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On the Magnitude of Laterality Effects and Sex Differences in Functional Lateralities

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In the last 20 years, the hypothesis that men and women differ in functional lateralities has been used to account for sex-related differences in verbal and spatial skills. However, this hypothesis has not been clearly supported, with some reviewers confirming it (McGlone, 1980 for example), and others rejecting it (Fairweather, 1982 for example). The purpose of the present study was to provide a definite test of this hypothesis and to estimate the magnitude of overall laterality effects by means of a meta-analytic procedure. A total of 396 significance levels from a variety of studies on functional asymmetries utilising auditory, visual, or tactile presentation of verbal or nonverbal stimuli were sampled. Results showed that laterality effects tend to be large and significant but that they are heterogeneous in the visual modality. Homogeneity was generally achieved by a partition of the studies in terms of the specific task used. The results also showed sex differences to be significant in two modalities (visual and auditory). The data indicated the presence of sex differences in favour of men in functional asymmetries. However, it appears that the findings are not resistant to the file drawer problem. The results are discussed with regard to their implications for explanations of individual differences in cognitive abilities. The relation between functional lateralities and anatomical asymmetries is also discussed.

INTRODUCTION

The belief that sex differences in specific intellectual abilities exist was quite common even before they were empirically documented (Ellis, 1928). A large number of studies have also been conducted to test this belief and most of these studies were summarised by Garai and Scheinfeld (1968). Since then, other reviews and meta-analyses have demonstrated that men and women differ on specific cognitive abilities (Harris, 1978; Linn & Petersen, 1985; Maccoby &

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Jacklin, 1974; Voyer, Voyer, & Bryden, 1995). Prominent in these reviews is the finding that sex differences are found in verbal and spatial abilities. Specifically, women generally perform better than men in verbal tasks, while the opposite pattern is found in spatial tasks. It is now commonly accepted that sex differences exist in verbal and spatial skills (but see Caplan, MacPherson, & Tobin, 1985, Feingold, 1988, and Hyde & Linn, 1988). Based on this conclusion, a number of hypotheses have been advanced to account for sex differences in these areas (see McGee, 1979, for a review). Several of these hypotheses have been tested but few have been conclusively rejected. Thus, a number of theories still remain.

One theory makes use of the fact that the areas of intellectual abilities where sex differences are found seem to parallel what is conventionally recognised as the basic hemispheric specialisation of the brain. A number of reviewers and theorists agree that the left hemisphere of the brain appears to be specialised for verbal functions, whereas the right hemisphere seems to be dominant for nonverbal, visuo-spatial functions (Bryden, 1982; Geschwind & Galaburda, 1987; Iaccino, 1993). Based on the parallel between the abilities where sex differences are found and the respective specialisation of the cerebral hemispheres, one obvious explanation of sex differences in verbal and spatial skills is that they may be due to sex differences in cerebral organisation. However, the exact nature of the differences in cerebral organisation is not clear.

One attempt to clarify this issue was made by Levy (1969) who proposed that people with bilateral representation of language functions should perform relatively poorly on spatial tasks. Following this idea, Levy (1971) suggested that men are better than women on spatial tasks because they are more lateralised for spatial functions. According to Levy (1971), verbal functions are bilaterally represented in women, and they overlap with spatial functions. This presumably would explain why women perform better than men on verbal tasks, while men are superior to women on spatial tasks. The overlap of verbal functions with spatial functions impairs women's spatial abilities and increases their verbal abilities, while greater asymmetry of spatial functions in men leaves their spatial abilities intact, as compared to those of women. Levy (1971) explained the negative effect of overlapping of verbal functions on spatial functions as resulting from interference. That is, bilateral language competence produces a specific interference with spatial functions by forcing the two hemispheres to process information using a verbal approach. This interpretation implies that asymmetry of function allows an optimal performance of the right (spatial) hemisphere, as in this case, verbal functions do not interfere with spatial functions. This hypothesis predicts that men should show hemispheric advantages on specific tasks, whereas women should not.

Other attempts have been made to clarify the nature of sex differences in cerebral organisation. For instance, Buffery and Gray (1972) proposed a hypothesis that is virtually the opposite to that of Levy (1971). They claimed

that women are more lateralised than men. In men, spatial functions are hypothesised to be bilaterally represented. This produces superior spatial abilities in men. These authors argued that bilateral representation is an advantage in one case (spatial functions), and asymmetry an advantage in another case (verbal functions). According to Buffery and Gray (1972, p.144), verbal skills require quick associations and serial ordering, which presumably demands fast and intricate neural mechanisms. "Such mechanisms could benefit from being subserved by specific structures with a clearly lateralized and localized cerebral representation". On the other hand, the three-dimensional nature of spatial tasks benefits from a bilateral cerebral representation. In this case, the authors assume a positive effect of bilateral representation of spatial functions. This is in obvious contradiction with Levy's (1971) model, as it predicts that women should show laterality effects on specific tasks, whereas men should not. It is worth noting that a more recent model of the development of cerebral lateralisation proposed by Geschwind and Galaburda (1987) also implies that men should be less lateralised than women in a variety of tasks.

One way to test these hypotheses is to look at anatomical sex differences in the brain. The rationale guiding this type of research is that if there are sex differences in functional asymmetries, they should have anatomical substrates. A number of anatomical studies have revealed the presence of sex differences in the shape and size of specific structures in the corpus callosum (Clarke, Kraftsik, Van der Loos, & Innocenti, 1989; de Lacoste-Utamsing & Holloway, 1982; Reinartz, Coffman, Smoker, & Godersky, 1988). In general, results show that women tend to have a larger corpus callosum than men. De Lacoste-Utamsing & Holloway (1982) claimed that their findings were compatible with the hypothesis that the female brain is less well lateralised than the male brain, a conclusion that can be applied to all the studies reporting a larger corpus callosum in women than in men. This means that it is plausible to assume that women engage in more interhemispheric communication than men do and, for that reason, they require a larger corpus callosum (Iaccino, 1993). Increased interhemispheric communication in women thus suggests that they are more likely than men to show bilateral representation of verbal and nonverbal functions, and this fits very well with Levy's hypothesis.

