

STAT 431 — Applied Bayesian Analysis — Course Notes

# Probability Review

Spring 2019

# Sample Space and Events

**sample space:** all possible outcomes (fully specified)

**event:** subset of sample space

Eg: Two coin flips

$S = \text{sample space} = \{HH, HT, TH, TT\}$

event  $A = \text{both flips same} = \{HH, TT\}$

[ Illustrate ... ]

The usual set operations apply to events:

$$\begin{aligned} A \cup B &= \textbf{union} \text{ of } A \text{ and } B \\ &= \text{outcomes in either (or both)} \end{aligned}$$

$$\begin{aligned} A \cap B &= \textbf{intersection} \text{ of } A \text{ and } B \\ &= \text{outcomes in both} \end{aligned}$$

$$\begin{aligned} \overline{A} &= \textbf{complement} \text{ of } A \\ &= \text{outcomes not in } A \end{aligned}$$

Events  $A$  and  $B$  are **disjoint** if

$$A \cap B = \emptyset \quad (\text{the null set})$$

# Probability

**probability:** assigns to each event  $A$  a number  $P(A)$ , with such properties as

- ▶  $0 \leq P(A) \leq 1$
- ▶  $P(\emptyset) = 0$  ( $\emptyset$  is the “null event”)
- ▶  $P(S) = 1$  ( $S$  is the sample space)
- ▶ if  $A$  and  $B$  are disjoint,

$$P(A \cup B) = P(A) + P(B)$$

- if  $A_1, \dots, A_n$  are disjoint,

$$P\left(\bigcup_{i=1}^n A_i\right) = \sum_{i=1}^n P(A_i)$$

- $P(\overline{A}) = 1 - P(A)$

Eg: two fair coin flips

$A$  = both heads

$B$  = both tails

$$P(A) = ?$$

$$P(B) = ?$$

$$P(\overline{A}) = ?$$

$$P(A \cup B) = ?$$

$$P(A \cap B) = ?$$

# Conditioning

If  $P(B) \neq 0$ , the **conditional probability** of  $A$  given  $B$  is

$$P(A \mid B) = \frac{P(A \cap B)}{P(B)}$$

( $P(A)$  is sometimes called the **marginal probability** of  $A$ )

Conditional probabilities behave like ordinary probabilities:

$$\text{e.g. } 0 \leq P(A \mid B) \leq 1$$

$$\text{e.g. } P(\overline{A} \mid B) = 1 - P(A \mid B)$$

Note:

$$\begin{aligned}P(A \cap B) &= P(B) P(A \mid B) && (P(B) \neq 0) \\&= P(A) P(B \mid A) && (P(A) \neq 0)\end{aligned}$$

In general,

$$\text{joint} = \text{marginal} \times \text{conditional}$$



Eg: two fair coin flips

$$A = \{HH, TT\} \quad B = \{HH, HT, TH\}$$

$$P(A) = ? \quad P(B) = ? \quad P(A \cap B) = ?$$

$$P(A \mid B) = ?$$

$$P(B \mid A) = ?$$

[ Interpret ... ]

Bayesian perspective:

$M$  = a particular model for the data

$D$  = (event of) the data

$P(M)$  = probability of  $M$  if we have no other information  
= “prior”

$P(D \mid M)$  = probability given to  $D$  when  $M$  is true  
= “likelihood”

$P(M \mid D)$  = probability of  $M$  after observing  $D$   
= “posterior”

# Bayes' Rule

[ Illustrate sample space ... ]

Notice:

$$P(A) = P((A \cap B) \cup (A \cap \overline{B}))$$

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$$\begin{aligned}P(A) &= P((A \cap B) \cup (A \cap \overline{B})) \\&= P(A \cap B) + P(A \cap \overline{B})\end{aligned}$$

# Bayes' Rule

[ Illustrate sample space ... ]

Notice:

$$\begin{aligned}P(A) &= P((A \cap B) \cup (A \cap \overline{B})) \\&= P(A \cap B) + P(A \cap \overline{B}) \\&= P(B) P(A | B) + P(\overline{B}) P(A | \overline{B})\end{aligned}$$

(provided  $0 < P(B) < 1$ )

**Bayes' Rule** (simple form):

$$P(B \mid A) = \frac{P(A \cap B)}{P(A)}$$

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$$\begin{aligned} P(B \mid A) &= \frac{P(A \cap B)}{P(A)} \\ &= \frac{P(B) P(A \mid B)}{P(B) P(A \mid B) + P(\overline{B}) P(A \mid \overline{B})} \end{aligned}$$

(provided all conditional probabilities exist)

Eg: Say the pop. of Cyprus is 80% Greek, 20% Turkish.

Suppose English is spoken by 90% of the Greeks and 50% of the Turks.

What's the prob. and English-speaking Cypriot is Greek?

