MATLAB

else ... end

% Function declaration function [output1, ...] = function name(input1, ...) % Iterate over the vector v for i = v ... end % Repeat while condition is true while condition ... end % Execute first true alternative if condition_1 ... elseif condition_2 ...

Partial Fraction Decomposition

Given the LHS in the denominator, substitute the RHS.

$$(ax+b)^k \to \frac{A_1}{ax+b} + \dots + \frac{A_k}{(ax+b)^k}$$
$$(ax^2 + bx + c)^k \to$$
$$\frac{A_1x + B_1}{ax^2 + bx + c} + \dots + \frac{A_kx + B_k}{(ax^2 + bx + c)^k}$$

Differential Equations Mechanical Systems

$$\sum F = \frac{\mathrm{d}p}{\mathrm{d}t}, \quad \sum M = I \frac{\mathrm{d}^2\theta}{\mathrm{d}t^2}$$

where momentum p = mv, moment M =Second-Order ODEs Fx and inertia $I = mr^2$.

$$\begin{split} F_T &= (c-v)\,d_f \quad F_g = -mg \\ F_s &= -kx \qquad \qquad F_f = -bv^2 \text{ (or } -bv) \end{split}$$

Electrical Circuits

$$i = \frac{\mathrm{d}q}{\mathrm{d}t}, \quad \sum v_{loop} = 0, \quad \sum i_{node} = 0$$
 Voltage drop across elements:

$$v_R=iR, \quad v_C=\frac{q}{C}, \quad v_L=L\frac{\mathrm{d}i}{\mathrm{d}t}$$

Separable ODEs

For
$$\frac{dy}{dt} = p(t)q(y)$$
:

$$\int \frac{1}{q(y)} \frac{dy}{dt} dt = \int p(t) dt.$$

Linear ODEs

For
$$\frac{\mathrm{d}y}{\mathrm{d}t} + p(t)y = q(t)$$
, use the integrating factor: $I(t) = e^{\int p(t)\mathrm{d}t}$, so that
$$y(t) = \frac{1}{I(t)} \int I(t)q(t)\,\mathrm{d}t.$$

Linearisation

$$\begin{split} f(t) &\approx f(t_0) + f'(t_0)(t-t_0) \\ f\big(y(t)\big) &\approx f\big(y(t_0)\big) \\ &+ f'\big(y(t_0)\big)\big(y(t) - y(t_0)\big) \end{split}$$

Euler's Method

For
$$\mathbf{y}' = \mathbf{f}(t, \mathbf{y})$$
:
 $\mathbf{y}_{n+1} = \mathbf{y}_n + h\mathbf{f}(t_n, \mathbf{y}_n)$

Modified Euler's Method

$$\begin{split} & \boldsymbol{y}_{n+1} = \\ & \boldsymbol{y}_n + \frac{h}{2} \left(\boldsymbol{f}(t_n, \ \boldsymbol{y}_n) + \boldsymbol{f}(t_{n+1}, \ \boldsymbol{y}_{n+1}) \right) \\ & \text{where } \boldsymbol{f}(t_{n+1}, \ \boldsymbol{y}_{n+1}) \text{ is determined using } \\ & \text{Euler's method.} \end{split}$$

$$ay'' + by' + cy = F(t)$$

General Solution

$$y(t) = y_H(t) + y_P(t)$$

Homogeneous Solution

$$y_H(t)=e^{\lambda t}$$

To solve for λ , substitute homogeneous form into the homogeneous

ODE.

Real distinct roots:

$$y_H(t) = c_1 e^{\lambda_1 t} + c_2 e^{\lambda_2 t}$$

Real repeated roots:

$$\overline{y_H(t) = c_1 e^{\lambda t} + c_2 t e^{\lambda t}}$$

Complex conjugate roots:
$$\lambda = \alpha \pm \beta i$$

$$y_H(t) = e^{\alpha x} (c_1 \cos(\beta t) + c_2 \sin(\beta t))$$

Particular Solution

See table below. Substitute y_P into the nonhomogeneous ODE, and solve the undetermined coefficients.

