APMA 1655 Honors Statistical Inference I

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Homework 3

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• You are strongly encouraged to work in groups, but solutions must be written independently.

1 Review

To help you better answer the questions in HW 3, we review the example of Bernoulli distributions as follows:

- The experiment of interest is flipping a fair coin;
- the sample space corresponding to this experiment is $\Omega = \{\text{heads}, \text{tails}\}$;
- the probability \mathbb{P} is defined by $\mathbb{P}(A) = \frac{\#A}{\#\Omega}$, i.e., $\mathbb{P}(\{\texttt{heads}\}) = \mathbb{P}(\{\texttt{tails}\}) = \frac{1}{2}$;
- the random variable X is defined by

$$X(heads) = 1, X(tails) = 0.$$

The CDF of X is

$$F_X(x) = \begin{cases} 0, & \text{if } x < 0, \\ \frac{1}{2}, & \text{if } 0 \le x < 1, \\ 1, & \text{if } x \ge 1. \end{cases}$$
 (1)

Proof:

- $1. \ \mathrm{When} \ x<0, \ \mathrm{we \ have} \ A_x=\{\omega\in\Omega \, : \, X(\omega)\leq x\}=\emptyset; \ \mathrm{then}, \ F_X(x)=\mathbb{P}(A_x)=\mathbb{P}(\emptyset)=0.$
- 2. When $0 \le x < 1$, we have $A_x = \{\omega \in \Omega : X(\omega) \le x\} = \{\text{tails}\}$; then, $F_X(x) = \mathbb{P}(A_x) = \mathbb{P}(\{\text{tails}\}) = \frac{1}{2}$.
- $3. \ \ \mathrm{When} \ x \geq 1, \ \mathrm{we \ have} \ A_x = \{\omega \in \Omega \, : \, X(\omega) \leq x\} = \Omega; \ \mathrm{then}, \ F_X(x) = \mathbb{P}(A_x) = \mathbb{P}(\Omega) = 1.$

The proof is completed.

In addition, the Wikipedia page on random variables is nice material for learning the concept of random variables.

2 Problem Set

- 1. Let (Ω, \mathbb{P}) be a probability space. Suppose B is an event and $0 < \mathbb{P}(B) < 1$. Please prove the following:
 - (a) (1 point) If A and B are independent, then A and B^c are also independent.

Proof. Using (b), we have that
$$\mathbb{P}(B^c | A) = 1 - \mathbb{P}(B | A) = 1 - \mathbb{P}(B) = \mathbb{P}(B^c)$$
.

(b) (1 point) $\mathbb{P}(A|B) + \mathbb{P}(A^c|B) = 1$.

Proof. We will first prove that $\mathbb{P}(A \cap B) = \mathbb{P}(B) - \mathbb{P}(A^c \cap B)$.

$$\begin{split} \mathbb{P}(A \cap B) &= \mathbb{P}(B \cap A) & \text{(commutativity)} \\ &= 1 - \mathbb{P}((B \cap A)^c) & \text{(def of complement)} \\ &= 1 - \mathbb{P}(B^c \cup A^c) & \text{(De Morgan's Law)} \\ &= 1 - \mathbb{P}(B^c \cup A^c \cap \Omega) & \text{($E \cap \Omega = E$)} \\ &= 1 - \mathbb{P}(B^c \cup A^c \cap (B \cup B^c)) & \text{(def of complement)} \\ &= 1 - \mathbb{P}(B^c \cup (A^c \cap B) \cup (A^c \cap B^c)) & \text{(distributive law)} \\ &= 1 - \mathbb{P}(B^c \cup (A^c \cap B^c) \cup (A^c \cap B)) & \text{(commutativity)} \\ &= 1 - \mathbb{P}(B^c \cup (A^c \cap B)) & \text{(def of \cup)} \\ &= 1 - \mathbb{P}(B^c) - \mathbb{P}(A^c \cap B) & \text{(additivity)} \\ &= \mathbb{P}(B) - \mathbb{P}(A^c \cap B) & \text{(additivity)} \end{split}$$

Now, we can use this relation to show that $\mathbb{P}(A|B) + \mathbb{P}(A^c|B) = 1$.

$$\mathbb{P}(A \mid B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} \qquad \text{(conditional probability)}$$

$$= \frac{\mathbb{P}(B) - \mathbb{P}(A^c \cap B)}{\mathbb{P}(B)} \qquad \text{(substitute from above)}$$

$$= \frac{\mathbb{P}(B)}{\mathbb{P}(B)} - \frac{\mathbb{P}(A^c \cap B)}{\mathbb{P}(B)} \qquad \text{(distributive prop.)}$$

$$= 1 - \mathbb{P}(A^c \mid B) \qquad \text{(conditional probability)}$$

2. (2 points) Let \mathfrak{n} be a positive integer, and $\Omega \stackrel{\mathrm{def}}{=} \{1,2,\ldots,\mathfrak{n}\}$. Suppose \mathbb{P} is a function of subsets of Ω defined as follows

$$\mathbb{P}(A) \stackrel{\text{def}}{=} \frac{\#A}{\#\Omega}$$
, for all $A \subset \Omega$.

You have proved in HW 1 that (Ω, \mathbb{P}) is a probability space.

We define a random variable X as follows

$$X(\omega) = \omega$$
, for all $\omega \in \Omega = \{1, 2, ..., n\}$.

Please derive the CDF of the random variable X defined above. Please present your answer using a formula like the one in Eq. (1).

