

# 概率论

## 第一章 事件及其概率

### 1. 概率的运算:

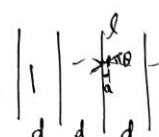
$$P(A) = 1 - P(\bar{A})$$

$$A \cap B = \phi, P(A \cup B) = P(A+B) = P(A) + P(B)$$

$$P(A \cup B) = P(A) + P(B) - P(AB), P(\bigcup_{k=1}^m A_k) = \sum_{k=1}^m P(A_k) - \sum_{k < l} P(A_k A_l) + \dots + (-1)^{m-1} P(\bigcap_{k=1}^m A_k)$$

### 2. 古典概型

### 3. N何概型

布封投针:   $l \leq d, 0 \leq \theta \leq \frac{\pi}{2}, a < \frac{d}{2}$

$$P(A) = P(a \leq \frac{l}{2} \sin \theta) = \frac{\int_0^{\frac{\pi}{2}} \frac{l}{2} \sin \theta d\theta}{\frac{d}{2} \cdot \frac{\pi}{2}} = \frac{2l}{\pi d}$$

约会问题: 甲乙相约 7:00-8:00 见面, 到达时间等可能, 人到达等待 20 分钟.



$$P = \frac{60^2 - 40^2}{60^2} = \frac{5}{9}$$

### 4. 概率的公理化体系:

匹配问题: n 封信随机投入 n 信箱, 至少一封信投入正确.

令  $A_i$  表示第 i 封信装入正确的信箱

$$P(A_i) = \frac{(n-1)!}{n!} = \frac{1}{n}, P(A_i A_j) = \frac{(n-2)!}{n!} = \frac{1}{n(n-1)}, i \neq j$$

$$\dots, P(\bigcap_{i=1}^n A_i) = \frac{1}{n!}$$

$$P(\bigcup_{i=1}^n A_i) = \sum_{i=1}^n P(A_i) - \sum_{i < j} P(A_i A_j) + \dots + (-1)^{n-1} P(\bigcap_{i=1}^n A_i)$$

$$= 1 - \frac{1}{2!} + \dots + (-1)^{n-1} \cdot \frac{1}{n!} \rightarrow 1 - \frac{1}{e}$$

### 5. 条件概率:

定义:  $P(A|B) = \frac{P(AB)}{P(B)}, P(AB) = P(A|B)P(B) = P(B|A)P(A)$

全概率公式:  $P(A) = \sum_{k=1}^N P(A|B_k)P(B_k), P(A) = P(A|B)P(B) + P(A|\bar{B})P(\bar{B})$

贝叶斯公式:  $P(B_k|A) = \frac{P(A|B_k)P(B_k)}{\sum_{k=1}^N P(A|B_k)P(B_k)}, P(B|A) = \frac{P(A|B)P(B)}{P(A)}$

例:  $P(A|B) = 0.95, P(\bar{A}|\bar{B}) = 0.9, P(B) = 0.0009, \text{求 } P(B|A)$

$$P(A|\bar{B}) = 0.1, P(A) = P(A|B)P(B) + P(A|\bar{B})P(\bar{B}) = 0.10034$$

$$P(B|A) = \frac{P(A|B)P(B)}{P(A)} = 0.0038$$

考试: 占比 60%.

七个大题, 多计算少证明, 没有概念.

(23题-题送分)

拿枪模型, 噪模型

隔板法, 多还少处理(容斥)

独立性: 定义:  $P(AB) = P(A)P(B)$   $P(A|B) = P(A)$ .

$$P(A \cup B) = P(A) + P(B).$$

$$\text{三个事件独立: } \begin{cases} P(ABC) = P(A)P(B)P(C) \\ P(AB) = P(A)P(B), P(AC) = P(A)P(C), P(BC) = P(B)P(C) \end{cases}$$

## 第二章 随机变量与分布函数

1. 随机变量与分布函数:

(i). 离散型随机变量:

(1). 退化分布.  $X \sim \begin{pmatrix} 1 \\ 1 \end{pmatrix}$

(2). 两点分布.  $X \sim \begin{pmatrix} 1 & 0 \\ p & q \end{pmatrix}$   $p+q=1$ .

(3). 二项分布  $X \sim \begin{pmatrix} 0 & 1 & \dots & k & \dots & n \\ \binom{n}{k} p^k q^{n-k} & \dots & \binom{n}{n} p^n & \dots & \binom{n}{0} q^n & \dots \end{pmatrix}$   $p+q=1$

(4). Poisson 分布  $X \sim \begin{pmatrix} 0 & 1 & \dots & k & \dots \\ e^{-\lambda} \lambda^k k! & \dots & e^{-\lambda} \lambda^k k! & \dots \end{pmatrix}$   $\lambda > 0$ .  $X \sim P(\lambda)$ .

$$\sum_{k=0}^{\infty} \frac{\lambda^k}{k!} e^{-\lambda} = 1, P(X \geq k) \approx P(X = k). \quad (k \text{ 足够大})$$

(5). N 何分布.  $X \sim \begin{pmatrix} 1 & 2 & \dots & k & \dots \\ p & p^2 & \dots & p^k & \dots \end{pmatrix}$ .

