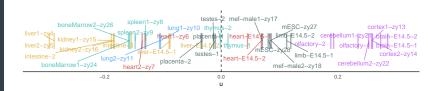
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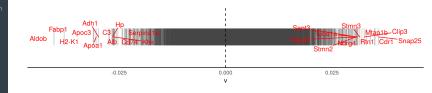
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Top PC1 genes

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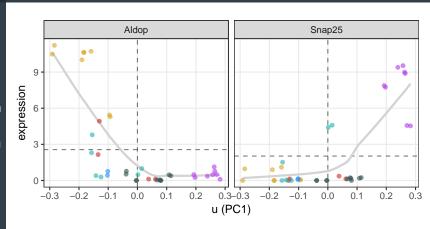
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Top PC1 genes

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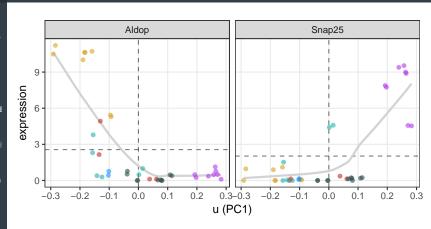
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Expression values of genes with highest PC1 *loadings* (=entries in vector **v** here)

will tend to correlate with sample PC1 loadings (=entries in vector u).

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Given a gene expression matrix (x_{ig})

▶ with samples *i* in rows and genes *g* in columns

PC1 model for random variable X_{ig} is:

$$X_{ig} - \mathbb{E}[X_{g}] = u_{i}dv_{g} + \epsilon_{ig}$$

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assuming that $\epsilon_{ig} \sim \mathcal{N}(0, \sigma^2)$ and constraining

$$\sum_{i} u_i^2 = \sum_{g} v_g^2 = 1$$

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$$\sum_{i} u_i^2 = \sum_{g} v_g^2 = 1$$

Model says that:

- gene g with $v_g > 0$ expected to have higher expression
- ▶ in sample *i* than in sample *j* if $u_i > u_i$
 - by an amount $(u_i u_i)v_g$.

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That is: PC1 model assigns

- \blacktriangleright to each sample *i* a score u_i
- \blacktriangleright and to each gene g a score v_g such that
 - ► high scoring genes highly expressed in high scoring samples

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assuming that $\epsilon_{\mathit{ig}} \sim \mathcal{N}(0, \sigma^2)$ and constraining

$$\sum_{i} u_i^2 = \sum_{g} v_g^2 = 1$$

 $\mathbb{E}\left[X_{\cdot g}\right]$ is mean expression of gene g in full population.

When doing PCA:

- ▶ use estimated $\tilde{x}_{ig} = x_{ig} \frac{1}{n} \sum_{i} x_{ig}$ in place of $x_{ig} \mathbb{E}[X_{g}]$
- principled approach should account for this estimation. . .

PCA plot

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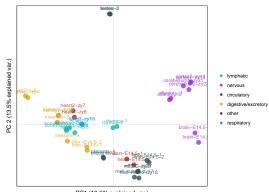
Feature Selection

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It seems unreasonable to expect a single linear pattern to explain everything in our data, though...let's try two:

$$\tilde{x}_{ig} = (u_{i1}d_{11}v_{g1} + u_{i2}d_{22}v_{g2}) + \epsilon_{ig}$$

and again select $u_{iq},\ d_{qq},\ v_{gq},$ and ϵ_{ig} so as to minimize $\sum\limits_{i,g}\epsilon_{ig}^2$



PC1 (18.5% explained var.)

PCA by SVD

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Can continue this process to obtain n principal components (assuming $n \le p$).

These can be calculated via the singular value decomposition (SVD) of $\underline{\tilde{X}}$ (note $\underline{\tilde{X}}$ is data matrix, not random variable)

$$\underline{\tilde{\mathbf{X}}} = \underline{\mathbf{U}}\underline{\mathbf{D}}\underline{\mathbf{V}}^T$$

where $\underline{\mathbf{U}}$ and $\underline{\mathbf{V}}$ are orthogonal matrices and $\underline{\mathbf{D}}$ is an $n \times n$ diagonal matrix with the diagonal sorted in descending order.