System of ODEs

Given y' = Ay,

$${m y}_H(t) = {m q} e^{\lambda t}.$$

 λ_i are the eigenvalues of **A** that satisfy $\det\left(\mathbf{A} - \lambda \mathbf{I}\right) = 0.$

 q_i are the associated eigenvectors that satisfy

$$(\mathbf{A} - \lambda_i \mathbf{I}) \, \boldsymbol{q}_i = \mathbf{0}.$$

For real distinct roots:

$$\overline{\boldsymbol{y}_H(t) = c_1 \boldsymbol{q}_1 e^{\lambda_1 t} + c_2 \boldsymbol{q}_2 e^{\lambda_2 t}}$$

Higher-Order ODEs

$$y^{(n)} + a_1 y^{(n-1)} + \dots + a_{n-1} y' + a_n y = 0$$

Let
$$y_1=y,\ y_2=y',\ \dots,\ y_n=y^{(n-1)}$$
 so that ${\pmb y}=\langle y_1,\,y_2,\,\dots,\,y_n\rangle.$ Then ${\pmb y}'={\bf A}{\pmb y},$ where

$$\mathbf{A} = \begin{bmatrix} 0 & 1 & 0 & \cdots & 0 \\ 0 & 0 & 1 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & 1 \\ -a_n & -a_{n-1} & -a_{n-2} & \cdots & -a_1 \end{bmatrix}$$

$$F(t) \qquad \qquad y_P(t)$$
 a constant
$$A$$
 a polynomial of degree n
$$\sum_{i=0}^n A_i t^i$$

$$e^{kt} \qquad \qquad Ae^{kt}$$

$$\cos{(\omega t)} \text{ or } \sin{(\omega t)} \qquad \qquad A_0 \cos{(\omega t)} + A_1 \sin{(\omega t)}$$
 a combination of the above linearly dependent to $y_H(t)$ multiply $y_P(t)$ by t until linearly independent

Probability

Events

$$\begin{split} \Pr\left(A\right) &= \sum_{x \in A} p(x) \\ \Pr\left(A^{C}\right) &= \Pr\left(\overline{A}\right) = 1 - \Pr\left(A\right) \end{split}$$

Unions

$$\begin{split} \Pr\left(A \cup B\right) &= \Pr\left(A\right) + \Pr\left(B\right) \\ &- \Pr\left(A \cap B\right) \\ &= 1 - \Pr\left(A^C \cap B^C\right) \end{split}$$

Bayes' Theorem

$$\Pr(B \mid A) = \frac{\Pr(A \mid B) \Pr(B)}{\Pr(A)}$$

Disjoint Events

Events don't have outcomes in common. For disjoint events B_i :

$$A \cap B = \emptyset$$

$$\Pr(A \cap B) = 0$$

Independent Events

Outcome of events do not influence each other. Joint probability:

$$\Pr(A \cap B) = \Pr(A) \Pr(B)$$

Dependent (Conditional) Events

Outcome of event depends on the Variance outcome of the other. Joint probability of A given B:

$$\Pr(A \mid B) = \frac{\Pr(A \cap B)}{\Pr(B)}$$

Total Probability

$$A = \bigcup_{i=1}^{n} (A \cap B_i)$$

$$\Pr(A) = \sum_{i=1}^{n} \Pr(A \cap B_i)$$

Expectation

Average output or mean:

$$\mu = \mathrm{E}(X)$$
.

Measure of spread from the mean:

$$\sigma^2 = \mathrm{E}\left(X^2\right) - \mathrm{E}\left(X\right)^2.$$

Probability Distributions Discrete Random Variables

countably many outcomes. Distributed with a probability mass function p(x).

Continuous Random Variables

Has an infinite number of outcomes. Distributed with a probability density function f(x).

