Claim

# CS 6150: $\rm HW4-Graphs,\ Randomized\ algorithms$

Submission date: Wednesday, Nov 10, 2021 (11:59 PM)

This assignment has 5 questions, for a total of 50 points. Unless otherwise specified, complete and reasoned arguments will be expected for all answers.

Question	Points	Score
QuickSelect	6	
Sampling from a stream	6	
Walking on a path	12	
Birthdays and applications	12	
Checking matrix multiplication	14	
Total:	50	

**Instructions.** For all problems in which you are asked to develop an algorithm, write down the pseudocode, along with a rough argument for correctness and an analysis of the running time (unless specified otherwise). Failure to do this may result in a penalty. If you are unsure how much detail to provide, please contact the instructors on Piazza.

Recall that given an (unsorted) array of **distinct** integers A[0, 1, ..., n-1] and a parameter  $1 \le k \le n$ , the Selection problem asks to find the kth smallest entry of A. In class, we saw an algorithm that used a randomized implementation of ApproximateMedian, and showed that it leads to an O(n) time algorithm. Let us now consider a different procedure, that is similar to QuickSort.

PROCEDURE QUICKSELECT(A, k)

- 1. If |A| = 1, return the only element
- 2. Select x from A uniformly at random
- 3. Form arrays B and C, containing the elements of A that are  $\langle x \rangle$  and  $\langle x \rangle$  respectively
- 4. If |B| = (k-1), return x, else if |B| < (k-1), return QUICKSELECT(C, k-|B|-1), else return QUICKSELECT(B, k)

Let T(n) be defined as the **expected running time** of QuickSelect on an array of length n. Using the law of conditional expectation, prove that

$$T(n) \le n + \sum_{j=1}^{n} \frac{1}{n} \max\{T(j-1), T(n-j)\}.$$

Using this along with T(1) = 1, prove that  $T(n) \le 4n$ . Write down a description of all the events you use when you use conditional expectation.

(For the purposes of this question, you may ignore the additional O(1) time for steps (1-2) and (4) of the procedure above.) [Hint: Follow the analysis for QuickSort seen in class, use induction.]

**Side note.** It is interesting to see that the constant term (the 4 in 4n) above is much better than what we had for the deterministic algorithm we saw before. It turns out that there's a way of improving the constant further: instead of choosing x uniformly at random, we pick a small sample from the array and pick the sample median.

#### Solution.

We define  $X_n$  = running time of Quickselect on array of length n. We define events  $E_j = j$ 'th smallest element is chosen as pivot at first step. We define events  $T(n) = \mathbb{E}[X_n]$ . Then we have

$$X_n | E_j \le n + \max\{X_{n-j}, X_{j-1}\} \tag{1}$$

When we apply the law of conditional expectation,

$$\mathbb{E}[X_n] \le \sum_{j=1}^n \frac{1}{n} (n + \max\{\mathbb{E}[X_{n-j}], \mathbb{E}[X_{j-1}]\})$$

$$T(n) \le n + \sum_{j=1}^n \frac{1}{n} (\max\{T(n-j), T(j-1)\})$$
(2)

We can prove that  $T(n) \leq 4n$  by induction. We assume that it is true for number less than n.

$$T(n) \le n + \sum_{j=1}^{n} \frac{1}{n} (\max\{T(n-j), T(j-1)\})$$

$$\le n + \sum_{j=n/2}^{n} \frac{2}{n} T(j)$$

$$\le n + 4 * \frac{2}{n} \frac{n/2 + n}{2} n/2$$

$$< 4n$$
(3)

If you have an array of n elements, sampling one at random is easy: you choose an index i at random in  $\{0, 1, \ldots, n-1\}$  and return the ith element. Now suppose you have a stream of elements  $a_1, a_2, \ldots$  (suppose they are <u>all distinct</u> for simplicity), and you don't know how many will arrive beforehand. Your goal is the following: at the end of the stream, you should output a random element from the stream.

The trivial algorithm is to store all the elements in an array (say a dynamic array), and in the end, output a random element. But it turns out that this can be done with very little memory.

Consider the following procedure: we maintain a special variable x, initialized to the first element of the array. At time t, upon seeing  $a_t$ , we set  $x = a_t$  with probability 1/t, otherwise we keep x unchanged.

Prove that in the end, the variable x stores a uniformly random sample from the stream. (In other words, if the stream had N elements,  $\Pr[x=a_i]=1/N$  for all i.)

[Hint: try doing a direct computation.]