(6). 超几何分布.  $N$  件产品  $M$  件次品. 抽取  $n$  件. 次品数  $X$ .

$$P(X=k) = \frac{C_M^k C_{N-M}^{n-k}}{C_N^n}.$$

(ii). 连续型随机变量:

密度函数  $f(x) \geq 0$ .  $\int_{-\infty}^{+\infty} f(x) dx = 1$ .

(1). 均匀分布.  $X \sim p(x) = \begin{cases} \frac{1}{b-a}, & x \in (a, b) \\ 0, & \text{其他} \end{cases}$   $X \sim U(a, b)$ .

(2). 指数分布.  $p(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & \text{其他} \end{cases}$   $X \sim \exp(\lambda)$ .

$$P(X > x) = e^{-\lambda x}, \quad x > 0.$$

(3). 正态分布.  $p(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$   $X \sim N(\mu, \sigma^2)$ .

$$P(X > \mu + \sigma x) = \int_{\mu + \sigma x}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} du = \int_x^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}} du \rightarrow \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}, \quad x \rightarrow \infty$$

(iii). 一般随机变量:

定义: 分布函数:  $F(x) = P(X \leq x)$ . 注意范围

$$F(-\infty) = 0, F(+\infty) = 1.$$

$F(x)$  单调不减

$$P(X < x) = F(x-0)$$

$F(x)$  左极限存在, 右连续.

$$P(X = x) = F(x) - F(x-0).$$

离散型随机变量的分布函数, 阶跃函数.

连续型随机变量的分布函数.  $F(x) = P(X \leq x) = \int_{-\infty}^x p(u) du$

$$F(x) = p(x).$$

$$(1). \text{均匀分布. } F(x) = \begin{cases} 0, & x < a \\ \frac{x-a}{b-a}, & a \leq x < b \\ 1, & x \geq b. \end{cases}$$

$$(2). \text{指数分布. } X \sim \exp(\lambda). \quad p(x) = \begin{cases} 0, & x < 0 \\ \lambda e^{-\lambda x}, & x \geq 0 \end{cases}$$

$$F(x) = \begin{cases} 0, & x < 0 \\ 1 - e^{-\lambda x}, & x \geq 0 \end{cases}$$

$$(3). \text{正态分布. } X \sim N(0, 1).$$

$$F(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}} du = \Phi(x).$$

2. 随机变量的变换:

$$\text{例: } X \sim \exp(\lambda). \quad f(x) = \begin{cases} 0, & x \leq 0 \\ \log x, & x > 0. \end{cases} \quad Y = f(X).$$

$$P(Y \leq y) = P(\log X \leq y) = P(X \leq e^y) = 1 - e^{-e^y}$$

$$\text{例: } X \sim N(0, 1). \quad Y = |X| \text{ 的分布.}$$

$$P(Y \leq y) = P(-y \leq X \leq y) = P(X \leq y) - P(X \leq -y) = \Phi(y) - \Phi(-y).$$

$$Y \sim p(y) = \Phi(y) + \Phi(-y) = \frac{\sqrt{2}}{\sqrt{\pi}} e^{-\frac{y^2}{2}}, \quad y \geq 0.$$

$$\text{例: } X \sim N(\mu, \sigma^2). \quad Y = \frac{X - \mu}{\sigma} \text{ 的分布.}$$

$$P(Y \leq y) = P(X \leq \mu + \sigma y) = \int_{-\infty}^{\mu + \sigma y} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(u-\mu)^2}{2\sigma^2}} du = \int_{-\infty}^y \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}} du, \quad Y \sim N(0, 1).$$

一般地,  $X \sim p_X(x)$ .  $f(x)$  有反函数  $f^{-1}(y)$ .  $Y = f(X)$ .

$$Y \sim p_Y(y) = p_X(f^{-1}(y)) |f^{-1}(y)|.$$

3. 随机向量与联合密度函数.

$$(i). \text{离散型随机向量. } p_{ij} = P(X = x_i, Y = y_j).$$

$$\text{边缘分布. } X \sim \begin{pmatrix} x_1 & x_2 & \dots & x_i & \dots \\ p_{i1} & p_{i2} & \dots & p_{ii} & \dots \end{pmatrix}. \quad Y \sim \begin{pmatrix} y_1 & y_2 & \dots & y_j & \dots \\ p_{1j} & p_{2j} & \dots & p_{ij} & \dots \end{pmatrix}.$$

$$\text{条件分布. } P(Y = y_j | X = x_i) = \frac{P(X = x_i, Y = y_j)}{P(X = x_i)} = \frac{p_{ij}}{p_{i.}}$$

$$Y | X = x_i \sim \begin{pmatrix} y_1 & \dots & y_j & \dots \\ \frac{p_{i1}}{p_{i.}} & \dots & \frac{p_{ij}}{p_{i.}} & \dots \end{pmatrix}.$$

$$\text{独立性: } P(X = x_i, Y = y_j) = P(X = x_i) P(Y = y_j) \quad \forall i, j. \quad p_{ij} = p_{i.} \cdot p_{.j}$$

(ii). 连续型随机变量:

联合密度函数  $p(x, y)$ .

$$P(X \leq x, Y \leq y) = \int_{-\infty}^x \int_{-\infty}^y p(u, v) du dv.$$

$$\text{边缘分布: } X \sim p_X(x) = \int_{-\infty}^{+\infty} p(x, y) dy.$$

$$Y \sim p_Y(y) = \int_{-\infty}^{+\infty} p(x, y) dx.$$

$$\text{联合正态分布: } p(x, y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}\left(\frac{(x-\mu_1)^2}{\sigma_1^2} - \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2}\right)}$$

$$(X, Y) \sim N(\mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \rho).$$

$$\begin{aligned} & \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}(x^2 - 2\rho xy + y^2)} dx dy \\ &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}(x-\rho y)^2} e^{-\frac{y^2}{2}} dx dy \\ &= \int_{-\infty}^{+\infty} 1 \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy = 1 \end{aligned}$$

$$\text{边缘分布: } X \sim p_X(x) = \int_{-\infty}^{+\infty} p(x, y) dy$$

$$\begin{aligned} &= \int_{-\infty}^{+\infty} \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}\left(\frac{(x-\mu_1)^2}{\sigma_1^2} - \frac{2\rho(x-\mu_1)(y-\mu_2)}{\sigma_1\sigma_2} + \frac{(y-\mu_2)^2}{\sigma_2^2}\right)} dy \\ &= \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi(1-\rho^2)}\sigma_2} e^{-\frac{1}{2(1-\rho^2)\sigma_2^2}\left(y-\mu_2 - \frac{\sigma_2\rho(x-\mu_1)}{\sigma_1}\right)^2} dy \\ &= \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} \end{aligned}$$

$$\text{条件分布: } P(Y \leq y | X=x) = \lim_{\varepsilon \rightarrow 0} P(Y \leq y | x-\varepsilon < X \leq x+\varepsilon) = \frac{\int_{-\infty}^y p(x, v) dv}{p_X(x)}.$$

$$p_{Y|X}(y|x) = \frac{p(x, y)}{p_X(x)}, \quad p_{X|Y}(x|y) = \frac{p(x, y)}{p_Y(y)}$$

$$\text{例: } (X, Y) \sim N(\mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \rho)$$

$$p_{X|Y}(x|y) = \frac{p(x, y)}{p_Y(y)} = \frac{1}{\sqrt{2\pi(1-\rho^2)}\sigma_1} e^{-\frac{(x-\mu_1 - \frac{\rho\sigma_1}{\sigma_2}(y-\mu_2))^2}{2\sigma_1^2(1-\rho^2)}}$$

$$X|Y=y \sim N\left(\mu_1 + \frac{\rho\sigma_1}{\sigma_2}(y-\mu_2), \sigma_1^2(1-\rho^2)\right)$$

$$Y|X=x \sim N\left(\mu_2 + \frac{\rho\sigma_2}{\sigma_1}(x-\mu_1), \sigma_2^2(1-\rho^2)\right)$$

$$\text{独立性: } p(x, y) = p_X(x)p_Y(y), \quad \forall x, y.$$

$$(X, Y) \sim N(\mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \rho), \quad X, Y \text{ 独立} \Leftrightarrow \rho=0$$

(iii). 一般随机变量:

$$F(x, y) = P(X \leq x, Y \leq y).$$

$$F(-\infty, y) = F(x, -\infty) = 0, \quad F(+\infty, +\infty) = 1.$$

$F(x, y)$  关于  $x, y$  单调不减, 左极限存在, 右连续.

$$P(a < X \leq b, c < Y \leq d) = F(b, d) - F(a, d) - F(b, c) + F(a, c)$$

$$\text{边缘分布: } F_X(x) = F(x, +\infty), \quad F_Y(y) = F(+\infty, y).$$

$$\text{条件分布: } P(Y \leq y | X=x) = \lim_{\varepsilon \rightarrow 0} \frac{F(x+\varepsilon, y) - F(x-\varepsilon, y)}{F_X(x+\varepsilon) - F_X(x-\varepsilon)}$$

多元随机变量.

$$\text{独立性: } F(x, y) = F_X(x)F_Y(y), \quad \forall x, y \in \mathbb{R}.$$



4. 随机向量的运算:

(1). 加法: (i). 离散型随机向量:

$$P(X=x_i, Y=y_j) = p_{ij} \quad Z = X+Y$$

$$P(Z=z_k) = \sum_{i,j: x_i+y_j=z_k} p_{ij}$$

(ii). 连续型随机向量:

$$F_Z(z) = P(X+Y \leq z) = \int_{x+y \leq z} p(x,y) dx dy$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{z-x} p(x,y) dy dx$$

$$= \int_{-\infty}^z \int_{-\infty}^{\infty} p(x, y-x) dx dy$$

$$p_Z(z) = \int_{-\infty}^{\infty} p(x, z-x) dx \xrightarrow{x, y \text{ 独立}} \int_{-\infty}^{\infty} p_X(x) p_Y(z-x) dx$$