$$\Pr\left(x_1 \leq X \leq x_2\right) = \int_{x_1}^{x_2} f(x) \,\mathrm{d}x$$

$$F(x)=\Pr\left(X\leq x\right)$$

Valid probabilities

Cumulative probability

Expectation

Variance

General Linear Combinations

$$\begin{split} \operatorname{E}\left(aX \pm b\right) &= a\operatorname{E}\left(X\right) \pm b\\ \operatorname{E}\left(aX \pm bY\right) &= a\operatorname{E}\left(X\right) \pm b\operatorname{E}\left(Y\right)\\ \operatorname{Var}\left(aX \pm b\right) &= a^{2}\operatorname{Var}\left(X\right)\\ \operatorname{Var}\left(aX \pm bY\right) &= a^{2}\operatorname{Var}\left(X\right)\\ &+ b^{2}\operatorname{Var}\left(Y\right)\\ &\pm 2ab\operatorname{Cov}\left(X,\,Y\right)\\ \operatorname{Cov}\left(X,\,Y\right) &= \operatorname{E}\left(XY\right) - \operatorname{E}\left(X\right)\operatorname{E}\left(Y\right)\\ &= \rho_{XY}\sqrt{\operatorname{Var}\left(X\right)\operatorname{Var}\left(Y\right)} \end{split}$$

a constant that describes the statistical Uniform (Continuous) relationship between X and Y. $-1 \le$ $\rho_{XY} \leq 1$.

Continuous

$$\begin{split} \operatorname{E}\left(aX \pm b\right) &= a \operatorname{E}\left(X\right) \pm b \\ \operatorname{E}\left(aX \pm bY\right) &= a \operatorname{E}\left(X\right) \pm b \operatorname{E}\left(Y\right) \\ \operatorname{Var}\left(aX \pm b\right) &= a^{2} \operatorname{Var}\left(X\right) \\ \operatorname{Var}\left(aX \pm bY\right) &= a^{2} \operatorname{Var}\left(X\right) \\ &+ b^{2} \operatorname{Var}\left(Y\right) \\ &\pm 2ab \operatorname{Cov}\left(X, \, Y\right) \\ \operatorname{Cov}\left(X, \, Y\right) &= \operatorname{E}\left(XY\right) - \operatorname{E}\left(X\right) \operatorname{E}\left(Y\right) \\ &= \rho_{XY} \sqrt{\operatorname{Var}\left(X\right) \operatorname{Var}\left(Y\right)}. \end{split}$$

The correlation $\operatorname{Corr}(X, Y)$ or ρ_{XY} is average rate λ .

Binomial (Discrete)

Probability of x successes out of nindependent trials, each with chance p.

Poisson (Discrete)

Probability of observing x events over an interval t, where events occur at an average rate λ . $\mu = \lambda t$.

Exponential (Continuous)

The time t between events, where the events are independent and occur at an

The probability of any value $x \in [a, b]$ is constant.

Normal (Continuous)

Events occur more frequently near μ and less frequently further away from μ .

Standardised Normal (Continuous)

$$Z = \frac{X - \mu}{\sigma} \sim \mathcal{N}\left(0, \, 1\right)$$

Distribution Probability		Cumulative Probability	μ	σ^2
Binomial: $X \sim \text{bin}(n, p)$	$\binom{n}{x}p^x\left(1-p\right)^{n-x}$	$See\ table\ above.$	np	$np\left(1-p\right)$
Poisson: $X \sim \text{Pois}(\mu)$	$e^{-\mu}\mu^x/x!$	$See\ table\ above.$	λt	λt
Exponential: $T \sim \exp\{(\lambda)\}$	$\lambda e^{-\lambda t}$	$1 - \lambda e^{-\lambda t}$	$1/\lambda$	$1/\lambda^2$
Uniform: $X \sim U(a, b)$	$1/\left(b-a\right)$	(x-a)/(b-a)	(a+b)/2	$\left(b-a\right)^2/12$
Normal: $X \sim N(\mu, \sigma^2)$	_	_	μ	σ^2

Sample Statistics

Given n samples x_i :

$$\begin{aligned} \text{Mean:} \quad & \overline{x} = \frac{1}{n} \sum_{i=1}^{n} x_i \\ \text{Variance:} \quad & s^2 = \frac{1}{n-1} \sum_{i=1}^{n} \left(x_i - \overline{x} \right)^2 \\ & = \frac{1}{n-1} \left(\sum_{i=1}^{n} x_i^2 - n \overline{x}^2 \right) \end{aligned}$$

t Distribution

sample mean is distributed according to the t distribution where

$$\begin{split} T &= \frac{\overline{X} - \mu}{s/\sqrt{n}} \sim t_{n-1} \\ \Pr\left(T \leq t_{n-1,\,1-\alpha/2}\right) &= 1 - \frac{\alpha}{2} \\ \Pr\left(T > t_{n-1,\,\alpha/2}\right) &= \frac{\alpha}{2} \end{split}$$