A number of reviews have examined sex differences in laterality. Bryden (1979), Fairweather (1982), Harris (1978), and McGlone (1980) are the most commonly cited ones. In three of these reviews (Bryden, 1979; Harris, 1978; McGlone, 1980) support was found for Levy's hypothesis. However, Fairweather (1982) reported that sex differences in laterality do not exist.

This inconsistency in the conclusions reached by different reviewers concerning the existence of sex differences in laterality is puzzling. A basic difference between Fairweather's review and the other three is that Fairweather only reviewed studies using a divided visual field paradigm. This might have produced a biased sample of studies and resulted in a false conclusion. However,

the presence of a systematic bias is unlikely. A more likely possibility relates to the fact that these authors used a qualitative approach in reviewing the studies. Even though qualitative reviews allow a good understanding of conceptual and methodological issues which may account for differences in results among studies, they do not permit the quantification of the magnitude of sex differences in laterality. In examining the question of whether or not sex differences in laterality are significant, qualitative reviews only allow a count of the studies showing a significant effect, as compared to studies showing a non-significant effect. In the long run, this paradigm produces arbitrary decisions as to whether studies are more likely to produce one outcome rather than another.

Meta-analysis has the advantage of allowing precise quantification of the magnitude of the effect in a series of studies (Rosenthal, 1991). Hence, it was used in the present study. However, it is important to keep in mind that a meta-analysis is only as good as the studies entering it, as it may be contaminated by poorly designed studies. To minimise the influence of this problem, an attempt was made in the present analysis to include only studies published in refereed journals. This approach was based on the assumption that methodologically flawed studies are less likely to appear in such publications. Studies published in non-refereed journals were included only when they were cited in the reviews previously mentioned (i.e. Bryden, 1979; Fairweather, 1982; Harris, 1978; McGlone, 1980).

Even when it is conducted carefully, a meta-analysis is not necessarily the perfect way to summarise literature, because it has its own flaws (see Sohn, 1995 and Eysenck, 1995, for examples). Nevertheless, meta-analysis provides a clear summary of the effects of interest and it also provides a definite test of significance of these effects. Therefore, based on the claim that its advantages outweigh its disadvantages (see Lipsey & Wilson, 1995), meta-analysis was used to provide a definite test of the hypothesis that there are sex differences in functional lateralities. Because the data were readily available along with the data pertaining to sex differences, the magnitude of the main effect of laterality was also examined.

A careful analysis of sex differences in laterality has recently been conducted by Hiscock et al. (1994). However, it concerned only auditory effects and it did not use a conventional meta-analytic approach. Specifically, these authors used a systematic counting procedure which allowed them to tabulate the frequency of significant sex differences in auditory laterality in studies drawn from a selective sample of neuropsychology journals. This procedure had both the advantages and disadvantages of a literature review, and it did not allow the authors to quantify the magnitude of sex differences in laterality. Nevertheless, it supported the notion that there are weak sex differences in laterality, indicating greater auditory lateralisation in men. Thus, based on the fact that laterality effects are consistently reported and that a number of reviews suggest the existence of sex differences, the hypotheses of

the present analysis are that (a) overall laterality effects will be significant, and (b) sex differences in laterality will be significant.

META-ANALYTIC PROCEDURE

Selection Criteria for Inclusion in the Meta-analysis

The present meta-analysis includes only published studies presenting results obtained with dichotic listening, divided visual field, or dichhaptic (tactile) stimulation. Results obtained with samples of brain-damaged subjects were not included in the present analysis. The present study is thus concerned only with functional lateralities observed in brain-intact individuals. Another requirement was that the sex of subjects had to be considered as an independent variable in the data analysis of an individual study to warrant its inclusion in the sample of the present analysis. This is an important distinction, because in extracting the relevant information from the studies included in Fairweather's (1982) literature review, it was found that three studies were reported as showing no sex differences (Butler & Miller, 1979; Day, 1977; Martin, 1978), when the sex of subjects effect was in fact not tested in their analysis. Relevant published studies included in the reviews conducted by Bryden (1979), Fairweather (1982), Harris (1978), and McGlone (1980) were selected for inclusion in the present analysis. As previously mentioned, only studies published in refereed journals were included in the analysis. The only studies published in non-refereed journals which were included in the analysis were the ones cited in the Bryden (1979), Fairweather (1982), Harris (1978), and McGlone (1980) reviews. Furthermore, refereed journals likely to publish laterality studies were reviewed. Thus, all the issues of *Brain and Cognition*, *Brain and Language*, *Canadian Journal of Psychology*, *Cortex*, *Journal of Clinical and Experimental Neuropsychology*, *Journal of Experimental Psychology: Human Perception and Performance*, *Neuropsychologia*, and *Perception and Psychophysics* were systematically searched to extract laterality studies using sex of subject as a variable in their design. CD-ROM Psych-Lit searches were also conducted, and the reference list of recent papers was searched for relevant studies. Thus, the present meta-analysis provides a nearly exhaustive review of the published literature on sex-related differences in functional laterality in brain-intact subjects up to 1993. In most of those studies, the primary purpose was not to test the existence of sex differences in laterality. Sex of subjects was examined only because it was easily observable.

The studies selection procedure resulted in the sampling of 396 significance levels drawn from 266 studies. A list of the significance levels entered in the present analysis is available from the author on request. The studies included in the meta-analysis are listed in the Appendix.

The "file drawer problem" (see Rosenthal, 1979)—the possibility that published studies are a biased sample of the studies that are actually carried out,

as it is presumed that only experiments with significant results are published—is one problem associated with the inclusion of only published studies in a meta-analysis. The exclusive use of published studies is likely to result in an overestimation of the effects under study. The number of studies averaging null results that would be necessary to offset the significance of the findings at the 0.05 level (fail-safe number, see Rosenthal, 1991) is usually computed to cope with this problem. In the present study, the fail-safe number will be calculated to assess the resistance of the meta-analytic results to the file drawer hypothesis. The fail-safe number will be compared with the criterion of $5 \times$ the number of the sampled studies + 10 ($5k + 10$), suggested by Rosenthal (1980, 1991), to evaluate its significance and allow rejection of the hypothesis that significant results are due to the exclusive sampling of published studies.