In terms of the components \tilde{x}_{ig} :

$$\tilde{x}_{ig} = \sum_{q} u_{iq} d_{qq} v_{gq}$$

PCA biplot

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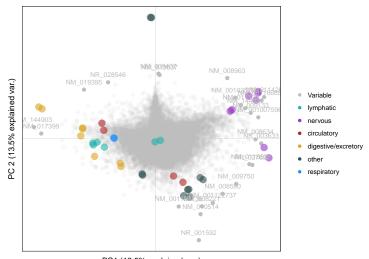
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Useful to also plot the points (v_{g1}, v_{g2}) for those genes g with large contributions to the first two principal components:



PC1 (18.5% explained var.)

PCA and eigendecomposition

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The SVD of $\underline{\tilde{\mathbf{X}}}$ also has the following interesting properties:

- $\underline{\underline{D}}$ the diagonal of $\underline{\underline{D}}$ contains the square roots of the eigenvalues of $\underline{\tilde{X}}\underline{\tilde{X}}^T$ (and thus also of $\underline{\tilde{X}}^T\underline{\tilde{X}}$).
- $\underline{\underline{U}}$ the columns of $\underline{\underline{U}}$ are the eigenvectors of the matrix $\underline{\tilde{X}}\underline{\tilde{X}}^T$, so that $\underline{\tilde{X}}\underline{\tilde{X}}^T\underline{\underline{U}} = \underline{\underline{U}}\underline{\underline{D}}^2$.
- \underline{V} the columns of \underline{V} are the n eigenvectors of the matrix $\underline{\tilde{X}}^T\underline{\tilde{X}}$ with non-zero eigenvalues, so that $\underline{\tilde{X}}^T\underline{\tilde{X}}V=\underline{V}\underline{D}^2$.

Note that:

 $\frac{1}{n-1}\tilde{\mathbf{X}}^T\tilde{\mathbf{X}}$ is the estimated gene-gene covariance matrix, and

 $\frac{1}{p-1}\tilde{\mathbf{X}}\tilde{\mathbf{X}}^T$ is the estimated sample-sample covariance matrix.

So SVD relates the eigendecompositions of the estimated geneand sample-covariance matrices.

Probabilistic PCA and factor analysis

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PCA using k PCs can be derived from small σ limit of model:

$$\mathbb{P}\left(\tilde{\mathbf{X}} = \tilde{\mathbf{x}} \mid M, [w_{ga}], (z_a)\right) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2\sigma^2} \sum_{g=1}^{p} \left(\tilde{x}_g - \sum_{a=1}^{k} w_{ga} z_a\right)^2\right]$$

where the k-vector $\mathbf{z} \sim \mathcal{N}(0, I)$ is a latent (unobserved) variable.

Probabilistic PCA and factor analysis

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where the k-vector $\mathbf{z} \sim \mathcal{N}(0, I)$ is a latent (unobserved) variable.

If instead of assuming that $\widetilde{X}_g - \sum_a w_{ga} Z_a \sim \mathcal{N}(0, \sigma^2)$ for some common fixed infinitesimal σ we allow:

$$E_g = \widetilde{X}_g - \sum_{a=1}^k w_{ga} Z_a \sim \mathcal{N}(0, \psi_g^2)$$

- where ψ_g is no longer assumed small
- ▶ but we retain assumption of between-gene independence, we derive the standard *factor analysis* model (Roweis & Ghahramani (1999)).

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Instead of Euclidean distances (\propto sums of squares), stochastic neighbor embedding attempts to preserve

- similarities scores analagous to probabilities
- while dramatically reducing dimensionality.

The similarity scores in full-dimensional space are derived from

$$p_{j|i} = \frac{\exp\left(-\frac{||\mathbf{x}_i - \mathbf{x}_j||^2}{2\sigma_i^2}\right)}{\sum\limits_{k \neq i} \exp\left(-\frac{||\mathbf{x}_i - \mathbf{x}_k||^2}{2\sigma_i^2}\right)}$$

which is then symmetrized to a joint probability analogue

$$p_{ij} = \frac{p_{i|j} + p_{j|i}}{2n}$$

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Instead of Euclidean distances (\propto sums of squares), stochastic neighbor embedding attempts to preserve

- similarities scores analagous to probabilities
- while dramatically reducing dimensionality.

