Uncovering Subjective Models from Survey Expectations

Chenyu (Sev) Hou [1] Tao Wang [2]

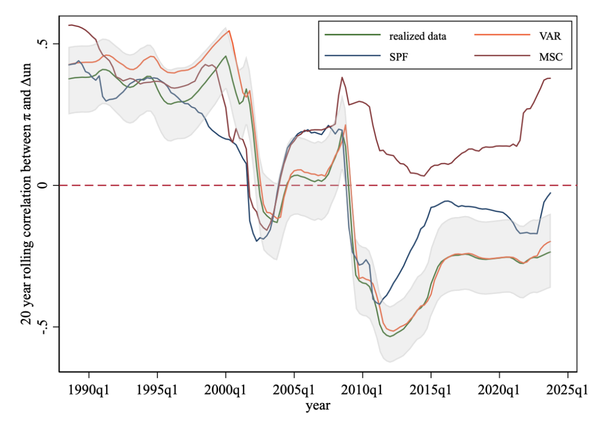
August 21, 2025

ESWC 2025, Seoul

[1]Simon Fraser University

[2]Bank of Canada

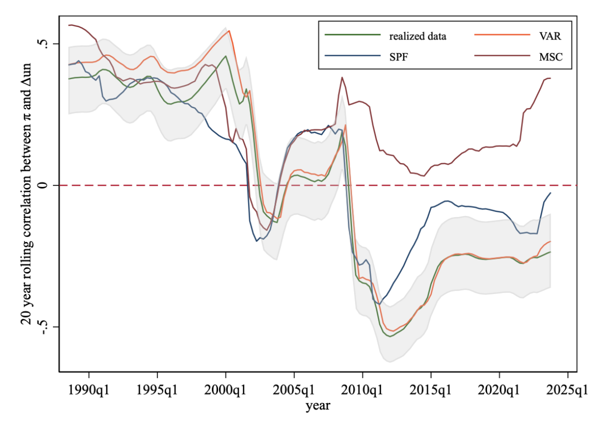
π_t and ΔU_t: Actual Data v.s. Expected



Correlation using 20-year rolling window with 90% CI, 1969-2023. Data from FRED, SPF, and Michigan Survey of Consumers (MSC).

- Households perceive $corr(\mathbb{E}\pi, \mathbb{E}u) > 0$, different from data and professionals (Bhandari et al. (2025) and Candia et al. (2020))

π_t and ΔU_t: Actual Data v.s. Expected



Correlation using 20-year rolling window with 90% CI, 1969-2023. Data from FRED, SPF, and Michigan Survey of Consumers (MSC).

- Households perceive $corr(\mathbb{E}\pi, \mathbb{E}u) > 0$, different from data and professionals (Bhandari et al. (2025) and Candia et al. (2020))
- A robust pattern in survey expectations
  - All the time. Cross-section
  - Across all groups. Group
  - Not due to individual or time fixed effects. Panel
- Growing literature with many explanations: observational equivalence in data.

This paper

We use survey expectations to differentiate two types of models that generate cross-variable expectation correlations: incomplete information versus subjective model

- Methodology
  1. Test survey data patterns against model predictions under the two assumptions
     - a joint sign restriction on $\mathbb{E}$ correlation and between-variable serial correlation in forecast errors
     - $\rightarrow$ incomplete information cannot solely lead to a positive correlation of π and u $\mathbb{E}$s
     - $\rightarrow$ subjective model is necessary to generate such expectation patterns

This paper

We use survey expectations to differentiate two types of models that generate cross-variable expectation correlations: incomplete information versus subjective model

- Methodology
  1. Test survey data patterns against model predictions under the two assumptions
     - a joint sign restriction on $\mathbb{E}$ correlation and between-variable serial correlation in forecast errors
     - $\rightarrow$ incomplete information cannot solely lead to a positive correlation of π and u $\mathbb{E}$s
     - $\rightarrow$ subjective model is necessary to generate such expectation patterns
  2. Structurally estimate an expectation-realization VAR to uncover the subjective model
     - $\rightarrow$ reject the null hypothesis of the subjective model $=$ the objective model

This paper

We use survey expectations to differentiate two types of models that generate cross-variable expectation correlations: incomplete information versus subjective model

- Methodology
  1. Test survey data patterns against model predictions under the two assumptions
     - a joint sign restriction on $\mathbb{E}$ correlation and between-variable serial correlation in forecast errors
     - $\rightarrow$ incomplete information cannot solely lead to a positive correlation of π and u $\mathbb{E}$s
     - $\rightarrow$ subjective model is necessary to generate such expectation patterns
  2. Structurally estimate an expectation-realization VAR to uncover the subjective model
     - $\rightarrow$ reject the null hypothesis of the subjective model $=$ the objective model
- Additional Findings
  - Asymmetry: the perceived correlation goes from π to u:
     - Overpredicted π in $t-1$ $\rightarrow$ overpredicted un in t. Not the opposite
     - Supporting evidence from self-reported news exposure and newspaper narratives

This paper

We use survey expectations to differentiate two types of models that generate cross-variable expectation correlations: incomplete information versus subjective model

- Methodology
  1. Test survey data patterns against model predictions under the two assumptions
     - a joint sign restriction on $\mathbb{E}$ correlation and between-variable serial correlation in forecast errors
     - $\rightarrow$ incomplete information cannot solely lead to a positive correlation of π and u $\mathbb{E}$s
     - $\rightarrow$ subjective model is necessary to generate such expectation patterns
  2. Structurally estimate an expectation-realization VAR to uncover the subjective model
     - $\rightarrow$ reject the null hypothesis of the subjective model $=$ the objective model
- Additional Findings
  - Asymmetry: the perceived correlation goes from π to u:
    - Overpredicted π in $t-1$ $\rightarrow$ overpredicted un in t. Not the opposite
    - Supporting evidence from self-reported news exposure and newspaper narratives
- Macro implication
  - An amplification (dampening) of supply (demand) shock responses in a textbook NK model

A Formal Test of Joint Learning

Modeling framework

- Building on the test of information rigidity in expectation formation (Coibion and Gorodnichenko, 2012; Andrade and Le Bihan, 2013)
  - A Noisy Information model with linear Gaussian noises (Lucas (1976); Woodford (2001))
  - Can be a result of Rational Inattention (Sims, 2003; Maćkowiak et al., 2018)
- Our extensions
  - Univariate→multivariate expectation formation ("Joint learning")
  - Perceived law of motion $\neq$ actual law of motion (" Subjective model")
- Focus on bi-variate case: π and u.

Multivariate environment

Actual data is generated by:

$$\boldsymbol{L}_{t+1,t} = \underbrace{A}_{\text{Actual law of motion}} \boldsymbol{L}_{t,t-1} + w_{t+1,t}$$

$$w_{t+1,t} \sim N(0, \underbrace{Q}_{\text{var-cov of innovations to the state}})$$

Multivariate environment + noisy information

$$\boldsymbol{L}_{t+1,t} = \underbrace{A}_{\text{Actual law of motion}} \boldsymbol{L}_{t,t-1} + w_{t+1,t}$$

$$w_{t+1,t} \sim N(0, \underbrace{Q}_{\text{var-cov of innovations to the state}})$$

Agents observe a noisy signal

$$\boldsymbol{s}_t = \underbrace{G}_{\text{Signal mixture}} \boldsymbol{L}_{t,t-1} + \eta_t$$

$$\eta_t \sim N(0, \underbrace{R}_{\text{var-cov of noise on signal}})$$

Multivariate environment + noisy information + subjective model

$$\boldsymbol{L}_{t+1,t} = \underbrace{A}_{\text{Actual law of motion}} \boldsymbol{L}_{t,t-1} + w_{t+1,t} \tag{1}$$

$$w_{t+1,t} \sim N(0, \underbrace{Q}_{\text{var-cov of innovations to the state}}) \tag{2}$$

$$\overbrace{\boldsymbol{s}_t}^{\text{Noisy signal}} = \underbrace{G}_{\text{Signal mixture}} \boldsymbol{L}_{t,t-1} + \eta_t \tag{3}$$

$$\eta_t \sim N(0, \underbrace{R}_{\text{var-cov of noise on signal}}) \tag{4}$$

Agent may also subjectively believe that:

$$\boldsymbol{L}_{t+1,t} = \underbrace{\hat{A}}_{\text{Perceived law of motion}} \boldsymbol{L}_{t,t-1} + w_{t+1,t} \tag{5}$$

Expectations dynamics

Given the noisy signal and subjective model, expectation about $\boldsymbol{L}_{t+1,t}$ is

$$\underbrace{\boldsymbol{L}_{t+1,t|t}}_{\equiv \mathbb{E}_t(L_{t+1,t})} = \underbrace{\hat{A}}_{\text{Perceived law of motion}} \left((I - KG)\boldsymbol{L}_{t,t-1|t-1} + \underbrace{K}_{\text{Kalman gain}} \boldsymbol{s}_t \right)$$

Under **FIRE**:

- $\hat{A} = A$, $R = \mathbf{0}$, $KG = I$.
- $\rightarrow$ Expectation
$$\boldsymbol{L}^{FIRE}_{t+1,t|t} = A\boldsymbol{L}_{t,t-1}$$
- $\rightarrow$ Correlation between $\boldsymbol{L}^{FIRE}_{t+1,t|t}$ is the same as the correlation of $\boldsymbol{L}_{t,t-1}$. FIRE

Different causes of $corr(\mathbb{E}\pi, \mathbb{E}u) > 0$

$$\underbrace{\boldsymbol{L}_{t+1,t|t}}_{\text{Expectation at t regarding t+1}} = \hat{A}(\boldsymbol{L}_{t,t-1|t-1} + K(\boldsymbol{s}_t - G\boldsymbol{L}_{t,t-1|t-1}))$$

$$\boldsymbol{s}_t = G\boldsymbol{L}_{t,t-1} + \epsilon_t$$

The positive correlation between elements in $\boldsymbol{L}_{t+1,t|t}$ stems from various possibilities:

- Information friction:
  - Non-diagonal G: correlated signals (e.g., Kamdar (2019) through R.I.).
  - Non-diagonal R: correlated noises (e.g., pessimistic heuristics and sentiment).