(2). 减法: 连续型随机向量:  $Z = Y - X$

$$p_Z(z) = \int_{-\infty}^{\infty} p(x, z+x) dx$$

例:  $X \sim P(\lambda), Y \sim P(\mu), X, Y \text{ 独立}, Z = X+Y$

$$P(Z=k) = P(X+Y=k) = \sum_{l=0}^k P(X=l) P(Y=k-l)$$

$$= \sum_{l=0}^k \frac{\lambda^l}{l!} e^{-\lambda} \cdot \frac{\mu^{k-l}}{(k-l)!} e^{-\mu} = \frac{(\lambda+\mu)^k}{k!} e^{-(\lambda+\mu)}$$

例:  $X, Y \sim U(0,1), X, Y \text{ 独立}, Z = X+Y$

$$p_Z(z) = \int_{-\infty}^{\infty} p_X(x) p_Y(z-x) dx = \begin{cases} \int_0^z dx = z, & 0 < z \leq 1 \quad (z-x \geq 0, x \geq 0) \\ \int_{z-1}^1 dx = z-z, & 1 < z < 2 \quad (z-x \leq 1, x \leq 1) \end{cases}$$

例:  $X \sim N(\mu_1, \sigma_1^2), Y \sim N(\mu_2, \sigma_2^2), X, Y \text{ 独立}, Z = X+Y$

$$\begin{aligned} p_Z(z) &= \int_{-\infty}^{\infty} p_X(x) p_Y(z-x) dx \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} \cdot \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{(z-x-\mu_2)^2}{2\sigma_2^2}} dx \\ &= \frac{1}{\sqrt{2\pi(\sigma_1^2+\sigma_2^2)}} e^{-\frac{(z-(\mu_1+\mu_2))^2}{2(\sigma_1^2+\sigma_2^2)}} \end{aligned}$$

(3). 乘法: 连续型随机向量:  $Z = XY$   $p_Z(z) = \int_{-\infty}^{\infty} \frac{1}{|x|} p(x, \frac{z}{x}) dx$

(4). 除法: 连续型随机向量:  $Z = \frac{Y}{X}$   $p_Z(z) = \int_{-\infty}^{\infty} |x| p(x, xz) dx$

例:  $X, Y \sim U(0,1)$  相互独立,  $Z = XY, Z = \frac{Y}{X}$

$$p_Z(z) = \int_{-\infty}^{\infty} \frac{1}{x} p(x, \frac{z}{x}) dx = \int_z^1 \frac{1}{x} dx = |\ln z|, \quad (x, \frac{z}{x} \in [0,1])$$

$$p_Z(z) = \int_{-\infty}^{\infty} x p(x, xz) dx = \begin{cases} \int_0^{\frac{1}{z}} x dx = \frac{1}{2z^2}, & z \geq 1 \\ \int_0^1 x dx = \frac{1}{2}, & 0 < z < 1 \end{cases}$$

例:  $X, Y \sim N(0,1)$  相互独立,  $Z = \frac{Y}{X}$

$$\begin{aligned} p_Z(z) &= \int_{-\infty}^{\infty} |x| p(x, xz) dx = \int_{-\infty}^{\infty} |x| \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2 x^2}{2}} dx \\ &= \frac{1}{\pi} \int_0^{\infty} x e^{-\frac{x^2(1+z^2)}{2}} dx \\ &= \frac{1}{\pi} \cdot \frac{1}{1+z^2} \quad \text{Cauchy 分布} \end{aligned}$$

★(15). 变换:  $X, Y$  是连续型随机变量, 联合密度  $p_{X,Y}(x,y)$

$$\begin{cases} U = f_1(X, Y) \\ V = f_2(X, Y) \end{cases} \text{ 求 } (U, V) \text{ 的分布.}$$

$$\text{解: } \begin{cases} X = g_1(u, v) \\ Y = g_2(u, v) \end{cases} \quad J = \begin{vmatrix} \frac{\partial g_1}{\partial u} & \frac{\partial g_1}{\partial v} \\ \frac{\partial g_2}{\partial u} & \frac{\partial g_2}{\partial v} \end{vmatrix}$$

$$p_{U,V}(u,v) = p_{X,Y}(g_1(u,v), g_2(u,v)) |J|.$$

例:  $X, Y \sim \exp(1)$  相互独立.  $U = X+Y, V = \frac{Y}{X}$ .

$$\begin{cases} U = X+Y \\ V = \frac{Y}{X} \end{cases} \Rightarrow \begin{cases} X = \frac{U}{1+V} \\ Y = \frac{UV}{1+V} \end{cases} \quad J = \frac{U}{(1+V)^2} \quad p_{X,Y}(x,y) = e^{-(x+y)}$$

$$p_{U,V}(u,v) = e^{-\left(\frac{u}{1+V} + \frac{uv}{1+V}\right)} \frac{u}{(1+V)^2} = ue^{-u} \cdot \frac{1}{(1+V)^2} \quad \cdot I\{u,v>0\}$$