### Solution.

If we select  $a_i$ , it means that we set  $x = a_t$  and we won't set x later.

$$Pr[x = a_i] = \frac{1}{t} \cdot \frac{t}{t+1} \cdot \frac{t+1}{t+2} \dots \frac{N-1}{N} = \frac{1}{N}$$
 (4)

Consider a path of length n comprising vertices  $v_0, v_1, \ldots, v_n$ . A particle starts at  $v_0$  at t = 0, and in each time step, it moves to a **uniformly random neighbor** of the current vertex. Thus if it is at  $v_s$  at time t for some s > 0, then at time (t + 1), it moves to  $v_{s+1}$  or  $v_{s-1}$  with probability 1/2 each. (If it is at  $v_0$ , the only neighbor is  $v_1$  and so it moves there.) The particle gets "absorbed" once it reaches  $v_n$  and the walk stops.

Define T(i) as the expected number of time steps taken by a particle starting at i to reach  $v_n$ . By definition, T(n) = 0.

- (a) [5] Prove that T(0) = 1 + T(1), and further, that for any 0 < s < n,  $T(s) = 1 + \frac{T(s-1) + T(s+1)}{2}$ .
- (b) [5] Use this to prove that T(s) = (2s+1) + T(s+1) for all  $0 \le s < n$ , and then find a closed form for T(0). [Hint: Use induction.]
- (c) [2] Give an upper bound for the probability that the particle walks for  $> 4n^2$  steps without getting absorbed.

#### Solution.

(a) We define  $X_s$  = running time taken from node s to reach  $v_n$ . We define events E = left or right as the particle choose to move to left node or right node. Then

$$X_0|[E = right] = 1 + X_1$$
  
 $X_0 = X_0|[E = right] = 1 + X_1$  (5)  
 $T(0) = 1 + T(1)$ 

Similarly, for starting from node s,

$$X_{s}|[E = right] = 1 + X_{s+1}$$

$$X_{s}|[E = left] = 1 + X_{s-1}$$

$$X_{s} = \frac{1}{2}(X_{s}|[E = right] + X_{s}|[E = left])$$

$$T(s) = 1 + \frac{T(s-1) + T(s+1)}{2}$$
(6)

(b) It is true when s = 0. We assume it is true for  $j \le s - 1$ , then we would prove it with induction,

$$T(s) = 1 + \frac{T(s-1) + T(s+1)}{2}$$

$$T(s) = 1 + \frac{2s - 1 + T(s) + T(s+1)}{2}$$

$$T(s) = 2s + 1 + T(s+1)$$
(7)

So it is also true for j = s. For T(0), we have,

$$T(0) = 2 * 0 + 1 + T(1) = 2 * 0 + 1 + 2 * 1 + 1 + T(2) = \dots = \sum_{j=0}^{n-1} (2 * j + 1) + T(n)$$

$$= n^{2}$$
(8)

(c)  $\mu = n^2$ 

$$Pr(T > 4\mu) < \frac{1}{4} \tag{9}$$

- (a) [5] What is the expected number of pairs (i, j) with i < j such that person i and person j have the same birthday? For what value of n (as a function of m) does this number become 1?
- (b) [7] This idea has some nice applications in CS, one of which is in estimating the "support" of a distribution. Suppose we have a radio station that claims to have a library of one million songs, and suppose that the radio station plays these songs by picking, at each

step a uniformly random song from its library (with replacement), playing it, then picking the next song, and so on.

Suppose we have a listener who started listening when the station began, and noticed that among the first 200 songs, there was a repetition (i.e., a song played twice). Prove that the probability of this happening (conditioned on the library size being a million songs) is < 0.05. Note that this gives us "reasonable doubt" about the station's claim that its library has a million songs.

Hint: Compute the probability of the complementary event —that all songs would be distinct— and prove that it must be large. You may use the inequality  $(1-x)^n \ge 1-nx$  (for x > 0 and a positive integer n) without proof.

[This idea has many applications in CS, for estimating the size of sets without actually enumerating them.]

#### Solution.

(a) We define  $X_i$  is a binary random variable of 1 or 0 which denotes the persons in pair i have the same birthday. We define  $X = \sum X_i$ . The number of pairs is,

$$N = \binom{n}{2} = \frac{n(n-1)}{2} \tag{10}$$

The expectation of  $X_i$  is,

$$\mathbb{E}[X_i] = 0 * Pr(X_i = 0) + 1 * Pr(X_i = 1) = Pr(X_i = 1)$$

$$= {m \choose 1} (\frac{1}{m})^2$$

$$= \frac{1}{m}$$
(11)

Based on the linearity of expectation, we can yield that,

$$\mathbb{E}[X] = \sum \mathbb{E}[X_i] = \frac{n(n-1)}{2} \cdot \frac{1}{m} = \frac{n(n-1)}{2m}$$
(12)

When  $n = \frac{1+\sqrt{8m+1}}{2}$ , the expectation becomes 1.