Different causes of $corr(\mathbb{E}\pi, \mathbb{E}u) > 0$

$$\underbrace{\boldsymbol{L}_{t+1,t|t}}_{\text{Expectation at t regarding t+1}} = \hat{A}(\boldsymbol{L}_{t,t-1|t-1} + K(\boldsymbol{s}_t - G\boldsymbol{L}_{t,t-1|t-1}))$$

$$\boldsymbol{s}_t = G\boldsymbol{L}_{t,t-1} + \epsilon_t$$

The positive correlation between elements in $\boldsymbol{L}_{t+1,t|t}$ stems from various possibilities:

- Information friction:
  - Non-diagonal G: correlated signals (e.g., Kamdar (2019) through R.I.).
  - Non-diagonal R: correlated noises (e.g., pessimistic heuristics and sentiment).
- Subjective model:
  - Non-diagonal $\hat{A}$: $\hat{A}$ has positive off-diagonal elements.

Different causes of $corr(\mathbb{E}\pi, \mathbb{E}u) > 0$

$$\underbrace{\boldsymbol{L}_{t+1,t|t}}_{\text{Expectation at t regarding t+1}} = \hat{A}(\boldsymbol{L}_{t,t-1|t-1} + K(\boldsymbol{s}_t - G\boldsymbol{L}_{t,t-1|t-1}))$$

$$\boldsymbol{s}_t = G\boldsymbol{L}_{t,t-1} + \epsilon_t$$

The positive correlation between elements in $\boldsymbol{L}_{t+1,t|t}$ stems from various possibilities:

- Information friction:
  - Non-diagonal G: correlated signals (e.g., Kamdar (2019) through R.I.).
  - Non-diagonal R: correlated noises (e.g., pessimistic heuristics and sentiment).
- Subjective model:
  - Non-diagonal $\hat{A}$: $\hat{A}$ has positive off-diagonal elements.
- These assumptions are observational equivalent in terms of expectation correlation

Serial correlations of forecast errors (FE)

Resolution: Forecast error tests (Coibion and Gorodnichenko, 2012) extended to multi-variate case:

$$
\begin{aligned}
FE_{t+1,t|t} &\equiv \boldsymbol{L}_{t+1,t} - \boldsymbol{L}_{t+1,t|t} \\
&= \hat{A}(I - KG)FE_{t,t-1|t-1} + \underbrace{M}_{(A-\hat{A}KG-\hat{A}(I-KG))} \boldsymbol{L}_{t,t-1} + w_{t+1,t} - \hat{A}K\epsilon_t
\end{aligned}
$$

Serial correlations of forecast errors (FE)

Resolution: Forecast error tests (Coibion and Gorodnichenko, 2012) extended to multi-variate case:

$$
\begin{aligned}
FE_{t+1,t|t} &\equiv \boldsymbol{L}_{t+1,t} - \boldsymbol{L}_{t+1,t|t} \\
&= \hat{A}(I - KG)FE_{t,t-1|t-1} + \underbrace{M}_{(A - \hat{A}KG - \hat{A}(I-KG))} \boldsymbol{L}_{t,t-1} + w_{t+1,t} - \hat{A}K\epsilon_t
\end{aligned}
$$

- Diagonal terms of $\hat{A}(I - KG)$: auto-correlation (Coibion and Gorodnichenko, 2012)
    - Non-zero diagonals indicate information rigidity

Serial correlations of forecast errors (FE)

Resolution: Forecast error tests (Coibion and Gorodnichenko, 2012) extended to multi-variate case:

$$
\begin{aligned}
FE_{t+1,t|t} &\equiv \mathbf{L}_{t+1,t} - \mathbf{L}_{t+1,t|t} \\
&= \hat{A}(I - KG)FE_{t,t-1|t-1} + \underbrace{M}_{(A-\hat{A}KG-\hat{A}(I-KG))} \mathbf{L}_{t,t-1} + w_{t+1,t} - \hat{A}K\epsilon_t
\end{aligned}
$$

- Diagonal terms of $\hat{A}(I - KG)$: auto-correlation (Coibion and Gorodnichenko, 2012)
    - Non-zero diagonals indicate information rigidity
- Off-diagonal terms: between-correlation (Our focus)
    - their signs depend on off-diagonals of $\hat{A}$ and G

Serial correlations of forecast errors (FE)

Resolution: Forecast error tests (Coibion and Gorodnichenko, 2012) extended to multi-variate case:

$$
\begin{aligned}
FE_{t+1,t|t} &\equiv \boldsymbol{L}_{t+1,t} - \boldsymbol{L}_{t+1,t|t} \\
&= \hat{A}(I-KG)FE_{t,t-1|t-1} + \underbrace{M}_{(A-\hat{A}KG-\hat{A}(I-KG))} \boldsymbol{L}_{t,t-1} + w_{t+1,t} - \hat{A}K\epsilon_t
\end{aligned}
$$

- Diagonal terms of $\hat{A}(I-KG)$: auto-correlation (Coibion and Gorodnichenko, 2012)
    - Non-zero diagonals indicate information rigidity
- Off-diagonal terms: between-correlation (Our focus)
    - their signs depend on off-diagonals of $\hat{A}$ and G
- Special case of FIRE: $A=\hat{A}$ and $KG=I \rightarrow \hat{A}(I-KG)=\mathbf{0}$

Scenario 1: correlated signals, i.e. G is non-diagonal

$$\hat{A}(I - KG) = \begin{pmatrix} \rho_1 & 0 \\ 0 & \rho_2 \end{pmatrix} \begin{pmatrix} \frac{g_2^2\sigma_2^2+\sigma_s^2}{m} & -\frac{g_1 g_2 \sigma_1^2}{m} \\ -\frac{g_1 g_2 \sigma_2^2}{m} & \frac{g_1^2\sigma_1^2+\sigma_s^2}{m} \end{pmatrix}$$
$$= \begin{pmatrix} \rho_1 \frac{g_2^2\sigma_2^2+\sigma_s^2}{m} & -\rho_1 \frac{g_1 g_2 \sigma_1^2}{m} \\ -\rho_2 \frac{g_1 g_2 \sigma_2^2}{m} & \rho_2 \frac{g_1^2\sigma_1^2+\sigma_s^2}{m} \end{pmatrix} \tag{6}$$

- $m = g_1^2\sigma_1^2 + g_2^2\sigma_2^2 + \sigma_s^2$
- $G = [g_1, g_2]$: the vector of signals (due to "optimal signal selection")
- When signals go in the same direction, $g_1 g_2 > 0$, the cross terms are negative.

Scenario 2: subjective model

$$\hat{A}(I - KG) = \begin{pmatrix} \rho_1 & m_1 \\ m_2 & \rho_2 \end{pmatrix} \times \begin{pmatrix} \frac{\sigma_{1,s}^2}{\sigma_1^2+\sigma_{1,s}^2} & 0 \\ 0 & \frac{\sigma_{2,s}^2}{\sigma_2^2+\sigma_{2,s}^2} \end{pmatrix}$$

$$= \begin{pmatrix} \frac{\sigma_{1,s}^2 \rho_1}{\sigma_1^2+\sigma_{1,s}^2} & \frac{\sigma_{2,s}^2 m_1}{\sigma_2^2+\sigma_{2,s}^2} \\ \frac{\sigma_{1,s}^2 m_2}{\sigma_1^2+\sigma_{1,s}^2} & \frac{\sigma_{2,s}^2 \rho_2}{\sigma_2^2+\sigma_{2,s}^2} \end{pmatrix} \tag{7}$$

- $G = I_2$: no signal correlation (can be any diagonal matrix)
- The signs of cross terms (the between-variable serial correlation of FEs) are the same as the perceived correlation

Joint-learning tests for π and un

$$\begin{pmatrix} fe^{\pi}_{t+1,t|t} \\ fe^{un}_{t+1,t|t} \end{pmatrix} = \boldsymbol{\beta}_0 + \underbrace{\begin{pmatrix} \beta_{11} & \beta_{12} \\ \beta_{21} & \beta_{22} \end{pmatrix}}_{\equiv \hat{A}(I-KG)} \begin{pmatrix} fe^{\pi}_{t,t-1|t-1} \\ fe^{un}_{t,t-1|t-1} \end{pmatrix} + \boldsymbol{\theta} X_{t,t-1} + e_t \tag{8}$$

- β_{12} and β_{21}: between-variable serial correlations of forecast errors
- **Predictions:** if only correlated signals but not subjective model, β_{12} and β_{21} are both negative.
- Two complications:
    - In MSC (and most household's survey), $\mathbb{E}u$ is qualitative. We impute them following Bhandari et al. (2025). Imputation
    - Expectations are year-ahead measures, so we derive a year-ahead version of (8). Year-ahead test

Joint-learning tests with consensus expectations

- Joint learning: **subjective model suggesting** $\pi \to un$.
- Bottom line: **information friction cannot be the only reason**.

