

HW3 Solutions

STAT431 Spr19

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1 Conjugate joint prior density

1.1 Values of parameters in Normal-inverse gamma prior

The four parameters in the Normal Inverse-gamma prior are $\mu_0, \kappa, \alpha, \beta$. For the conditional normal part, we want a prior mean of 0 and an equivalent prior sample size of 1, so we set $\mu_0 = 0$ and $\kappa = 1$ for the conditional normal part.

For the inverse-gamma portion, we want σ^2 to marginally have prior mean $\frac{\beta}{\alpha-1} = 1$ and prior variance $\frac{\beta^2}{(\alpha-1)^2(\alpha-2)} = 1$. We then solve for the values of α, β .

$$\frac{\beta}{\alpha - 1} = 1 \Rightarrow \frac{\beta^2}{(\alpha - 1)^2} = 1 \quad (1)$$

$$\Rightarrow \frac{1}{(\alpha - 2)} = 1 \quad (2)$$

$$\Rightarrow \alpha = 3 \quad (3)$$

$$\Rightarrow \beta = 2 \quad (4)$$

We set $\alpha = 3$ and $\beta = 2$ for the inverse-gamma parameters, and we set $\mu_0 = 0$ and $\kappa = 1$ for the conditional normal part.

1.2 Marginal posterior for μ

The marginal posterior density for μ is the t distribution, with mean 4.6188, scale parameter 0.1868, and degrees of freedom 16.

```
n = length(Y)
my_kappa = 1
mu_0 = 0
my_alpha = 3
my_beta = 2
```

```
n_0 = my_alpha * 2
sigmasq_0 = my_beta * 2 / n_0
```

```
# Mean
post_mean = (my_kappa*mu_0 + n*mean(Y)) / (my_kappa + n)
```

```
# Scale
post_scale = (n_0*sigmasq_0 + (n-1)*var(Y) + (my_kappa*n*(mean(Y) - mu_0)^2)/(my_kappa +
  ((my_kappa + n)*(n_0 + n))
```

```
# Degrees of Freedom
post_df = n + n_0
```

1.3 95% Credible interval for μ

The posterior 95% equal-tailed credible interval for μ is equal to (3.702, 5.535).

```
(post_mean + qt(c(0.025, 0.975), post_df) * sqrt(post_scale)) %>% round(3)
```

1.4 Marginal posterior for σ^2

The marginal posterior distribution for σ^2 is Inverse-gamma, with parameters shape=8, scale=16.445.

```
# Shape
my_alpha + n/2

# Scale
my_beta + (n-1)/2*var(Y) + (my_kappa*n*(mean(Y) - mu_0)^2) / (2*(my_kappa+n))
```

1.5 95% Credible interval for σ^2

The posterior 95% equal-tailed credible interval for σ^2 is equal to (1.1402, 4.761).

```
1 / qgamma(c(0.975, 0.025), post_shape, post_scale)
```

2 Standard noninformative prior

2.1 Marginal posterior for μ

The marginal posterior distribution for μ is the t -distribution, with parameters mean $\bar{Y} = 5.0807$, scale parameter $\frac{s^2}{n} = 0.06025$, and degrees of freedom $n - 1 = 9$.

```
post_mean = mean(Y)
post_scale = var(Y) / n
post_df = n-1
```

2.2 95% Credible interval for μ

```
(post_mean + qt(c(0.025, 0.975), post_df) * sqrt(post_scale)) %>% round(3)
```

The posterior 95% equal-tailed credible interval for μ is equal to (4.525, 5.636).

2.3 Marginal posterior for σ^2

The marginal posterior distribution for σ^2 is Inverse-gamma, with parameters shape=4.5, scale=2.711.

```
post_shape = (n-1) / 2
post_scale = (n-1)*var(Y) / 2
```

2.4 95% Credible interval for σ^2

The posterior 95% equal-tailed credible interval for σ^2 is equal to (0.285, 2.008).

```
1 / qgamma(c(0.975, 0.025), post_shape, post_scale)
```

3 GRADUATE ONLY - Deriving Jeffrey's prior

3.1 Information matrix derivations

$$\ln L(\mu, \sigma^2 | Y) = -\frac{n}{2} \ln(2\pi) - \frac{n}{2} \ln(\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (Y_i - \mu)^2 \quad (5)$$

$$\frac{d}{d\sigma^2} \ln L(\mu, \sigma^2 | Y) = -\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n (Y_i - \mu)^2 \quad (6)$$

$$\frac{d}{d\mu} \frac{d}{d\sigma^2} \ln L(\mu, \sigma^2 | Y) = \frac{d}{d\mu} \left(\frac{1}{2(\sigma^2)^2} \sum_{i=1}^n (Y_i^2 - 2\mu Y_i + \mu^2) \right) \quad (7)$$

$$= \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n (-2Y_i + 2\mu) \quad (8)$$

$$\mathbf{E} \left[\frac{d^2}{d\mu d\sigma^2} \ln L(\mu, \sigma^2 | Y) \right] = \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n \mathbf{E}(-2Y_i + 2\mu) \quad (9)$$

$$= 0 \quad (10)$$

The last line follows because $\mathbf{E}(Y_i) = \mu$ by definition.

3.2 Jeffrey's prior derivation

In part (a) we found that $I_{12}(\mu, \sigma^2) = 0$. Therefore the determinant of the information matrix is only the product of the diagonal elements I_{11}, I_{22} .

$$\det(I) = I_{11}I_{22} = \frac{n}{\sigma^2} \frac{n}{2(\sigma^2)^2} = \frac{n^2}{2(\sigma^2)^3} \quad (11)$$

Therefore the Jeffreys prior density is proportional to $\sqrt{\det(I)} = \frac{n}{\sqrt{2}(\sigma^2)^{3/2}}$