(b) We set  $n = 10^6$  and m = 200, we have the following inequality,

$$Pr(X \ge t \,\mathbb{E}[x]) \le \frac{1}{t}$$
 (13)

By setting t = 50,  $t \mathbb{E}[x] < 1$ , we could have,

$$Pr(X \ge 1) \le Pr(X \ge 50 \,\mathbb{E}[x]) \le \frac{1}{t} = 0.02$$
 (14)

So we could prove that the probability is < 0.05.

The best known algorithms here are messy and take time  $O(n^{2.36...})$ . However, the point of this exercise is to prove a simpler statement. Suppose someone gives a matrix C and claims that C = AB, can we quickly verify if the claim is true?

- (a) [5] First prove a warm-up statement: suppose a and b are any two 0/1 vectors of length n, and suppose that  $a \neq b$ . Then, for a random binary vector  $x \in \{0, 1\}^n$  (one in which each coordinate is chosen uniformly at random), prove that  $\Pr[\langle a, x \rangle \neq \langle b, x \rangle \pmod{2}] = 1/2$ . [In other words, with a probability 1/2, we can "catch" the fact that  $a \neq b$ .]
- (b) [6] Now, design an  $O(n^2)$  time algorithm that tests if C = AB and has a success probability  $\geq 1/2$ . (You need to bound both the running time and probability.)
- (c) [3] Show how to improve the success probability to 7/8 while still having running time  $O(n^2)$ .

#### Solution.

(a) Suppose  $a_i \neq b_i$ . Let  $\alpha = a_i x_i$ ,  $\beta = b_i x_i$ ,  $\bar{\alpha} = \sum_{i \neq j} a_j x_j \pmod{2}$ ,  $\bar{\beta} = \sum_{i \neq j} b_j x_j \pmod{2}$ , We assume all following calculation will apply modulo 2 implicitly.

$$\Pr[\langle a, x \rangle \neq \langle b, x \rangle \pmod{2}] = \Pr[\bar{\alpha} = \bar{\beta} | \alpha \neq \beta] \Pr[\alpha \neq \beta] + \Pr[\bar{\alpha} \neq \bar{\beta} | \alpha \neq \beta] \Pr[\alpha = \beta]$$

$$(15)$$

We know  $\alpha = \beta$  if  $x_i = 0$  and  $\alpha \neq \beta$  if  $x_i = 1$ . So  $\Pr[\alpha \neq \beta] = \Pr[\alpha = \beta] = \frac{1}{2}$ . Besides,  $\bar{\alpha}$  is independent of  $\alpha$  and  $\bar{\beta}$  is independent of  $\beta$ . So  $\Pr[\bar{\alpha} \neq \bar{\beta} | \alpha \neq \beta] = \Pr[\bar{\alpha} \neq \bar{\beta}]$ . Then we would have,

$$\Pr[\langle a, x \rangle \neq \langle b, x \rangle \pmod{2}] = \Pr[\bar{\alpha} = \bar{\beta}] \Pr[\alpha \neq \beta] + \Pr[\bar{\alpha} \neq \bar{\beta}] \Pr[\alpha = \beta]$$

$$= \frac{1}{2} (\Pr[\bar{\alpha} \neq \bar{\beta}] + \Pr[\bar{\alpha} = \bar{\beta}])$$

$$= \frac{1}{2}$$
(16)

(b) The algorithm is in Algorithm 1.

**Running time** This algorithm take  $n^2$  time as as we y = Bx first and then do T1 = Ay. Each matrix-vector product takes only  $O(n^2)$  time.

**Success probability**. If C = AB, this algorithm will always output the correct answer. If  $C \neq AB$ , at least one row in C and AB are different. This algorithm will have 0.5 probability to be successful.

$$\Pr[success] = 1 * \Pr[C = AB] + \frac{1}{2}\Pr[C \neq AB] = \frac{1}{2} + \Pr[C = AB] \ge \frac{1}{2}$$
 (17)

(c) We can boost the performance by repeating. Each time the error probability is  $\leq \frac{1}{2}$ , and we can repeat the algorithm 3 times, the error probability is less than  $\frac{1}{2}^3 = \frac{1}{8}$  and the success probability is greater than  $\frac{7}{8}$ .

## Algorithm 1 Matrix comparison

```
Input A, B
Output Matrix-comparison(A, B)
function Matrix-comparison(A, B)
Pick a vector x \in {}_R\{0,1\}^n
T1 = ABx, T2 = Cx \triangleright T1 = ABx takes n^2 time as we y = Bx first and then do T1 = Ay
if T1 == T2 then
return True
else
return False
```