For t-SNE, the (symmetric) similarity scores in low-dimensional space are

$$q_{ij} = \frac{\left(1 + ||\mathbf{y}_i - \mathbf{y}_j||^2\right)^{-1}}{\sum\limits_{k \neq I} \left(1 + ||\mathbf{y}_k - \mathbf{y}_I||^2\right)^{-1}}$$

How best to preserve full-dimensional similarities?

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Generally cannot perfectly match q_{ij} to p_{ij} .

Instead minimize Kullback-Leibler divergence between the two:

$$D_{\mathsf{KL}}(\underline{\mathbf{P}} \, || \, \underline{\mathbf{Q}}) = \sum_{i,j} p_{ij} \, (\log p_{ij} - \log q_{ij})$$

 D_{KL} is a key concept in information theory

- related to efficiency of encoding information about events
 - the true probabilities of which are <u>P</u>
 - using a code built on (incorrect) probabilities $\underline{\mathbf{Q}}$.

How to set σ_i

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Idea here is to fix constant *perplexity* value for all samples i

ightharpoonup by choosing σ_i such that, for each sample i

$$\prod_{j \neq i} p_{j|i}(\sigma_i)^{p_{j|i}(\sigma_i)} = \mathsf{desired}$$
 perplexity value

Intuition:

Require each sample to have similar # of "stochastic neigbors"

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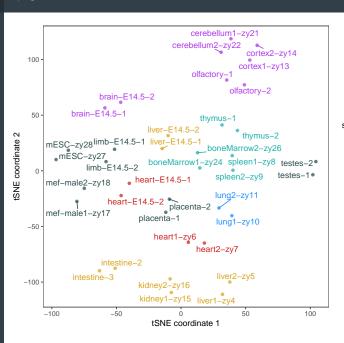
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system

- circulatory
- digestive/excretory
- lymphatic
- nervous
- other
- respiratory

What is a classifier?

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Learning: t-SNE

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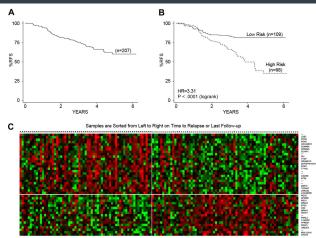
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A 38-gene expression classifier predictive of relapse-free survival (RFS) could distinguish 2 groups with differing relapse risks: low (4-year RFS, 81%, n = 109) versus high (4-year RFS, 50%, n = 98; P < .001).

Taken from Kang et al. (2010).

Classification by gene expression

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Goal:

Given sample i, assign class label y_i

ightharpoonup using measured gene expression levels x_{ig} .

Vector \mathbf{x}_i contains all gene measurements x_{ig} for sample i.

For simplicity, consider only two-class problems $y_i \in \{0, 1\}$.

Define random variables X and Y

 \triangleright of which x_i and y_i will be regarded as particular realizations.

Probalistic models yield $\mathbb{P}(Y = y \mid \mathbf{X} = \mathbf{x})$.

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Select modeling strategy M

- ightharpoonup apply algorithm to fit parameters heta
- using a set S_{train} of samples

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Select modeling strategy M

- ightharpoonup apply algorithm to fit parameters heta
- ightharpoonup using a set S_{train} of samples
- possibly by maximizing $\prod_{i \in S_{\text{train}}} \mathbb{P}_{M,\theta}(Y = y_i \mid \mathbf{X} = \mathbf{x}_i)$

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Select modeling strategy M

- lacktriangle apply algorithm to fit parameters $oldsymbol{ heta}$
- using a set S_{train} of samples
- ▶ possibly by maximizing $\prod_{i \in S_{train}} \mathbb{P}_{M,\theta}(Y = y_i \mid \mathbf{X} = \mathbf{x}_i)$
- lacktriangledown or maybe it should be $\prod_{i\in\mathcal{S}_{\mathsf{train}}}\mathbb{P}_{M}(m{\theta}\mid Y=y_{i},\mathbf{X}=\mathbf{x}_{i})$?

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- lacktriangledown or maybe it should be $\prod_{i \in S_{\mathsf{train}}} \mathbb{P}_M(\theta \mid Y = y_i, \mathbf{X} = \mathbf{x}_i)$?