(16). 极值:  $X_{(1)}$  为极小值,  $X_{(k)}$  为第  $k$  小值,  $X_{(n)}$  为极大值.  $X_1, \dots, X_n$  相互独立

$$\begin{aligned} \text{极大值: } F_{X_{(n)}}(x) &= P(X_{(n)} \leq x) = P(\max_{1 \leq i \leq n} X_i \leq x) = P(\bigcap_{i=1}^n \{X_i \leq x\}) \\ &= \prod_{i=1}^n P(X_i \leq x) = F^n(x). \end{aligned}$$

$$p_{X_{(n)}}(x) = nF^{n-1}(x)f(x)$$

$$\begin{aligned} \text{极小值: } P(X_{(1)} > x) &= P(\min_{1 \leq i \leq n} X_i > x) = P(\bigcap_{i=1}^n \{X_i > x\}) \\ &= \prod_{i=1}^n P(X_i > x) = (1-F(x))^n \end{aligned}$$

$$F_{X_{(1)}}(x) = 1 - P(X_{(1)} > x) = 1 - (1-F(x))^n$$

$$p_{X_{(1)}}(x) = n(1-F(x))^{n-1}f(x).$$

求极大与极小的联合密度: 关于两变量的条件分布

$$\star \text{第 } k \text{ 小值: } p_{X_{(k)}}(x) = (n-k+1)C_n^{k-1}F^{k-1}(x)f(x)(1-F(x))^{n-k}$$

$$p(x,y) = \frac{\partial^2 F(x,y)}{\partial x \partial y}$$

### 第三章 数字特征.

1. 数学期望:

(i). 离散型随机变量:  $EX = \sum_{k=1}^N x_k p_k$

$$(1). \text{Poisson 分布: } EX = \sum_{k=0}^{\infty} k P(X=k) = \sum_{k=0}^{\infty} k \frac{\lambda^k e^{-\lambda}}{k!} = \lambda e^{-\lambda} \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} = \lambda.$$

$$(2). \text{几何分布: } EX = \sum_{k=1}^{\infty} k P(X=k) = \sum_{k=1}^{\infty} k p(1-p)^{k-1} = \frac{1}{p}.$$

(ii). 连续型随机变量:  $EX = \int_{-\infty}^{\infty} x p(x) dx$

$$(1). \text{正态分布: } EX = 0.$$

$$(2). \text{指数分布: } EX = \int_{-\infty}^{\infty} x p(x) dx = \int_0^{\infty} x \lambda e^{-\lambda x} dx = \frac{1}{\lambda} \int_0^{\infty} x e^{-x} dx = \frac{1}{\lambda}.$$

性质: (1). 线性性:  $E(a+bX) = a + bEX$ .

$$(2). \text{加法原理: } E(X+Y) = EX + EY$$

$$(3). \text{两数变换: (i). 离散型: } P(X=x_k) = p_k, E f(X) = \sum_{k=1}^N f(x_k) p_k.$$

$$(ii). \text{连续型: } E f(X) = \int_{-\infty}^{\infty} f(x) p(x) dx = \int_{-\infty}^{\infty} f(x) dF(x)$$

2. 3差:  $Var X = E(X-EX)^2 = EX^2 - (EX)^2$ .  $EX^2 = \int_{-\infty}^{+\infty} x^2 p(x) dx$ .  $Var X = \int_{-\infty}^{+\infty} (x-EX)^2 p(x) dx$ .

(1). Poisson分布:  $P(X=k) = \frac{\lambda^k e^{-\lambda}}{k!}$ .  $EX = \lambda$ .

$$EX^2 = \sum_{k=0}^{\infty} k^2 \frac{\lambda^k e^{-\lambda}}{k!} = \sum_{k=2}^{\infty} \frac{\lambda^k e^{-\lambda}}{(k-2)!} + \lambda = \lambda^2 + \lambda$$

$$Var X = \lambda.$$

(2). 指数分布:  $p(x) = \lambda e^{-\lambda x}$ .  $EX = \frac{1}{\lambda}$

$$EX^2 = \int_{-\infty}^{+\infty} x^2 p(x) dx = \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx = \frac{2}{\lambda^2}.$$

$$Var X = \frac{1}{\lambda^2}.$$

(3). 正态分布:  $p(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$ .  $EX = 0$ .

$$EX^2 = \int_{-\infty}^{+\infty} x^2 p(x) dx = \int_{-\infty}^{+\infty} x^2 \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = 1.$$

$$Var X = 1.$$

性质: (1). 线性性:  $Var(a+bX) = b^2 Var X$ .

(2). 加法:  $Var(X+Y) = Var X + Var Y + 2(EXY - EXEY)$ .  $Var(\sum_{i=1}^n X_i) = \sum_{i=1}^n Var X_i + 2 \sum_{1 \leq i < j \leq n} Cov(X_i, X_j)$

(3). 独立性: 若  $X, Y$  独立  $Var(X+Y) = Var X + Var Y$ .

