

## HW 3

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### 1

Given  $X_1, X_2, \dots, X_n$  are  $n$  i.i.d. Bernoulli random variables following:

$$P(X = x; p) = p^x(1 - p)^{1-x}, \quad x = 0, 1$$

where  $p$  is the probability of  $X$  being 1. We want to find the MLE of  $p$  using the given samples. The likelihood function is given by:

$$L(p) = \prod_{i=1}^n P(X_i = x_i; p) = \prod_{i=1}^n p^{x_i}(1 - p)^{1-x_i}$$

To find the MLE, we take the natural logarithm of the likelihood function:

$$\ell(p) = \ln L(p) = \sum_{i=1}^n (x_i \ln p + (1 - x_i) \ln(1 - p))$$

Next, we differentiate  $\ell(p)$  with respect to  $p$  and set it to zero to find the critical points:

$$\frac{d\ell(p)}{dp} = \sum_{i=1}^n \left( \frac{x_i}{p} - \frac{1 - x_i}{1 - p} \right) = 0$$

This simplifies to:

$$\sum_{i=1}^n \frac{x_i}{p} = \sum_{i=1}^n \frac{1 - x_i}{1 - p}$$

Multiplying both sides by  $p(1 - p)$  gives:

$$(1 - p) \sum_{i=1}^n x_i = p \sum_{i=1}^n (1 - x_i)$$

Let  $S = \sum_{i=1}^n x_i$ , then we can rewrite the equation as:

$$\begin{aligned}(1-p)S &= p(n-S) \\ S - pS &= pn - pS \\ S &= pn \\ p &= \frac{S}{n} \therefore \\ \hat{p} &= \frac{\sum_{i=1}^n x_i}{n}\end{aligned}$$

Thus, the MLE of  $p$  is the sample mean of the observed data.

## 2

We are given a binary classification problem with outputs  $y \in \{0, 1\}$ , and assume that  $p(y|x)$  is Bernoulli  $\hat{p}(x; \theta^*)$  for some parameter vector  $\theta^* \in \mathbb{R}^d$ , where for each  $\theta$ ,  $\hat{p}(x; \theta)$  returns a number between  $[0, 1]$ .

- (a) We know the cross-entropy loss for each parameter  $\theta$  for a set of  $n$  examples with predictions  $\hat{p}(x_i; \theta)$  and true labels  $y_i$  is:

$$CE(\theta) = -\frac{1}{n} \sum_{i=1}^n (y_i \ln \hat{p}_i + (1 - y_i) \ln(1 - \hat{p}_i))$$

We can derive this loss function from the principle of the MLE:

$$\begin{aligned}\hat{\theta} &= \arg \max_{\theta} P(y_1, y_2, \dots, y_n | x_1, x_2, \dots, x_n; \theta) \\ &= \arg \max_{\theta} \prod_{i=1}^n P(y_i | x_i; \theta) \\ &= \arg \max_{\theta} \prod_{i=1}^n \hat{p}(x_i; \theta)^{y_i} (1 - \hat{p}(x_i; \theta))^{1-y_i} \\ &= \arg \max_{\theta} \sum_{i=1}^n (y_i \ln \hat{p}(x_i; \theta) + (1 - y_i) \ln(1 - \hat{p}(x_i; \theta))) \\ &= \arg \min_{\theta} -\frac{1}{n} \sum_{i=1}^n (y_i \ln \hat{p}(x_i; \theta) + (1 - y_i) \ln(1 - \hat{p}(x_i; \theta))) \\ &= \arg \min_{\theta} CE(\theta)\end{aligned}$$

Thus, minimizing the cross-entropy loss is equivalent to maximizing the likelihood of the observed data.

- (b) Let's instantiate the prediction as the sigmoid function applied to a linear combination of input features  $\hat{p}(x; \theta) = \sigma(\theta^T x)$ , where  $\sigma(z) = \frac{1}{1+e^{-z}}$ .