$A$  = speaks English       $B$  = is Greek

Expression for the desired probability?

$$P(B) = ? \qquad P(\overline{B}) = ?$$

$$P(A | B) = ? \qquad P(A | \overline{B}) = ?$$

Answer?



Now generalize ...

Suppose  $B_1, B_2, B_3, \dots$  form a **partition** of  $S$ :

- ▶ all are disjoint
- ▶  $\bigcup_{\text{all } j} B_j = S$  (exhaustive)

Also, assume  $P(B_j) \neq 0$ , all  $j$ .

**Law of Total Probability:**

$$P(A) = \sum_{\text{all } j} P(B_j) P(A \mid B_j)$$

[ Illustrate ... ]

**Bayes' Rule** (for probabilities):

If  $B_1, B_2, \dots$  is a partition,

$$P(B_i | A) = \frac{P(B_i) P(A | B_i)}{\sum_{\text{all } j} P(B_j) P(A | B_j)}$$

(The previous special case had  $B_1 = B$ ,  $B_2 = \overline{B}$ .)

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So

$$P(B_i | A) \propto P(B_i) P(A | B_i)$$

(since the denominator doesn't depend on  $i$ )

Bayesian application:

$M_1, M_2, \dots$  = distinct models for the data

$D$  = (event of) the data

By Bayes' Rule,

$$P(M_i | D) \propto P(M_i) P(D | M_i)$$

posterior  $\propto$  prior  $\times$  likelihood

The (inverse) proportionality constant

$$\sum_{\text{all } j} P(M_j) P(D | M_j)$$

is called the **normalizing constant**.

Eg: Waldo (revisited)



# Independent Events

Events  $A$  and  $B$  are **independent** if

$$P(A \cap B) = P(A) P(B)$$

(otherwise **dependent**)

If  $P(B) \neq 0$ , this is the same as

$$P(A | B) = P(A)$$

(conditional = marginal)

Events  $A$  and  $B$  are **conditionally independent** given  $C$  if

$$P(A \cap B \mid C) = P(A \mid C) P(B \mid C)$$

Note:  $A$  and  $B$  are not necessarily independent!

Often there are data events that are independent conditional on the model:

$$P(D_1 \cap D_2 \mid M) = P(D_1 \mid M) P(D_2 \mid M)$$

That is, the likelihood may factor.



# Random Variables and Distributions

**random variable:** real-valued function on the sample space

May be ...

- ▶ **discrete:** takes values in a countable set  
e.g. binomial, geometric, Poisson
- ▶ **continuous:** takes values on a continuum  
e.g. normal, exponential, gamma

The **distribution** of a random variable  $X$  is characterized by its **density**:

- ▶ Discrete density:

$$p(x) = P(X = x)$$

(sometimes called a “mass function”)

- ▶ Continuous density:  $p(x)$  such that

$$\int_G p(x) dx = P(X \in G)$$

(often called a p.d.f.)

The **joint distribution** of random variables  $X$  and  $Y$  can often be characterized by a **joint density**

$$p(x, y)$$

► Both discrete:

$$p(x, y) = \mathbf{P}(X = x, Y = y) = \mathbf{P}(\{X = x\} \cap \{Y = y\})$$

► Jointly continuous:

$$\int_{G_1} \int_{G_2} p(x, y) \, dy \, dx = \mathbf{P}(X \in G_1, Y \in G_2)$$

The individual densities of  $X$  and  $Y$  are their **marginal densities**, which define their **marginal distributions**.

Eg:

$$p(x) = \begin{cases} \sum_{\text{all } y} p(x, y), & Y \text{ discrete} \\ \int p(x, y) dy, & Y \text{ continuous} \end{cases}$$

# Conditioning

The **conditional distribution** of  $X$  given  $Y$  is characterized by the **conditional density**

$$p(x \mid y) = \frac{p(x, y)}{p(y)} \quad (\text{wherever } p(y) > 0)$$

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The **conditional distribution** of  $X$  given  $Y$  is characterized by the **conditional density**

$$p(x \mid y) = \frac{p(x, y)}{p(y)} \quad (\text{wherever } p(y) > 0)$$

Note:

$$p(x, y) = p(y) p(x \mid y) = p(x) p(y \mid x)$$

is another example of the general form

$$\text{joint} = \text{marginal} \times \text{conditional}$$

This idea can be used to **define** the joint density when  $X$  and  $Y$  are of different types.

For example, if  $X$  is continuous and  $Y$  is discrete, let

$$p(x, y) = p(y) p(x | y) = p(x) p(y | x)$$

where

$$p(x | y) = \text{a continuous density for each } y$$

$$p(y | x) = \text{a discrete density for each } x$$

(Use whichever of these is most convenient.)