Table 1: Aggregate Test on Joint Learning, MSC v.s. SPF

	MSC		SPF	
	1984-2023	1990-2018	1984-2023	1990-2018
	(1)	(2)	(3)	(4)
β_{11}	0.64^{***}	0.65^{***}	0.79^{***}	0.76^{***}
	(0.080)	(0.085)	(0.064)	(0.093)
β_{12}	-0.11	-0.02	0.19	-0.08
	(0.076)	(0.095)	(0.117)	(0.199)
β_{21}	$\mathbf{0.13}^{***}$	$\mathbf{0.21}^{***}$	0.05	0.06
	(0.033)	(0.063)	(0.034)	(0.049)
β_{22}	0.71^{***}	0.50^{***}	0.63^{***}	0.51^{***}
	(0.044)	(0.092)	(0.060)	(0.097)
Observations	152	116	152	116

* The first and third columns are using the full sample 1984-2023; the second and fourth columns are results for the sub-sample 1990-2018. Newey-West standard errors are reported in brackets.

Uncovering the Subjective Model

Uncovering the subjective model using an expectation-realization VAR

- Previously, we used sign restrictions to test whether $\hat{A}$ is diagonal, relying on some additional assumptions
- Here, we directly uncover $\hat{A}$

$$\underbrace{Y_{t+1}}_{\equiv \begin{pmatrix} \boldsymbol{L}_{t,t-1|t-1} \\ \boldsymbol{L}_{t,t-1} \end{pmatrix}} = \underbrace{\begin{pmatrix} \hat{A}(I-KG) & \hat{A}KG \\ \mathbf{0}_{2\times 2} & A \end{pmatrix}}_{:=\Phi} \cdot Y_t + \underbrace{\begin{pmatrix} \hat{A}K & \mathbf{0}_{2\times 2} \\ \mathbf{0}_{2\times 2} & I_{2\times 2} \end{pmatrix}}_{F} \cdot \begin{pmatrix} \eta_t \\ w_{t+1,t} \end{pmatrix} \tag{9}$$

- Φ estimated from an (un)restricted VAR
- $\hat{A} = \Phi_{1,1} + \Phi_{1,2}$ and $\Phi_{2,2} = A$
- Identification does not rely on K and G
- Then test $\hat{A} = A$ elementwise

The uncovered subjective model of households and professionals

Table 2: Estimates of Joint Learning Model (9)

	MSC, quarterly, Q1 1984 - Q4 2023		SPF, quarterly, Q1 1984 - Q4 2023	
Parameters	Estimates	Standard Errors	Estimates	Standard Errors
A	$\begin{bmatrix} 0.836 & -0.058 \\ 0.034 & 0.617 \end{bmatrix}$	$\begin{bmatrix} 0.053 & 0.057 \\ 0.042 & 0.095 \end{bmatrix}$	$\begin{bmatrix} 0.837 & -0.056 \\ 0.014 & 0.751 \end{bmatrix}$	$\begin{bmatrix} 0.061 & 0.074 \\ 0.041 & 0.093 \end{bmatrix}$
$\hat{A}$	$\begin{bmatrix} 0.741 & -0.149 \\ 0.137 & 0.831 \end{bmatrix}$	$\begin{bmatrix} 0.050 & 0.082 \\ 0.044 & 0.048 \end{bmatrix}$	$\begin{bmatrix} 0.955 & -0.038 \\ 0.040 & 0.495 \end{bmatrix}$	$\begin{bmatrix} 0.019 & 0.016 \\ 0.035 & 0.239 \end{bmatrix}$
T-test:	test-stat	p-val	test-stat	p-val
$\hat{A}_{21} > A_{21}$	1.581	0.057	0.546	0.293

The table reports the estimates and their NW standard errors from the GMM estimation of the 4-variable VAR model. Iterative weighting matrix are used in the GMM estimation.

Alternative sample With feedback loop

Additional Evidence

Expectations conditional on the type of news heard

Table 3: Panel Regression with Self-reported News

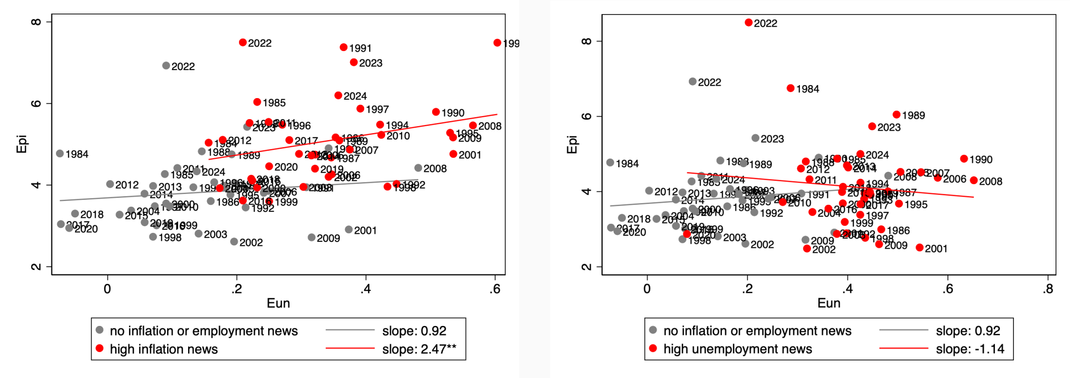
Expectation on:	$E\pi$	Eun	$E\pi$	Eun
news on:	(1)	(2)	(3)	(4)
Inflation fav	-0.21*	-0.06***	-0.21*	-0.05***
	(0.117)	(0.017)	(0.118)	(0.017)
Inflation unfav	0.43***	0.06***	0.42***	0.05***
	(0.085)	(0.010)	(0.085)	(0.010)
Employment fav	-0.03	-0.14***	-0.01	-0.13***
	(0.056)	(0.009)	(0.057)	(0.009)
Employment unfav	0.05	0.10***	0.04	0.09***
	(0.054)	(0.007)	(0.054)	(0.007)
Interest rate fav	-0.03	-0.06***	-0.01	-0.04***
	(0.071)	(0.012)	(0.072)	(0.012)
Interest rate unfav	0.02	0.11***	0.02	0.10***
Observations	169304	189158	169304	189158
R^2	0.673	0.677	0.673	0.681
Time F.E.	Y	Y	Y	Y
Individual F.E.	Y	Y	Y	Y
Full set news control	N	N	Y	Y

Consensus expectations conditional on the news exposure

Figure 1: Consensus expectations conditional on news heard

(a) With unfavorable inflation news

(b) With unfavorable employment news



Notes: Scatter plot for consensus expected inflation and unemployment each year from 1984-2023. Gray dots in all panels are expectations for individuals without employment or inflation news. Left panel: red dots are expectations conditional on hearing high inflation news. Right panel: red dots are expectations conditional on hearing high unemployment news.

Inflation-unemployment associations in newspapers

- $P(\mathbb{I}_{i,t}(\text{Joint mention}) = 1) = \Phi(\beta_0 + \sum_k \beta_k D_{k,I,t} + \beta_\pi \pi_t + \beta_u u_t)$

- Probability of news report making association between inflation and unemployment increases significantly with realized inflation.

