## Stat 432 Homework 5

Assigned: Feb 23, 2019; Due: 11:59pm Mar 2, 2019

Question 1 (linear regression)

[2 points] Sorry for the confusion, I meant to say page 21. You get 2 points for free for this problem. The irreducible error is  $E[\|\mathbf{y}^* - \mathbf{X}\boldsymbol{\beta}\|^2] = n\sigma^2$ , which comes from the model assumption and cannot be reduced further; the bias (of the estimator  $\mathbf{X}\hat{\boldsymbol{\beta}}$ ) is  $\mathbf{X}\boldsymbol{\beta} - E[\mathbf{X}\hat{\boldsymbol{\beta}}] = 0$ , and the variance (of the estimator) is  $E[\|\mathbf{X}\hat{\boldsymbol{\beta}} - E[\mathbf{X}\boldsymbol{\beta}\|^2] = p\sigma^2$ .

[2 points] This is actually written on page 15 of lecture note penalized. Note that in the ridge regression, we introduce some bias  $\operatorname{Bias}(\widehat{\beta}_j^{\operatorname{ridge}}) = \frac{-\lambda}{1+\lambda}\beta_j$  to trade off for the reduced variance  $\operatorname{Var}(\widehat{\beta}_j^{\operatorname{ridge}}) = \frac{1}{(1+\lambda)^2}\operatorname{Var}(\widehat{\beta}_j^{\operatorname{ols}})$ . Therefore, the overall the prediction error will be smaller than that of OLS:

$$\operatorname{Bias}^{2}(\widehat{\beta}_{j}^{\operatorname{ridge}}) + \operatorname{Var}(\widehat{\beta}_{j}^{\operatorname{ridge}}) = \frac{\lambda^{2}}{(1+\lambda)^{2}} \beta_{j}^{2} + \frac{1}{(1+\lambda)^{2}} \operatorname{Var}(\widehat{\beta}_{j}^{\operatorname{ols}}) < \operatorname{Var}(\widehat{\beta}_{j}^{\operatorname{ols}})$$

which is always satisfied for some  $\lambda$  because the quadratic curve of  $\lambda$  passes the origin. 1 point for specifying bias and variance in the ridge regression setting; 1 point for discussing the bias-variance trade-off.

Question 2 (model selection criteria)

The Boston Housing data is a classical dataset that models the median house values medv of different areas of Boston. Because a lot of variables exhibit an asymmetry, we will use some transformations.

```
data(Boston, package="MASS")
   head(Boston)
       crim zn indus chas
                            nox
                                       age
                                              dis rad tax ptratio black lstat medv
                                   rm
## 1 0.00632 18 2.31
                        0 0.538 6.575 65.2 4.0900
                                                    1 296
                                                             15.3 396.90 4.98 24.0
## 2 0.02731 0 7.07
                        0 0.469 6.421 78.9 4.9671
                                                    2 242
                                                             17.8 396.90 9.14 21.6
## 3 0.02729 0 7.07
                        0 0.469 7.185 61.1 4.9671
                                                  2 242
                                                             17.8 392.83 4.03 34.7
## 4 0.03237 0 2.18
                        0 0.458 6.998 45.8 6.0622
                                                   3 222
                                                             18.7 394.63 2.94 33.4
## 5 0.06905 0 2.18
                        0 0.458 7.147 54.2 6.0622
                                                   3 222
                                                             18.7 396.90 5.33 36.2
## 6 0.02985 0 2.18
                        0 0.458 6.430 58.7 6.0622
                                                    3 222
                                                             18.7 394.12 5.21 28.7
   useLog = c(1,3,5,6,8,9,10,14)
   Boston[,useLog] = log(Boston[,useLog])
   Boston[,2] = Boston[,2] / 10
   Boston[,7] = Boston[,7]^2.5 / 10^4
   Boston[,11] = \exp(0.4 * Boston[,11])/1000
   Boston[,12] = Boston[,12] / 100
   Boston[,13] = sqrt(Boston[,13])
```

## part a)

[1 point]

We fit the following the linear model and summarize the outputs

```
lmfit = lm(medv~., Boston)
summary(lmfit)$coefficients

## Estimate Std. Error t value Pr(>|t|)
```

```
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 4.176874035 0.379016727 11.0202895 2.158941e-25
## crim -0.014606367 0.011650102 -1.2537545 2.105267e-01
```