Analysis Procedure

The present analysis followed the procedure presented by Rosenthal (1991). This author presented meta-analytic techniques designed for the combination and comparison of significance levels. Comparison of the significance levels allows for the conclusion that the results in a specific sample of studies are homogeneous and were drawn from the same population. In other words, homogeneity of the significance levels indicates that the studies included in a specific meta-analysis can be considered as replications of each other, and that a pooled estimate of the results provides a valid summary of the findings from the sample of studies. However, when heterogeneity is detected, it is likely that the pooled estimate is not representative of the state of affairs in a sample. When this is the case, further partitioning of the studies is required.

Following Rosenthal's (1991) recommendations, significance testing was conducted by obtaining the standard normal deviate (Z) associated with each significance level. Z s in disagreement with the bulk of the results were given a negative sign. Following this procedure, a positive Z value represented a left hemisphere advantage for verbal tasks and a right hemisphere advantage for nonverbal tasks, while a negative Z value indicated a right hemisphere advantage for verbal tasks and a left hemisphere advantage for nonverbal tasks in the analysis of the main effect of laterality. In the analysis of sex differences in laterality, a positive Z value represented a greater lateralisation in men than in women, while a negative Z represented a greater lateralisation in women. Studies that reported non-significant effects and provided no statistics concerning these effects were assigned a Z value of zero, as recommended by Rosenthal (1991). There are several reasons why this approach was used. One such reason is that the direction of the effect remains undefined under these conditions. The logic underlying this approach makes the assumption that the negative and positive outcomes are equally represented in cases where relevant statistics are not provided. Therefore, assigning a value of zero to the

corresponding Z s reflects the average of all these outcomes. This approach is also the most conservative procedure and it is least likely to lead to erroneous conclusions. Finally, the mere exclusion of these cases would contribute to the file drawer problem.

The test of significance of the combined Z values was computed using a modified Stouffer's method (Rosenthal, 1980, 1991), in which each significance level was weighted by its error degrees of freedom ($N-2$). This approach clearly gives more weight to the results obtained in studies with a large sample of participants than to the results of small sample studies (Rosenthal, 1991).

Cohen's d (Cohen, 1977) was used as an estimate of effect size. This index represents the standardised difference between the mean of the groups under study. It was chosen because it allows a direct interpretation of the results in standard deviation units. Z values were transformed into d s by using the formula provided by Rosenthal (1991).

The present analysis followed an hierarchical approach in evaluating the magnitude of the main effect of laterality and sex differences in functional asymmetries. In both cases, an analysis was first conducted separately for each modality (auditory, visual, tactile) to compare and combine the significance levels in these broad groupings. Following this procedure, the significance levels were partitioned into type of task (verbal, nonverbal) when heterogeneity was detected. Tasks were defined as verbal or nonverbal by the present author, on the basis of the type of cognitive processing involved. When homogeneity was still not achieved with those clusters, a partition based on specific task was performed. The classification of specific tasks was based on the description provided in the individual studies. Tasks with similar descriptions were classified under the same label by the present author.

The analysis of the main effect of laterality was performed first, followed by the analysis of sex differences in laterality. Because homogeneity of significance levels tends to be difficult to achieve in a meta-analysis (Linn & Petersen, 1985), a criterion similar to that used by Voyer, Voyer, and Bryden (1995) was used to classify significance levels as close to homogeneity. This means that when the obtained homogeneity statistic (chi-square) was significant at the 0.05 level, but not at the 0.005 level, significance levels were considered to be close to homogeneous. This range is arbitrary, but sets clear limits to what can be considered as satisfactorily homogeneous.

META-ANALYTIC RESULTS

Laterality Effects

The meta-analytic results for the laterality effects are presented in Table 1. In the visual modality, laterality effects were highly significant. The fail-safe analysis indicated that 65,447 studies with non-significant results would be needed to offset the significance of the results at the 0.05 level. However the significance

TABLE 1
Meta-analytic Results for Laterality Effects as a Function of the Testing Modality and the Nature of the Task

Partitioning Factor	N	Weighted Estimator of Effect Size (<i>d</i>)	Test of Significance (<i>Z</i>) ^c	Fail-Safe	Homogeneity Statistic (χ^2)
Visual Modality	235	0.531	21.84	65,447	486.71*
Verbal Tasks	123	0.674	19.94	25,170	241.66*
Nonverbal Tasks	112	0.388	10.94	9340	216.38*
Auditory Modality	120	0.499	17.72	27,102	143.73 ^a
Verbal Tasks	94	0.519	17.04	19,298	79.58 ^a
Nonverbal Tasks	26	0.391	4.96	621	50.21*
Tactile Modality	41	0.353	5.20	650	81.46*
Verbal Tasks	17	0.257	2.65	18	21.24 ^a
Nonverbal Tasks	24	0.390	4.59	389	37.32 ^b

**P* < 0.05
^a Homogeneity achieved
^b Close to homogeneity
^c All *P* < 0.01

levels were not homogeneous. This means that the studies using the visual modality included in the present analysis are not all drawn from the same population.

In the auditory modality, laterality effects were also highly significant. In this case, the fail-safe analysis indicated that 27,012 studies with non-significant results would be needed to offset the significance of the results at the 0.05 level. In this modality, the significance levels were found to be homogeneous. This means that the studies included in the auditory modality are basically replications of each other, in spite of the variety of tasks used in the studies included in the sample. It is thus legitimate to state that laterality effects are clearly significant and homogeneous in the auditory modality, regardless of the nature of the task or the specific task used.

Finally, in the tactile modality, the analysis revealed that laterality effects were significant. The fail-safe value indicated that 650 studies with non-significant results would be needed to offset the significance of the results at the 0.05 levels. However, as was the case in the visual modality, the significance levels were found to be heterogeneous. This means that the studies included in the tactile modality sample are not all drawn from the same population and require further partitioning.