	(1)	(2)	(3)
economy	1.07***	1.07***	1.07***
	(0.03)	(0.03)	(0.03)
fed	0.22***	0.21***	0.21***
	(0.03)	(0.03)	(0.03)
growth	0.60***	0.61***	0.61***
	(0.03)	(0.03)	(0.03)
oil price	0.24***	0.24***	0.24***
	(0.05)	(0.05)	(0.05)
recession	0.48***	0.47***	0.47***
	(0.03)	(0.03)	(0.03)
uncertainty	0.14***	0.15***	0.15***
	(0.05)	(0.05)	(0.05)
π_t		**3.73*****	**3.62*****
		(0.93)	(0.96)
u_t	-0.01		-0.00
	(0.01)		(0.01)
N	150465	150465	150465

Macro Implications of Subjective Model

Shock propagation in a textbook NK model

3-Equation NK

$$
\begin{aligned}
\pi_t &= \beta \mathbb{E}_t \pi_{t+1} + \kappa y_t + s_t \\
y_t &= \mathbb{E}_t y_{t+1} - \frac{1}{\sigma}(i_t - \mathbb{E}_t \pi_{t+1} - \rho) + d_t \\
i_t &= \rho + \phi_\pi \pi_t + \phi_y y_t
\end{aligned}
$$

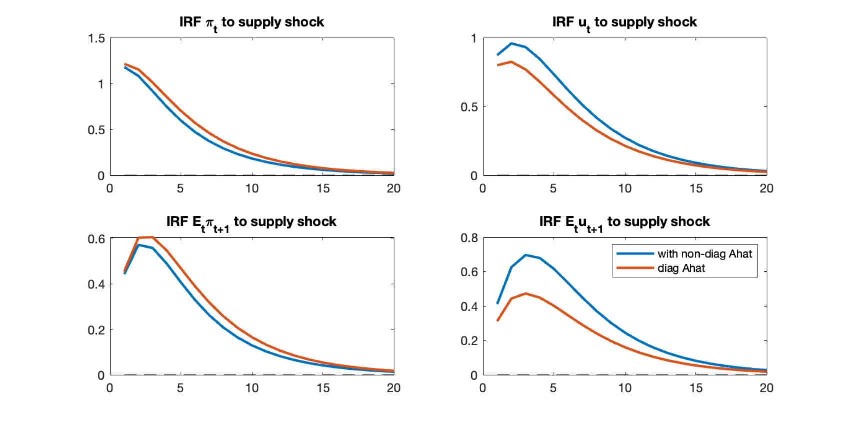
Okun's Law

$$u_t = -\chi y_t$$

Expectation formation

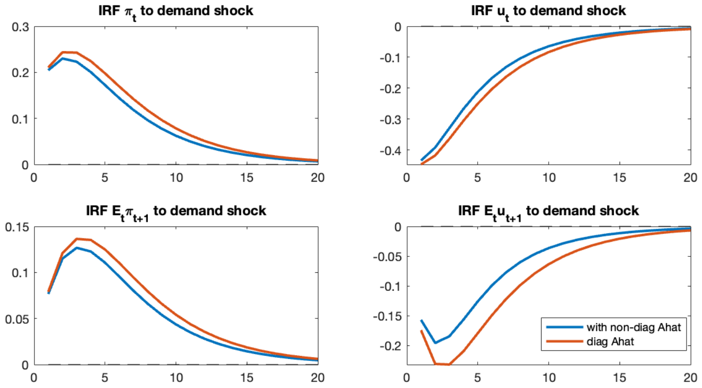
$$
\begin{aligned}
L_{t+1,t|t} &= \hat{A}(I - KG)L_{t,t-1|t-1} + \hat{A}KGL_{t,t-1} + \hat{A}K\eta_t \\
L_{t+1,t|t} &\equiv \begin{pmatrix} \mathbb{E}_t \pi_{t+1} \\ \mathbb{E}_t u_{t+1} \end{pmatrix} \\
L_{t,t-1} &\equiv \begin{pmatrix} \pi_t \\ u_t \end{pmatrix}
\end{aligned}
$$

Figure 2: IRF in Response to Supply Shock



Demand shock

Figure 3: IRF in Response to Demand Shock



Conclusion

- Households think about macroeconomic variables jointly
- $\mathbb{E}(\pi)\uparrow\ \rightarrow \mathbb{E}(un)\uparrow$
- Formal tests + structural estimation suggest:
  - HH's subjective model differs from objective one.
  - Correlated expectation is not only due to information friction.
- Asymmetric reaction to inflation and unemployment (real activity) news.
  - π news triggers associations of π and *un* in expectations
  - ... as well as newspapers' narratives
- Implications for monetary policy
  - Caution on expectation management policy – may have unintended contractionary effects.
  - Dampened response to demand shocks and amplified responses to supply shocks.

Appendix

Cross-correlation: MSC

Table 4: Correlation MCS: more variables

	(1)	(2)	(3)	(4)	(5)
(1) inflation ($E\pi_{t+4,t}$)	1.00	0.31***	−0.13	−0.43***	−0.51***
(2) unemp change ($E\Delta un_{t+4,t}$)		1.00	−0.41***	−0.64***	−0.28***
(3) interest rate change ($E\Delta i_{t+4,t}$)			1.00	0.40***	0.07*
(4) Busi Condition change ($E\Delta y_{t+4,t}$)				1.00	0.77***
(5) real income change ($E\Delta w_{t+4,t}$)					1.00

* **** means significant at 1%,** means 5 % and * means 10%, data in use are quarterly 1978q1-2018q4 from MSC.*

Cross-correlation: FRED

Table 5: Correlation FRED: more variables

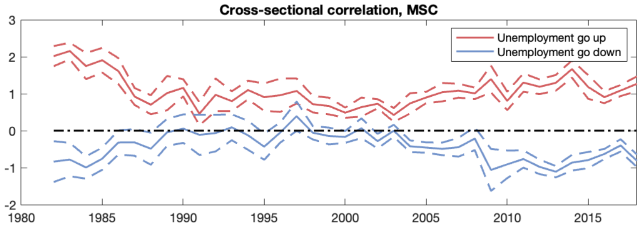
	(1)	(2)	(3)	(4)	(5)
(1) CPI	1.00	0.11	0.38^{***}	-0.03	-0.32^{***}
(2) Δun		1.00	-0.52^{***}	-0.79^{***}	-0.77^{***}
(3) ΔFFR			1.00	0.43^{***}	0.26^{***}
(4) $\Delta RGDP$				1.00	0.79^{***}
(5) Δw					1.00

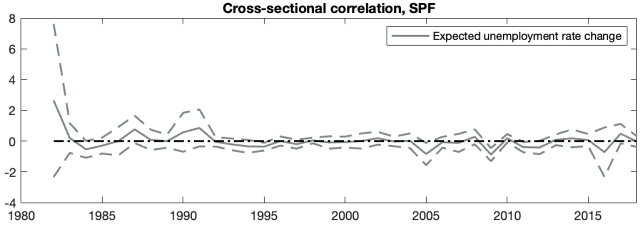
* **** means significant at 1%,** means 5 % and * means 10%, data in use are quarterly 1978q1-2018q4 from FRED.*

Back

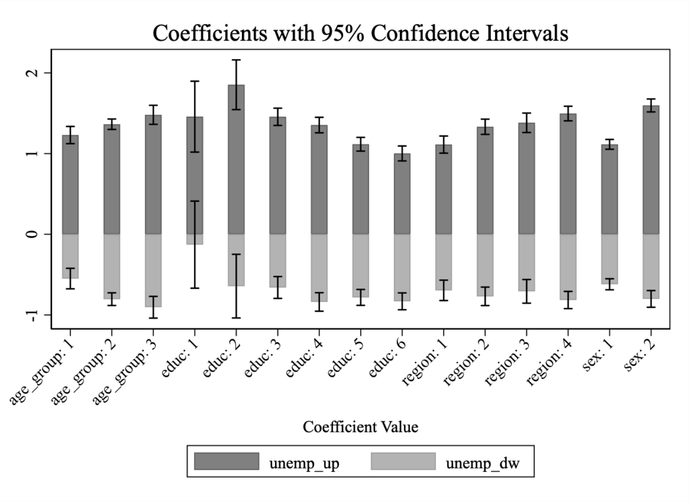
Time variations of the perceived correlation in consensus expectations

Estimate $E_{i,t}\pi_{t+12,t} = \beta_0 + \beta_1 E_{i,t} un_{t+12,t} + \theta X_{i,t} + \epsilon_{i,t}$. Parameter of interest is β_1: Back





Regression by group



Cross-sectional correlation across groups

Back

Controlling for individual FE and time FE

$$E_{i,t}\pi_{t+12,t} = \beta_0 + \beta_1 E_{i,t} un_{t+12,t} + \beta_2 E_{i,t} i_{t+12,t} + \theta X_{i,t} + D_t + \mu_i + \epsilon_{i,t}$$

Table 6: FE Panel Regression

	MSC		SCE		SPF
Unemployment up	0.30^{***}	$\hat{\beta}_1$	0.012^{***}	$\hat{\beta}_1$	-0.17^{***}
	(0.05)		(0.002)		(0.06)
Unemployment down	-0.22^{***}				
	(0.05)				
FE	Y		Y		Y
Time dummy	Y		Y		Y

* Controlling for individual and time-varying characteristics, individual fixed effect, and time-fixed effect. Standard errors are adjusted for heteroscedasticity and autocorrelation.

Back

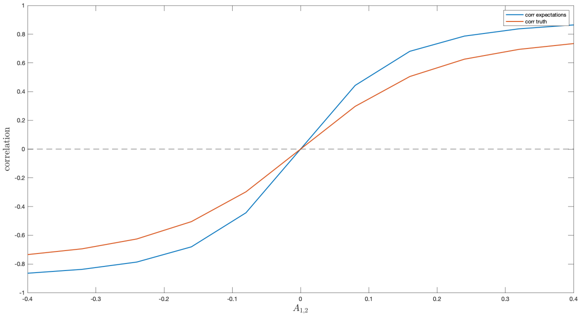
FIRE

Correlation of Expectations under FIRE:

- For our case $\boldsymbol{L} = \begin{pmatrix} \pi_t \\ un_t \end{pmatrix}$ Empirical estimates of A from 1984-2023:

$$\begin{bmatrix} 0.87 & -0.05 \\ (0.05) & (0.06) \\ 0.02 & 0.67 \\ (0.06) & (0.11) \end{bmatrix}$$

- NW s.e. in brackets and BIC select 1 lag of VAR.