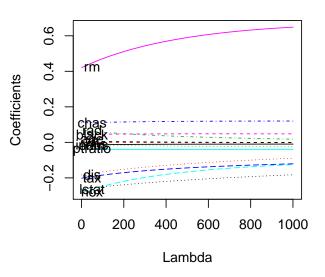
```
## zn
               0.001391943 0.005638624 0.2468585 8.051207e-01
               -0.012709368 0.022311989 -0.5696206 5.691950e-01
## indus
## chas
               0.109980144 0.036633710 3.0021569 2.817052e-03
               -0.283111884 0.105340487 -2.6875885 7.440816e-03
## nox
## rm
               0.421107840 0.110174650 3.8221845 1.492369e-04
               0.006403368 0.004863019 1.3167476 1.885362e-01
## age
              -0.183154286 0.036803637 -4.9765268 8.966317e-07
## dis
## rad
               0.068361590 0.022473189
                                         3.0419176 2.476295e-03
## tax
               -0.201832385 0.048432167 -4.1673209 3.641214e-05
## ptratio
              -0.040017441 0.008091477 -4.9456285 1.043293e-06
## black
               0.044471934 0.011455971
                                          3.8819872 1.177364e-04
## lstat
               -0.262615094 0.016091240 -16.3203760 3.598708e-48
part b)
[3 points, 1 point for each subproblem.] * We can use leaps package to calculate the Mallow's Cp statistics.
# load leaps and calcualte the best subset fit
library(leaps)
RSSleaps=regsubsets(as.matrix(Boston[,-14]),Boston[,14],nvmax=13)
sumleaps=summary(RSSleaps,matrix=T)
msize=apply(sumleaps$which,1,sum)
n=dim(Boston)[1]
p=dim(Boston)[2]
# calcualte the Mallow's Cp
Cp=sumleaps$rss/(summary(lmfit)$sigma^2) + 2*msize - n
# Compare the results with the return by regsubsets function
cbind(msize, Cp, sumleaps$cp)
##
      msize
                    Ср
## 1
          2 166.568016 166.568016
## 2
          3 110.304269 110.304269
          4 90.220657 90.220657
## 3
## 4
          5 61.263595
                       61.263595
## 5
          6 45.547551 45.547551
## 6
          7 27.940433 27.940433
## 7
          8 21.385925 21.385925
## 8
         9 14.968721 14.968721
## 9
         10
            9.737754
                        9.737754
## 10
         11 10.346297 10.346297
## 11
         12 10.584968 10.584968
## 12
         13 12.060939 12.060939
## 13
         14 14.000000 14.000000
  • Now we calcualt AIC and BIC based on their definitions.
# calcualt AIC and BIC based on their definitions
AIC = n*log(sumleaps$rss/n) + 2*msize
as.matrix(AIC)
##
           [,1]
## 1 -1479.977
## 2 -1524.114
## 3
     -1540.763
## 4 -1566.119
## 5 -1580.467
## 6 -1597.197
```

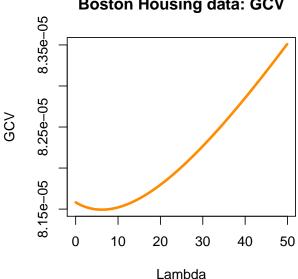
```
-1603.591
## 8
     -1609.989
## 9 -1615.317
## 10 -1614.739
## 11 -1614.545
## 12 -1613.084
## 13 -1611.146
BIC = n*log(sumleaps$rss/n) + msize*log(n)
# compare it with the return by regsubsets function
cbind(BIC, sumleaps$bic+n*log(sum((Boston[,14] - mean(Boston[,14]))^2/n)))
##
            BIC
## 1
     -1471.524 -1471.524
      -1511.434 -1511.434
## 3 -1523.857 -1523.857
## 4 -1544.986 -1544.986
## 5 -1555.107 -1555.107
## 6 -1567.611 -1567.611
## 7 -1569.779 -1569.779
## 8 -1571.950 -1571.950
## 9 -1573.051 -1573.051
## 10 -1568.247 -1568.247
## 11 -1563.827 -1563.827
## 12 -1558.139 -1558.139
## 13 -1551.975 -1551.975
  • We select the best models using step based on AIC and BIC respectively. To use Mallow's Cp returned
     by regsubsets function, we select the model that has C_p closest to p, namely the 9th model from the
     first subproblem.
# backward with AIC
stepaic<-step(lmfit, direction="backward", trace=0)</pre>
paste(variable.names(stepaic),collapse = ' + ')
## [1] "(Intercept) + chas + nox + rm + dis + rad + tax + ptratio + black + lstat"
# backward with BIC
stepbic<-step(lmfit, direction="backward", k=log(n), trace=0)</pre>
paste(variable.names(stepbic),collapse = ' + ')
## [1] "(Intercept) + chas + nox + rm + dis + rad + tax + ptratio + black + lstat"
# paste(colnames(sumleaps$which)[sumleaps$which[which.min(sumleaps$bic),]],collapse=' + ')
# best result based on Mallow's Cp
paste(colnames(sumleaps$which)[sumleaps$which[9,]],collapse=' + ')
## [1] "(Intercept) + chas + nox + rm + dis + rad + tax + ptratio + black + lstat"
In this example, all the three selection criteria give the same conclusion (0.5 points off for missing conclusion.).
     Question 3 (ridge regression)
[2 points] Now use lm.ridge function to fit the same dataset. We choose \lambda in a sequence between 0 and
100 and use geneneralized cross-validiation (GCV) to choose the best value (1 point). We plot shrinking
coefficients and the GCV values as below.
library(MASS)
# ridge regression
```

```
ridge.fit = lm.ridge(medv~., Boston, lambda=seq(0,100,by=0.1))
# plot shrinking coefficients and the GCV values
par(mfrow=c(1,2))
matplot(coef(ridge.fit)[, -1], type = "l", xlab = "Lambda", ylab = "Coefficients")
text(rep(50, p-1), coef(ridge.fit)[1,-1], colnames(Boston)[1:p-1])
title("Boston Housing data: Ridge Coefficients")
plot(ridge.fit$lambda[1:500], ridge.fit$GCV[1:500], type = "1", col = "darkorange",
     ylab = "GCV", xlab = "Lambda", lwd = 3)
title("Boston Housing data: GCV")
```

## **Boston Housing data: Ridge Coefficient**

## **Boston Housing data: GCV**





Now we use GCV to choose the best  $\lambda$  and report the penalized regression coefficients (1 point).

```
# use GCV to select the best lambda
ridge.fit$lambda[which.min(ridge.fit$GCV)]
```

```
## [1] 6.2
```

```
# ridge regression coefficients
round(coef(ridge.fit)[which.min(ridge.fit$GCV), ], 4)
```

## crim zn indus chas nox rm age dis rad tax ptratio black lstat ## 4.0428 -0.0134 0.0009 -0.0147 0.1118 -0.2594 0.4538 0.0053 -0.1719 0.0610 -0.1902 -0.0401 0.0454 -0.2544