In an attempt to achieve homogeneity, significance levels were partitioned based on the nature of the task material (verbal, nonverbal). Results of this

analysis are also presented in Table 1. Note that the significance levels were partitioned as a function of the nature of the task in the auditory modality as well, even though the overall analysis revealed homogeneous effects in this modality. This partition was performed to allow comparison with the results obtained in the other two modalities. When the results obtained in the visual and tactile clusters are examined, Table 1 shows that homogeneity of the significance levels was obtained in the tactile modality after partitioning by nature of the task. It thus appears that the tactile modality produces significant and homogeneous laterality effects when verbal and nonverbal tasks are considered separately. In the visual modality, however, a partition by nature of the task did not allow for homogeneity to be achieved. This means that even though verbal and nonverbal tasks used in the visual modality show significant laterality effects, these effects are not all drawn from the same population. For this reason, they required further partitioning.

Significance levels in the visual modality were partitioned by specific task to assess the effect of this division on homogeneity. Results of this partition are shown in Table 2. This analysis showed that most tasks demonstrated significant laterality effects. In fact, all the verbal tasks showed significant effects (at least when a one-tailed test of significance was used), while only dot localisation and mental rotation did not show significant effects in nonverbal tasks. This partition achieved homogeneous significance levels in most clusters. Only lexical decision studies and study-specific tasks (classified as others) showed heterogeneous effects in the verbal tasks, while studies using faces did not achieve homogeneity in the nonverbal tasks. Further partitioning in terms of dependent variable (accuracy, reaction time, etc.), scoring procedure, handedness, age, or selectivity of the participants (see Hedges & Olkin, 1985) did not achieve homogeneity in these clusters. Nevertheless, it is legitimate to conclude that laterality effects are generally significant and homogeneous when the studies are partitioned in terms of the specific task used in the visual modality.

Inter-group Differences. The analyses reported so far only concerned the magnitude and significance of laterality effects. They did not address the question of whether or not differences in the magnitude of laterality effects exist between modalities and types of task. One might suggest that a factorial analysis of variance (ANOVA) would provide an answer to this question. However, Hedges and Becker (1986) argued that ANOVA is not appropriate for effect sizes, and they suggested an analogue to analysis of variance for effect sizes. Unfortunately, in its present form, this approach does not allow a test of the interaction between variables. For this reason, the weighted regression procedure proposed by Hedges and Becker (1986) was used. The unbiased estimate of effect size (d) was the dependent variable, and testing modality (visual, auditory, tactile) and type of task (verbal, nonverbal) were the independent variables. Because testing modality and type of task are categorical

TABLE 2

Meta-analytic Results for Laterality Effects as a Function of Specific Task in the Visual Modality

Partitioning Factor	N	Weighted Estimator of Effect Size (<i>d</i>)	Test of Significance (<i>Z</i>)	Fail-Safe	Homogeneity Statistic (χ^2)
<i>Verbal Tasks</i>					
Lexical Decision	65	0.579	12.75*	5106	120.99*
Letters	19	0.653	6.45*	547	31.47 ^b
Naming	25	1.040	13.77*	2310	28.05 ^a
Category Judgment	2	0.285	1.73 ^c	0	2.09 ^a
Digits	6	0.629	4.82*	23	11.30 ^b
Picture-Word Matching	4	0.500	2.84*	17	5.65 ^a
Others	2	0.672	1.90 ^c	4	4.50*
<i>Nonverbal Tasks</i>					
Dot Detection	7	0.261	1.98*	9	9.11 ^a
Dot Localisation	6	<0.001	<0.001	0	<0.001 ^a
Dot Enumeration	8	0.649	2.05*	82	9.66 ^a
Shapes	6	0.362	2.66*	21	13.50 ^b
Faces	56	0.486	9.67*	3293	106.09*
Mental Rotation	7	0.211	1.33	0	16.23 ^b
Objects	7	0.295	2.19*	25	17.79 ^b
Line Orientation	8	0.339	3.33*	23	11.65 ^a
Others	7	0.529	2.74*	24	9.22 ^a

* $P < 0.05$

^a Homogeneity achieved

^b Close to homogeneity

^c $P < 0.05$, one-tailed

variables, they were transformed by the use of effect coding (Pedhazur, 1982), to make them adequate for regression analysis. The interaction vectors were computed by cross-multiplying the vectors representing the task and modality variables. The regression weights were the same values that were used to calculate the weighted estimates of effect sizes presented in Table 1. These weights are the inverse of the variance associated with each effect size ($1/v_i$). A computational formula for these weights can be found in Hedges and Becker (1986). The regression procedure involved the forced entry of the main effect of modality (represented by two vectors), followed by the main effect of type of task. The interaction vectors were entered last.

According to Hedges and Becker (1986), the conventional test of significance reported for regression coefficients (t-test) is incorrect. For this reason, the calculation of a Z score, representing a corrected test of significance (Hedges & Becker, 1986) was required in the present analysis. This approach revealed a significant effect of modality, $Z = 3.01$, $P > 0.01$, and a significant effect of type

of task, $Z=6.43$, $P<0.01$. These main effects were qualified by an interaction between modality and type, $Z=3.76$, $P<0.01$. Simple main effect analyses indicated that the effect of task was significant both in the visual and auditory modality ($Z=7.04$, $P<0.01$, and $Z=2.14$, $P<0.05$, respectively). In both cases this finding reflected the fact that effect sizes were significantly larger in verbal tasks than in nonverbal tasks (see Table 1). In the tactile modality, the effect of type of task was not significant, $Z=-1.14$, ns.