Proof. Consider the following cases for the CDF of the random variable X:

- When x < 1, we have that $F_X(x) = \mathbb{P}(X \le x) = \mathbb{P}(\emptyset) = 0$.
- $\bullet \ \ \mathrm{When} \ 1 \leq x < n, \ \mathrm{we \ have \ that} \ F_X(x) = \mathbb{P}(X \leq x) = \mathbb{P}(\{1, \dots, \lfloor x \rfloor\}) = \frac{\lfloor x \rfloor}{n}.$
- When $x \ge n$, we have that $F_X(x) = \mathbb{P}(X \le x) = \mathbb{P}(\Omega) = 1$.

Therefore,
$$F_X(x) = \begin{cases} 0 & \text{if } x < 1, \\ \frac{\lfloor x \rfloor}{n} & \text{if } 1 \le x < n, \\ 1 & \text{if } x \ge n. \end{cases}$$

3. (2 points) Let X be a random variable defined on the probability space (Ω, \mathbb{P}) . Suppose X satisfies the following

Please derive the CDF of the random variable X. Please present your answer using a formula like the one in Eq. (1).

Proof. Consider the following cases for $F_X(x)$:

- When x < 1, we have that $F_X(x) = \mathbb{P}(X \le x) = \mathbb{P}(\emptyset) = 0$.
- When $1 \le x < 2$, we have that $F_X(x) = \mathbb{P}(X \le x) = 0 + 1/2 = 1/2$.
- When $2 \le x < 3$, we have that $F_X(x) = \mathbb{P}(X \le x) = 1/2 + 1/4 = 3/4$.
- When $3 \le x < 4$, we have that $F_X(x) = \mathbb{P}(X \le x) = 3/4 + 1/8 = 7/8$.
- When $4 \le x < 5$, we have that $F_X(x) = \mathbb{P}(X \le x) = 7/8 + 1/16 = 15/16$.
- When $x \ge 5$, we have that $F_X(x) = \mathbb{P}(X \le x) = 15/16 + 1/16 = 1$.

$$\mathrm{Therefore,}\ F_X(x) = \begin{cases} 0 & \mathrm{if}\ x < 1, \\ 1/2 & \mathrm{if}\ 1 \leq x < 2, \\ 3/4 & \mathrm{if}\ 2 \leq x < 3, \\ 7/8 & \mathrm{if}\ 3 \leq x < 4, \\ 15/16 & \mathrm{if}\ 4 \leq x < 5, \\ 1 & \mathrm{if}\ x \geq 5. \end{cases} \ \square$$

4. (2 points) Let X be a random variable defined on the probability space (Ω, \mathbb{P}) . Suppose X satisfies the following

$$\mathbb{P}(\{\omega \in \Omega : X(\omega) = 0\}) = 1.$$

Please derive the CDF of the random variable X. Please present your answer using a formula like the one in Eq. (1).

Proof. We have that $\mathbb{P}(X=0)=1=\mathbb{P}(\Omega) \implies \{\omega \in \Omega \mid X(\omega)=0\}=\Omega$. Thus,

- When x < 0, we have that $F_X(x) = \mathbb{P}(X \le x) = \mathbb{P}(\emptyset) = 0$.
- When $x \ge 0$, we have that $F_X(x) = \mathbb{P}(X \le x) = \mathbb{P}(\Omega) = 1$.

Therefore,
$$F_X(x) = \begin{cases} 0 & \text{if } x < 0, \\ 1 & \text{if } x \ge 1. \end{cases}$$

5. (2 points) Let X be a random variable defined on the probability space (Ω, \mathbb{P}) . Suppose the CDF of X is the following

$$F_X(x) = \begin{cases} 0, & \text{if } x < 1; \\ \log x, & \text{if } 1 \le x < e; \\ 1, & \text{if } e \le x. \end{cases}$$

Please compute the values of the following:

(a) $\mathbb{P}(\{\omega \in \Omega : X(\omega) < 2\});$

Proof.

$$\begin{split} \mathbb{P}(X < 2) &= \mathbb{P}(X \le 2) - \mathbb{P}(X = 2) & \text{(additivity)} \\ &= \mathbb{P}(X \le 2) & \text{(}\mathbb{P}(X = 2) = 0\text{)} \\ &= F_X(2) & \text{(def of CDF)} \\ &= \log(2) \end{split}$$

(b) $\mathbb{P}(\{\omega \in \Omega : 0 < X(\omega) < 3\})$;

Proof.

$$\begin{split} \mathbb{P}(0 < X \leq 3) &= \mathbb{P}(X \leq 3) - \mathbb{P}(X \leq 0) & (\mathrm{additivity}) \\ &= F_X(3) - F_X(0) & (\mathrm{def \ of \ } CDF) \\ &= 1 - 0 & (\mathrm{def \ of \ } F_X(x)) \\ &= 1 & (\mathrm{subtraction}) \end{split}$$

(c) $\mathbb{P}(\{\omega \in \Omega : 2 < X(\omega) < 2.5\})$.

Proof.

$$\begin{split} \mathbb{P}(2 < X < 2.5) &= \mathbb{P}(X \le 2.5) - \mathbb{P}(X = 2.5) - \mathbb{P}(X \le 2) \\ &= \mathbb{P}(X \le 2.5) - \mathbb{P}(X \le 2) \\ &= F_X(2.5) - F_X(2) \\ &= \log(2.5) - \log(2) \\ &= \log(1.25) \end{split} \qquad \text{(additivity)}$$

Remark: For simplicity, many textbooks suppress the ω and represent $\mathbb{P}(\{\omega \in \Omega : X(\omega) < 2\})$, $\mathbb{P}(\{\omega \in \Omega : 0 < X(\omega) \leq 3\})$, and $\mathbb{P}(\{\omega \in \Omega : 2 < X(\omega) < 2.5\})$ as $\mathbb{P}(X < 2)$, $\mathbb{P}(0 < X < 3)$, and $\mathbb{P}(2 < X < 2.5)$, respectively. When you read those textbooks, this remark helps you understand what they mean.