

# Probability and Random Processes (15B11MA301)

## Lecture-17



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# Course Content

- Uniform Distribution
- Exponential Distributions
- Statistical Properties of Uniform and Exponential Distribution
- Solved Examples
- Practice Questions

# Continuous Distributions

## (i) Uniform distribution

If a random variable  $X$  is uniformly distributed over an interval  $(a, b)$ , then its probability density function (pdf)  $f(x)$  is constant in that interval, i.e.,

$$f(x) = \begin{cases} k(\text{constant}) & \text{when } a < x < b. \\ 0, & \text{otherwise.} \end{cases}$$

As  $f(x)$  is a valid probability density function (pdf), therefore

$$\int_a^b f(x)dx = 1.$$

$$\Rightarrow k \int_a^b 1 \cdot dx = k(b - a) \text{ or } k = \frac{1}{b - a}.$$

$$\text{Hence, } f(x) = \begin{cases} \frac{1}{b - a}, & a < x < b \\ 0, & \text{otherwise.} \end{cases}$$

**Example 1:** If a random variable  $X$  is uniformly distributed over the interval  $[0, 5]$ . compute the following:

- (i)  $P(1 < X < 3)$       (ii)  $P(X < 4)$       (iii)  $P(X > 1)$   
 (iv)  $P(2 < X < 5 / X < 5)$     (v)  $P(0 < X < 5 / X > 1)$ .

**Solution**

Here,  $a = 0, b = 5$   $f(x) = \frac{1}{5}, 0 \leq x \leq 5$      $\left( \because f(x) = \frac{1}{b-a} \right)$

$$(i) \quad P(1 < X < 3) = \int_1^3 f(x) dx = \int_1^3 \frac{1}{5} dx = 2/5 = 0.40$$

$$(ii) \quad P(X < 4) = \int_0^4 f(x) dx = \int_0^4 \frac{1}{5} dx = 4/5 = 0.80$$

$$(iii) \quad P(X > 1) = \int_1^5 f(x) dx = \int_1^5 \frac{1}{5} dx = 4/5 = 0.80$$

$$(iv) \quad P(2 < X < 5 / X < 5) = \frac{P(2 < X < 5)}{P(X < 5)} = \frac{3/5}{1} = 0.60$$

$$(v) \quad P(0 < X < 4 / X > 1) = \frac{P(1 < X < 4)}{P(X > 1)} = \frac{3/5}{4/5} = 0.75.$$

## Mean and Variance of Uniform Distribution

$$\begin{aligned} \text{(i) Mean of } X &= E(X) = \int_a^b xf(x)dx = \int_a^b x \frac{1}{b-a} dx \\ &= \frac{1}{b-a} \left( \frac{x^2}{2} \right)_a^b = \frac{a+b}{2}. \end{aligned}$$

$$\text{(ii) Variance of } X = Var(X) = E(X^2) - E(X)^2$$

$$\text{We have } E(X) = \frac{a+b}{2} \text{ and}$$

$$\begin{aligned} E(X^2) &= \int_a^b x^2 f(x)dx = \frac{1}{3(b-a)} \left[ x^3 \right]_a^b \\ &= \frac{(b^3 - a^3)}{3(b-a)} = \frac{b^2 + ab + a^2}{3}. \end{aligned}$$

$$\text{Var}(X) = \frac{(a-b)^2}{12}$$

## Moment Generating Function (MGF) of Uniform Distribution

- Let  $X$  be a uniformly distributed random variable in the interval  $[a, b]$ , then its moment generating function is defined by

$$\begin{aligned} M_X(t) &= E(e^{tX}) = \int_a^b f(x)e^{tx} dx \\ &= \frac{1}{b-a} \int_a^b e^{tx} dx = \frac{1}{b-a} \left( \frac{e^{tx}}{t} \right)_a^b \\ &= \frac{1}{b-a} \left( \frac{e^{tb}}{t} - \frac{e^{ta}}{t} \right) = \frac{e^{tb} - e^{ta}}{(b-a)t}. \\ &= \frac{1}{(b-a)} \sum_{r=0}^{\infty} \frac{t^r}{(r+1)!} (b^{r+1} - a^{r+1}). \end{aligned}$$

- Moments from MGF of Uniform Distribution

The  $r$ th moment of  $X$  about origin = Coefficient of  $\frac{t^r}{r!}$  in  $M_X(t)$ .

$$\text{That is, we have } E(X^r) = \frac{b^{r+1} - a^{r+1}}{(b-a)(r+1)}.$$

$$\therefore \text{Mean} = E(X) = \frac{b^2 - a^2}{(b-a)(1+1)} = (b+a)/2.$$

$$E(X^2) = \frac{b^3 - a^3}{3(b-a)} = \frac{a^2 + ab + b^2}{3}.$$

$$\therefore \text{Variance}(X) = E(X^2) - \{E(X)\}^2 = \frac{(b-a)^2}{12}.$$

- The Characteristic function of Uniform Distribution

$$\phi_X(\omega) = E(e^{i\omega X}) = \frac{1}{b-a} \int_a^b e^{i\omega x} dx = \frac{e^{i\omega b} - e^{i\omega a}}{i\omega(b-a)}.$$

**Example 2:** Find the characteristic function of a random variable  $X$  which is uniformly distributed over the interval  $[1, 6]$  and hence find the first three moments about origin from it.