Sex Differences in Laterality¹

The same approach used to analyse laterality effects was applied to the analysis of sex differences in laterality. Results of this analysis are presented in Table 3. In the visual modality, this analysis revealed a significant mean weighted d of

TABLE 3
Meta-analytic Results for Sex Differences in Functional Asymmetries as a Function of Testing Modality and Nature of the Task

Partitioning Factor	N	Weighted Estimator of Effect Size (d)	Test of Significance (Z)	Fail-Safe	Homogeneity Statistic (χ^2)
Visual Modality	235	0.067	2.49*	1155	269.27 ^a
Verbal Tasks	123	0.058	1.47	0	136.56 ^a
Nonverbal Tasks	112	0.076	2.06*	302	132.41 ^a
Auditory Modality	120	0.063	2.73*	207	98.42 ^a
Verbal Tasks	94	0.062	2.26*	119	86.34 ^a
Nonverbal Tasks	26	0.070	1.78 ^c	0	12.06 ^a
Tactile Modality	41	0.148	1.80 ^c	136	81.46*
Verbal Tasks	17	0.129	0.89	0	40.50*
Nonverbal Tasks	24	0.155	1.59	0	40.09 ^b

* $P<0.05$

^a Homogeneity achieved

^b Close to homogeneity

^c $P<0.05$, one-tailed

¹ Preliminary results obtained in the present analysis were presented in Voyer (1994), *Brain and Cognition*, 26, pp. 211–216. The meta-analytic procedure used in this preliminary report was based on an approach to cumulation of effect sizes that resulted in a conservative estimation of significance tests. Because of this, the initial procedure revealed non-significant sex differences in laterality. Problems were noted with this approach as it did not provide an accurate test of significance. Specifically, the likelihood of a Type II error increased with the size of the sample of studies (I.C. McManus, personal communication, November 18, 1994). The approach used in the present analysis corrected this problem and provided an adequate test of significance.

0.067, demonstrating that, based on these figures, men tend to be more lateralised than women in this modality. The fail-safe analysis showed that 1155 studies with non-significant results would be needed to offset the significance of the present findings at the 0.05 level. Furthermore, the significance levels were found to be homogeneous. This means that the pooled estimate of effect size is representative of the state of affairs in this sample. Sex differences are thus clearly significant in this modality, regardless of the task used.

In the auditory modality, a significant mean weighted d of 0.063 was found, indicating the presence of sex differences in laterality in this modality as well. The fail-safe analysis revealed that 207 studies with non-significant results would be required to offset the significance of the results at the 0.05 level. Once again, the significance levels were found to be homogeneous. Sex differences in laterality are thus significant and homogeneous in this cluster.

Finally, in the tactile modality, the mean weighted d of 0.148 only achieved significance when based on a one-tailed test. The fail-safe analysis showed that 136 studies with non-significant results would be needed to offset the significance of the effect at the 0.05 level. The significance levels were heterogeneous in this cluster and required further partitioning.

Significance levels in the tactile modality were partitioned based on the nature of the task in an attempt to achieve homogeneity. Results of this partition are also presented in Table 3. Note that, as was done in the analysis of laterality effects, the significance levels were partitioned as a function of the nature of the task in the visual and auditory modalities as well, even though the significance levels were homogeneous in these clusters. This allowed comparison with the results obtained in the tactile modality. Examination of the tactile cluster in Table 3 shows that sex differences were not significant in verbal and nonverbal tasks, while they remained heterogeneous in the verbal cluster. Further partitioning by specific task was thus warranted in the tactile modality. The results of this partitioning are presented in Table 4. It appears that homogeneity of significance levels was achieved (or nearly achieved) in the tactile modality when partitioning was based on the specific task used. In most clusters, sex differences were non-significant in spite of the relatively large d s reported in Table 4 but they were homogeneous. Line orientation was the only partition that demonstrated nearly significant sex differences ($P=0.08$, two-tailed). However, with only three significance levels sampled in this cluster, the fail-safe value was negative and indicated essentially non-significant results. Finally, the digits grouping was the only one that showed heterogeneous effects. The small sample size in this cluster made further partitioning meaningless.

Inter-group Differences. As was the case with the analysis of the main effect of laterality, the analyses reported so far did not examine differences in the magnitude of laterality effects between modalities and types of task. The weighted regression procedure proposed by Hedges and Becker (1986) was used

TABLE 4

Meta-analytic Results for Sex Differences in Functional Asymmetries as a Function of Specific Task in the Tactile Modality

<i>Partitioning Factor</i>	<i>N</i>	<i>Weighted Estimator of Effect Size (d)</i>	<i>Test of Significance (Z)</i>	<i>Homogeneity Statistic (χ^2)</i>
<i>Verbal Tasks</i>				
Letters	11	0.212	1.14	20.36 ^b
Digits	6	-0.039	-0.20	18.12*
<i>Nonverbal Tasks</i>				
Shapes	19	0.129	1.30	35.37 ^b
Line Orientation	3	0.403	1.75 ^c	2.56 ^a
Bisection	2	0.416	1.39	1.92 ^a

* $P < 0.05$

^a Homogeneity achieved

^b Close to homogeneity

^c $P < 0.05$, one-tailed

once again to investigate this aspect. As before, the unbiased estimate of effect size (d) was the dependent variable, and testing modality (visual, auditory, tactile) and type of task (verbal, nonverbal) were the independent variables. Testing modality and type of task were once again coded by the use of effect coding (Pedhazur, 1982), to make them appropriate for regression analysis, and the interaction vectors were computed by cross-multiplying the vectors representing the task and modality variables. The regression weights were the same values that were used in the regression analysis reported in relation to the main effect of laterality.

In the analysis of sex differences in laterality, this approach revealed no significant effects and no interaction (all P s > 0.10). This means that the magnitude of sex differences did not vary in a significant way between the different modalities and types of tasks.

DISCUSSION

The results of the present analysis suggest conclusions relevant to laterality effects in general, and sex differences in laterality in particular. Even though the findings concerning these two types of effects are related, they will be discussed separately.

Laterality Effects

Concerning laterality effects, it appears that they tend to be highly significant, particularly when verbal tasks are used in the visual and auditory modalities.

Furthermore, they are homogeneous when the auditory modality is used. In the tactile modality, a partition in terms of the nature of the task (verbal, nonverbal) produced homogeneous effects, whereas a partition by specific task used produced a satisfactory level of homogeneity in the visual modality. Only two clusters (dot localisation and mental rotation) showed non-significant laterality effects. It is thus possible to conclude that laterality effects are clearly significant and homogeneous in most contexts in which they are studied. This conclusion supports the first hypothesis of the present study stating that laterality effects are significant. However, it is worth noting that in several clusters—i.e. tactile verbal in Table 1; digits, picture–word matching, dot detection, shapes, objects, line orientation, and the two study-specific tasks (others) categories in Table 2—the fail-safe value is much smaller than the criterion of $5k + 10$ suggested by Rosenthal (1991). This means that the results obtained in these clusters are not resistant to the file drawer problem.

