

Multivariate Normal distⁿ

$\underline{X} = (x_1 \ x_2 \ \dots \ x_n)^T$ is vector which is an r.v. following Normal

$$f_{\underline{X}}(\underline{x}) = \frac{1}{(2\pi)^{n/2} \sqrt{|\Sigma|}} e^{-\frac{1}{2} (\underline{x} - \underline{\mu})^T \Sigma^{-1} (\underline{x} - \underline{\mu})}$$

$$\underline{x} \in \mathbb{R}^N$$

Here, $E(\underline{X}) = \underline{\mu}$

and Σ is a variance covariance matrix of elements of \underline{X}

$$\Sigma = \begin{pmatrix} V(x_1) & \text{cov}(x_1, x_2) & \dots & \text{cov}(x_1, x_n) \\ & V(x_2) & \dots & \text{cov}(x_2, x_n) \\ & & \dots & \\ & & & V(x_n) \end{pmatrix}$$

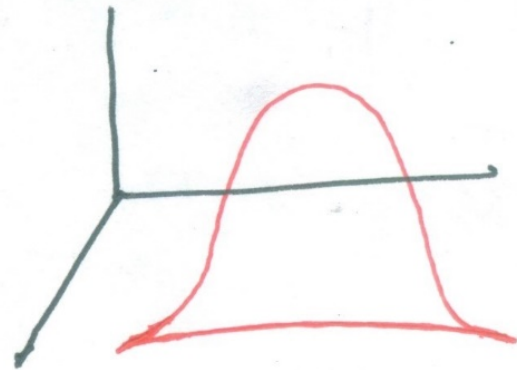
Example $\underline{X} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}$

$$f_{\underline{X}}(\underline{x}) = \frac{1}{(2\pi) \sqrt{|\Sigma|}} e^{-\frac{1}{2} (\underline{x} - \underline{\mu})^T \Sigma^{-1} (\underline{x} - \underline{\mu})}$$

Exponent $\left[\begin{pmatrix} x_1 \\ x_2 \end{pmatrix} - \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} \right]^T \begin{pmatrix} v(x_1) & \text{cov}(x_1, x_2) \\ & v(x_2) \end{pmatrix}^{-1} \left[\begin{pmatrix} x_1 \\ x_2 \end{pmatrix} - \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} \right]$

$$\left[(x_1 - \mu_1) \quad (x_2 - \mu_2) \right] \begin{pmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{pmatrix}^{-1} \begin{pmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{pmatrix}$$

$$\underline{X} \sim N_n(\underline{\mu}, \Sigma)$$



Results:

1. $\underline{X} \sim N_n(\underline{\mu}, \underline{\Sigma})$

Then for any \underline{l} , $\underline{l}'\underline{X} \sim N_1(\underline{l}'\underline{\mu}, \underline{l}'\underline{\Sigma}\underline{l})$

$$E(\underline{l}'\underline{X}) = \underline{l}'E(\underline{X}) = \underline{l}'\underline{\mu}$$

$$V(\underline{l}'\underline{X}) = \underline{l}'V(\underline{X})\underline{l} = \underline{l}'\underline{\Sigma}\underline{l}$$

→ For a multivariate Normal, each component separately is univariate Normal.

$$\underline{l} = (0 \ 0 \ \dots \ 0 \ 1 \ 0 \ \dots \ 0)$$

So, If $\exists \underline{\mu}$ and $\underline{\Sigma}$ such that for any \underline{l} , $\underline{l}'\underline{X} \sim N(\underline{l}'\underline{\mu}, \underline{l}'\underline{\Sigma}\underline{l})$
then $\underline{X} \sim N_n(\underline{\mu}, \underline{\Sigma})$.

Variance Covariance matrix

$$\text{If } \Sigma = \text{diag}(\sigma_1 \ \sigma_2 \ \dots \ \sigma_n)$$

then • x_1, x_2, \dots, x_n are independent

$$\bullet \text{cov}(x_i, x_j) = 0$$

- x_1, x_2, \dots, x_n are n ^{independently} Normally distributed r.v.s having means $\mu_1, \mu_2, \dots, \mu_n$ and variance 1, then

$$\underline{X} = (x_1, \dots, x_n)^T \sim N_n(\underline{\mu}, \underline{I})$$

Again if $\underline{\mu} = \underline{0}$

$$\text{then } \underline{X} \sim N_n(\underline{0}, \underline{I})$$

So, $x_1 \sim N(0,1)$
 $x_2 \sim N(0,1)$
 \vdots
 $x_n \sim N(0,1)$ \rightarrow independent then $\underline{X} = \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{pmatrix} \sim N_n(\underline{0}, \underline{I})$

(3)

$$\text{Let } \underline{U} = (x_1 \ x_2 \ \dots \ x_r)^T$$

$$\text{and } \underline{V} = (x_{r+1} \ x_{r+2} \ \dots \ x_n)^T$$

$$\text{Thus, } \underline{X} = \begin{pmatrix} \underline{U} \\ \underline{V} \end{pmatrix}$$

$$\underline{X} \sim N_n(\underline{\mu}, \Sigma)$$

$$\text{Then } \underline{U} \sim N_r(\underline{\mu}_1, \Sigma_1)$$

$$\text{and } \underline{V} \sim N_{n-r}(\underline{\mu}_2, \Sigma_2)$$

$$\text{Here, } \underline{\mu}_1 = (\mu_1 \ \mu_2 \ \dots \ \mu_r)^T$$

$$\text{and } \underline{\mu}_2 = (\mu_{r+1} \ \mu_{r+2} \ \dots \ \mu_n)^T \quad \text{Thus, } \underline{\mu} = \begin{pmatrix} \underline{\mu}_1 \\ \underline{\mu}_2 \end{pmatrix}$$