Machine Learning point of view on classification: fit model should accurately classify samples $j \notin S_{\text{train}}$

 \blacktriangleright whose true classifications y_j may not already be known.

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- ▶ possibly by maximizing $\prod_{i \in S_{\text{train}}} \mathbb{P}_{M,\theta}(Y = y_i \mid \mathbf{X} = \mathbf{x}_i)$
- or maybe it should be $\prod_{i \in S_{\text{train}}} \mathbb{P}_{M}(\theta \mid Y = y_{i}, \mathbf{X} = \mathbf{x}_{i})?$

Machine Learning point of view on classification:

fit model should accurately classify samples $j \notin S_{\mathsf{train}}$

ightharpoonup whose true classifications y_j may not already be known.

Generally $(M, oldsymbol{ heta})$ less accurate for $j
otin S_{\mathsf{train}}$ than for $i \in S_{\mathsf{train}}$.

Thus useful to apply $(M, oldsymbol{ heta})$ to $j \in S_{\mathsf{test}}$ where $S_{\mathsf{test}} \cap S_{\mathsf{train}} = \emptyset$

▶ but where the $\{y_i \mid j \in S_{test}\}$ are still known.

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Perhaps simplest approach to classification:

k-nearest neighbors

Given feature vector \mathbf{x} (e.g., expression levels x_g)

▶ with *k* nearest training vectors

$$\{\mathbf{x}_j \mid j \in \mathsf{NN}_k\},\$$

with $\|\mathbf{x}_j - \mathbf{x}\| \le \|\mathbf{x}_i - \mathbf{x}\|$ if $j \in NN_k$ and $i \notin NN_k$:

$$\mathbb{P}(Y=1 \mid X=x) = \frac{1}{|\mathsf{NN}_k|} \sum_{j \in \mathsf{NN}_k} y_j$$

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$$\mathbb{P}(Y=1 \mid X=x) = \frac{1}{|\mathsf{NN}_k|} \sum_{j \in \mathsf{NN}_k} y_j$$

Best when feature space low-dimensional with natural metric.

k-nearest-neighbors is implemented as:

R class::knn

Python sklearn.neighbors.KNeighborsClassifier

```
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```

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```
# R:
knnTest = knn(
    train = xtrain,
    test = xtest,
    cl = ytrain,
    k = 3
)
nCorrect = sum(diag(table(knnTest, ytest)))
```

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Unsupervised Learning: t-SNF

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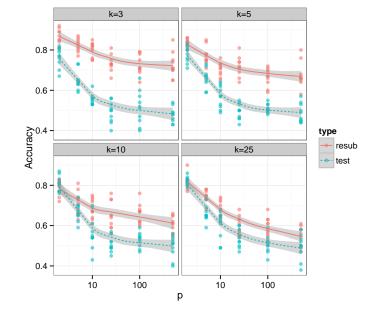
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knn and the curse of dimensionality

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Volume of p-dimensional hypersphere of radius r is

$$V_p(r) \propto r^p$$

For x to have many neighbors nearer than r

▶ must be many $x_i \in S_{train}$ in volume $V_p(r)$ centered at x.

knn and the curse of dimensionality

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For x to have many neighbors nearer than r

▶ must be many $x_i \in S_{train}$ in volume $V_p(r)$ centered at x.

If the dimensionality p is large and r is small, this is very unlikely.

So must use points far away to guess what's going on at x.

Not surprisingly this doesn't always work . . .

knn and the curse of dimensionality

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Not surprisingly this doesn't always work . . .

May be better to do feature selection or feature extraction

- and then fit model using reduced feature set
- will return to this idea later.

