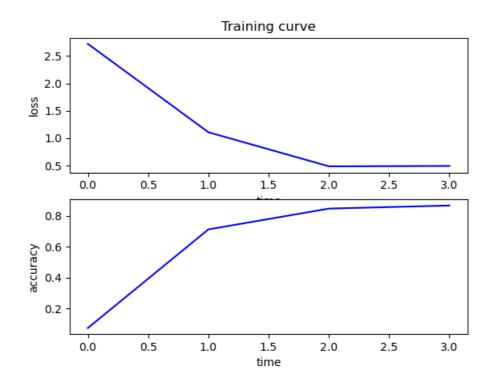
CS 229, Fall 2018

Problem Set #4 Solutions: EM, DL, & RL

1.



2.

(a)

If $\hat{\pi}_0 = \pi_0$, then

$$\mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} R(s,a) = \mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \frac{\pi_1(s,a)}{\pi_0(s,a)} R(s,a)$$

$$= \sum_{(s,a)} \frac{\pi_1(s,a)}{\pi_0(s,a)} R(s,a) p(s) \pi_0(s,a)$$

$$= \sum_{(s,a)} R(s,a) p(s) \pi_1(s,a)$$

$$= \mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_1(s,a)}} R(s,a)$$

(b)

If $\hat{\pi}_0 = \pi_0$, then

$$\mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} = \mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \frac{\pi_1(s,a)}{\pi_0(s,a)}$$

$$= \sum_{(s,a)} p(s,a) \frac{\pi_1(s,a)}{\pi_0(s,a)}$$

$$= \sum_{(s,a)} p(s) \pi_0(s,a) \frac{\pi_1(s,a)}{\pi_0(s,a)}$$

$$= \sum_{(s,a)} p(s) \pi_1(s,a)$$

$$= \sum_{(s,a)} p(s) \pi_1(s,a)$$

$$= 1$$

$$\frac{\mathbb{E} \frac{s \sim p(s)}{s \sim \pi_0(s,a)} \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} R(s,a)}{\mathbb{E} \frac{s \sim p(s)}{a \sim \pi_0(s,a)} \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)}} = \mathbb{E} \frac{s \sim p(s)}{a \sim \pi_1(s,a)} R(s,a)$$

(c)

$$\frac{\mathbb{E} \sum\limits_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} R(s,a)}{\mathbb{E} \sum\limits_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)}} = \frac{\sum_{(s,a)} p(s,a) \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} R(s,a)}{\sum_{(s,a)} p(s,a) \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)}}$$

If there is only a single data element in the dataset, then

$$rac{\sum_{(s,a)} p(s,a) rac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} R(s,a)}{\sum_{(s,a)} p(s,a) rac{\pi_1(s,a)}{\hat{\pi}_0(s,a)}} = rac{p(s,a) rac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} R(s,a)}{p(s,a) rac{\pi_1(s,a)}{\hat{\pi}_0(s,a)}} = R(s,a)$$

So if $\pi_0 \neq \pi_1$, then

$$\mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} R(s,a) \neq \mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_1(s,a)}} R(s,a)$$

(d)

i.

If $\hat{\pi}_0 = \pi_0$, then

$$\mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \left((\mathbb{E}_{a \sim \pi_1(s,a)} \, \hat{R}(s,a)) = \mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_1(s,a)}} \, \hat{R}(s,a) \right)$$

$$\mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \left(\frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} (R(s,a) - \hat{R}(s,a)) \right) = \mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_1(s,a)}} \left(R(s,a) - \hat{R}(s,a) \right)$$

$$\mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s,a)}} \left((\mathbb{E}_{a \sim \pi_1(s,a)} \, \hat{R}(s,a)) + \frac{\pi_1(s,a)}{\hat{\pi}_0(s,a)} (R(s,a) - \hat{R}(s,a)) \right) = \mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_1(s,a)}} R(s,a)$$

ii.

If $\hat{R}(s,a)=R(s,a)$, then

$$\mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_0(s, a)}} \left(\left(\mathbb{E}_{a \sim \pi_1(s, a)} \, \hat{R}(s, a) \right) + \frac{\pi_1(s, a)}{\hat{\pi}_0(s, a)} (R(s, a) - \hat{R}(s, a)) \right) = \mathbb{E}_{\substack{s \sim p(s) \\ a \sim \pi_1(s, a)}} R(s, a)$$

(e)

i.

Importance sampling estimator. Estimating $\hat{\pi}_0$ is easier in this situation.

ii.

Regression estimator. Estimating $\hat{R}(s,a)$ is easier in this situation.

3.

$$\begin{split} f_u(x) &= \arg\min_{v \in \mathcal{V}} ||x - v||^2 = \frac{uu^T x}{u^T u} = uu^T x \\ \arg\min_{u:u^T u = 1} \sum_{i = 1}^m ||x^{(i)} - f_u(x^{(i)})||_2^2 &= \arg\min_{u:u^T u = 1} \sum_{i = 1}^m ||x^{(i)} - uu^T x^{(i)}||_2^2 \\ &= \arg\min_{u:u^T u = 1} \sum_{i = 1}^m (x^{(i)} - uu^T x^{(i)})^T (x^{(i)} - uu^T x^{(i)}) \\ &= \arg\min_{u:u^T u = 1} \sum_{i = 1}^m x^{(i)^T} x^{(i)} - x^{(i)^T} uu^T x^{(i)} \\ &= \arg\max_{u:u^T u = 1} \sum_{i = 1}^m u^T x^{(i)} x^{(i)^T} u \\ &= \arg\max_{u:u^T u = 1} \sum_{i = 1}^m u^T x^{(i)} x^{(i)^T} u \\ &= \arg\max_{u:u^T u = 1} u^T \left(\sum_{i = 1}^m x^{(i)} x^{(i)^T}\right) u \end{split}$$

4.

