

Assignment 2

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Github Link

<https://github.com/Tarandeep97/AI5030>

1 PROBLEM

(52) Suppose $X \sim \text{Cauchy}(0,1)$. Then the distribution of $\frac{1-X}{1+X}$ is?

2 SOLUTION - METHOD I

A continuous random variable X follows **Cauchy distribution** with parameters μ and λ if its pdf is given by

$$f(x) = \begin{cases} \frac{\lambda}{\pi} \cdot \frac{1}{\lambda^2 + (x-\mu)^2}, & -\infty < x < \infty; \\ 0, & -\infty < \mu < \infty, \lambda > 0; \\ \text{Otherwise.} \end{cases}$$

The parameter μ and λ are location and scale parameters respectively.

When $\mu=0$ and $\lambda=1$, then the distribution is called **Standard Cauchy Distribution**. The pdf of standard Cauchy distribution is

$$f(x) = \begin{cases} \frac{1}{\pi} \cdot \frac{1}{1+x^2}, & -\infty < x < \infty; \\ 0, & \text{Otherwise.} \end{cases}$$

Let,

$$Y = g(X) = \frac{1-X}{1+X} \quad (2.0.1)$$

Range of Y

Since $g(X)$ is self-inverse of itself, it is bijective in nature. Self-inverse nature of $g(X)$ is verified below,

$$g\left(\frac{1-X}{1+X}\right) = \frac{1 - \frac{1-X}{1+X}}{1 + \frac{1-X}{1+X}} = X \quad (2.0.2)$$

So, $g(X) \in (-\infty, \infty)$.

Expression for CDF of Y

$$F_Y(y) = P(Y \leq y) \quad (2.0.3)$$

$$= P\left(\frac{1-X}{1+X} \leq y\right) \quad (2.0.4)$$

If $X \in (-\infty, -1)$, $(1+X) < 0$. So, inequality should be reversed.

$$= P((1-X) \geq y(1+X)) \quad (2.0.5)$$

$$= P((1-X) \geq (y+yX)) \quad (2.0.6)$$

$$= P((1-y) \geq (X+yX)) \quad (2.0.7)$$

$$= P\left(\frac{1-y}{1+y} \geq X\right) = P\left[-\infty < X \leq \left(\frac{1-y}{1+y}\right)\right] \quad (2.0.8)$$

$$= F_X\left(\frac{1-y}{1+y}\right) - F_X(-\infty) \quad (2.0.9)$$

If $X \in (-1, \infty)$, $(1+X) > 0$. The expression can be rewritten as

$$= P\left(\frac{1-y}{1+y} \leq X\right) \quad (2.0.10)$$

$$= P\left[\left(\frac{1-y}{1+y}\right) \leq X < \infty\right] \quad (2.0.11)$$

$$= F_X(\infty) - F_X\left(\frac{1-y}{1+y}\right) \quad (2.0.12)$$

Now, using above equation CDF of Y will be,

$$F_Y(y) = P[-\infty < X \leq -1] + P\left[\left(\frac{1-y}{1+y}\right) \leq X < \infty\right] \quad (2.0.13)$$

$$= (F_X(-1) - F_X(-\infty)) + \left(F_X(\infty) - F_X\left(\frac{1-y}{1+y}\right)\right) \quad (2.0.14)$$

$$= \frac{3}{4} - \frac{1}{\pi} \cdot \tan^{-1}\left(\frac{1-y}{1+y}\right) \quad (2.0.15)$$

Expression for PDF of Y

$$f_Y(y) = \frac{dF_Y(y)}{dy} = -\frac{1}{\pi} \cdot \frac{d\left(\tan^{-1}\left(\frac{1-y}{1+y}\right)\right)}{dy} \quad (2.0.16)$$

$$= -\frac{1}{\pi} \cdot \frac{1}{1 + \left(\frac{1-y}{1+y}\right)^2} \cdot \frac{d\left(\frac{1-y}{1+y}\right)}{dy} \quad (2.0.17)$$

$$= -\frac{1}{\pi} \cdot \frac{1}{1 + \left(\frac{1-y}{1+y}\right)^2} \cdot \left(-\frac{2}{(y+1)^2}\right) \quad (2.0.18)$$

$$= \frac{1}{\pi} \cdot \frac{1}{1 + y^2} \quad (2.0.19)$$

Hence, $Y \sim \text{Cauchy}(0,1)$.

3 SOLUTION - METHOD II

Theorem: If $U \sim \text{Normal}(0,1)$, $V \sim \text{Normal}(0,1)$ are independent then $X = \frac{U}{V}$ is a Cauchy Distribution.

$$Y = \frac{1-X}{1+X} = \frac{1 - \frac{U}{V}}{1 + \frac{U}{V}} = \frac{U-V}{U+V} \quad (3.0.1)$$

$U-V$ and $U+V$ are also independent and normally distributed.

Hence, $Y \sim \text{Cauchy}(0,1)$.

4 APPENDIX

Theorem: If $U \sim \text{Normal}(0,1)$, $V \sim \text{Normal}(0,1)$ are independent then $X = \frac{U}{V}$ is a Cauchy Distribution.