3. Chebyshev不等式:  $\forall \varepsilon > 0, P(|X-EX| > \varepsilon) \leq \frac{Var X}{\varepsilon^2}$ .

推广:  $\forall \varepsilon > 0, P(X > \varepsilon) \leq \frac{E f(X)}{f(\varepsilon)}$  单调不减,  $f > 0$

应用:  $S_n \sim B(n, p)$

$$P\left(\left|\frac{S_n}{n} - p\right| > \varepsilon\right) \leq \frac{Var(\frac{S_n}{n})}{\varepsilon^2} = \frac{np(1-p)}{n\varepsilon^2}$$

4. 协方差:  $Cov(X, Y) = E(X-EX)(Y-EY) = EXY - EXEY$ .

$$\text{协方差阵 } \Sigma = \begin{pmatrix} Var X & Cov(X, Y) \\ Cov(X, Y) & Var Y \end{pmatrix}$$

$Cov(X, Y) = 0$  表示  $X, Y$  不相关  $\iff$  独立.

例:  $\theta \sim U(0, 2\pi)$ .  $X = \sin \theta, Y = \cos \theta$ .  $V = 0$ .

对二元联合正态分布:  $(X, Y) \sim N(\mu_1, \sigma_1^2, \mu_2, \sigma_2^2, \rho)$   $Cov(X, Y) = \rho \sigma_1 \sigma_2$ .

相关系数:  $\rho = \frac{Cov(X, Y)}{\sqrt{Var X \cdot Var Y}}$

5. 条件期望:

(i). 离散型随机变量:  $P(X=x_i | Y=y_j) = \frac{p_{ij}}{p_{.j}}$

$$E(X | Y=y_j) = \sum_{i=1}^{\infty} x_i P(X=x_i | Y=y_j)$$

(ii). 连续型随机变量:  $P(X=x | Y=y) = \frac{p(x, y)}{p_Y(y)} = p_{X|Y}(x|y)$

$$E(X | Y=y) = \int_{-\infty}^{+\infty} x p_{X|Y}(x|y) dx$$

全期望公式:  $E(E(X|Y)) = EX$

$$E(E(Y|X)) = EY$$

6. 矩: 定义  $k$  阶矩:  $EX^k$

$$k$$
 阶中心矩:  $E(X-EX)^k$

例:  $X \sim N(0, \sigma^2)$

$$EX^{2k} = (2k-1)!! \sigma^{2k} \quad EX^{2k+1} = 0, \quad k \geq 1$$

7. 特征函数:  $X \sim F_X(x)$   $\varphi(t) = Ee^{itX} = \int_{-\infty}^{+\infty} e^{itx} dF_X(x)$

(i) 离散型随机变量:

(1) 退化分布:  $P(X=c)=1$   $\varphi(t) = e^{itc}$

(2) 两点分布:  $P(X=1)=p, P(X=0)=1-p$   $\varphi(t) = pe^{it} + 1-p$

(3) 二项分布:  $X \sim B(n, p)$

$$\varphi(t) = Ee^{itX} = \sum_{k=0}^n e^{itk} C_n^k p^k (1-p)^{n-k} = (1-p + pe^{it})^n$$

(4) Poisson 分布:  $X \sim P(\lambda)$

$$\varphi(t) = Ee^{itX} = \sum_{k=0}^{\infty} e^{itk} \frac{\lambda^k}{k!} e^{-\lambda} = e^{\lambda(e^{it}-1)}$$

(ii) 连续型随机变量:

(1) 均匀分布:  $X \sim U(a, b)$

$$\varphi(t) = Ee^{itX} = \int_a^b e^{itx} \frac{1}{b-a} dx = \frac{e^{itb} - e^{ita}}{it(b-a)}$$

(2) 指数分布:  $X \sim \exp(\lambda)$

$$\varphi(t) = Ee^{itX} = \int_0^{\infty} e^{itx} \lambda e^{-\lambda x} dx = \frac{\lambda}{\lambda - it}$$

$$X \sim N(\mu, \sigma^2) \\ \varphi(t) = e^{it\mu - \frac{\sigma^2 t^2}{2}}$$

(3) 正态分布:  $X \sim N(0, 1)$

$$\varphi(t) = Ee^{itX} = \int_{-\infty}^{+\infty} e^{itx} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = \int_{-\infty}^{+\infty} \cos tx \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx + \int_{-\infty}^{+\infty} i \sin tx \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = \dots = e^{-\frac{t^2}{2}}$$

性质:  $\varphi(0) = 1$   $|\varphi(t)| \leq 1$   $\varphi(-t) = \overline{\varphi(t)}$   $\varphi(t)$  在  $\mathbb{R}$  上一致连续.

(1)  $Ee^{it(ax+c)} = e^{itc} \varphi_X(at)$  例:  $Y \sim N(\mu, \sigma^2) \Leftrightarrow Y = \sigma X + \mu, X \sim N(0, 1)$

$$\varphi_Y(t) = Ee^{it(\sigma X + \mu)} = e^{it\mu - \frac{\sigma^2 t^2}{2}}$$

(2)  $Z = X + Y$  则  $\varphi_Z(t) = \varphi_X(t) \varphi_Y(t)$  ( $X, Y$  相互独立) (= 二项分布 = 两点分布<sup>n</sup>)

(3) 唯一性定理:  $\varphi_X(t) = \varphi_Y(t) \Leftrightarrow X \stackrel{d}{=} Y, F_X(x) = F_Y(y)$