The finding that verbal tasks produce larger laterality effects than nonverbal tasks in the visual and auditory modalities suggests that the type of tasks is relevant if one wants to maximise the likelihood of detecting laterality effects. Specifically, the present results indicate that the use of a verbal task would increase this likelihood in the visual and auditory modalities. This finding is compatible with the notion that nonverbal tasks produce effects that are not as consistent as those obtained with verbal tasks (Bryden, 1982).

Sex Differences in Laterality

The results of the present analysis suggest that sex differences in laterality are significant in the visual and auditory modalities, and in the tactile modality when a one-tailed test of significance is used. In fact, the prevalence of positive effect sizes indicates that men generally obtain larger laterality effects than women. The second hypothesis of the present study, which predicted that sex differences in laterality would be significant, is therefore confirmed. The present results thus extend the conclusion reached by Hiscock et al. (1994) to the visual and tactile modalities. Furthermore, the results of the present analysis argue for the plausibility of an explanation of sex differences in cognitive abilities in terms of sex differences in cerebral organisation. Specifically, the present findings provide evidence in agreement with Levy's (1971) hypothesis, but they contradict predictions made on the basis of the models proposed by Buffery and Gray (1972), and Geschwind and Galaburda (1987). It thus appears that sex differences in verbal and spatial abilities can be explained, at least in part, by the fact that men tend to be more lateralised than women.

The results of the present study also suggest that reported anatomical sex differences in the brain seem to relate to functional brain asymmetries. However, even though a number of studies have found sex differences in the size and shape of specific structures in the corpus callosum (Clarke et al., 1989; de

Lacoste-Utamsing & Holloway, 1982; Reinarz et al., 1988), many authors reported a failure to replicate this finding (Demeter, Ringo, & Doty, 1985; Kertesz, Polk, Howell, & Black, 1987; Oppenheim, Lee, Nass, & Gazzaniga, 1987; Witelson, 1985). It is worth noting that the study conducted by de Lacoste-Utamsing and Holloway (1982), often cited as demonstrating sex differences in the size of the splenium, only found a marginal difference between men and women ($P=0.08$). An alternative hypothesis stating that within-sex individual differences may be more striking than between-sex differences (Bleier, Houston, & Byne, 1986; Byne, Bleier, & Houston, 1988) has been proposed to account for inconsistent findings. Peters (1988) also suggested that brain size, not sex of subjects, might be a critical variable, as sex differences in brain size are consistently reported. It thus appears that, even though the question of whether or not sex differences exist in the anatomy of the corpus callosum has not yet been answered, sex differences in functional asymmetries are small but significant. The small effect sizes observed for sex differences in functional lateralities can be conceived of as reflecting the inconsistencies in the findings of anatomical studies.

Inconsistencies in the findings concerning sex differences in brain anatomy suggest that asymmetries in brain anatomy and functional asymmetries are not necessarily related. Specifically, studies investigating the relation between the morphometry of the corpus callosum and laterality effects have obtained mixed results. Hines et al. (1992) reported a significant correlation between posterior callosal size and language lateralisation. However, these authors admitted that their finding did not provide strong evidence because the correlation between callosal size and laterality was significant only when based on a one-tailed test, and it was not consistent across the different analyses they conducted. In a similar study, Clarke, Lufkin, and Zaidel (1993) failed to obtain a significant correlation between the size of the corpus callosum and a measure of language lateralisation. The lack of strong evidence argues for the notion that there is not necessarily a link between anatomical measures of laterality and measures of functional asymmetries. This suggests the possibility that functional asymmetries and anatomical asymmetries are two unrelated phenomena. From this perspective, it is plausible to believe that functional asymmetries reflect physiological rather than anatomical asymmetries. It might prove fruitful to investigate the relation between functional lateralities and physiological asymmetries, as revealed in studies examining evoked response potential, cerebral blood flow, positron emission tomography, etc. Another possibility is that functional asymmetries may reflect processing or strategy differences between the hemispheres (Bryden, 1980). In this case, the absence of a relation between anatomical and functional asymmetries would be expected.

Unfortunately, in the context of a meta-analysis it is not possible to determine which of these speculations is the correct explanation for the lack of a relation between functional and anatomical asymmetries. This underlines the fact that

meta-analysis does not allow causal statements, and emphasises the claim that a meta-analysis cannot replace a well designed empirical study in providing answers to theoretically guided questions (Sohn, 1995). This suggests that much empirical research remains to be done before a clear understanding of the relation between anatomical asymmetries and functional lateralities can be achieved.

The present analysis indicates that sex differences in laterality are significant. However, it is important to be cautious in interpreting tests of significance drawn from a sample of exclusively published studies. In fact, when the fail-safe values associated with the significant effects are examined, they fail to reach the criterion proposed by Rosenthal (1991). The fail-safe numbers associated with sex differences in the visual, auditory, and tactile modality are 1155, 207, and 136, respectively. In each of these clusters, the fail-safe is smaller than the criterion of $5k + 10$ (criterion: 1185 in the visual modality, 610 in the auditory modality, and 215 in the tactile modality). The criterion set by Rosenthal (1991) is admittedly arbitrary, and it could be argued that it is unlikely that 1155 studies of sex differences in visual laterality with null results are crammed in file drawers. Nevertheless, because it is difficult to determine what constitutes an unlikely fail-safe value, it is important to be aware that sex differences in laterality as reported in the present study are not resistant to the file drawer problem.

Another way to represent the file drawer problem is to use a plot of sample size (N) as a function of effect size (d). In this type of plot, when the overall effect size is significantly different from zero, unretrieved or unpublished significant effect sizes tend to truncate the distribution on one side. This plot can be found in Fig. 1 for laterality effects and in Fig. 2 for sex differences in laterality. A comparison of Figs. 1 and 2 clearly reflects the findings obtained with the fail-safe values. Specifically, effect sizes for the main effect of laterality, as seen in Fig. 1, are grouped above the zero point and the distribution is clearly truncated on the negative (left) side. This reflects the large fail-safe values associated with laterality effects. Concerning sex differences in laterality, the distribution found in Fig. 2 is not as clear-cut as the one in Fig. 1. Even though there are more effect sizes on the positive (right) side of the distribution, quite a few studies obtained negative effect sizes. The distribution of effect sizes for sex differences in laterality is therefore not as clearly truncated, reflecting the relatively small fail-safe values obtained in this analysis.

