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THE MEAN NUMBER OF MAXIMA ABOVE HIGH LEVELS IN GAUSSIAN RANDOM FIELDS

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Abstract

An asymptotic formula for the mean number of maxima above a level of an n -dimensional stationary Gaussian field has been given by Nosko without proof. In this note a short general proof of this formula is given.

RANDOM FIELDS; MAXIMA ABOVE A LEVEL; GAUSSIAN PROCESSES

Introduction

Belyayev [1] attributes to Nosko [2] the following asymptotic result for $M_\xi(S)$, the mean number of maxima above a level ξ of an n -dimensional stationary Gaussian field $X(t)$, $t = t_1, \dots, t_m$, with mean zero, in the Lebesgue measurable set S of measure $w(S)$:

$$(1) \quad M_\xi(S)/w(S) = (2\pi)^{-\frac{1}{2}(n+1)} \mu^{-n+\frac{1}{2}} |A| \xi^{n-1} \exp(-\xi^2/2\mu) [1 + O(1/\xi)], \dots$$

where $|A|$ is the determinant of the covariance matrix of the partial derivatives $\partial X/\partial t_i$ and μ is the variance of X .

However, a perusal of Nosko's paper reveals that the result is stated without proof. A thorough literature search carried out by the author and some colleagues failed to produce a published proof of the general result. Only a partial proof is given in Nosko [3], applicable only to a two-dimensional isotropic Gaussian field, and the method used cannot be generalized to more than two dimensions.

In this note a short proof of (1) for the general case is given.

The exact mean number of maxima above a level

The following exact result is given by Belyayev [1]. Let $X(t)$, ($t = t_1, \dots, t_n$) be a separable, stationary Gaussian field with mean zero, twice differentiable in quadratic mean. Let $X_i(t) = \partial X(t)/\partial t_i$, $X_{ij}(t) = \partial^2 X(t)/\partial t_i \partial t_j$.

Let the following conditions be satisfied:

$$\max_{i,j} E |X_{ij}(t+h) - X_{ij}(t)|^2 \leq \frac{C}{|\log |h||^{1+\epsilon}}$$

for $C > 0$, $\epsilon > 0$; $i, j = 1, \dots, n$.

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We introduce the following notation. Let $\Delta = (z_{ij})$ be a symmetric $n \times n$ matrix. Then $z = \text{Vec}(z_{ij})$ will denote the column vector of length $n(n+1)/2$ obtained from the symmetric matrix (z_{ij}) by placing the successive columns on and above the main diagonal under one another. Let $y = (y_i)$ be a vector of length n . Let $\varphi(x, y, z)$ be the joint density function of $X(t)$, $X_i(t)$, $X_{ij}(t)$. Let S be a Lebesgue measurable set in R^n , and $M_\xi(S)$ the mean number of maxima of $X(t)$ above ξ in S . Let $w(S)$ be the Lebesgue measure of S . Let $|\Delta|$ represent the determinant of Δ . Then

$$(2) \quad m_\xi = M_\xi(S)/w(S) = (-1)^n \int_L |\Delta| \varphi(x, \mathbf{0}, z) dx dz, \dots$$

where L is the set of points (x, z) for which Δ is negative definite, and $x > \xi$.

Proof of the asymptotic formula

The basic idea of the proof is simply that as $\xi \rightarrow \infty$, the z -set for which Δ is negative definite extends to the whole z -space.

To prove this we analyse more closely the function φ . First we carry out a number of simplifying transformations.

(1) It is easy to see that there always exists an orthogonal transformation of the field coordinates t which will make the derivatives X_i uncorrelated.

(2) We change the scale of the t_i as well as the scale of X to make the variances of X and the X_i equal to unity.

By considering the spectral representation of X , it is then trivial to reach the following conclusions

- (1) X and X_i are uncorrelated for all i ;
- (2) X and X_{ij} are uncorrelated for $i \neq j$;
- (3) X_i and X_j are uncorrelated for $i \neq j$;
- (4) X_i and X_{jk} are uncorrelated for all i, j, k .

Thus we can write

$$\varphi(x, y, z) = \varphi_1(x) \varphi_2(y) \varphi_3(z | x),$$

and we note that

$$\varphi_2(\mathbf{0}) = (2\pi)^{-\frac{1}{2}n} = (-1)^n \alpha, \text{ say.}$$

Thus

$$m_\xi = \alpha \int_\xi^\infty \varphi_1(x) dx \int_P |\Delta| \varphi_3(z | x) dz,$$

where P is the set of z for which Δ is negative definite.

On account of (2) above, the only effect of conditioning the variables X_{ij} on

$X = x$ will be to shift the mean of the X_{ij} by an amount $-x$, and to change the covariance matrix in a way independent of x . If we make the change of variables $z_{ij} = v_{ij} - x\delta_{ij}$, where δ_{ij} is the Kronecker delta, we find

$$m_\xi = \alpha \int_\xi^\infty \varphi_1(x) dx \int_{U(x)} |V - xI| \varphi_4(v) dv,$$

where $V = (v_{ij})$, $v = \text{Vec}(v_{ij})$, $\varphi_4(v)$ is a multivariate normal distribution independent of x , and $U(x)$ is the set of points v for which the matrix $V - xI$ is negative definite.

We now claim that by making x sufficiently large, we can make $U(x)$ contain the sphere $\sigma(R)$ centred at the origin with arbitrary radius R . In fact, since V is symmetric, we can diagonalize it by means of an orthogonal matrix P . Thus

$$P'(V - xI)P = P'VP - xI = \text{diag}(w_1 - x, \dots, w_n - x),$$

say.

Thus Δ will be negative definite provided $x > \max_i w_i$. But $\max_i w_i \leq \|V\|$, where $\|\cdot\|$ denotes any matrix norm. Thus, in particular

$$\max_i w_i \leq \max_j \sum_i |v_{ij}| \leq n \max_{i,j} |v_{ij}| \leq n \|v\|,$$

and $U(x)$ will contain the sphere $\sigma(R)$ provided $x > nR$.

Next we note that $|V - xI|$ is the characteristic polynomial of the matrix V , so that the coefficient of x^n is $(-1)^n$. It is now easy to estimate that the relative error made in replacing $U(x)$ by the whole space of v and the determinant $|\Delta|$ by $(-1)^n x^n$ is $O(1/\xi)$. Thus

$$m_\xi = \alpha \int_\xi^\infty (-1)^n x^n \varphi_1(x) dx [1 + O(1/\xi)].$$

And, since $\int_\xi^\infty x^n \varphi_1(x) dx = \xi^{n-1} \varphi_1(\xi) [1 + O(1/\xi)]$, we obtain the asymptotic formula

$$m_\xi = (2\pi)^{-\frac{1}{2}(n+1)} \xi^{n-1} \exp(-\xi^2/2) [1 + O(1/\xi)].$$

Finally, reverting to the original coordinates of t and X , we obtain Nosko's formula (1).

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