**Solution:** Here  $a = 1, b = 6$  and  $f(x) = \frac{1}{b-a} = \frac{1}{5}$

$$\begin{aligned}\phi_X(\omega) &= E(e^{i\omega x}) = \int_1^6 f(x)e^{i\omega x} dx \\ &= \frac{1}{5} \int_1^6 e^{i\omega x} dx = \frac{e^{i\omega 6} - e^{i\omega}}{5i\omega}. \\ &= \frac{1}{5} \sum_{r=0}^{\infty} \frac{(i\omega)^r}{(r+1)!} (6^{r+1} - 1).\end{aligned}$$

$$E(X^r) = \text{coefficient of } \frac{(i\omega)^r}{r!} = \frac{6^{r+1} - 1}{r+1}$$

Putting ,  $r = 1, 2, 3$ , we obtain

$$E(X) = 3.50, \quad E(X^2) = 215 / 15 = 14.33 \quad \text{and} \quad E(X^3) = 64.75.$$



## (ii) Exponential Distribution

- A random variable  $X$  is said to be exponentially distributed if its probability density function is given by

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x \geq 0 \\ 0 & \text{otherwise,} \end{cases}$$

where  $\lambda > 0$  is a parameter.

- Examples. Life time of a machine or its components; Arrival times of telephone calls or arrival times of train at a railway station etc.
- If the occurrences of events over nonoverlapping intervals are independent, then the waiting time distribution of these events can be shown to be exponential

## Mean and Variance of Exponential Distribution

(i) Mean of  $X = E(X) = \int_0^{\infty} x\lambda e^{-\lambda x} dx$

$$= \lambda \left[ x \left( \frac{e^{-\lambda x}}{-\lambda} \right) - \left( \frac{e^{-\lambda x}}{\lambda^2} \right) \right]_0^{\infty} = \frac{1}{\lambda}.$$

(ii) Variance of  $X = Var(X) = E(X^2) - \{E(X)\}^2$

$$E(X^2) = \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx$$

$$= \lambda \left[ x^2 \left( \frac{e^{-\lambda x}}{-\lambda} \right) - 2x \left( \frac{e^{-\lambda x}}{\lambda^2} \right) + 2 \left( \frac{e^{-\lambda x}}{-\lambda^3} \right) \right]_0^{\infty} = \frac{2}{\lambda^2}.$$

$$\therefore Var(X) = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}.$$

## The $r$ th Moment About Origin

$$\mu'_r = E(X^r) = \int_0^{\infty} x^r \lambda e^{-\lambda x} dx$$

Put  $\lambda x = z$ , we obtain

$$\mu'_r = E(X^r) = \frac{1}{\lambda^r} \int_0^{\infty} z^{(r+1)-1} e^{-z} dz = \frac{\Gamma(r+1)}{\lambda^r}.$$

[Use the definition of Gamma function, i.e.,  $\int_0^{\infty} x^{n-1} e^{-x} dx = \Gamma(n)$ ]

- On putting  $r = 1, 2, 3, \dots$ , all the moments about origin can be obtained.

## Moment Generating Function (MGF) of Exponential Distribution

- Let  $X$  be an exponential random variable, then its moment generating function is defined by

$$\begin{aligned}M_X(t) &= E(e^{tX}) = \int_0^{\infty} e^{tx} (\lambda e^{-\lambda x}) dx \\&= \lambda \int_0^{\infty} e^{-x(\lambda-t)} dx = \lambda \left( \frac{e^{-x(\lambda-t)}}{t-\lambda} \right)_0^{\infty} \\&= \frac{\lambda}{\lambda-t} = \left( 1 - \frac{t}{\lambda} \right)^{-1} = 1 + \frac{t}{\lambda} + \left( \frac{t}{\lambda} \right)^2 + \left( \frac{t}{\lambda} \right)^2 + \dots\end{aligned}$$

The  $r$ th moment of  $X$  about origin = Coefficient of  $\frac{t^r}{r!}$  in  $M_X(t)$ .