Correlation between expected variables and realized variables when $A(1,2)$ changes under FIRE

Joint Learning Test

When it's not FIRE, maintain a simplification restriction:

Assumption 1

The variance-covariance matrix of prior $\mathbf{L}^i_{t,t-1|t-1}$ is diagonal and common to each individual:

$$\Sigma := diag(\{\sigma_j^2\})$$

Consider different scenarios afore-mentioned:

1. When $\hat{A}$ is diagonal, consider different G and R: Independent learning.
2. When $\hat{A}$ non-diagonal: Joint learning.
    - When R and G are diagonal.

Joint Learning Test I

Proposition 1

(Independent Learning) *If* $\hat{A} = diag(\{a_i\}_{i=1}^{n})$*, denote the off-diagonal elements of* $\hat{A}(I - KG)$ *as* β_{ij} *with* $i \neq j$*. We have:*
(1) $\beta_{ij} = 0$ *if* G *and* R *are diagonal.*
(2) $\beta_{ij} = \beta_{ji} = 0$ *or* $\beta_{ij}\beta_{ji} > 0$ *if* G *or* R *is non-diagonal.*

- Out test coefficient (2-d case):

$$\hat{A}(I - KG) = \begin{pmatrix} \beta_{11} & \beta_{12} \\ \beta_{21} & \beta_{22} \end{pmatrix}$$

Back

Joint Learning Test I

Proposition 1

(Independent Learning) *If* $\hat{A} = diag(\{a_i\}_{i=1}^{n})$, *denote the off-diagonal elements of* $\hat{A}(I - KG)$ *as* β_{ij} *with* $i \neq j$. *We have:*
(1) $\beta_{ij} = 0$ *if* G *and* R *are diagonal.*
(2) $\beta_{ij} = \beta_{ji} = 0$ *or* $\beta_{ij}\beta_{ji} > 0$ *if* G *or* R *is non-diagonal.*

- Independent learning + separated signals (Coibion and Gorodnichenko, 2012; Andrade and Le Bihan, 2013): zero between-correlation, non-zero auto-correlation of F.E.

$$\hat{A}(I - KG) = \begin{pmatrix} \beta_{11} \neq 0 & \beta_{12} = 0 \\ \beta_{21} = 0 & \beta_{22} \neq 0 \end{pmatrix}$$

Back

Joint Learning Test I

Proposition 1

(Independent Learning) *If* $\hat{A} = diag(\{a_i\}_{i=1}^n)$*, denote the off-diagonal elements of* $\hat{A}(I - KG)$ *as* β_{ij} *with* $i \neq j$*. We have:*
(1) $\beta_{ij} = 0$ *if* G *and* R *are diagonal.*
(2) $\beta_{ij} = \beta_{ji} = 0$ *or* $\beta_{ij}\beta_{ji} > 0$ *if* G *or* R *is non-diagonal.*

- Independent learning + mixed signals (R.I. like in Kamdar (2019)): same signs on between-correlation of F.E.

$$\hat{A}(I - KG) = \begin{pmatrix} \beta_{11} \neq 0 & \beta_{12} <> 0 \\ \beta_{21} <> 0 & \beta_{22} \neq 0 \end{pmatrix}$$

Back

Joint Learning Test II: non-diagonal R

Corollary 1

(Non-diagonal R: correlated noises) *If $\hat{A}$ and G are diagonal and $R = \begin{pmatrix} \sigma_{1,s}^2 & \rho \\ \rho & \sigma_{2,s}^2 \end{pmatrix}$, the off-diagonal elements of $\hat{A}(I - KG)$ have the same signs as ρ.*

$$\hat{A}(I - KG) = \begin{pmatrix} \beta_{11} \neq 0 & sgn(\beta_{12}) = sgn(\rho) \\ sgn(\beta_{21}) = sgn(\rho) & \beta_{22} \neq 0 \end{pmatrix}$$

Illustration Back

Joint Learning Test II: non-diagonal G

Corollary 2

(Non-diagonal G: correlated signals) *If $\hat{A}$ is diagonal, $R = \begin{pmatrix} \sigma_{1,s}^2 & 0 \\ 0 & \sigma_{2,s}^2 \end{pmatrix}$ is diagonal, and $G = \begin{pmatrix} g_1 & g_2 \\ 0 & g_4 \end{pmatrix}$, the off-diagonal elements of $\hat{A}(I - KG)$ have the opposite signs as $g_1 g_2$.*

$$\hat{A}(I - KG) = \begin{pmatrix} \beta_{11} \neq 0 & sgn(\beta_{12}) = -sgn(g_1 g_2) \\ sgn(\beta_{21}) = -sgn(g_1 g_2) & \beta_{22} \neq 0 \end{pmatrix}$$

Illustration Back

Joint Learning Test III

Proposition 2

(Joint Learning) *If off-diagonal elements of $\hat{A}(I-KG)$ are not both zeros and of different signs, then $\hat{A}$ is non-diagonal, regardless whether G and R are diagonal or not.*

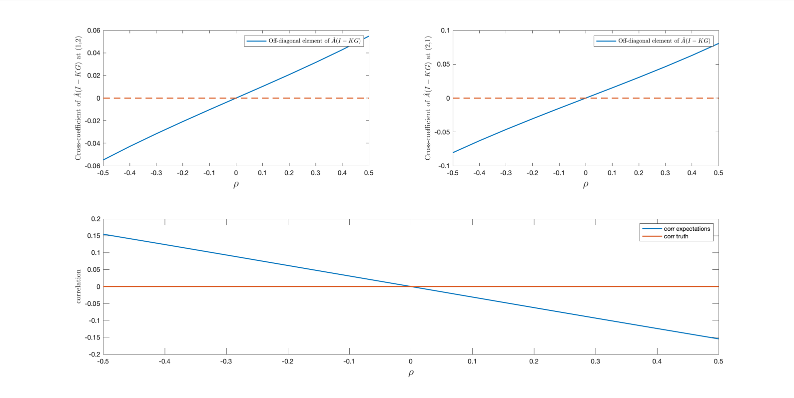
Proposition 3

(Joint Learning with separate signals) *If both G and R are diagonal and $\hat{A} = (a_{ij})_{n\times n}$ is non-diagonal, denote $\hat{A}(I-KG) = (\beta_{ij})_{n\times n}$. The signs of these off-diagonal elements are the same as their counterparts in $\hat{A}$, i.e. $\beta_{ij}a_{ij} > 0$.*

$$\hat{A}(I-KG) = \begin{pmatrix} \beta_{11} \neq 0 & sgn(\beta_{12}) = sgn(a_{12}) \\ sgn(\beta_{21}) = sgn(a_{21}) & \beta_{22} \neq 0 \end{pmatrix}$$

Illustration Back

Example I: non-diagonal R



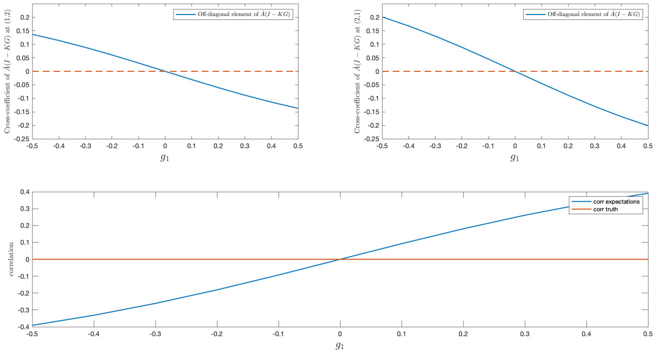
Off-diagonal elements of $\hat{A}(I-KG)$ and correlation between expectations.

- A working example for illustration:

$$A = \hat{A} = \begin{bmatrix} 0.7 & 0 \\ 0 & 0.9 \end{bmatrix}, \quad \Sigma = \begin{bmatrix} 7/4 & 0 \\ 0 & 2 \end{bmatrix}, \quad R = \begin{bmatrix} 1.5 & \rho \\ \rho & 1 \end{bmatrix}, \quad G = \begin{bmatrix} 1 & 0 \\ 0 & 1/3 \end{bmatrix},$$

Back

Example II: non-diagonal G



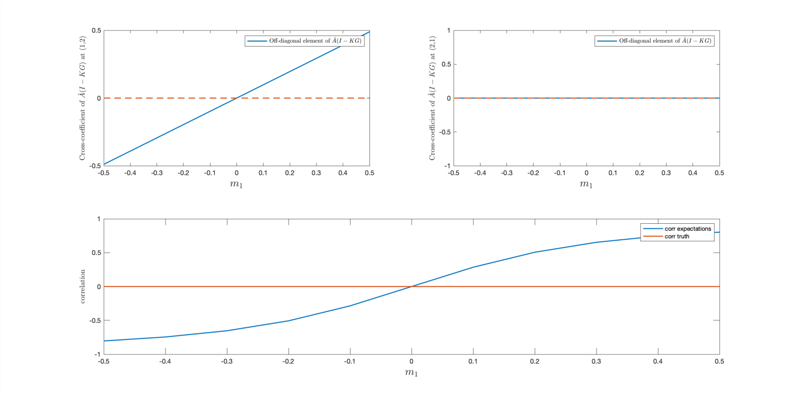
Off-diagonal elements of $\hat{A}(I-KG)$ and correlation between expectations.