A general process for working with the joint distribution of  $X$  and  $Y$ :

1. Specify the **marginal** density of  $X$
2. Specify the **conditional** density of  $Y$  given  $X$
3. Use the product of these densities as their joint density



## Example: Uniform-Binomial

$$X \sim \text{uniform}(0, 1)$$

$$Y \mid X = x \sim \text{binomial}(n, x)$$

( $n$  is a given “number of trials”,  $x$  is “success prob.”)

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( $n$  is a given “number of trials”,  $x$  is “success prob.”)

So

$$p(x) = \begin{cases} 1 & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$p(y \mid x) = \binom{n}{y} x^y (1 - x)^{n-y} \quad y = 0, \dots, n$$

... and the “joint density” is

$$p(x)p(y \mid x) = \begin{cases} \binom{n}{y} x^y (1-x)^{n-y} & 0 < x < 1, y = 0, \dots, n \\ 0 & \text{otherwise} \end{cases}$$

The **marginal** density for  $X$  is (of course)

$$p(x) = \begin{cases} 1 & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

The marginal density for  $Y$  is (for  $y = 0, \dots, n$ )

$$\begin{aligned} p(y) &= \int p(x) p(y \mid x) dx \\ &= \int_0^1 \binom{n}{y} x^y (1-x)^{n-y} dx \\ &= \binom{n}{y} \int_0^1 \underbrace{x^y (1-x)^{n-y}}_{\text{"kernel" of a beta density}} dx \end{aligned}$$

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Recall density of  $\text{beta}(\alpha, \beta)$  distribution:

$$\frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} \quad 0 < x < 1$$

(see Cowles, Table A.2)

Thus, for  $y = 0, \dots, n$ ,

$$p(y) = \binom{n}{y} \frac{\Gamma(y+1)\Gamma(n-y+1)}{\Gamma(n+2)} \cdot \int_0^1 \underbrace{\frac{\Gamma(n+2)}{\Gamma(y+1)\Gamma(n-y+1)} x^y (1-x)^{n-y}}_{\text{beta}(y+1, n-y+1) \text{ density}} dx$$

Thus, for  $y = 0, \dots, n$ ,

$$\begin{aligned}
 p(y) &= \binom{n}{y} \frac{\Gamma(y+1)\Gamma(n-y+1)}{\Gamma(n+2)} \\
 &\quad \cdot \underbrace{\int_0^1 \frac{\Gamma(n+2)}{\Gamma(y+1)\Gamma(n-y+1)} x^y (1-x)^{n-y} dx}_{\text{beta}(y+1, n-y+1) \text{ density}} \\
 &= \binom{n}{y} \frac{y! (n-y)!}{(n+1)!} \cdot 1
 \end{aligned}$$

Thus, for  $y = 0, \dots, n$ ,

$$\begin{aligned}
 p(y) &= \binom{n}{y} \frac{\Gamma(y+1)\Gamma(n-y+1)}{\Gamma(n+2)} \\
 &\quad \cdot \underbrace{\int_0^1 \frac{\Gamma(n+2)}{\Gamma(y+1)\Gamma(n-y+1)} x^y (1-x)^{n-y} dx}_{\text{beta}(y+1, n-y+1) \text{ density}} \\
 &= \binom{n}{y} \frac{y! (n-y)!}{(n+1)!} \cdot 1 \\
 &= \frac{n!}{y! (n-y)!} \frac{y! (n-y)!}{(n+1) \cdot n!} = \frac{1}{n+1}
 \end{aligned}$$



So the marginal distribution of  $Y$  is a “discrete uniform” distribution:

$$p(y) = \begin{cases} \frac{1}{n+1} & y = 0, \dots, n \\ 0 & \text{otherwise} \end{cases}$$

[ Illustrate ... ]

# Bayes' Rule

**Bayes' Rule** (for densities):

$$p(y \mid x) = \frac{p(y) p(x \mid y)}{C}$$

where

$$C = p(x) = \begin{cases} \sum_{\text{all } y} p(y) p(x \mid y), & Y \text{ discrete} \\ \int p(y) p(x \mid y) dy, & Y \text{ continuous} \end{cases}$$

$C$  is the **normalizing constant**.

Bayesian application:

Suppose the model is (fully) defined by a **parameter**  $\theta$ .

Let  $y$  be the observed data.

Then

$$\begin{aligned} p(\theta \mid y) &\propto p(\theta) \cdot p(y \mid \theta) \\ \text{posterior} &\propto \text{prior} \times \text{likelihood} \end{aligned}$$

where the proportionality is in  $\theta$  (for fixed  $y$ ).

(The likelihood is sometimes written as  $L(\theta; y)$ .)

## More Variables

Densities can be extended to three or more variables.