(a)

$$\begin{split} \nabla_{W}\ell(W) &= \nabla_{W} \sum_{i=1}^{n} \left(\log |W| + \sum_{j=1}^{d} \log \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2} (w_{j}^{T} x^{(i)})^{2} \right) \right) \\ &= m(W^{-1})^{T} - \sum_{i=1}^{n} \nabla_{W} \sum_{j=1}^{d} \frac{1}{2} (w_{j}^{T} x^{(i)})^{2} \\ &= m(W^{-1})^{T} - \sum_{i=1}^{n} W x^{(i)} x^{(i)}^{T} \\ &= m(W^{-1})^{T} - W X^{T} X \\ &= 0 \end{split}$$

$$W^T W = m(X^T X)^{-1}$$

Let R be an arbitrary orthogonal matrix, and let W'=RW. Then

$$W'^T W' = (RW)^T RW = W^T R^T RW = W^T W$$

So if W is a solution, then any W' is also a solution.

(b)

$$egin{aligned}
abla_W \ell(W) &=
abla_W \Big(\log |W| + \sum_{j=1}^d \log rac{1}{2} \mathrm{exp} ig(- |w_j^T x^{(i)}| ig) \Big) \ &= (W^{-1})^T -
abla_W \sum_{j=1}^d |w_j^T x^{(i)}| \ &= (W^T)^{-1} - \mathrm{sign}(W x^{(i)}) x^{(i)}^T \ &W := W + lpha \Big((W^T)^{-1} - \mathrm{sign}(W x^{(i)}) x^{(i)}^T \Big) \end{aligned}$$

5.

(a)

$$egin{aligned} \|B(V_1) - B(V_2)\|_{\infty} &= \gamma igg\| \max_{a \in A} \sum_{s' \in S} P_{sa}(s') ig[V_1(s') - V_2(s')ig] igg\|_{\infty} \ &= \gamma \max_{s' \in S} igg| \max_{a \in A} \sum_{s' \in S} P_{sa}(s') ig[V_1(s') - V_2(s')ig] igg| \ &\leq \gamma \|V_1 - V_2\|_{\infty} \end{aligned}$$

The inequality holds because for any $lpha,x\in\mathbb{R}^n$, if $\sum_ilpha_i=1$ and $lpha_i\geq 0$, then $\sum_ilpha_ix_i\leq \max_ix_i$

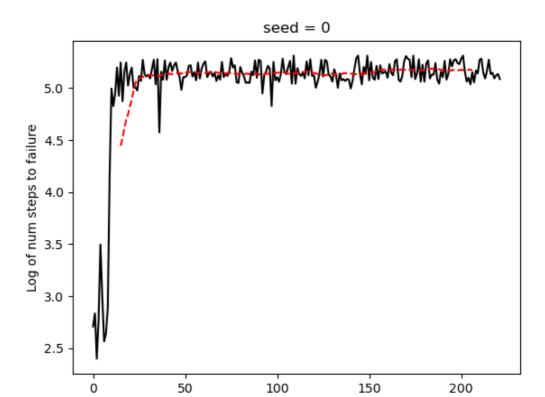
(b)

Assume that V_1 and V_2 are both fixed points, i.e., $B(V_1) = V_1, B(V_2) = V_2$.

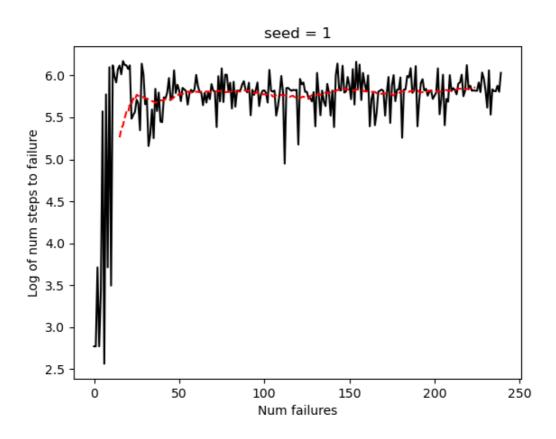
$$\|V_1-V_2\|_{\infty}=\|B(V_1)-B(V_2)\|_{\infty}\leq \gamma \|V_1-V_2\|_{\infty}$$
 $\|V_1-V_2\|_{\infty}=0$ $V_1=V_2$

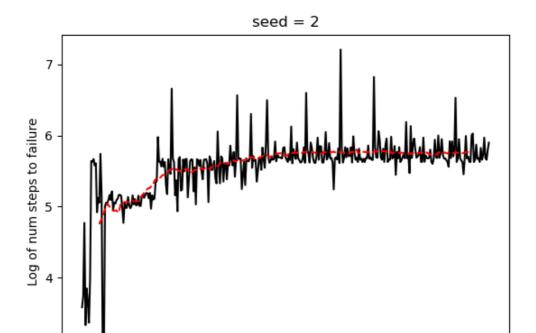
So *B* has at most one fixed point.

6.



Num failures





Num failures

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