$$\Sigma = \begin{pmatrix} \sigma_1 & \sigma_{12} & \sigma_{13} & \dots & \sigma_{1r} & \sigma_{1r+1} & \dots & \sigma_{1n} \\ & \sigma_2 & \sigma_{23} & \dots & \sigma_{2r} & \sigma_{2r+1} & \dots & \sigma_{2n} \\ & & & & \vdots & & & \\ & & & & \sigma_r & \sigma_{rr+1} & \dots & \sigma_{rn} \\ \hline & & & & & \sigma_{r+1} & \dots & \sigma_{r+1n} \\ & & & & & & \ddots & \\ & & & & & & & \sigma_n \end{pmatrix}$$

$$\begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}$$

If $\underline{U}, \underline{V}$ are independent then $\Sigma_{12} = 0$

Let $C_{m \times n}$ matrix

Define $\underline{Y}_{m \times 1} = C_{m \times n} \underline{X}_{n \times 1}$

$$\text{Then } E(\underline{Y}) = C_{m \times n} E(\underline{X}) \\ = C \underline{\mu}$$

$$\text{and } V(\underline{Y}) = C \Sigma C^T$$

$$\Rightarrow \underline{Y} \sim N_m(C \underline{\mu}, C \Sigma C^T)$$

Consider a linear fcn $\underline{l}^T \underline{Y}$

$$\underline{l}^T \underline{Y} \sim N_1(\underline{l}' C \underline{\mu}, \underline{l}' C \Sigma C' \underline{l})$$

[$\underline{l}' C \underline{X}$ is also a linear fcn defined on \underline{X}
 $\underline{P}' \underline{X} \sim N_1(\underline{P}' \underline{\mu}, \underline{P}' \Sigma \underline{P})$ or $N_1(\underline{l}' C \underline{\mu}, \underline{l}' C \Sigma C' \underline{l})$ $\underline{P} = C^T \underline{l}$]

$$\begin{aligned} \underline{U}_1 &\sim N_n(\underline{\mu}_1, \Sigma_1) \\ \underline{U}_2 &\sim N_n(\underline{\mu}_2, \Sigma_2) \end{aligned} \bigg\} \text{indep}$$

Then, $\underline{U}_1 + \underline{U}_2 = \underline{U} \sim N_n(\underline{\mu}, \Sigma)$

$$\underline{\mu} = (\underline{\mu}_1 + \underline{\mu}_2) \quad \Sigma = \begin{pmatrix} \cancel{\Sigma_1} & \cancel{0} \\ \cancel{0} & \cancel{\Sigma_2} \end{pmatrix} \Sigma_1 + \Sigma_2$$

Similarly

$$\begin{aligned} X_1 &\sim N_1(\mu_1, \sigma_1^2) \\ X_2 &\sim N_1(\mu_2, \sigma_2^2) \end{aligned} \bigg\} \text{indep.}$$

$$Y = X_1 + X_2 \sim N_1(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

Mixture of Normal ?

Important Result (CLT)

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X_1, X_2, \dots, X_n are iid r.v.s with $E(x_i) = \mu$
and $V(x_i) = \sigma^2$

Define $S_n = \frac{\sum_i x_i}{n}$

$$\frac{S_n - \mu}{\sigma/\sqrt{n}} \sim N(0,1) \quad \text{for large value of } n.$$

Example $X_i \sim \text{Rec}(0,1)$

$$E(x_i) = \frac{1}{2} \quad \text{and} \quad V(x_i) = \frac{1}{12}$$

$$\begin{aligned} S_n &= \frac{1}{n} \sum_{i=1}^n x_i, \quad S_n - \mu = \frac{1}{n} \sum_{i=1}^n x_i - \frac{1}{2} \\ &= \frac{1}{n} \left(\sum_{i=1}^n x_i - \frac{n}{2} \right) \end{aligned}$$

5.1

$$\begin{aligned}\frac{\sqrt{n}(S_n - \mu)}{\sigma} &= \sqrt{12}\sqrt{n} \frac{1}{n} \left(\sum_{i=1}^n x_i - \frac{n}{2} \right) \\ &= \frac{\sqrt{12}}{\sqrt{n}} \left(\sum_{i=1}^n x_i - \frac{n}{2} \right) \\ &= \sqrt{\frac{12}{n}} \left(\sum_{i=1}^n x_i - \frac{n}{2} \right)\end{aligned}$$

For $n=12$

$$\left(\sum_{i=1}^{12} x_i - 6 \right) \sim N(0,1)$$

Note: 1. Draw 12 random samples from $\text{Uni}(0,1)$ $\left[\begin{array}{l} \equiv \text{random numbers} \\ [0,1] \end{array} \right]$

2. $(\text{Sum of random No.s} - 6) \sim N(0,1)$

3. 2 provides one sample of standard Gaussian

$$\begin{array}{l} X_1 \sim \text{Uni}(0,1) \\ X_2 \sim \text{Uni}(0,1) \end{array} \bigg\rangle \text{indep} \left[\begin{array}{l} \equiv 2 \text{ random numbers} \\ [0,1] \end{array} \right] \quad 6$$

$$\begin{array}{l} U_1 = \sqrt{-2 \ln X_1} \cos 2\pi X_2 \\ U_2 = \sqrt{-2 \ln X_2} \sin 2\pi X_2 \end{array} \bigg\rangle \text{indep } N(0,1)$$

Note: 2 random numbers from $[0,1]$ \rightarrow 2 random samples from $N(0,1)$.