- Same working example for illustration:

$$A = \hat{A} = \begin{bmatrix} 0.7 & 0 \\ 0 & 0.9 \end{bmatrix}, \quad \Sigma = \begin{bmatrix} 7/4 & 0 \\ 0 & 2 \end{bmatrix}, \quad R = \begin{bmatrix} 1.5 & 0 \\ 0 & 1 \end{bmatrix}, \quad G = \begin{bmatrix} 1 & g_1 \\ 0 & 1/3 \end{bmatrix},$$

Back

Example III: non-diagonal $\hat{A}$ and diagonal R and G



Off-diagonal elements of $\hat{A}(I-KG)$ and correlation between expectations.

- Same working example as before except:

$$\hat{A} = \begin{bmatrix} 0.7 & m_1 \\ 0 & 0.9 \end{bmatrix}$$

Back

Complication I: impute *Eun*

Assumption 2

At each period t, survey respondent i forms a belief $x_{i,t}$ that indicates the change of asked variable x, this belief follows a normal distribution:

$$x_{i,t} \sim N(\mu_t, \sigma_t^2)$$

The survey respondent will respond in categorical fashion:

$$category_{i,t} = \begin{cases} increase & x_{it} > b + a \\ decrease & x_{it} < b - a \\ same & x_{it} \in [-a + b, b + a] \end{cases}$$

Complication I: impute *Eun*

We want to recover μ_t, we can observe fraction of people responding "increase" (f_t^u) and "decreasing" (f_t^d). From normality:

$$\sigma_t = \frac{2a}{\Phi^{-1}(1 - f_t^u) - \Phi^{-1}(f_t^d)} \tag{10}$$

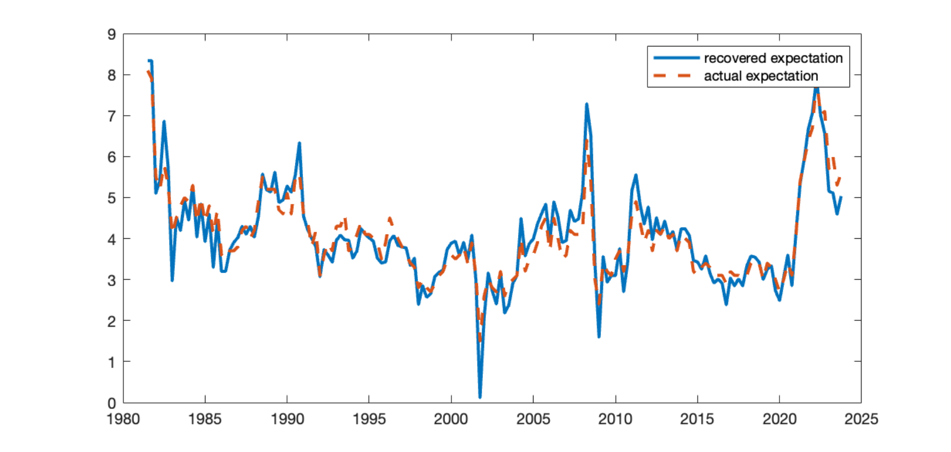
$$\mu_t = a + b - \sigma_t \Phi^{-1}(1 - f_t^u) \tag{11}$$

- Get a and b using average σ_t and μ_t approximated by SPF. (Bhandari et al., 2025)
- Can test with $E\pi_t$ data.

Back

Complication I: impute *Eun*

Figure 4: Recovered Expected Inflation v.s. Actual



Correlation between imputed and actual: 0.99. Back

Complication II: Year ahead Expectation

- The baseline test is derived with quarter-to-quarter changes, whereas data on expectations are year-ahead. We can iterate the forecasting error equation forward:

$$
\begin{aligned}
FE_{t+4,t|t} = \hat{W}\hat{A}(I-KG)\hat{W}^{-1}FE_{t+3,t-1|t-1} + (I-\hat{W}\hat{A}(I-KG)\hat{W}^{-1})\boldsymbol{L}_{t+3,t-1} \\
-(I+\hat{W}\hat{A}KG)\boldsymbol{L}_{t,t-1} + A\boldsymbol{L}_{t+3,t+2} - \hat{W}\hat{A}K\eta_t + w_{t+4,t+3}
\end{aligned} \tag{12}
$$

- With $\hat{W} = \hat{A}^3 + \hat{A}^2 + \hat{A} + I$.
- The test results hold true as the quarterly specification. Shown with Monte Carlo.

Complication II: Year ahead test

Table 13: Simulation Results: FIRE or Independent Learning with Uncorrelated Signals

	FIRE or Independent Learning: $\hat{A} = A$, $g_2 = 0$, $\rho = 0$							
	FIRE				Independent Learning			
	Y-ahead Spec (10)		Q-ahead Spec (6)		Y-ahead Spec (10)		Q-ahead Spec (6)	
	Truth	Test	Truth	Test	Truth	Test	Truth	Test
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
β_{11}	0	-0.01	0	0.04	0.54	0.51^{***}	0.54	0.47^{***}
	-	(0.03)	-	(0.09)	-	(0.09)	-	(0.09)
β_{12}	0	0.03	0	0.15	0	-0.14	0	-0.14
	-	(0.04)	-	(0.11)	-	(0.010)	-	(0.10)
β_{21}	0	0.01	0	0.10	0	-0.03	0	-0.09
	-	(0.02)	-	(0.09)	-	(0.04)	-	(0.11)
β_{22}	0	-0.00	0	0.18	0.43	0.49^{***}	0.43	0.61^{***}
	-	(0.05)	-	(0.12)	-	(0.07)	-	(0.11)

* ***,**,*: Significance at 1%,5% and 10% level. Columns (2) and (6) are estimation results for year-ahead joint-learning test (10), and columns (4) and (8) are for quarter-ahead specification (6). Newey-West standard errors are reported in brackets.

Complication II: Year ahead test

Table 14: Simulation Results: Independent Learning with Correlated Signals

	Independent Learning when G or R are non-diagonal							
	G non-diagonal:				R non-diagonal:			
	$m_1 = 0,\ g_2 = 0.5,\ \rho = 0$				$m_1 = 0,\ g_2 = 0,\ \rho = -2$			
	Y-ahead spec (10)		Q-ahead spec (6)		Y-ahead spec (10)		Q-ahead spec (6)	
	Truth	Test	Truth	Test	Truth	Test	Truth	Test
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
β_{11}	0.57	0.56^{***}	0.57	0.52^{***}	0.49	0.43^{***}	0.49	0.37^{***}
	–	(0.05)	–	(0.08)	–	(0.05)	–	(0.09)
β_{12}	-0.14	-0.28^{***}	-0.10	-0.26^{***}	-0.17	-0.25^{***}	-0.13	-0.24^{***}
	–	(0.09)	–	(0.10)	–	(0.09)	–	(0.09)
β_{21}	-0.07	-0.10^{***}	-0.10	-0.20^{**}	-0.09	-0.11^{***}	-0.12	-0.17
	–	(0.04)	–	(0.10)	–	(0.04)	–	(0.11)
β_{22}	0.40	0.46^{***}	0.40	0.55^{***}	0.39	0.49^{***}	0.39	0.63^{***}
	–	(0.07)	–	(0.11)	–	(0.07)	–	(0.11)

Complication II: Year ahead test

Table 15: Simulation Results: Joint Learning

	Joint Learning: $m_1 = 0.5$, G and R are diagonal			
	Year-ahead spec (10)		Quarter-ahead spec (6)	
	Truth	Test	Truth	Test
	(1)	(2)	(3)	(4)
β_{11}	0.54	0.48***	0.54	0.44***
	-	(0.08)	-	(0.08)
β_{12}	0.32	0.49**	0.31	0.35***
	-	(0.22)	-	(0.10)
β_{21}	0	-0.02	0	-0.08
	-	(0.04)	-	(0.09)
β_{22}	0.43	0.54***	0.43	0.70***
	-	(0.12)	-	(0.14)

* ***,**,*: Significance at 1%,5% and 10% level. Column (2)