The magnitude of the estimates of effect sizes associated with sex differences in laterality suggests that a further distinction needs to be made between significance and meaningfulness. The largest significant effect size is found in the visual modality when nonverbal tasks are used, and it only represents a difference of 0.076 standard deviation units between the laterality effect of men and women. According to Cohen (1977) a d of 0.076 represents a small effect size. It is significant in the present study because meta-analytic procedures are

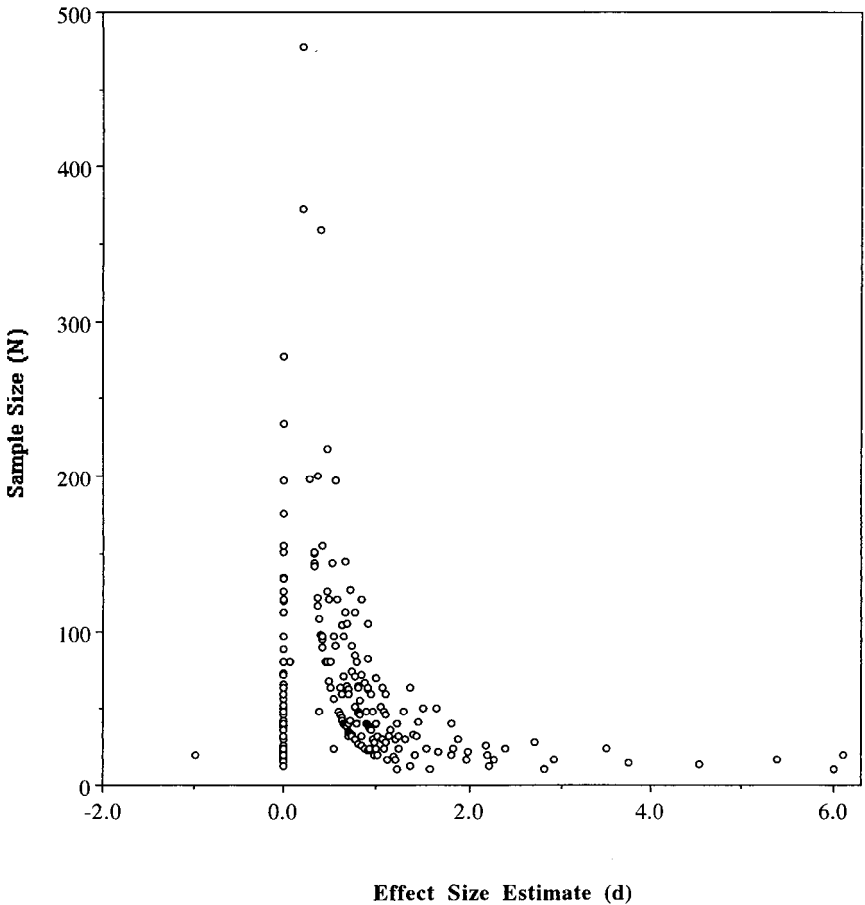


FIG. 1. Plot of sample size (N) as a function of effect size estimate (d) for laterality effects.

very sensitive and provide a powerful test of the null hypothesis, especially when a large number of studies is sampled. In fact if one was to obtain an effect size of 0.076 in a single study, a sample of about 2244 participants would be required to achieve significance at the 0.05 level with 95% power. None of the single studies sampled in the present analysis included such a large sample. For the purpose of comparison, 29 participants would be required to achieve significance at the 0.05 level with 95% power if the largest effect size obtained with laterality effects (0.674 in verbal tasks presented visually; see Table 1) was observed in a single study. It thus appears that the statistical significance of small effect sizes in a meta-analysis is due to the fact that it allows evaluation of the effects based on a procedure that virtually cumulates sample sizes as well as

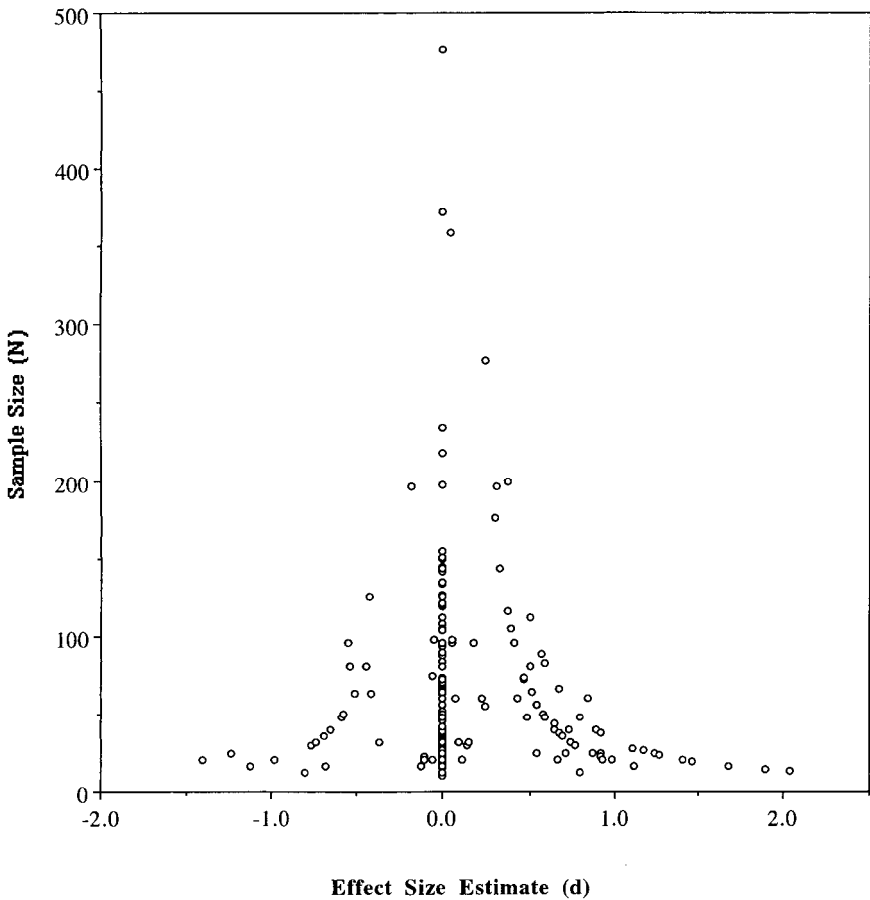


FIG. 2. Plot of sample size (N) as a function of effect size estimate (d) for sex differences in laterality.