Linear Regression

Given a set of n points (x_i, y_i) that are assumed to have a linear relationship, the model for y_i is

$$\begin{aligned} y_i &= \beta_0 + \beta_1 x_i + \varepsilon_i \\ \text{where } \varepsilon_i &\sim \text{N}\left(0,\,\sigma^2\right) \text{ is the residual. } s \\ \text{(RMS error) gives an estimate for } \sigma &\longrightarrow \\ \text{how close the data is to the model.} \end{aligned}$$

Estimates

$$T_{\text{test}}: \frac{\hat{\beta}_0 - \beta_0}{s_{\hat{\beta}_0}} \sim t_{n-2}, \quad \frac{\hat{\beta}_1 - \beta_1}{s_{\hat{\beta}_1}} \sim t_{n-2}$$

Confidence Intervals

 $\begin{array}{ll} 0 \leq p(x) \leq 1 & \overline{f(x) \geq 0} \\ \sum_{u \leq x} p(u) & \int_{-\infty}^{x} f(u) \, \mathrm{d}u \\ \sum_{\Omega} x p(x) & \int_{\Omega} x f(x) \, \mathrm{d}x \\ \underline{\sum}_{\Omega} \left(x - \mu\right)^{2} p(x) & \int_{\Omega} \left(x - \mu\right)^{2} f(x) \, \mathrm{d}x \end{array}$

Discrete

Given the confidence level c:

$$c = 1 - \alpha$$
.

The confidence interval for
$$\overline{x}$$
:
$$CI_c = \overline{x} \pm t_{n-1,\,1-\alpha/2} \frac{s}{\sqrt{n}}.$$

Hypothesis Testing

Hypothesis testing assesses the likelihood of observing the sample $\underline{i}\underline{f}$ the null hypothesis was true.

Hypothesis

$$H_0: \mu = \mu_0 \quad \text{vs} \quad H_1: \mu \neq \mu_0$$
 Generally there is no evidence supporting H_0 , but only evidence or lack of evidence

for rejecting H_0 .

Test Statistic

The measure of the distance that the proposed mean is from the sample mean:

$$T_{\text{test}} = \frac{\overline{x} - \mu_0}{s / \sqrt{n}}$$

p-Value

Strength of evidence against H_0 . The p-value is the α value that satisfies:

$$\begin{split} |T_{\mathrm{test}}| &= t_{n-1,\,1-\alpha/2}.\\ \Pr\left(T \leq |T_{\mathrm{test}}|\right) &= 1 - \frac{\alpha}{2}\\ \Pr\left(T > |T_{\mathrm{test}}|\right) &= \frac{\alpha}{2} \end{split}$$

Strength of evidence against H_0

very strong	strong		some	weak	insuffi	cient
0.0	001 (0.01	0.0	05	0.1	

Inferences about β_0 and β_1

$$CI_c = \hat{\beta}_0 \pm t_{n-2, 1-\alpha/2} s_{\hat{\beta}_0}$$
$$CI_c = \hat{\beta}_1 \pm t_{n-2, 1-\alpha/2} s_{\hat{\beta}}$$

 $CI_c = \hat{\beta}_1 \pm t_{n-2,\,1-\alpha/2} s_{\hat{\beta}_1}$ where $s_{\hat{\beta}_0}$ and $s_{\hat{\beta}_1}$ are the standard errors (SE) for $\hat{\beta}_0$ and $\hat{\beta}_1$. $H_{0,\,\beta_0}:\beta_0=0$ tests whether the model crosses the origin. $H_{0,\beta_1}:\beta_1=0$ tests whether the model is constant.

The percentage of the observed variance in y that is explained by the model.

Residual Plots

Test the following two assumptions:

- 1. The relationship between X and Yis best modelled linearly — clear indication of a non-linear trend suggests assumption is not valid.
- 2. The variance of residuals is the same for all observations, (not affected by y_i) — uneven width of residual suggests assumption is not valid. This may lead to inaccurate inferences.