Joint Estimation: Alternative Sample

MSC, quarterly				
	Q1 1984 - Q4 2019		**Q1 1990 - Q4 2018**	
Parameters	Estimates	Standard Errors	Estimates	Standard Errors
A	$\begin{bmatrix} 0.807 & -0.070 \\ 0.062 & 0.922 \end{bmatrix}$	$\begin{bmatrix} 0.059 & 0.114 \\ 0.022 & 0.072 \end{bmatrix}$	$\begin{bmatrix} 0.781 & -0.060 \\ 0.059 & 0.930 \end{bmatrix}$	$\begin{bmatrix} 0.068 & 0.145 \\ 0.031 & 0.082 \end{bmatrix}$
$\hat{A}$	$\begin{bmatrix} 0.663 & -0.096 \\ 0.189 & 0.807 \end{bmatrix}$	$\begin{bmatrix} 0.063 & 0.089 \\ 0.057 & 0.056 \end{bmatrix}$	$\begin{bmatrix} 0.663 & -0.081 \\ 0.271 & 0.769 \end{bmatrix}$	$\begin{bmatrix} 0.080 & 0.094 \\ 0.064 & 0.057 \end{bmatrix}$
T-test:	test-stat	p-val	test-stat	p-val
$\hat{A}_{21} > A_{21}$	2.094	0.018	2.999	0.001
SPF, quarterly				
	Q1 1984 - Q4 2019		**Q1 1990 - Q4 2018**	
Parameters	Estimates	Standard Errors	Estimates	Standard Errors
A	$\begin{bmatrix} 0.788 & -0.070 \\ 0.048 & 0.906 \end{bmatrix}$	$\begin{bmatrix} 0.070 & 0.100 \\ 0.024 & 0.071 \end{bmatrix}$	$\begin{bmatrix} 0.749 & -0.047 \\ 0.042 & 0.920 \end{bmatrix}$	$\begin{bmatrix} 0.079 & 0.113 \\ 0.030 & 0.077 \end{bmatrix}$
$\hat{A}$	$\begin{bmatrix} 0.951 & 0.004 \\ 0.026 & 0.787 \end{bmatrix}$	$\begin{bmatrix} 0.018 & 0.041 \\ 0.016 & 0.041 \end{bmatrix}$	$\begin{bmatrix} 0.937 & -0.027 \\ 0.026 & 0.806 \end{bmatrix}$	$\begin{bmatrix} 0.021 & 0.030 \\ 0.031 & 0.044 \end{bmatrix}$
T-Test	test-stat	p-val	test-stat	p-val
$\hat{A}_{21} > A_{21}$	-0.883	0.811	-0.410	0.659

Back

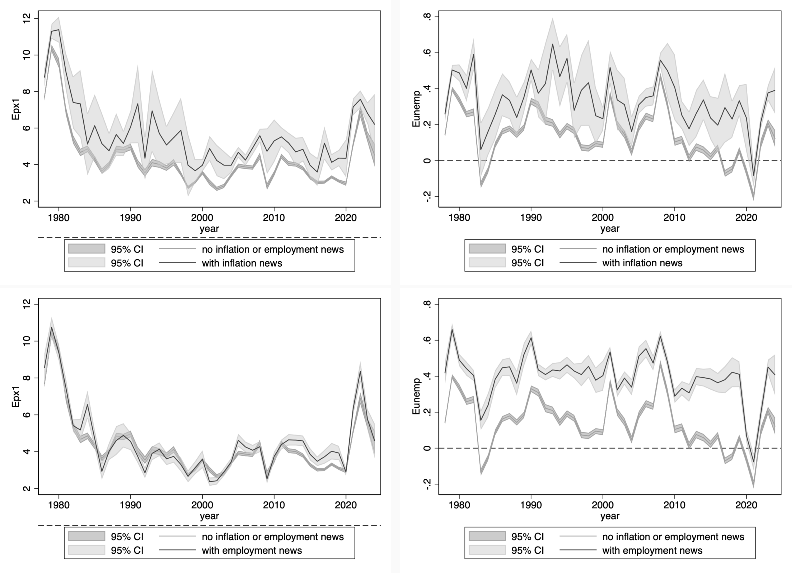
Joint Estimation: with Feedback loop

MSC, quarterly				
	Q1 1984 - Q4 2019		**Q1 1990 - Q4 2018**	
Parameters	Estimates	Standard Errors	Estimates	Standard Errors
A	$\begin{bmatrix} 0.863 & 0.021 \\ -0.003 & 0.751 \end{bmatrix}$	$\begin{bmatrix} 0.073 & 0.162 \\ 0.042 & 0.074 \end{bmatrix}$	$\begin{bmatrix} 0.863 & 0.051 \\ -0.017 & 0.721 \end{bmatrix}$	$\begin{bmatrix} 0.078 & 0.169 \\ 0.042 & 0.076 \end{bmatrix}$
$\hat{A}$	$\begin{bmatrix} 0.663 & -0.096 \\ 0.189 & 0.807 \end{bmatrix}$	$\begin{bmatrix} 0.063 & 0.089 \\ 0.057 & 0.056 \end{bmatrix}$	$\begin{bmatrix} 0.663 & -0.081 \\ 0.271 & 0.769 \end{bmatrix}$	$\begin{bmatrix} 0.080 & 0.094 \\ 0.064 & 0.057 \end{bmatrix}$
T-test:	test-stat	p-val	test-stat	p-val
$\hat{A}_{21} > A_{21}$	2.227	0.013	3.112	0.001

SPF, quarterly				
	Q1 1984 - Q4 2019		**Q1 1990 - Q4 2018**	
Parameters	Estimates	Standard Errors	Estimates	Standard Errors
A	$\begin{bmatrix} 0.696 & -0.091 \\ 0.021 & 0.792 \end{bmatrix}$	$\begin{bmatrix} 0.078 & 0.090 \\ 0.031 & 0.072 \end{bmatrix}$	$\begin{bmatrix} 0.678 & -0.062 \\ 0.019 & 0.785 \end{bmatrix}$	$\begin{bmatrix} 0.086 & 0.107 \\ 0.034 & 0.089 \end{bmatrix}$
$\hat{A}$	$\begin{bmatrix} 0.951 & 0.004 \\ 0.026 & 0.787 \end{bmatrix}$	$\begin{bmatrix} 0.018 & 0.041 \\ 0.016 & 0.041 \end{bmatrix}$	$\begin{bmatrix} 0.937 & -0.027 \\ 0.026 & 0.806 \end{bmatrix}$	$\begin{bmatrix} 0.021 & 0.030 \\ 0.031 & 0.044 \end{bmatrix}$
T-Test	test-stat	p-val	test-stat	p-val
$\hat{A}_{21} > A_{21}$	0.136	0.446	0.1534	0.439

Back

π news drives expectations across domains but *un* news is domain-specific

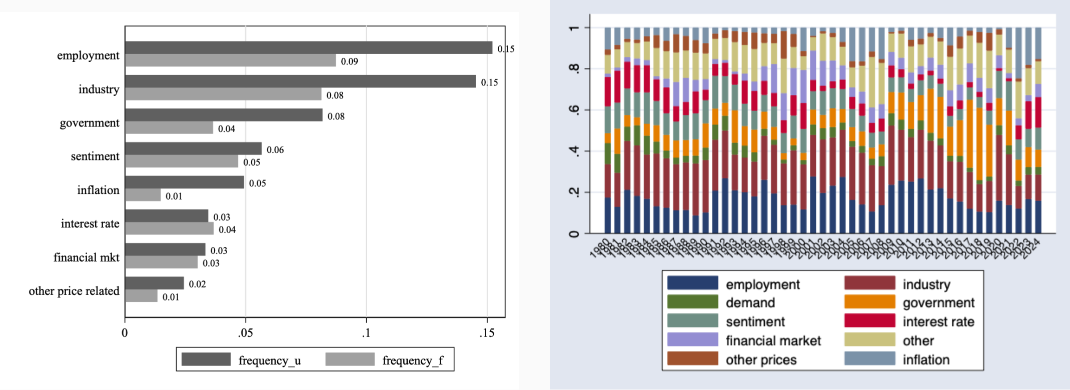


Reported news in MSC

Figure 5: Type of news

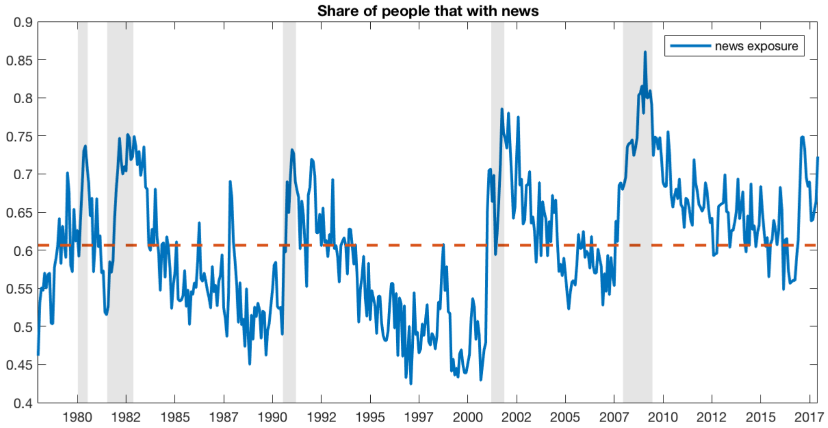
(a) Fraction of fav and unfav news

(b) Type of news across time



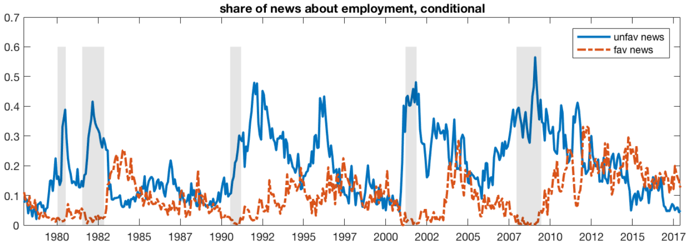
Notes: Panel (a): fractions of favorable and unfavorable news reported by individuals with news in MSC. Panel (b): shares of different types of news out of total news reported each year.