significance levels. The importance of this factor is illustrated by the finding that an effect size of 0.076 represents an effect that accounts for approximately 0.1% of the total variance in a given sample (the effect size of 0.674 in the analysis of laterality effects accounts for about 10% of the total variance). Such a small percentage of variance accounted for requires a very large sample size to achieve significance. This effect reaches significance in the present analysis because it includes results obtained with a total of 21,197 participants (10,630 men, 10,567 women). This demonstrates that meta-analysis is a very powerful instrument. However, the fail-safe value can be considered its way of warning the meta-analyst against a Type I error. This suggests that sex differences might

be significant in the context of a meta-analysis, but not meaningful at the level of an individual study.

The purpose of this discussion is to warn researchers to use caution in interpreting the present results. Even though the present analysis indicates that sex differences in laterality are significant, it does not mean that each individual study will produce a significant effect. Caution in interpreting meta-analytic results is particularly recommended when small fail-safe values are obtained.

It is interesting to note that the largest effect sizes for sex differences in laterality were found in the tactile modality, but they only achieved one-tailed statistical significance. Fortunately, the directionality of the hypothesis predicting the presence of sex differences in laterality makes the use of one-tailed significance legitimate. The significance of sex differences in the tactile modality is compatible with the results of the weighted regression analysis, which indicated that the magnitude of sex differences in laterality is not affected by testing modality and type of task. This suggests that sex differences in laterality in the tactile modality are comparable to those in the other modalities. However, the relatively small sample of studies in this grouping and the generally small number of subjects used in tactile studies did not provide enough power for two-tailed significance to be detected. This lack of power is especially important when one considers that only one of the partitions required to achieve homogeneity in the tactile modality achieved one-tailed significance (line orientation, see Table 4). This reflects the problems inherent in the need to partition effect sizes in small homogeneous groupings, because as groupings become smaller, power is reduced proportionally.

Another relevant problem concerns the apparently arbitrary decision to attribute a Z value of zero when authors reported the effects as nonsignificant and provided no further information. Even though this is the approach recommended by Rosenthal (1991), one may argue that this procedure is too conservative. However, the prevalence of this type of findings in the analysis of sex differences in laterality, where 271 (68.4%) of the Z values were equal to zero, reflects the conclusions reached in the present discussion. Given such a large number of studies with null results, it can be argued that this type of finding is typical. Therefore, it is legitimate to state that sex differences in laterality are significant but small in the context of meta-analysis, but that null results are prevalent at the level of individual studies. It is worth noting that the frequency of Z values of zero was much smaller in the analysis of laterality effects. Of the studies analysed, 118 (29.8%) produced this type of results. Null results are therefore not as frequent when laterality effects are concerned and this reflects the relatively large effect sizes and fail-safe values obtained when these effects were analysed.

CONCLUSIONS

The present meta-analysis indicates that laterality effects are clearly significant and relatively consistent. Specifically, the largest laterality effects are found in verbal tasks in the visual and auditory modality. Naming tasks produce the largest laterality effect in all the tasks sampled in the present analysis. However, it is important to remember that the conclusions of the present study concerning laterality effects apply only to studies that included sex of subjects as an independent variable in their design. It is probable that the results obtained in the present study in relation to the main effect of laterality can be extrapolated to studies that did not investigate the influence of sex of subjects on their results. Precise estimates of effect sizes for these studies remain to be determined.

Sex differences in functional asymmetries are also significant and they are relatively consistent, given the ease with which homogeneity of significance levels was obtained when they were partitioned in meaningful clusters. However, the present results indicate that sex differences in laterality are susceptible to the file drawer problem. This suggests that caution is recommended when attempting to use the present findings for the purpose of theory elaboration. Even though it is plausible to argue that a valid theory of lateralisation needs to account for sex differences in laterality, the limitations of the present analysis must be taken into account. When a sample of exclusively published studies is analysed, the small fail-safe numbers take on a very important meaning. It is quite probable that most researchers interested in laterality have a large number of unpublished studies showing non-significant sex differences gathering dust in their file drawers. Of course, this number would be very difficult to determine, but it would probably be much larger than the fail-safe values obtained in the present meta-analysis if all the laterality researchers across the world were considered. In view of this problem, it might be more fruitful to investigate why sex differences in laterality are found in some studies but not in others. A partial answer to this question may reside in the fact that many studies involve samples of participants that are too small to allow detection of sex differences. One can only wonder why such low sample sizes are used when the effects under study are so small. It is quite likely that, because most researchers are not interested primarily in the detection of sex differences, they use a sample size sufficient for the main effect of laterality to reach significance. Given that the magnitude of laterality effects is much larger than that of sex differences in laterality, this results in studies that are powerful enough to detect laterality effects but not sex differences. From this perspective, it might prove useful for future researchers to present power calculations for sex differences in laterality and to include the sample size required for this effect to be significant when it fails to reach conventional significance. Finally, the presence of sex differences in laterality argues for the plausibility of hypotheses stating that sex differences in cognitive abilities stem from sex differences in

lateralities. Specifically, the results of the present study argue for the plausibility of Levy's (1971) hypothesis, because they support the notion that men are more lateralised than women. Thus, sex differences in lateralities might be responsible for cognitive sex differences. However, a meta-analysis does not allow definite causal claims to be made. For this reason, more empirical work is needed to test the hypothesis that there is a causal link between lateralities and cognitive skills.

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APPENDIX

Note: The following studies were included in the meta-analysis.

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