Then, minimizing the cross-entropy loss is equivalent to minimizing the familiar logistic loss given by:

$$L(\theta) = \frac{1}{n} \sum_{i=1}^n \ln(1 + e^{-\tilde{y}_i \theta^T x_i}) \text{ where } \tilde{y}_i = \begin{cases} 1 & \text{if } y_i = 1 \\ -1 & \text{if } y_i = 0 \end{cases}$$

We can show this by substituting  $\hat{p}(x_i; \theta) = \sigma(\theta^T x_i)$  into the cross-entropy loss:

$$\begin{aligned} CE(\theta) &= -\frac{1}{n} \sum_{i=1}^n (y_i \ln \sigma(\theta^T x_i) + (1 - y_i) \ln(1 - \sigma(\theta^T x_i))) \\ &= -\frac{1}{n} \sum_{i=1}^n \left( y_i \ln \frac{1}{1 + e^{-\theta^T x_i}} + (1 - y_i) \ln \left( 1 - \frac{1}{1 + e^{-\theta^T x_i}} \right) \right) \\ &= -\frac{1}{n} \sum_{i=1}^n \left( y_i \ln \frac{1}{1 + e^{-\theta^T x_i}} + (1 - y_i) \ln \frac{1}{1 + e^{\theta^T x_i}} \right) \\ &= -\frac{1}{n} \sum_{i=1}^n \left( -y_i \ln(1 + e^{-\theta^T x_i}) - (1 - y_i) \ln(1 + e^{\theta^T x_i}) \right) \\ &= \frac{1}{n} \sum_{i=1}^n \left( y_i \ln(1 + e^{-\theta^T x_i}) + (1 - y_i) \ln(1 + e^{\theta^T x_i}) \right) \end{aligned}$$

We know that:

$$\tilde{y}_i = \begin{cases} 1 & \text{if } y_i = 1 \\ -1 & \text{if } y_i = 0 \end{cases}$$

Thus, we can define  $\tilde{y}_i$  such that:

$$\begin{aligned} \tilde{y}_i &= 2y_i - 1 \therefore \\ y_i &= \frac{1 + \tilde{y}_i}{2}, 1 - y_i = \frac{1 - \tilde{y}_i}{2} \end{aligned}$$

Substituting this back into the expression for  $CE(\theta)$  gives:

$$CE(\theta) = \frac{1}{n} \sum_{i=1}^n \left( \frac{1 + \tilde{y}_i}{2} \ln(1 + e^{-\theta^T x_i}) + \frac{1 - \tilde{y}_i}{2} \ln(1 + e^{\theta^T x_i}) \right)$$

Thus, if we plug in  $\tilde{y}_i = 1$ , we get:

$$CE(\theta) = \frac{1}{n} \sum_{i=1}^n \ln(1 + e^{-\theta^T x_i})$$

And if we plug in  $\tilde{y}_i = -1$ , we get:

$$CE(\theta) = \frac{1}{n} \sum_{i=1}^n \ln(1 + e^{\theta^T x_i})$$

which is equivalent to the logistic loss function:

$$L(\theta) = \begin{cases} \frac{1}{n} \sum_{i=1}^n \ln(1 + e^{-\theta^T x_i}) & \text{if } \tilde{y}_i = 1 \\ \frac{1}{n} \sum_{i=1}^n \ln(1 + e^{\theta^T x_i}) & \text{if } \tilde{y}_i = -1 \end{cases}$$

**3**

**4**

(a) See attached Jupyter notebook for code and plots.

(b)

| Dataset | Gaussian Naive Bayes   | Linear Logistic Regression   |
|---------|--|--|
| 1       | Each feature has mixed overlap between 1 and 2. Because GNB assumes that each feature is independent, they are modeled separately and thus the boundary becomes a vertical line.   | Dataset is linearly separable so logistic regression performs really well, giving us a clean fit.  |
| 2       | The decision boundary becomes quadratic, so GNB can create a smooth circled boundary. However because the features are correlated in a slanted ellipse, naive bayes fails to recognize this and thus still somewhat poorly fits, but is still better than logistic regression. | Fails as data is not linearly separable but rather overlaps. Thus, it draws a line that minimizes classification error, but it cannot match the ellipse structure. |
| 3       | Each class is radially symmetric, and even though features are not independent (related to radius), GNB can still model it well in circles.  | Fails as data is not linearly separable. Thus again, it draws a line that minimizes error but cannot match circular shape as it is constrained to 1 dimension      |