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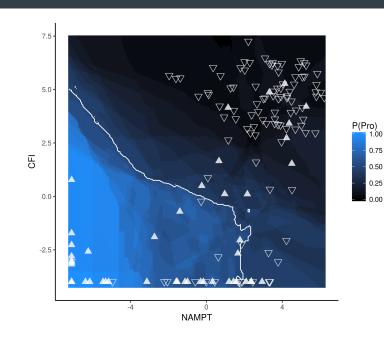
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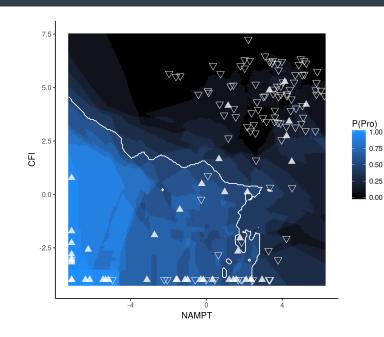
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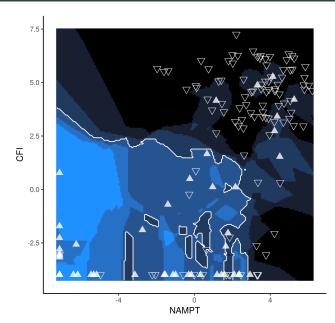
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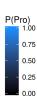
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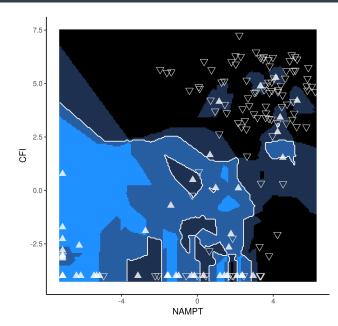
Overfitting

Cross-

Performanc

Metrics

Feature Selection





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Unsupervise Learning: PCA

Unsupervise Learning: t-SNE

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k-Neares Neighbor

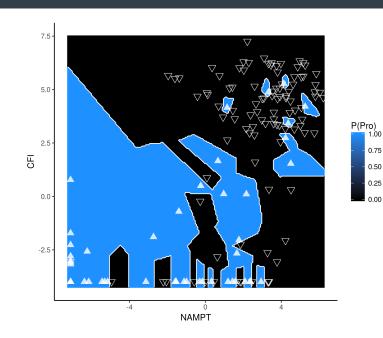
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Cross-Validation (CV)

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Evaluating performance by resubstitution suffers from bias.

But what if we don't have a test set S_{test} lying around?

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Evaluating performance by resubstitution suffers from bias.

But what if we don't have a test set S_{test} lying around?

Can split whatever sample set you have up into training+test set.

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Evaluating performance by resubstitution suffers from bias.

But what if we don't have a test set S_{test} lying around?

Can split whatever sample set you have up into training+test set.

If not many samples available: might split samples S into S_1 and S_2 and then try:

- 1. first train M on S_1 to obtain (M, θ_1) for testing on S_2 ;
- 2. then train on S_2 to model (M, θ_2) for testing on S_1 .

Unbiased performance estimate could then be obtained

- ▶ using the predictions $\mathbb{P}_{M,\theta_2}(Y \mid \mathbf{X})$ for samples in S_1 and
- ▶ predictions $\mathbb{P}_{M,\theta_1}(Y \mid X)$ for samples in S_2 .

K-Fold Cross-Validation

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Generalization: split S up into K subsets S_k for each of which:

- 1. a model (M, θ_{-k}) is trained using training set $S_{-k} = \bigcup_{q \neq k} S_q$
- 2. predictions $\mathbb{P}_{M,\theta_{-k}}(Y \mid \mathbf{X} = \mathbf{x}_i)$ are made for samples $i \in S_k$
- 3. performance estimates are made for each S_k based on $\mathbb{P}_{M,\theta_{-k}}(Y \mid \mathbf{X} = \mathbf{x}_i)$ and then averaged over all K folds.

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Very important:

Cross-validation valid only if all *supervised* steps in fitting model conducted separately within each of the *k* folds.

5-Fold Cross-Validation

Machine Learning Methods for Gene Expression Data

Cross-

Validation

```
# R:
library(caret)
knnCV = train(
    x = xtrain.
    v = vtrain.
    method = "knn",
    tuneGrid = data.frame(k=3),
    trControl = trainControl(method="cv", number=5)
cvAccuracyEstimate = knnCV$results[ , "Accuracy"]
```

Python:

```
from sklearn.neighbors import KNeighborsClassifier
from sklearn.cross_validation import cross_val_score
knnClass = KNeighborsClassifier(n neighbors=3)
cvAccs = cross_val_score(estimator = knnClass,
                         X = array(xtrain),
                         y = array(ytrain),
                         cv = 5
cvAccuracyEstimate = mean(cvAccs)
```