推论:  $\varphi(t)$  绝对可积,  $f(x) = \frac{1}{2\pi} \int_{-\infty}^{+\infty} e^{-itx} \varphi(t) dt$

$$\varphi(t) = \sum_{k=-\infty}^{+\infty} a_k e^{ikt}, \text{ 则 } P(X=k) = a_k, (\sum a_k = 1)$$



利用唯一性定理计算随机变量的分布

例1.  $y(t) = e^{-t}$

$$p(x) = \frac{1}{2\pi} \int_{-\infty}^{+\infty} e^{itx} e^{-t} dt = \frac{1}{2\pi} \int_0^{+\infty} e^{-t(1-ix)} dt + \frac{1}{2\pi} \int_{-\infty}^0 e^{t(1-ix)} dt$$

$$= \frac{1}{2\pi} \left( \frac{-1}{ix+1} e^{-t(ix+1)} \Big|_0^{+\infty} + \frac{1}{1-ix} e^{t(1-ix)} \Big|_{-\infty}^0 \right) = \frac{1}{2\pi} \cdot \left( \frac{1}{ix+1} + \frac{1}{1-ix} \right) = \frac{1}{\pi} \cdot \frac{1}{1+x^2}$$

例2.  $y(t) = \cos t = \frac{1}{2} e^{it} + \frac{1}{2} e^{-it}$   $p(x=1) = \frac{1}{2}$ ,  $p(x=-1) = \frac{1}{2}$ .

例3.  $X_k \sim N(\mu_k, \sigma_k^2)$ .  $Z = \sum_{k=1}^n X_k$ .

$$f_Z(t) = E e^{itZ} = \prod_{k=1}^n E e^{itX_k} = \prod_{k=1}^n e^{it\mu_k - \frac{\sigma_k^2 t^2}{2}} = e^{it \sum \mu_k - \frac{\sum \sigma_k^2 \cdot t^2}{2}}$$

$$\Rightarrow Z \sim N(\mu, \sigma^2). \mu = \sum \mu_k, \sigma^2 = \sum \sigma_k^2.$$

二维随机变量的特征函数:  $\phi(t_1, t_2) = E e^{i(t_1 X + t_2 Y)}$

当  $X, Y$  相互独立时:  $\phi(t_1, t_2) = \phi_X(t_1) \phi_Y(t_2)$

## 第四章 概率极限理论

1. 大数定律:

(1). 伯努利大数定律:  $S_n \sim B(n, p)$ .

$$P(|\frac{S_n}{n} - p| > \varepsilon) \rightarrow 0, n \rightarrow \infty.$$

(2). 切比雪夫大数定律: 设随机变量  $\{X_k\}$ ,  $E X_k = \mu$ ,  $S_n = \sum_{k=1}^n X_k$ .  $\frac{\text{Var } S_n}{n^2} \rightarrow 0$ .

(要求方差存在)  $\frac{S_n}{n} \xrightarrow{P} \mu, n \rightarrow \infty$ .

推广:  $E X_k = \mu_k$ ,  $\frac{\text{Var } S_n}{n^2} \rightarrow 0$

$$\frac{S_n}{n} - \frac{\sum \mu_k}{n} \xrightarrow{P} 0, n \rightarrow \infty.$$

(3). 辛钦大数定律:  $\{X_k\}$  是一列独立同分布的随机变量  $E X_k = \mu$ ,  $S_n = \sum_{k=1}^n X_k$ .

(要求独立同分布)  $\frac{S_n}{n} \xrightarrow{P} \mu$ .

2. 中心极限定理:

(1). 林德贝格-莱维中心极限定理:  $S_n \sim B(n, p)$ .  $\frac{S_n - np}{\sqrt{npq}} \xrightarrow{d} N(0, 1)$ .

(2). Poisson 极限定理:  $S_n \sim B(n, p_n)$ .  $np_n \rightarrow \lambda, 0 < \lambda < \infty$ .

$$\forall k=0, 1, 2, \dots, P(S_n=k) \rightarrow \frac{\lambda^k e^{-\lambda}}{k!}, n \rightarrow \infty \quad \text{二项分布的泊松逼近}.$$

(3). Levy-Feller 中心极限定理:  $\{X_k\}$  是一列独立同分布的随机变量,  $E X_k = \mu$ ,  $\text{Var } X_k = \sigma^2$ ,  $S_n = \sum_{k=1}^n X_k$ .

$$\forall x, P\left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x\right) \rightarrow \Phi(x). \text{ 即 } \frac{S_n - n\mu}{\sigma\sqrt{n}} \xrightarrow{d} N(0, 1).$$

(4). Lyapunov 中心极限定理:  $\{X_k\}$  是一列独立同分布的随机变量,  $E X_k = \mu_k$ ,  $\text{Var } X_k = \sigma_k^2$ ,  $S_n = \sum_{k=1}^n X_k$ ,  $B_n = \sum_{k=1}^n \sigma_k^2$ .

$$\text{若 } B_n \rightarrow \infty, E|X_k|^3 < \infty, \frac{\sum_{k=1}^n E|X_k|^3}{B_n^{\frac{3}{2}}} \rightarrow 0, n \rightarrow \infty.$$

$$\text{则对 } \forall x, P\left(\frac{\sum_{k=1}^n (X_k - \mu_k)}{\sqrt{B_n}} \leq x\right) \rightarrow \Phi(x). \text{ 即 } \frac{\sum_{k=1}^n (X_k - \mu_k)}{\sqrt{B_n}} \xrightarrow{d} N(0, 1).$$