Proof: PDF of U and V is given as,

$$f_U(u) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{u^2}{2}}, f_V(v) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{v^2}{2}} \quad (4.0.1)$$

where $-\infty < u, v < \infty$

Since U and V are independent, their joint PDF is given by,

$$f_{UV}(u, v) = \frac{1}{2\pi} \cdot e^{-\frac{u^2+v^2}{2}} \quad (4.0.2)$$

$$F_X(x) = P(X \leq x) = P\left(\frac{U}{V} \leq x\right) \quad (4.0.3)$$

$$= P(U \leq Vx, V \geq 0) + P(U \geq Vx, V < 0) \quad (4.0.4)$$

$$= \int_0^\infty \int_{-\infty}^{Vx} f_{uv}(u, v) du dv + \int_{-\infty}^0 \int_{-\infty}^{Vx} f_{uv}(u, v) du dv \quad (4.0.5)$$

Differentiating $F_X(x)$ to obtain PDF of X , $f_X(x)$
(Using Leibniz's Integral Rule)

$$f_X(x) = \int_0^\infty v \cdot f_{uv}(vx, v) dv + \int_{-\infty}^0 (-v) \cdot f_{uv}(vx, v) dv \quad (4.0.6)$$

$$f_X(x) = \int_0^\infty v \cdot \frac{1}{2\pi} \cdot e^{-\frac{(1+x^2)v^2}{2}} dv + \int_{-\infty}^0 (-v) \cdot \frac{1}{2\pi} \cdot e^{-\frac{(1+x^2)v^2}{2}} dv \quad (4.0.7)$$

Let,

$$u = \frac{(1+x^2)v^2}{2} \quad (4.0.8)$$

$$\frac{du}{dv} = (1+x^2)v \Rightarrow dv = \frac{du}{(1+x^2)v} \quad (4.0.9)$$

$$f_X(x) = \int_0^\infty \frac{1}{2\pi(1+x^2)} \cdot e^{-u} du - \int_{-\infty}^0 \frac{1}{2\pi(1+x^2)} \cdot e^{-u} du \quad (4.0.10)$$

$$f_X(x) = \frac{1}{\pi(1+x^2)} \quad (4.0.11)$$

Hence, $X \sim \text{Cauchy}(0,1)$

Theorem: If $U \sim \text{Normal}(0,1)$ and $V \sim \text{Normal}(0,1)$, then $(U-V)$ and $(U+V)$ are independent.

Proof: For $(U+V)$ and $(U-V)$ to be independent their co-variance must be 0.

$$= E[(U+V)(U-V)] - E[U+V]E[U-V] \quad (4.0.12)$$

$$= E[U^2] + E[V^2] - (E[U] + E[V])(E[U] - E[V]) \quad (4.0.13)$$

Hence,

$$\text{cov}(U+V, U-V) = 0 \quad (4.0.14)$$

Leibniz Integral Rule: For an integral form

$$\int_{a(x)}^{b(x)} f(x, t) dt$$

where $-\infty < a(x), b(x) < \infty$, the derivative of this integral is expressed as,

$$\frac{d}{dx} \left(\int_{a(x)}^{b(x)} f(x, t) dt \right) \quad (4.0.15)$$

$$= f(x, b(x)) \cdot \frac{d}{dx} b(x) - f(x, a(x)) \cdot \frac{d}{dx} a(x) \quad (4.0.16)$$

$$+ \int_{a(x)}^{b(x)} \frac{\partial}{\partial x} f(x, t) dt \quad (4.0.17)$$

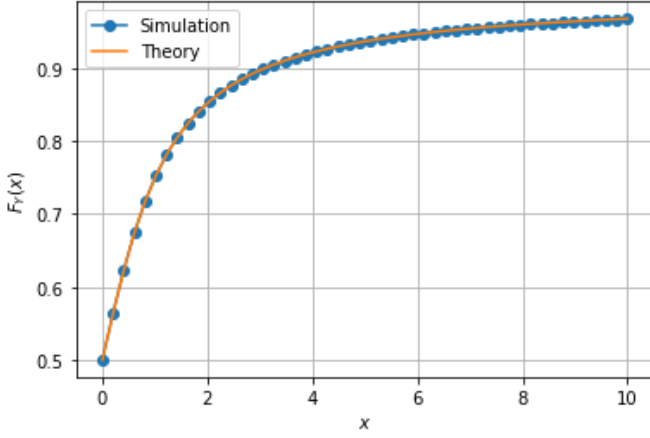


Fig 1. CDF of Y

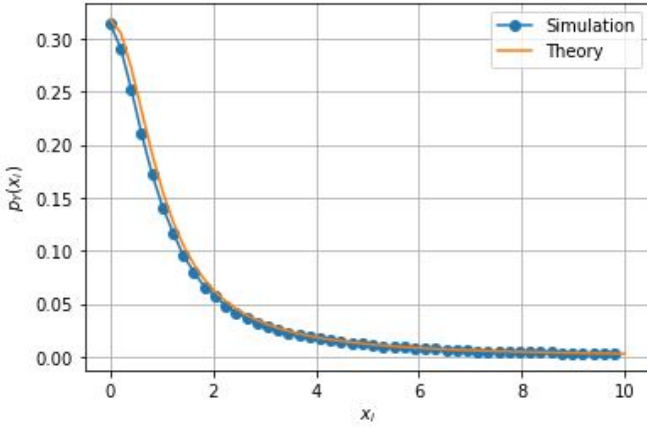


Fig 2. PDF of Y