5-Fold Cross-Validation

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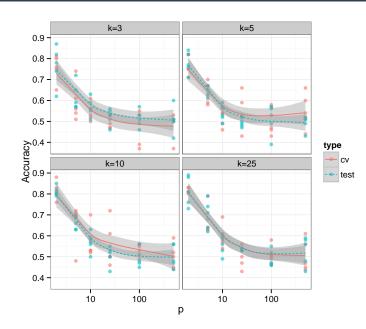
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Metrics—Binomial

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There are many ways to measure performance for classifiers; most are based only on the "discretized calls" \hat{y}

$$\hat{y}_{\mathcal{M}, oldsymbol{ heta}, \psi} = egin{cases} 1 & ext{if } \mathbb{P}_{\mathcal{M}, oldsymbol{ heta}}(Y = 1 \mid \mathbf{X} = \mathbf{x}) \geq \psi \\ 0 & ext{otherwise} \end{cases}$$

given some threshold ψ (e.g., $\psi = 0.5$).

Metrics—Binomial

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given some threshold ψ (e.g., $\psi = 0.5$).

Given a sample set S of size |S| = N composed of:

TP true positive samples: $y = \hat{y} = 1$

TN true negative samples: $y = \hat{y} = 0$

FP false positive samples: y = 0, $\hat{y} = 1$

FN false negative samples: $y = 1, \hat{y} = 0$,

define

Accuracy fraction of calls correct $\left(\frac{TP+TN}{N}\right)$

Sensitivity fraction of calls correct when y = 1 $\left(\frac{TP}{TP + FN}\right)$

Specificity fraction of calls correct when y = 0 $\left(\frac{TN}{TN + FP}\right)$

PPV fraction of calls correct when $\hat{y} = 1$ $\left(\frac{TP}{TP+FP}\right)$

NPV fraction of calls correct when $\hat{y} = 0$ $\left(\frac{TN}{TN + FN}\right)$.

Metrics—Receiver Operating Characteristic (ROC)

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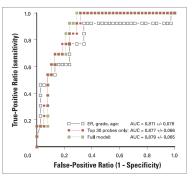


Fig 3. Receiver operating characteristic curves of three distinct pathologic complete response prediction models. The performance of the Diagonal Linear Discriminant Analysis-30 predictor and a predictor based on clinical variables and a combined clinical + pharmacogenomic prediction model are shown in the validation set fin = 511. ER extrogen receptor; AUC area under the curve.

Taken from Hess et al. (2006).

Consider binomial metrics over range of threshold values ψ .

Receiver operating characteristic (ROC) curve: sensitivity vs. specificity.

Metrics—Receiver Operating Characteristic (ROC)

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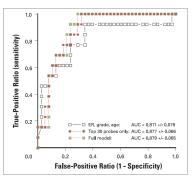


Fig 3. Receiver operating characteristic curves of three distinct pathologic complete response prediction models. The performance of the Diagonal Linear Discriminant Analysis—30 predictor and a predictor based on clinical variables and a combined clinical + pharmacogenomic prediction model are shown in the validation set (in = 51). ER actogen receptor; AUC area under the curve.

Taken from Hess et al. (2006).

Consider binomial metrics over range of threshold values ψ .

Receiver operating characteristic (ROC) curve: sensitivity vs. specificity.

Area under ROC curve (AUC) equals probability that

► score $\mathbb{P}(Y = 1 \mid \mathbf{X} = \mathbf{x})$ of a randomly chosen positive case (y = 1)

is higher than

score of a randomly chosen negative case (y = 0).

Feature selection

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Often assumed that expression patterns of most genes either:

- 1. uninformative or
- 2. redundant with a few maximally useful markers with respect to a particular classification task.

Feature selection attempts to identify optimal set of markers for inclusion in model.

Feature selection not always required but resulting simplification:

- 1. reduces computational workload,
- 2. can help to avoid overfitting (though feature selection can itself be susceptible to overfitting), and
- 3. facilitates model platform migration.

Taxonomy (adapted from Saeys et al. (2007))

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Filter Selection done before and independently of classifier construction. Can be univariate or multivariate.

Wrapper Embed classifier construction within feature selection process. Heuristic search methods compare models, favor adding or removing features based on optimization of some specified metric on resulting classifiers.