$$\therefore \mu'_r = E(X^r) = \frac{|r|}{\lambda^r}, \quad r = 1, 2, 3, \dots$$

- The Characteristic function of Exponential Distribution

$$\phi_X(\omega) = E(e^{i\omega X}) = \int_0^{\infty} e^{i\omega x} (\lambda e^{-\lambda x}) dx$$

$$= \lambda \int_0^{\infty} e^{-x(\lambda - i\omega)} dx = \lambda \left( \frac{e^{-x(\lambda - i\omega)}}{i\omega - \lambda} \right)_0^{\infty}$$

$$= \frac{\lambda}{\lambda - i\omega} = \left( 1 - \frac{i\omega}{\lambda} \right)^{-1}$$

$$= 1 + \frac{i\omega}{\lambda} + \left( \frac{i\omega}{\lambda} \right)^2 + \left( \frac{i\omega}{\lambda} \right)^2 + \dots + \left( \frac{i\omega}{\lambda} \right)^r + \dots$$

The  $r$ th moment of  $X$  about origin can be obtained as:

the coefficient of  $\frac{(i\omega)^r}{r!}$  in  $\phi_X(\omega)$  or  $\frac{1}{i^r} \left( \frac{d^r}{d\omega^r} (\phi_X(\omega)) \right)_{\omega=0}$

$$\therefore \mu'_r = E(X^r) = \frac{r!}{\lambda^r}, r = 1, 2, 3, \dots$$

**Example 3:** Let a random variable  $X$  follows an exponential distribution with mean of 20. Find (i)  $P(X < 30)$  (ii)  $P(X < 10 / X > 1)$  .

**Solution:** (i) We have mean  $= E(X) = \frac{1}{\lambda} = 20, \Rightarrow \lambda = 1 / 20 = 0.05$

$$P(X < 30) = \frac{1}{20} \int_0^{30} e^{-x/20} dx = \frac{1}{20} \left( \frac{e^{-x/20}}{-1/20} \right)_0^{30} = 1 - e^{-1.5}.$$

$$(ii) P(X < 10 / X > 1) = \frac{P(1 < X < 10)}{P(X > 1)} = \frac{1}{20} \int_1^{10} e^{-x/20} dx$$

$$P(1 < X < 10) = \frac{1}{20} \left( \frac{e^{-x/20}}{-1/20} \right)_1^{10} = e^{-0.05} - e^{-0.5}.$$

## Memoryless Property of Exponential Distribution

If  $X$  is exponentially distributed, then for any  $s, t > 0$

$$P(X > s + t \mid X > s) = P(X > t).$$

**Proof:** We have, for any  $k > 0$

$$P(X > k) = \int_k^{\infty} \lambda e^{-\lambda x} dx = \left( -e^{-\lambda x} \right)_k^{\infty} = e^{-\lambda k}.$$

$$\begin{aligned} \therefore P(X > s + t \mid X > s) &= \frac{P\{X > s + t \text{ and } X > s\}}{P\{X > s\}} \\ &= \frac{P(X > s + t)}{P(X > s)} = \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}} = e^{-\lambda t} = P(X > t). \end{aligned}$$

**Example 4:** The time  $X$  (in hours) to repair a mobile is exponentially distributed with mean  $1/10$ . Find the probability that (i) the repair time is less than 30 minutes, (ii) the repair time is between 1 to 2 hours (iii) a repair time is greater than 5 hours given that its duration exceeds 3 hours.

**Solution:** We have mean  $= E(X) = \frac{1}{\lambda} = \frac{1}{10}, \Rightarrow \lambda = 10$

$$(i) P(X > 0.5) = 10 \int_{0.5}^{\infty} e^{-10x} dx = 10 \left( \frac{e^{-10x}}{-10} \right)_{0.5}^{\infty} = e^{-5}.$$

$$(ii) P(1 < X < 2) = 10 \int_1^2 e^{-10x} dx = e^{-10} - e^{-20}.$$

$$\begin{aligned} (iii) P(X > 5 / X > 3) &= P(X > 3 + 2 / X > 3) \\ &= P(X > 2) = 10 \int_2^{\infty} e^{-10x} dx = e^{-20}. \end{aligned}$$



## Practice Questions

1. If  $X$  is uniformly distributed over the interval  $(-3, 3)$ , then evaluate the following:

(i)  $P(X > 1)$ , (ii)  $P(0 < X < 3)$ , (iii)  $\text{Var}(X)$ , (iv)  $P(|X| > 2)$ .

[Ans: (i)  $1/3$ , (ii)  $1/2$ , (iii)  $3$ , (iv)  $1/3$ ]

2. The time in hours required to repair a motor is exponentially distributed with parameter  $\lambda = 1/3$ . What is the probability that (i) the repair time exceeds 3 hours (ii) a repair takes 10 hours given that its duration exceeds 7 hours.

[Ans: (i)  $1/e$  (ii)  $1/e$ ]

3. If  $X$  follows exponential distribution with  $P(X < 1) = P(X > 1)$ , find the mean and variance of  $X$ .

[Ans: Mean = 1.45,  $\text{Var}(X) = 2.08$ ]

## References/Further Reading

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