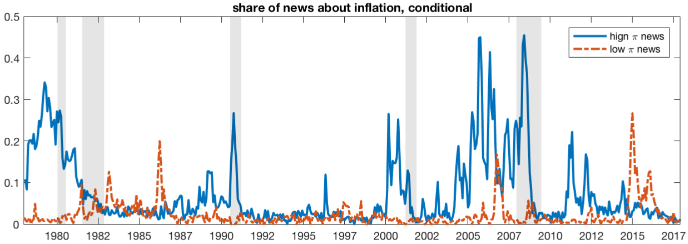
News measure



Share of people that report hearing any news across time. The dashed line represents on average 60% survey participants reported hearing about some news in the past few months.

News measure



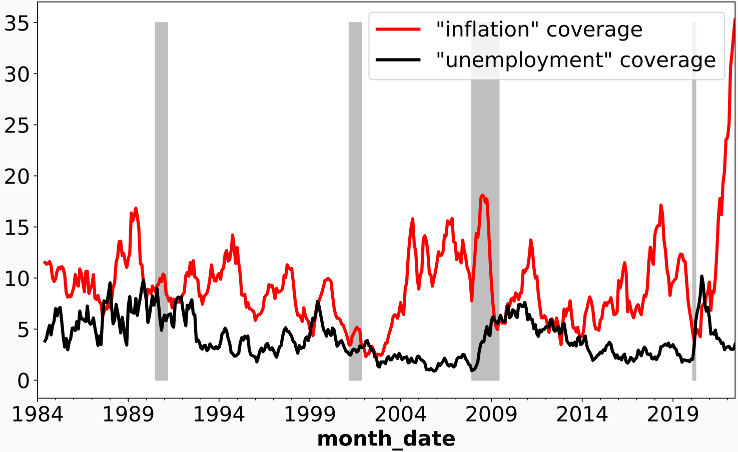


Correlation conditional on news

Table 7: Correlation Conditional on News Heard

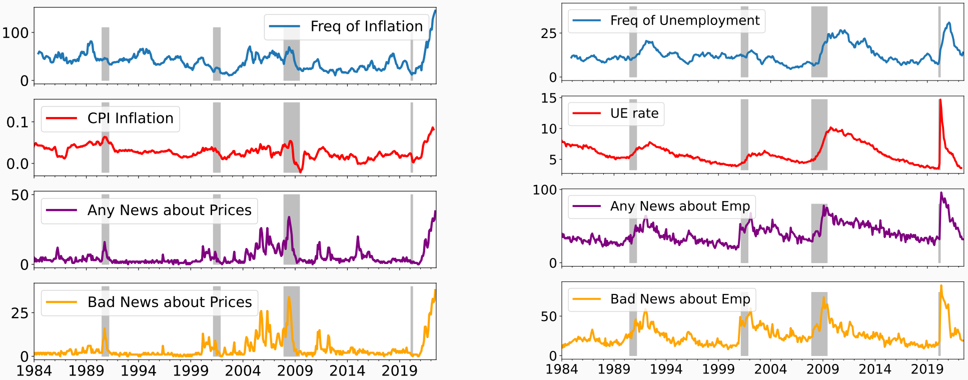
Dependent var:	$E\pi$	
	(1)	(2)
Eun	0.36***	0.38***
	(0.034)	(0.047)
Inflation fav $\times Eun$	0.17	0.16
	(0.164)	(0.164)
Inflation unfav $\times Eun$	0.36***	0.36***
	(0.117)	(0.118)
Employment fav $\times Eun$	0.03	0.03
	(0.089)	(0.090)
Employment unfav $\times Eun$	-0.20***	-0.16**
	(0.073)	(0.074)
Interest rate fav $\times Eun$	-0.23**	-0.24**
	(0.104)	(0.104)
Interest rate unfav $\times Eun$	-0.16	-0.16
	(0.114)	(0.115)
Industry fav $\times Eun$		0.06
		(0.092)
Industry unfav $\times Eun$		-0.23***
		(0.073)
Demand fav $\times Eun$		-0.14
		(0.145)
Demand unfav $\times Eun$		-0.57***
		(0.155)
Gov fav $\times Eun$		0.08
		(0.107)
Gov unfav $\times Eun$		0.01
		(0.079)
Sentiment fav $\times Eun$		0.01
		(0.112)
Sentiment unfav $\times Eun$		0.24**
		(0.113)
Stock fav $\times Eun$		-0.11
		(0.085)
Stock unfav $\times Eun$		0.06
		(0.115)
Other prices fav $\times Eun$		-0.01
		(0.152)
Other prices unfav $\times Eun$		-0.16
		(0.130)
Other real fav $\times Eun$		-0.11
		(0.168)
Other real unfav $\times Eun$		-0.21
		(0.157)
Wage fav $\times Eun$		-0.17
		(0.235)
Wage unfav $\times Eun$		0.00
		(0.224)
Observations	167346	167346
R^2	0.674	0.675
Time F.E.	Y	Y

Newspaper coverage of inflation and unemployment



The news coverage is defined as the sum of ratios of the word frequency divided by the total number of words in each article.

News on inflation and unemployment is domain-specific



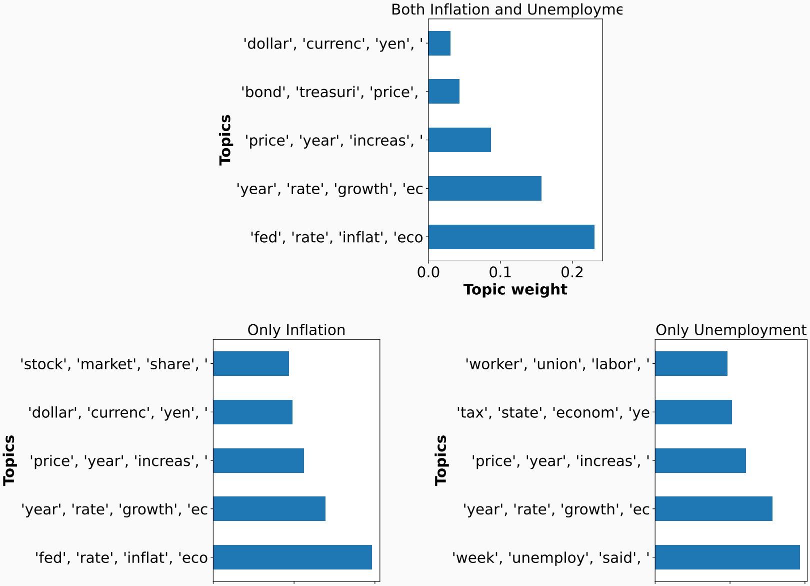
News coverage measured in the WSJ news archive.

Inflation news is always labeled as bad news

Table 7: News Coverage and Self-Reported News Exposure

Topic	Any News	Bad News	Good News
Inflation	0.605	0.627	-0.048
Unemployment	0.373	0.295	0.153

Topics in Inflation-Unemployment Narratives



Keywords in Different Inflation-Unemployment Narratives

Key words when
inflation, unemployment and
Fed are mentioned

Key words when
inflation, unemployment and
oil price are mentioned

Key words when
inflation, unemployment and
recession are mentioned

Key words when
inflation, unemployment and
growth are mentioned

References

Andrade, Philippe and Hervé Le Bihan, "Inattentive professional forecasters," *Journal of Monetary Economics*, 2013, *60* (8), 967–982.

Bhandari, Anmol, Jaroslav Borovička, and Paul Ho, "Survey data and subjective beliefs in business cycle models," *Review of Economic Studies*, 2025, *92* (3), 1375–1437.

Candia, Bernardo, Olivier Coibion, and Yuriy Gorodnichenko, "Communication and the Beliefs of Economic Agents," Working Paper 27800, National Bureau of Economic Research September 2020.

Coibion, Olivier and Yuriy Gorodnichenko, "What Can Survey Forecasts Tell Us about Information Rigidities?," *Journal of Political Economy*, 2012, *120* (1), 116 – 159.

Kamdar, Rupal, "The Inattentive Consumer: Sentiment and Expectations," 2019 Meeting Papers 647, Society for Economic Dynamics 2019.

Lucas, Robert E., "Econometric policy evaluation: A critique," *Carnegie-Rochester Conference Series on Public Policy*, 1976, *1*, 19 – 46.

Maćkowiak, Bartosz, Filip Matějka, and Mirko Wiederholt, "Dynamic rational inattention: Analytical results," *Journal of Economic Theory*, 2018, *176*, 650–692.

Sims, Christopher A., "Implications of rational inattention," *Journal of Monetary Economics*, 2003, *50* (3), 665 – 690. Swiss National Bank/Study Center Gerzensee Conference on Monetary Policy under Incomplete Information.

Woodford, Michael, "Imperfect Common Knowledge and the Effects of Monetary Policy," Working Paper 8673, National Bureau of Economic Research December 2001.