Embedded Feature selection is inherently built into some classifier construction methods.

Taxonomy (adapted from Saeys et al. (2007))

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Category	Advantages	Disadvantages	Examples
	Univariate		
Filter -	Fast Scalable Independent of classifier Multivariate	- feature dependencies - interaction w/classifier	t-test, ANOVA Wilcox test Rank Product
	+ feature dependencies Independent of classifier Intermediate complexity	Slower Less Scalable - interaction w/classifier	CFS Markov Blanket Filter
	Deterministic		
Wrapper –	Simple + interaction w/classifier + feature dependencies	Risk of over-fitting Greedy (local optima) Classifier dependent selec- tion	Forward Selection Backward Elimination Plus q minus r
	Randomized		
	Less prone to local optima + interaction w/classifier	High risk over-fitting Computationally intensive	Simulated Annealing Randomized Hill Climb- ing Genetic Algorithms
	+ feature dependencies	Classifier dependent selection	
Embedded	+ interaction w/classifier + feature dependencies Intermediate complexity	No modularity Restrict algorithms	Decision trees Weighted Naive Bayes LASSO regression

Linear Models for Univariate Feature Selection

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Most common univariate filter is *t*-test (or *F*-test if more than 2 groups) originating from model

$$\mathbb{P}(X_g = x \mid Y = y) = \frac{1}{\sqrt{2\pi\sigma_g^2}} \exp\left[\frac{(x - \mu_{yg})^2}{2\sigma_g^2}\right]$$

Now considering conditional probabilities of X_g given Y ...

- ... and considering each X_g separately!
 - "Univariate" analysis; not necessarily realistic, but tractable.

Linear Models

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Alternately, this model may be described by

$$X_{g} = \mu_{0g} + (\mu_{1g} - \mu_{0g})Y + \sigma_{g}\epsilon_{g}$$

with $\epsilon_{m{g}} \sim \mathcal{N}(0,1)$, implying

$$\mathbb{E}(X_g \mid Y = y) = \mu_{yg}$$

$$\mathbb{V}(X_g \mid Y = y) = \sigma_g^2$$

Can then rank features based on $|t_g|$ where:

$$t_{\mathsf{g}} = \frac{\hat{\mu}_{\mathsf{0g}} - \hat{\mu}_{\mathsf{1g}}}{\hat{\sigma}_{\mathsf{g}} \sqrt{\frac{1}{n_{\mathsf{0}}} + \frac{1}{n_{\mathsf{1}}}}}$$

Linear Models

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$$t_{g} = \frac{\hat{\mu}_{0g} - \hat{\mu}_{1g}}{\hat{\sigma}_{g}\sqrt{\frac{1}{n_{0}} + \frac{1}{n_{1}}}}$$

While X_g is linear in Y here,

lacktriangleright it is linearity in μ_{yg} that makes "linear model" in statistics.

Linear Models

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Alternately, this model may be described by

$$\mathbb{E}(X_g \mid Y = y) = \mu_{yg}$$

$$\mathbb{V}(X_g \mid Y = y) = \sigma_g^2$$

 $X_{\sigma} = \mu_{0\sigma} + (\mu_{1\sigma} - \mu_{0\sigma})Y + \sigma_{\sigma}\epsilon_{\sigma}$

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$$t_g = rac{\hat{\mu}_{0g} - \hat{\mu}_{1g}}{\hat{\sigma}_g \sqrt{rac{1}{n_0} + rac{1}{n_1}}}$$

While X_{σ} is linear in Y here,

ightharpoonup it is linearity in μ_{yg} that makes "linear model" in statistics.

With a few modifications linear models can model $Y \mid \mathbf{X}$ too.

knn Accuracy With t-Test Feature Selection



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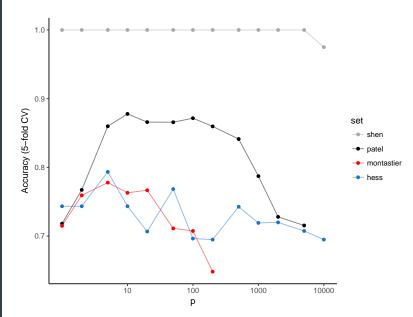
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