### 3. 收敛性:

(1). 依概率收敛: 对  $\forall \varepsilon > 0$ ,  $P(|X_n(\omega) - X(\omega)| > \varepsilon) \rightarrow 0, n \rightarrow \infty. \Rightarrow X_n \xrightarrow{P} X$ . 联系切比雪夫不等式 (如果  $\text{Var} X$  存在趋于 0)

判别法:  $\exists t > 0$ , s.t.  $E|X_n - X|^t \rightarrow 0, n \rightarrow \infty. \Rightarrow X_n \xrightarrow{P} X$ .

性质: 线性映射.

连续映射. 例:  $\{g_k\}$  独立同分布,  $g_k \sim U(0,1)$ ,  $\eta_n = (\prod_{k=1}^n g_k)^{\frac{1}{n}}$ . 求证:  $\eta_n \xrightarrow{P} c, n \rightarrow \infty$

$$\text{证: } \log \eta_n = \frac{1}{n} \sum_{k=1}^n \log g_k \xrightarrow{\text{大数}} E \log g_1 = \int_0^1 \log x dx = -1$$

$$\eta_n \xrightarrow{P} e^{-1}$$

(2). 依分布收敛: 对  $F$  的每个连续点  $x$ ,  $F_n(x) \rightarrow F(x), n \rightarrow \infty. \Rightarrow F_n \xrightarrow{d} F, X_n \xrightarrow{d} X$ .

$X_n \xrightarrow{d} X \Rightarrow X_n \xrightarrow{d} c$ . 反之不成立. 但  $X_n \xrightarrow{d} c \Rightarrow X_n \xrightarrow{P} c$ .

判别法: [Levy 连续性定理]  $X_n \xrightarrow{d} X \Leftrightarrow f_n(t) \rightarrow f(t)$

性质: 线性性

连续映射:

例:  $\{g_k\}$  是一列独立同分布随机变量  $P(g_k = \pm 1) = \frac{1}{2}$ .  $U_n = \prod_{k=1}^n \frac{g_k}{2^k}$ . 求证:  $U_n \rightarrow U \sim U(-1,1)$ .

$$\text{证: } f_n(t) = E e^{it U_n} = \prod_{k=1}^n E e^{it \frac{g_k}{2^k}}$$

$$= \prod_{k=1}^n \frac{1}{2} (e^{\frac{it}{2^k}} + e^{-\frac{it}{2^k}}) = \prod_{k=1}^n \cos \frac{t}{2^k}$$

$$f(t) = E e^{it U} = \int_{-1}^1 e^{itx} dx = \frac{1}{it} (e^{it} - e^{-it}) = \frac{\sin t}{t}$$

$$\text{证: } \prod_{k=1}^n \cos \frac{t}{2^k} \rightarrow \frac{\sin t}{t}, n \rightarrow \infty. \quad \sin 2\theta = 2 \sin \theta \cos \theta, \quad \frac{\sin \theta}{\theta} \rightarrow 1$$

$$\prod_{k=1}^n \cos \frac{t}{2^k} = \frac{1}{\sin \frac{t}{2^n}} \prod_{k=1}^n \cos \frac{t}{2^k} \sin \frac{t}{2^n} = \frac{1}{\sin \frac{t}{2^n}} \cdot \frac{\sin t}{2^n} \rightarrow \frac{1}{t} \cdot \frac{\sin t}{2^n} = \frac{\sin t}{t}$$

(3). 几乎处处收敛:  $\exists \Omega_0, P(\Omega_0) = 0$ . 且对  $\forall \omega \in \Omega \setminus \Omega_0, X_n(\omega) \rightarrow X(\omega), n \rightarrow \infty. \Rightarrow X_n \rightarrow X, \text{ a.s. (almost surely)}$

判别法:  $X_n \rightarrow X, \text{ a.s.} \Leftrightarrow \forall \varepsilon > 0, P(\bigcap_{n=1}^{\infty} \bigcup_{k=n}^{\infty} \{|X_k(\omega) - X(\omega)| > \varepsilon\}) = 0$ .

$$\Leftrightarrow \lim_{n \rightarrow \infty} P(\bigcup_{k=n}^{\infty} \{|X_k(\omega) - X(\omega)| > \varepsilon\}) = 0$$

Borel 大数律:  $\{g_k\}$  是一列独立同分布随机变量  $P(g_k = 1) = p, P(g_k = 0) = 1 - p. S_n = \sum_{k=1}^n g_k$ .

$$\text{证: } \frac{S_n}{n} \rightarrow p, \text{ a.s.}$$

柯尔莫哥洛夫强大数律:  $\{g_k\}$  是一列独立同分布随机变量  $E g_k = \mu. S_n = \sum_{k=1}^n g_k$ .

$$\text{证: } \frac{S_n}{n} \rightarrow \mu, \text{ a.s.}$$