E.g.  $X$ ,  $Y$ , and  $Z$  could have a joint density defined by

$$p(x, y, z) = p(x) p(y | x) p(z | x, y)$$

where conditioning on two variables is defined as, e.g.

$$p(z | x, y) = \frac{p(x, y, z)}{p(x, y)} \quad (\text{wherever } p(x, y) > 0)$$

The marginal densities would be denoted

$$\begin{array}{ccc} p(x, y), & p(x, z), & p(y, z), \\ p(x), & p(y), & p(z) \end{array}$$

Marginal densities are obtained by summing/integrating out the other variables, e.g.

$$p(x, z) = \begin{cases} \sum_{\text{all } y} p(x, y, z), & Y \text{ discrete} \\ \int p(x, y, z) dy, & Y \text{ continuous} \end{cases}$$

Similarly, joint conditionals can be defined as, e.g.

$$p(x, y \mid z) = \frac{p(x, y, z)}{p(z)} \quad (\text{wherever } p(z) > 0)$$

Certain rules for marginal densities extend to conditional densities, e.g.

$$p(x, y \mid z) = p(x \mid z) p(y \mid x, z)$$

# Independent Random Variables

$X$  and  $Y$  are **independent** when they have a joint density that factors into marginals:

$$p(x, y) = p(x) p(y)$$

Note: If  $X$  and  $Y$  are independent,

$$p(x \mid y) = p(x) \qquad p(y \mid x) = p(y)$$

Note: If  $p(x \mid y)$  doesn't depend on  $y$  (or if  $p(y \mid x)$  doesn't depend on  $x$ ), then  $X$  and  $Y$  are independent. (Why?)

Let  $Z$  be another random variable.

$X$  and  $Y$  are **conditionally independent given**  $Z = z$  if

$$p(x, y \mid z) = p(x \mid z) p(y \mid z)$$

In general, this does **not** imply that  $X$  and  $Y$  are (marginally) independent.



$X$  and  $Y$  are **conditionally independent given**  $Z$  if

$$p(x, y \mid z) = p(x \mid z) p(y \mid z) \quad \text{for all } z \quad (p(z) > 0)$$

This is (almost) equivalent to

$$p(x \mid y, z) = p(x \mid z)$$

and to

$$p(y \mid x, z) = p(y \mid z)$$

# Measures of Location and Spread

The **expected value** or **mean** of  $X$  is

$$E(X) = \begin{cases} \sum_{\text{all } x} x p(x), & X \text{ discrete} \\ \int x p(x) dx, & X \text{ continuous} \end{cases}$$

A **median**  $m_X$  of  $X$  satisfies

$$P(X < m_X) \leq 0.5 \quad \text{and} \quad P(X > m_X) \leq 0.5$$

A **mode** of  $X$  is a value maximizing  $p(x)$ . It need not exist or be unique.

An  $\alpha$ -**quantile**  $x_\alpha$  of  $X$  satisfies

$$P(X < x_\alpha) \leq \alpha \quad \text{and} \quad P(X > x_\alpha) \leq 1 - \alpha$$

If  $X$  is continuous,

$$P(X \leq x_\alpha) = \alpha$$

[ Illustrate ... ]

The **variance** of  $X$  is

$$\text{Var}(X) = \text{E}((X - \mu_X)^2)$$

where  $\mu_X = \text{E}(X)$ .

An **interquartile range (IQR)** of  $X$  is

$$x_{0.75} - x_{0.25}$$

(i.e. the difference between the first and third quartile)

The **conditional expected value** (or **conditional mean**) of  $X$  given  $Y = y$  is

$$E(X | Y = y) = \begin{cases} \sum_{\text{all } x} x p(x | y), & X \text{ discrete} \\ \int x p(x | y) dx, & X \text{ continuous} \end{cases}$$

The **conditional variance** of  $X$  given  $Y = y$  is

$$\text{Var}(X | Y = y) = E((X - \mu_{X|y})^2 | Y = y)$$

where  $\mu_{X|y} = E(X | Y = y)$ .

Notational note:

We sometimes write

$$E(X \mid y) \quad \text{for} \quad E(X \mid Y = y)$$

$$\text{Var}(X \mid y) \quad \text{for} \quad \text{Var}(X \mid Y = y)$$

Similarly, write

$$X \mid y \sim \dots \quad \text{for} \quad X \mid Y = y \sim \dots$$

# Transformation of Variables

Suppose  $X$  is continuous, with density  $p(x)$ , and let

$$Y = g(X)$$

where  $g$  has a differentiable inverse  $g^{-1}$ .

Then  $Y$  is continuous, with density

$$p(y) = p(x) \left| \frac{dx}{dy} \right|$$

where  $x$  is (implicitly) equal to  $g^{-1}(y)$ .

This **transformation-of-variables formula** is sometimes more explicitly written as

$$p_Y(y) = p_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right|$$

where  $p_X$  and  $p_Y$  are densities of  $X$  and  $Y$ .